

Point-by-point reply for the review of "GHOSH v1.0.0: a novel Gauss-Hermite High-Order Sampling Hybrid filter for computationally efficient data assimilation in geosciences"

We thank the three reviewers for the accurate reading of the manuscript, which has been deeply modified to account for the proposed corrections and improvements. In particular, the manuscript has been divided in two parts: Part 1 includes the twin experiment of GHOSH in the Lorenz96 testcase; Part 2 describes a realistic geoscientific application of GHOSH. Hereafter, our point-by-point responses to the Reviewer's comments are provided in bold green with line numbers referring to the updated version of the manuscript without track change.

Reply on Reviewer comment RC1

Summary: The paper present an interesting idea of using high-order sampling to improve the accuracy of higher statistical moments of the ensemble, therefore achieving better skill in ensemble prediction. I was expecting to learn some details about the new approach, but to my disappointment I am still quite confused after reading the paper. In my opinion, the paper tries to achieve too many things at once: introduce the method mathematically and show experiment results in two different models. No wonder that you don't have enough room to provide the details. In the methodology, many steps in the formula are stated as "can be proven" but with no guidance for readers to actually understand the logic (skipping too many steps in a proof). Ideally, when introducing a new method, I would expect a paper focused on the theory and illustrations using simple models; then followed by other paper(s) utilizing the method in bigger models describing how you solved any additional issues in implementation. The current paper did the introduction poorly, I suggest the authors to give it a thorough revision, at least address the following major issues.

We thank the Reviewer for the comments and suggestions, which highlight the need to state more clearly the goal and novelty of our manuscript, and of improving its readability.

We kept this in consideration in the revised version of the manuscript, where we clarified the main objective of the work, since it appears this was not clear enough. We are introducing a new method (GHOSH filter) that is focused on the use of a higher order of approximation to evaluate the ensemble mean in data assimilation applications.

As highlighted also by the Reviewer #3, we recognize that we have a lot of material, and possibly too much for a single paper. We received a positive answer from the Topic Editor about a request on the possibility to split the manuscript in two parts, and we prepared a revised version of the manuscript divided in two parts.

Major Issues:

1. It's unclear what is the actual size of the state vector now you are employing a weighted ensemble approach. When you state that you use $2r$ members to achieve 3rd-order sampling, such as the Wan & van der Merwe 2000 paper, the members are paired with mean $\pm r$ perturbations, does it mean you only need to store $r+1$ numbers?

Thank you for this remark on the size of the state vector. If the state vector has dimension N , then an ensemble method with ensemble size m must store m state vectors, for a total of m times N numbers. If the ensemble is weighted, also the weights must be considered, so the total becomes $mN + m$, as in the particle filter (PF) case. Differently from PF, the GHOSH weights are constant in time, thus the degrees of freedom are mN . In Wan & van der Merwe 2000, they use an ensemble size of $m=2r+1$ in the attempt of achieving a 3rd-order sampling. In comparison, ETKF and SEIK use $m=r+1$ ensemble members for their 2nd-order sampling. The GHOSH novel strategy tries to exploit the best of the two approaches: it uses only $m=r+1$ members (as the 2nd-order methods), but those members are chosen in such a way that in the principal components the order of the sampling is higher than 2. The simplest example of the last statement in the case of a 2-dimensional standard normal error is the following 2nd-order sampling, achieving 3rd-order in the x direction:

$$P1: (x1, y1) = (0, \sqrt{2}),$$

$$P2: (x2, y2) = (\sqrt{3/2}, -\sqrt{2}/2),$$

$$P3: (x3, y3) = (-\sqrt{3/2}, -\sqrt{2}/2).$$

In fact:

the 1st moment associated to x (i.e., the mean along x) is:

$$(x1 + x2 + x3)/3 = (0 + \sqrt{3/2} + (-\sqrt{3/2}))/3 = 0,$$

the 1st moment associated to y (i.e., the mean along y) is:

$$(y1 + y2 + y3)/3 = (\sqrt{2} + (-\sqrt{2}/2) + (-\sqrt{2}/2))/3 = 0,$$

the 2nd moment associated to x^2 (i.e., the variance of x) is:

$$(x1^2 + x2^2 + x3^2)/3 = (0 + 3/2 + 3/2)/3 = 1,$$

the 2nd moment associated to y^2 (i.e., the variance of y) is:

$$(y1^2 + y2^2 + y3^2)/3 = (2 + 1/2 + 1/2)/3 = 1,$$

the 2nd moment associated to xy (i.e., the covariance between x and y) is:

$$(x1 \cdot y1 + x2 \cdot y2 + x3 \cdot y3)/3 = (0 \cdot \sqrt{2} + \sqrt{3/2} \cdot (-\sqrt{2}/2) + (-\sqrt{3/2}) \cdot (-\sqrt{2}/2))/3 = 0,$$

the 3rd moment associated to x^3 is:

$$(x1^3 + x2^3 + x3^3)/3 = (0 + \sqrt{3/2}^3 + (-\sqrt{3/2})^3)/3 = 0.$$

Thus, the ensemble P1, P2, P3 is compliant with the statistical moments of a standard normal distribution up to order 2, and also have the same 0-skewness along the x-axis.

In the manuscript, we included a new Appendix C to guide the reader through some simple examples (similar to the one above) in order to improve the manuscript readability (see also answers to the Reviewer's next comments).

The weights are computed according to A3 which is a nonlinear system itself. What is the computational complexity of this step? You claimed that the GHOSH filter is no more costly than the 2nd-order EnKFs with same ensemble size, but can you provide some complexity arguments for each step in your filter? For example, given ensemble size m , obs count p and model state size n , the ETKF has complexity $m^2 p + m^3 + m^3 n$, according to Tippett 2003 (doi: 10.1175/1520-0493(2003)131<1485:ESRF>2.0.CO;2). Although with fixed ensemble size, the filter is doing additional steps (e.g. the weight computation). I'm not convinced that they are comparable.

The fact that GHOSH is no more costly than the 2nd-order EnKFs takes into account that the nonlinear system where weights are computed needs to be solved only once. Its computational cost depends on the chosen nonlinear solver. It should also be noted that, when available, analytical solutions should be preferred. During the initialization of the GHOSH filter, the provided python code (in the Code Availability section) uses an analytical constructive method to solve the system up to order 5 in $O(m^2)$ operations.

The GHOSH computational complexity is the same as ETKF, i.e., $O(m^2 p + m^3 + m^2 n)$, since the GHOSH filter requires a few more operations per step (e.g., resampling and eigenvalue decomposition) that are $O(m^3 + m^2 n)$ and do not add to the computational complexity. This means that, even if the GHOSH filter needs a little more computational time, it scales as the ETKF, the SEIK, or similar EnKFs. Further, in realistic applications, the EnKF computational cost is usually widely dominated by the model integration cost (and, secondly, by I/O), while forecast and analysis do not represent a concern.

In the revised manuscript, we followed the Reviewer's suggestion of evaluating explicitly the computational complexity (Part 1, L. 392, Section 3.1.6). In addition, we discussed timings for the two filters in the Lorenz96 twin experiment (according also to the Reviewer's comments). In particular, we modified the manuscript as follows (Part 1, L. 478):

“From the computational point of view, the whole experiment set needed around 9 computational hours. SEIK and GHOSH schemes used 12% and 16% of the total time respectively, while the rest was committed to model integration. The GHOSH filter executes more operations than SEIK (e.g., an eigenvalue decomposition) resulting in more computational time even if the asymptotic computational complexity of the two methods is the same (Section 3.1.6). However, even in the case of a relatively simple model like the Lorenz96, the time to solution is dominated by the model integration and the difference between GHOSH and SEIK only accounts for 4% of the total time.

Unexpectedly, the model integration time (averaged every 100 twin experiments) when applying the SEIK filter sometimes is longer than the GHOSH filter case, up to twice the time, depending on some settings and random factors.”

2. Fig. 1 provides some visual guide to understand the statement about "m sample points are required to represent a hth-order approximation of an error distribution in r-dimensional space". I can see that in 3D you need 4 points to span a volume (the covariance), and I can see that you need 4 degrees of freedom to provide a unique 3x3 covariance matrix. But from this point onward, I don't understand anymore.

First, the error distribution depicted is symmetric, meaning there is only the mean and identity covariance, there is not even correlation between variables in different dimensions, how can I visualize the higher-order moments (skewness)? Therefore, I don't see how the 4 points are giving accurate 3rd moment in 2D, there is no skewness anyway in the picture.

Considering this and the previous Reviewer's comments and comments from Reviewer #3, we described Fig. 1 in more detail in a novel appendix (Appendix C) adding the description (including equations and calculations) of the concepts summarised in Fig. 1 panels (e.g., computation of the moments).

The novel Appendix C shows all the detailed computations needed to ensure that a reader could follow each step. For example, considering 4 points (P1, P2, P3 and P4) giving accurate 3rd moments in 2D, the reader will be guided as follows:

$$P1: (x1, y1) = (\sqrt{2}, 0),$$

$$P2: (x2, y2) = (0, \sqrt{2}),$$

$$P3: (x3, y3) = (-\sqrt{2}, 0),$$

$$P4: (x4, y4) = (0, -\sqrt{2}).$$

The 1st moment associated to x (i.e., the mean along x) is:

$$(x1 + x2 + x3 + x4)/4 = (\sqrt{2} + 0 + (-\sqrt{2}) + 0)/4 = 0,$$

the 1st moment associated to y (i.e., the mean along y) is:

$$(y1 + y2 + y3 + y4)/4 = (0 + \sqrt{2} + 0 + (-\sqrt{2}))/4 = 0,$$

the 2nd moment associated to x^2 (i.e., the variance of x) is:

$$(x1^2 + \dots + x4^2)/4 = (2 + 0 + 2 + 0)/4 = 1,$$

the 2nd moment associated to y^2 (i.e., the variance of y) is:

$$(y1^2 + \dots + y4^2)/4 = (0 + 2 + 0 + 2)/4 = 1,$$

the 2nd moment associated to xy (i.e., the covariance between x and y) is:

$$(x1 \cdot y1 + \dots + x4 \cdot y4)/4 = (\sqrt{2} \cdot 0 + 0 \cdot \sqrt{2} + \sqrt{2} \cdot 0 + 0 \cdot \sqrt{2})/4 = 0,$$

the 3rd moment associated to x^3 is:

$$(x_1^3 + \dots + x_4^3)/4 = (\sqrt{2}^3 + 0 + (-\sqrt{2})^3 + 0)/4 = 0,$$

the 3rd moment associated to $x^2 y$ is:

$$(x_1^2 \cdot y_1 + \dots + x_4^2 \cdot y_4)/4 = (2 \cdot 0 + 0 \cdot \sqrt{2} + 2 \cdot 0 + 0 \cdot 2)/4 = 0,$$

the 3rd moment associated to $x y^2$ is:

$$(x_1 \cdot y_1^2 + \dots + x_4 \cdot y_4^2)/4 = (\sqrt{2} \cdot 0 + 0 \cdot 2 + \sqrt{2} \cdot 0 + 0 \cdot 2)/4 = 0,$$

the 3rd moment associated to y^3 is:

$$(y_1^3 + \dots + y_4^3)/4 = (0 + \sqrt{2}^3 + 0 + (-\sqrt{2})^3)/4 = 0.$$

Thus, the ensemble P1, P2, P3, P4 is compliant with the statistical moments of a standard normal distribution up to order 3.

Second, why does projection from high-dimensional to low-dimensional space relate to a higher-order sampling in the low-dimensional space? Somehow these 2 members project to the same member in lower dimension, which means they are paired in a way that projection will cancel out the perturbations and only leave the mean? But why is that equivalent to a weighted member, what is the weight in this 3-point case anyway? Linking back to the 4 points in 3D, are the positioning only allowed to be squares with no rotation? In the 3D case the tetrahedron is allowed to rotate arbitrarily. Here in 2D if the square is rotated they project to 4 points in 1D.

We totally agree with the Reviewer, each tetrahedron rotation implies a different shadow (i.e., projection). Only a few possible rotations of the tetrahedron project a perfect square in the xy plane, and even fewer rotations project a square that further projects into 3 points (instead of 4) in the x axis. The strength of GHOSH is exactly this: being able to catch the right rotation to achieve the best possible projections. Or, equivalently, the GHOSH sampling chooses ensemble members matching high order moments of the error distribution along some chosen directions (i.e., the principal components of the error), obtaining a better approximation of the mean along those directions.

Concerning the other Reviewer's comments about projections, please read the next response.

Finally, why does three weighted samples in 1D give a 5th order approximation? I don't see any moments in the sample distribution higher than the 2nd.

At this point, I would rather you don't show this figure at all, but replace Lines 89-108 with a more step-by-step reasoning, or formal proof, of the statement such as "2r samples give a 3rd-order approximation of r-dimensional space".

As discussed above, the addition of the Appendix C provides a step-by-step reasoning that will help the reader to understand the issues raised by the Reviewer.

The detailed description of Appendix C provides an explanation of what is shown in Fig. 1, including a practical example of a 5th order weighted ensemble in 1D composed by 3 points P1, P2 and P3, i.e.:

$$P1: x1 = \sqrt{3}; w1 = 1/6,$$

$$P2: x2 = -\sqrt{3}; w2 = 1/6,$$

$$P3: x3 = 0; w3 = 2/3.$$

The first moment (mean) is:

$$w1 \cdot x1 + w2 \cdot x2 + w3 \cdot x3 = 1/6 \cdot \sqrt{3} + 1/6 \cdot (-\sqrt{3}) + 2/3 \cdot 0 = 0,$$

the second moment (variance) is:

$$w1 \cdot x1^2 + w2 \cdot x2^2 + w3 \cdot x3^2 = 1/6 \cdot 3 + 1/6 \cdot 3 + 2/3 \cdot 0 = 1,$$

the third moment (skewness) is:

$$w1 \cdot x1^3 + w2 \cdot x2^3 + w3 \cdot x3^3 = 1/6 \cdot \sqrt{3}^3 + 1/6 \cdot (-\sqrt{3})^3 + 2/3 \cdot 0 = 0,$$

the fourth moment (kurtosis) is:

$$w1 \cdot x1^4 + w2 \cdot x2^4 + w3 \cdot x3^4 = 1/6 \cdot 9 + 1/6 \cdot 9 + 2/3 \cdot 0 = 3,$$

the fifth moment is:

$$w1 \cdot x1^5 + w2 \cdot x2^5 + w3 \cdot x3^5 = 1/6 \cdot \sqrt{3}^5 + 1/6 \cdot (-\sqrt{3})^5 + 2/3 \cdot 0 = 0.$$

Thus, the ensemble P1, P2, P3 is compliant with the statistical moments of a standard normal distribution up to order 5.

The property of this 1-dimensional 5th-order sampling, can be extended to a 2-dimensional ensemble, aimed to keep order 5 along the x-axis while achieving order 3 in the xy plane. Such 4-members ensemble is:

$$P1: (x1, y1) = (\sqrt{3}, 0); w1 = 1/6,$$

$$P2: (x2, y2) = (-\sqrt{3}, 0); w2 = 1/6,$$

$$P3: (x3, y3) = (0, \sqrt{3/2}); w3 = 1/3,$$

$$P4: (x4, y4) = (0, -\sqrt{3/2}); w4 = 1/3.$$

Note that all the moments along the x-axis are already computed above up to order 5, since this ensemble and the previous one are indistinguishable looking only at the x-coordinate. In fact, considering the projection on the x-axis, P3 and P4 behave as a unique member with weight $w3+w4$ because they have the same x-coordinate $x3=x4$. It remains to be checked if the moments involving y match the moments of standard normal distribution up to order 3, i.e.:

the 1st moment associated to y (i.e., the mean along y) is:

$$w1 \cdot y1 + w2 \cdot y2 + w3 \cdot y3 + w4 \cdot y4 =$$

$$= 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot \sqrt{3/2} + 1/3 \cdot (-\sqrt{3/2}) = 0,$$

the 2nd moment associated to y^2 (i.e., the variance of y) is:

$$w_1 \cdot y_1^2 + \dots + w_4 \cdot y_4^2 = 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot 3/2 + 1/3 \cdot 3/2 = 1,$$

the 2nd moment associated to xy (i.e., the covariance between x and y) is:

$$\begin{aligned} w_1 \cdot x_1 \cdot y_1 + \dots + w_4 \cdot x_4 \cdot y_4 &= \\ = 1/6 \cdot \sqrt{3} \cdot 0 + 1/6 \cdot (-\sqrt{3}) \cdot 0 + 1/3 \cdot 0 \cdot \sqrt{3/2} + 1/3 \cdot 0 \cdot (-\sqrt{3/2}) &= 0, \end{aligned}$$

the 3rd moment associated to $x^2 y$ is:

$$\begin{aligned} w_1 \cdot x_1^2 \cdot y_1 + \dots + w_4 \cdot x_4^2 \cdot y_4 &= \\ = 1/6 \cdot 3 \cdot 0 + 1/6 \cdot 3 \cdot 0 + 1/3 \cdot 0 \cdot \sqrt{3/2} + 1/3 \cdot 0 \cdot (-\sqrt{3/2}) &= 0, \end{aligned}$$

the 3rd moment associated to $x y^2$ is:

$$\begin{aligned} w_1 \cdot x_1 \cdot y_1^2 + \dots + w_4 \cdot x_4 \cdot y_4^2 &= \\ = 1/6 \cdot \sqrt{3} \cdot 0 + 1/6 \cdot (-\sqrt{3}) \cdot 0 + 1/3 \cdot 0 \cdot 3/2 + 1/3 \cdot 0 \cdot 3/2 &= 0, \end{aligned}$$

the 3rd moment associated to y^3 is:

$$w_1 \cdot y_1^3 + \dots + w_4 \cdot y_4^3 = 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot \sqrt{3/2}^3 + 1/3 \cdot (-\sqrt{3/2})^3 = 0.$$

The dimensions can be increased once more by building in a similar way a 3-dimensional ensemble, as shown in Fig. 1. This ensemble is indistinguishable from the previous one in the xy plane projection and capable of achieving 2nd order on the xyz space. Such ensemble is:

$$P_1: (x_1, y_1, z_1) = (\sqrt{3}, 0, -\sqrt{2}); w_1 = 1/6,$$

$$P_2: (x_2, y_2, z_2) = (-\sqrt{3}, 0, -\sqrt{2}); w_2 = 1/6,$$

$$P_3: (x_3, y_3, z_3) = (0, \sqrt{3/2}, \sqrt{2}/2); w_3 = 1/3,$$

$$P_4: (x_4, y_4, z_4) = (0, -\sqrt{3/2}, \sqrt{2}/2); w_4 = 1/3.$$

All the moments involving x and y are already checked in the previous calculation up to order 5 along x and up to order 3 in the xy plane. It remains to be checked if the moments involving z match the moments of standard normal distribution up to order 2, i.e.:

the 1st moment associated to z (i.e., the mean along z) is:

$$\begin{aligned} w_1 \cdot z_1 + w_2 \cdot z_2 + w_3 \cdot z_3 + w_4 \cdot z_4 &= \\ = 1/6 \cdot (-\sqrt{2}) + 1/6 \cdot (-\sqrt{2}) + 1/3 \cdot \sqrt{2}/2 + 1/3 \cdot \sqrt{2}/2 &= 0, \end{aligned}$$

the 2nd moment associated to z^2 (i.e., the variance of z) is:

$$w_1 \cdot z_1^2 + \dots + w_4 \cdot z_4^2 = 1/6 \cdot 2 + 1/6 \cdot 2 + 1/3 \cdot 1/2 + 1/3 \cdot 1/2 = 1,$$

the 2nd moment associated to xz (i.e., the covariance between x and z) is:

$$\begin{aligned} w_1 \cdot x_1 \cdot z_1 + \dots + w_4 \cdot x_4 \cdot z_4 &= \\ &= 1/6 \cdot \sqrt{3} \cdot (-\sqrt{2}) + 1/6 \cdot (-\sqrt{3}) \cdot (-\sqrt{2}) + 1/3 \cdot 0 \cdot \sqrt{2}/2 + 1/3 \cdot 0 \cdot \sqrt{2}/2 = 0, \end{aligned}$$

the 2nd moment associated to yz (i.e., the covariance between y and z) is:

$$\begin{aligned} w_1 \cdot y_1 \cdot z_1 + \dots + w_4 \cdot y_4 \cdot z_4 &= \\ &= 1/6 \cdot 0 \cdot (-\sqrt{2}) + 1/6 \cdot 0 \cdot (-\sqrt{2}) + 1/3 \cdot \sqrt{3}/2 \cdot \sqrt{2}/2 + 1/3 \cdot (-\sqrt{3}/2) \cdot \sqrt{2}/2 = 0. \end{aligned}$$

All the above proves that the weighted ensemble P1, P2, P3, P4 is a second-order ensemble achieving order 3 in the subspace of the x and y directions and order 5 along the x direction.

We are confident that these step-by-step calculations can help the Reader in following the manuscript arguments.

3. In your Lorenz96 experiments (Fig. 3), based on the ensemble spread compared to the error, the SEIK filter has basically diverged at around $t=5$. The climatological error is about 4 for the Lorenz96 system and you can see that after $t=5$ the SEIK filter solution is as bad as random draw from climatology. I would argue that you need to have a more stable setting for both SEIK and GHOSH here to establish a comparison, especially when we know that SEIK should work for the Lorenz96 system. For some cases when the SEIK is stable, can you still show that the new GHOSH filter will achieve higher accuracy or at least perform as well as the SEIK?

We thank the Reviewer for pointing out that the example in Fig. 3 was poorly chosen, since it shows a situation where SEIK does not improve with respect to climatology. Thanks to the above comment and to comments from other Reviewers, Figure 3 was changed in order to represent a better example (Fig. R1). Since SEIK and GHOSH present different behaviours depending on other parameters (e.g., ensemble size, observation frequency, forgetting factor), Fig. 4 resumes the whole set of 88000 Lorenz96 experiments that have been carried out by averaging blocks of 400 experiments with similar setup (but different random parameters, e.g. observations). Moreover, Fig. 4 shows that SEIK was tested also in stable and convergent conditions, i.e., green squares in the two top panels in Fig. 4. In addition, it is worth noting that the comparison between SEIK and GHOSH has also been tested considering the best tuning in terms of inflation for both filters (“best” line in each panel of Fig. 4) always resulting in GHOSH working as good as SEIK or better. Finally, consider that the set of Lorenz96 experiments was enlarged in the revised manuscript to include additional inflation (forgetting factor) values.

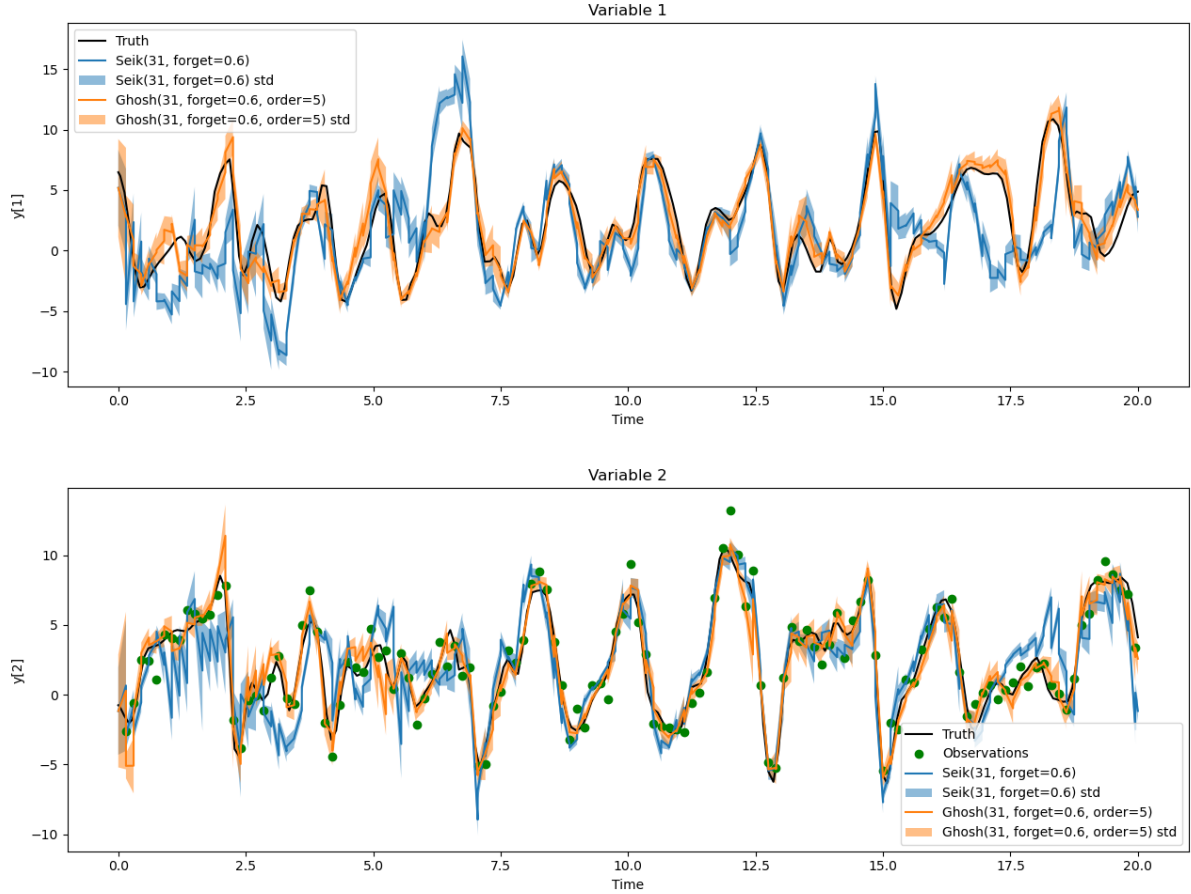


Fig. R1. Results from one twin experiment example: time between observation (green dots) is 0.15, forgetting factor is 0.6, ensemble size is 31. Shady areas around blue (SEIK) and orange (GHOSH) lines represent the ensemble standard deviation, truth is the black line. Top panel: time evolution of the first variable (non-assimilated); bottom panel: time evolution of the second variable (assimilated).

4. In the verification of filter performance for both models, I mainly see the use of RMSE and ensemble spread as error metrics. This means you are still evaluating the first two moments of error distributions. However, the GHOSH filter is motivated by its capability of capturing higher moments in error. Why didn't you show any proof that it is behaving as expected? For example, in the Lorenz96 case you can compute the skewness and kurtosis of the error distribution and compare the two filters. This will truly be a proof of concept, rather than implying the skill from the forecast error in ensemble mean (1st moment).

As discussed above, the main goal of the GHOSH filter is to achieve higher approximation order on the ensemble mean without focus on higher order moments. Thus, we think that RMSE is a sound metric to compare the GHOSH performances with respect to SEIK.

Minor Issues:

The use of nested parentheses in citation is a bit confusing. Try use (author year; author year)? Or is there a preferred format for GMD?

We thank the Reviewer for the suggestion. The GMD format does not require nested parentheses, and we updated all the citations accordingly.

Line 42: "Monte Carlo"?

We corrected the typo, thank you

Line 51: Is there a reference for this? Why is closer together samples more suitable for accuracy using second-order approximation?

The sentence is related to the fact that the nonlinear effects are smaller if the perturbation of the ensemble members is small and hence the second-order approximation holds better. In other words, if the true uncertainty is small, then a second order approximation is affected by a smaller error. According also to Reviewer #2 and #3 comments, we propose to modify the sentence focusing on uncertainty instead of ensemble spread (Part 1, L. 40):

“Furthermore, the second-order approximation is more effective the closer the ensemble members are to each other (i.e., small uncertainty), thus the higher the uncertainty the worse will be the approximation error in the mean computation. Since the state estimation is often affected by a relatively high uncertainty in geosciences data assimilation applications, this approximation error may be not negligible.”

Line 54: It's a bit not clear what you mean by "higher order of approximation" here. I think you refer to the higher moments of the distribution that the samples will exactly reproduce after model integration. If I'm not completely missing the point, this can be stated more clearly to help the readers follow.

Thanks to this Reviewer comment and to issues raised by Reviewer #3, we clarified the definition of order of approximation (Part 1, L. 29):

“The number of ensemble members can be reduced by adopting deterministic sampling methods (as opposed to stochastic EnKF methods, see e.g. Carrassi et al., 2018). Examples of deterministic EnKF using second-order sampling methods are SEIK (Pham, 2001) and ETKF (Bishop et al., 2001). Extending the definition of second-order exact methods in Pham (2001), with the term “order” we refer to the polynomial order of approximation of the filter or of its sampling method. Namely, in the case of a hth-order approximation, the ensemble has the property of providing a forecast mean with no error as long as the evolution function used for forecasting

(i.e., the model) is a polynomial of order h . As proven in Appendix A, this is achieved by sampling an ensemble that matches the first h statistical moments of the uncertainty probability density function (pdf) before evolution (also called prior in Bayesian statistics)."

Line 57: I can understand why you need $r+1$ samples to give mean and covariance (2nd order approximation), but it's a bit hard to see why $2r$ for 3rd order approximation. You say it can be proven, could you give a reference, or include a proof in appendix?

The details that are provided in the novel Appendix C now clarify the need of $2r$ samples for 3rd order approximation.

Line 59: apart from the unscented Kalman filter, other methods that preserve the 3rd moment of posterior pdf, such as the one by Hodyss 2011 (DOI: 10.1175/2011MWR3558.1), is also worth mentioning here.

Since the main aim of the GHOSH filter is a higher order of approximation of the mean and not to provide higher-than-second moments, we think that the citation is not useful in the sentence at former L. 59. However, it is perfectly suited for discussing the applicability of the GHOSH sampling to non-Gaussian priors and it was added to Section 6 of the manuscript in a new dedicated paragraph (Part 1, L. 535). We kindly thank the Reviewer for the suggested high-quality reference.

Line 95: Pham 1996 is the introduction paper for SEIK, are you sure this is the right citation for the second-order exact sampling?

Indeed, the second order exact sampling is the sampling method used in the SEIK filter, as explained in the Appendix of the cited paper. However, as noted also by Reviewer 3, this citation was improved to a more recent one (Pham, 2001).

Line 105-108: this paragraph is very vague, we achieve a fifth order approximation when the variance is "bigger" on x than y , can you be more precise exactly when do we achieve fifth order?

We thank the Reviewer for this comment, that helped us to make the manuscript clearer. The sampling is of order 5 in x , and if the variance on the x coordinate is bigger than on the other coordinates, then the best order of approximation is achieved along the direction that is affected by the largest error (as explained in comment to former L. 51). The sentence was modified to be more clear (Part 1, L. 90):

"Meaning that, if the z component of the distribution was less relevant than the x and y components (e.g., a non-standard normal distribution with smaller variance on the z coordinate), then the GHOSH third-order approximation mitigates the approximation error in the x and y dimensions where the spread is higher. And, if the variance is

larger on x than y , then the achieved fifth-order approximation reduces the error more efficiently where the spread is maximum (x coordinate).”

Line 112: h -order, shouldn't it be h th-order?

Thank you for the suggestion. We updated the term to h th-order.

Line 128: putting index as superscripts are very confusing, like for this equation, is $v_m^{2-x_i}$ raised to $2-x_i$ power or just indexed by $2-x_i$? Either put a parenthesis around them if they are indices, or just put the index as subscripts. It's also confusing to see a tuple $j_1, \dots, j_{x_i}$ ending up being a single index from 1 to s . Please provide more context in (1) about other equations, I guess the other equations in the set are matching a different moment indexed by $j_1, \dots, j_{x_i}$? How many equations are there in total?

According to the Reviewer comment, all the indexes are now subscripted in the updated manuscript. Moreover, according to comments provided also by Reviewer 3, the whole system (1) and the following paragraph were improved. More equations in (1) were explicitly written to help the manuscript readability and the text now guides more accurately the reader to understand equation (1).

Line 137: is there a formal relationship you can write down for r and s given h ? First, you will need $r > s$ to establish the principle components, otherwise there is rank deficiency. But what to expect given h ? Previously you stated that for $h=2$ you need $r \geq s+1$, and for $h=3$ you need $r \geq 2s$, in the 1D case you showed for $h=5$, you have $r \geq 4s$. I'm wildly guessing here that " $r \geq (h-1)s$ " might be what you imply?

Unfortunately, the general relation between r , h , and s is non-linear and we are not aware of any closed form to state it. On the other hand, it is true that certain values for r , h and s are valid if and only if system (1) admits a solution. The above relations for $h=2$ and $h=3$ holds (changing r with $r+1$, since probably the Reviewer was referring to the ensemble size). In the case $h=5$ we know that system (1) has a solution at least for $r \geq 2^{(s+1)} - 2$. This comes from the analytical constructive solution to system (1) implemented in the initialization of the GHOSH filter in the provided python code.

Line 179: why is the error covariance approximated not exactly given by the ensemble perturbation matrix?

We thank the Reviewer for pointing out the ambiguity of this statement. The mean and covariance of the analysis are approximations, since they are derived by the forecast mean and covariance, which are approximations affected by a sampling error. On the other hand, analysis mean and covariance after sampling are exact in the sense that

the ensemble is chosen to match those quantities. In order to remove any misunderstanding, we changed the superscripted “a” with “f” in equations (9)-(12).

Line 247: (26) is the same as (6), you can just remove (26) and say that C and Omega are computed using x^a in (4)-(7). By the way, what is the computational cost for this step? You have a eigenvalue decomposition here so at least $O(r^3)$?

Since the focus of the analysis is X^a , which is computed in (26), we prefer to keep that equation, even if it could be a little redundant. Also considering Reviewer #3 comments, former L. 248 was modified to make clear that (6) is not the only relevant equation (also equations (4)-(7) play a relevant role) (Part 1, L. 246). The computational cost of the eigenvalue decomposition depends on the adopted algorithm but it is safe to consider it $O(r^3)$. The resampling step has a computational complexity of $O(r^3 + N r^2)$, where the last term comes from the matrix multiplication with L. The addition of detailed consideration about the computational complexity now helps the reader to have an evaluation of the GHOSH costs (see major issue 1).

Line 296: what is l' ?

In the manuscript we called N the state vector dimension, and n the dimension of the observation vector. The localization process uses three operators. The first one reduces the dimension of the state vector from N to l. The second one reduces the dimension of the observation vector from n to l' . In general, l' is not equal to l, as N is not equal to n. The third one is the delocalization operator, used to build back a N-dimensional state vector.

Table 1 can be removed, your experiment name already encode the value of the parameters as listed.

We prefer to keep Table 1, since some experiment names only appear in the table. Moreover, Table 1 is useful for the reader to visualise all the configurations of the 3D experiments.

Line 365: have you defined the "forgetting factor" anywhere yet? Is it the same one from Pham 1996?

Yes, the forgetting factor is the same as Pham 1996, and it is defined at L. 264. To help the reader, we added a reference to the relevant equation at L. 394 (Part 1): “Both filters have the same forgetting factor (equation (32))”

Line 469: what range of values have you tested and how many different experiments have you run for the tuning process beyond the 18 you showed. This will provide an idea of the cost for tuning as well.

As observed by the Reviewer, the tuning was not limited to the 18 experiments listed in Table 1. However, some of the tuning simulations were very short or have been focused on validating GHOSH effects on a subset of the variables in the biogeochemical model. Moreover, tuning cost can depend on the typical characteristics of the realistic application. Thus, we think that adding the number of runs carried out for tuning the filter is not very informative for the reader. However, we slightly modify the sentence to better describe the tuning procedure (Part 2, L. 121):

“In the assimilation experiment, the GHOSH filter is configured to achieve an approximation order $h = 5$ in a subspace of dimension $s = 3$. A forecast, which includes inflation ($\rho = 0.9$), hybridization with Q and resampling, is performed with daily frequency. Since a tuning of the parametric uncertainty (represented in Q) and observation uncertainty (R) may be necessary to produce an effective assimilation when system uncertainties are not fully known (e.g., Cossarini et al., 2019; Ford and Barciela, 2017; Janjic´ et al., 2018; Mattern et al., 2017; Tandeo et al., 2020), the forgetting factor and the uncertainty variances have been chosen (after many short tests not included in this description) in order to provide effective assimilation increments without introducing degradation on biogeochemical variables.”

Line 479: Would it help to show the results for Lorenz96 immediately after the description of its experiment design? You can have section 4 focused on the Lorenz results and section 5 on the BGC model results.

This issue was solved by splitting the manuscript in two parts.

Reply on Reviewer comment RC2

This work proposes an ensemble filter (GHOSH) that conducts sampling in such a way that the resulting sample statistics can match the moments of a target distribution up to a specified order. The moment-matching trick is based on Hermite polynomial approximations to the underlying functions, whereas the target distribution is set to be Gaussian. The derived filter is tested in two examples, both indicating that the GHOSH outperforms an existing filter (SEIK) for the experiments conducted in the current work.

The manuscript is clearly written and reasonably organized in general. Below is a list of minor-to-moderate issues spotted in the current manuscript.

We really thank the Reviewer for the positive general comment and for the suggestions provided below. Hereafter, our point-by-point responses to the Reviewer's comments are provided in bold green with line numbers referring to the updated version of the manuscript without track change.

Firstly, it is worth noticing that, considering the suggestions of the other two Reviewers and having received a positive answer by the Topic Editor on this, we prepared a revised version of the manuscript divided in two parts.

Spotted issues

1. Page 1 –

^ Line 2: Consider replacing “one of” by “among” or something similar, since “algorithms” is the subject.

We propose to modify the sentence as follows (Part 1, L. 2):

“ensemble algorithms are among the most successful data assimilation approaches”.

^ Line 20: What does “a higher order of convergence” mean here?

2. Page 2 –

^ Line 42: “Montecarlo” → “Monte Carlo”.

We thank the Reviewer for spotting these typos. The manuscript was corrected changing “Montecarlo” into “Monte Carlo” (Part 1, L. 28), while “order of convergence” is no longer used in the revised abstract.

^ Line 51 – 52: Rephrase the sentence “the second order approximation is more effective the closer the ensemble members are to each other, thus, the larger the ensemble spread the worse will be the approximation error in the mean computation.”

According to this comment and to comments from Reviewer #1 and #3, this sentence was reformulated as follows:

“Most of the models used in geoscience applications are based on systems of differential equations that cannot be represented by a second-order polynomial and in all of these cases the second-order sampling methods provides a non-exact estimation of the forecast mean, which is affected by an error strictly related to the error made in approximating the model by a second-order polynomial. Furthermore, the second-order approximation is more effective the closer the ensemble members are to each other (i.e., small uncertainty), thus the higher the uncertainty the worse will be the approximation error in the mean computation. Since the state estimation is often affected by a relatively high uncertainty in geosciences data assimilation applications, this approximation error may be not negligible.”

3. Line 58 – 59, Page 3: The “ $2r + 1$ ensemble members” requirement does not appear exact. For the unscented transform, one can use either principal component analysis (PCA) or truncated singular value decomposition (TSVD) to reduce the number of ensemble members, see

<https://doi.org/10.1175/2008JAS2681.1>

<https://doi.org/10.1016/j.physd.2008.12.003>

It may be worth discussing the similarities and differences between the ideas used to control the number of ensemble members in the aforementioned works and the current manuscript.

In the mentioned works, the authors reduce the number of ensemble members, which is equal to $2r+1$, by reducing the subspace dimension r (by a PCA or a TSVD). We thank the Reviewer for suggesting the 2 relevant citations that were added to the introduction (Part 1, L. 49).

4. Line 109 – 117, Page 5: The discussion on the extension of GHOSH to more generic distributions makes sense. A missing part, however, is that the authors did not explain why they confine themselves to Gaussian distribution in the current work, and what could be the challenges for the generalization of GHOSH to more generic distributions

We thank the Reviewer for this suggestion. We added the following sentence in the Discussion:

“The statistical moments μ in equation (1) are key hyper-parameters that describe the uncertainty pdf moments for orders higher than 2. In our experiments we used the moments of a Gaussian distribution, but it is worth noting that the GHOSH algorithm does not enforce this choice and other pdfs could be used. However, Kalman filter’s analysis equations somehow prescribe a Gaussian approximation, thus the GHOSH filter keeps a link to Gaussianity, even if the GHOSH sampling does not. Interestingly, other filters, like Hodyss (2011), try to overcome this limitation and could represent good candidates to study the effects of a non-Gaussian GHOSH sampling.”

5. Eq. 38, Page 13: The notation w.r.t the Q_i component is somewhat confusing. I guess it should be $(Q_p)^{-1}$, but it looks like Q_p^{-1} .

Equation 38 was improved by adding parentheses as suggested.

6. In the experiments w.r.t the Lorenz96 model, localization does seem used. What is the reason behind this setting?

Localization is a strategy to reduce the degrees of freedom in order to successfully apply an ensemble filter when the number of ensemble members is much smaller than the dimension of the state vector. We did not use localization in the Lorenz96 model, since the number of variables was already comparable with the tested ensemble size (at least in the case of ensemble sizes 31 and 63, while the smallest ensemble size settings were specifically aimed to study the behaviour of the filters in undersampling conditions). We added the following sentence to the Filter setup section (Part 1, L. 400???):

“Localization has not been applied, since the number of variables N is relatively small and comparable with the ensemble size.”

7. Line 434 – 435, Page 19: Why “it implies that the PCA measures the Pearson correlation”?

The statement was misleading and is no more present in the updated manuscript.

8. Line 490, Page 21: If I’ve understood correctly, the “best” label corresponds to the configuration that leads to the best DA performance. If so, then in Figure 4 one should use one block to represent it, and I don’t see the point to use a single row for the representation.

Often, in realistic applications, the number of ensemble members and the observation frequency are not flexible parameters (the former is usually limited by the computational resources and the latter by data availability). On the other hand, the forgetting factor can be tuned to improve assimilation results. Thus, we believe that it could be informative for some readers to see the comparison between the two filters tuned with their best forgetting factor in a range of ensemble sizes and observation frequencies. Considering also a comment from Reviewer 3, we changed former L. 490 to better clarify that “best” means “best RMSE” (Part 1, L. 426):

“The first line in each colour-map, labelled “best”, represents the best result, in terms of lowest RMSE, obtained among the set of tested forgetting factors.”

9. Line 661, Page 33: “an higher” → “a higher”.

Thank you for spotting the typo, that was corrected.

Reply on Reviewer comment RC3

The manuscript introduces a higher-order ensemble sampling scheme and a corresponding ensemble Kalman filter analysis step. The sampling scheme utilizes ensemble weights to represent higher-order moments of the ensemble distribution without necessarily increasing the ensemble size. The analysis step is motivated by the error-subspace Kalman filter SEIK and utilizes the ensemble weights. Additional sampling steps are introduced during the forecast phase and after the analysis step in order to maintain the higher-order sampling to the prescribed order. After the introduction of the sampling and assimilation algorithm, the method is first assessed in comparison to the SEIK filter in twin experiments using the chaotic Lorenz-96 toy model, which is frequently used to assess data assimilation methods. Subsequently, the new method is used in a realistic model application in which satellite chlorophyll observations are assimilated into the OGSTM-BFM biogeochemical model. The manuscript concludes that the method can result in better state estimates than the SEIK filter for the same ensemble size in case of the Lorenz-96 model. Further it is concluded that using higher-order sampling is feasible in a realistic application and can result in better state estimates compared to using only second-order sampling in the biogeochemistry data assimilation application.

Before I come to scientific assessment, I like to point out that this manuscript does not seem to fit well into the scope of GMD. Presented is the development of a sampling and data assimilation method. This is neither a new model development, description, assessment or development of new assessment methods, which are the scope of GMD. The data assimilation is implemented and used with a toy model and a realistic ocean-biogeochemical model, but the focus is on the algorithm and no particular model-related developments are described. To this end, I recommend to transfer the manuscript to a journal with a better fitting scope. Within the journals of the EGU, this is Nonlinear Processes in Geophysics (NPG).

From the scientific viewpoint, the proposed algorithm is new and the experiments show that it has the potential to improve the skill of ensemble data assimilation by accounting for higher order moments in the ensemble distribution and hence improving the mean and covariance estimates used in the analysis step of the ensemble filter. However, the manuscript needs to be improved at different places. The description of the algorithm has to be refined and it has to be better explained. Further the numerical experiments have to be organized in a way that the benefit of the new algorithm becomes clear in typical settings. The application to the 3D ocean-biogeochemical model seems to overload the manuscript while it does not even compare the results to a standard algorithm. Overall, this requires a major revision of the manuscript.

We really thank the Reviewer for the very accurate reading of our manuscript and for having tested the algorithm. The comments and suggestions provided by the Reviewer have been very helpful to revise the manuscript contents, and we think that the manuscript quality has significantly increased after the completion of the revision. Hereafter, our point-by-point responses to the Reviewer's comments are provided in bold green with line numbers referring to the updated version of the manuscript without track change.

Firstly, we would like to clarify the main objective of the work, since it appears this was not clear enough. We are introducing a new method (GHOSH filter) that is focused on the use of a higher order of approximation to evaluate the ensemble mean in data assimilation applications. We are not evaluating standard deviation, kurtosis and higher statistical moments of the ensemble in the present work. On the other hand, it is true that the equality of the statistical moments up to order h implies an approximation of the mean up to order h , as demonstrated in Appendix A.

Concerning the option to move the manuscript in another journal, we chosen to submit the manuscript in GMD since “papers focussing on data assimilation are welcome” as “Development and technical papers” (https://www.geoscientific-model-development.net/about/manuscript_types.html#item2, last visit 27 March 2024). Indeed, we firstly submitted the manuscript as a “Development and technical paper” but the Topic Editor has then adjusted the manuscript type to “Model description paper”, but did not raise concerns about the relevance of the manuscript with respect to the GMD scope. More recently, when discussing the option to split the manuscript in two parts (see a later comment response), the Topic Editor confirmed that the work falls into the scope of GMD. Given the above reasons, being the review process in progress and having already taken quite a long time, we would prefer to stay in GMD.

Major comments:

On the derivation and mathematical treatment: At least for GMD, which is not a mathematical journal, the manuscript explains far too less of the algorithmic aspects. At several instances the authors just state 'can be proven', but omit attempts to explain why some equation holds. For me this is clearly insufficient and it is likely that readers cannot follow the descriptions. I do not recommend to include proofs (given that GMD is not a mathematical journal), but the authors should sufficiently explain why some relationship or equation holds. This does not need to be long, but will help the readers tremendously, while currently the authors force the readers to 'just believe' what they have written. (This comment will likely also hold should the manuscript be transferred for NPG, since also this is not a strictly mathematical journal).

Thank you for this comment that can help to make the manuscript more accessible to GMD readers. Some of the “can be proven” statements are explicitly addressed in the following point-by-point responses but other instances were addressed and modified in the new version of the manuscript.

Experiments with the Lorenz-96 model: The first set of numerical experiments is conducted using the Lorenz-96 model. The experiments show that the newly proposed GHOSH filter can yield smaller RMSEs by up to 66%. Given that this error reduction of the GHOSH filter relative to the SEIK filter is very large (I don't think that I have seen such magnitude even in comparisons of fully nonlinear particle filters with EnKFs), I recommend to be particularly careful in ensuring that the experiments are representative and comparable to previous

studies. In particular, one has to ensure that this advantage of the GHOSH filter is not an artifact of the particular model and data assimilation configuration. In this respect I see several weaknesses as the model configuration and implementation are rather untypical, the initialization of the ensemble is different from other studies even for the SEIK filters, and the experiments are far too short. I will comment in more detail on this further below.

We addressed all the suggestions on the Lorenz96 model experiments in the point-by-point responses, however we would like to thank the Reviewer, since the manuscript results have certainly improved by enlarging the set of toy-model numerical experiments.

Experiments with the realistic OGSTM-BFM model: In these experiments, in which real satellite chlorophyll data are assimilated, only different configurations of the GHOSH filter are assessed. Thus, a comparison to an EnKF like SEIK is missing. The result of the experiments is that using order 5 yields comparable RMSD for chlorophyll and phosphate but lower RMSD for nitrate, compared to using a lower order. The magnitude of this effect depends on the ensemble inflation, and for too strong inflation also the GHOSH filter leads to a deterioration of nitrate compared to the control run. To this end, the higher-order sampling can yield a better result in the multivariate update for a particular choice of the forgetting factor. Unfortunately, the manuscript does not provide insights into why this does happen.

Further, maps of the chlorophyll and phosphate as well as Hovmoeller plots of chlorophyll and phosphate are discussed. However, these discussions do not provide any further insight into the GHOSH filter, but merely show that this filter can be successfully applied to this model. These results might be interesting for researchers interested in chlorophyll assimilation, but for a methodological study on this new filter method, the experiments provide very little insight. Only the first part including Table 2 seems to be relevant, but here a comparison to the SEIK filter is missing.

Given the complexity of the model and the expert knowledge a reader needs to appreciate the discussion on the effects on chlorophyll, nitrogen and phosphate, I recommend to remove this experiment from the manuscript. In order to assess multivariate assimilation (which cannot be done with the Lorenz-96 model) I rather recommend to utilize some other idealized multivariate model. Perhaps, the 2-scale Lorenz-2005 model could be a possible choice. For the methodological study it would be very important to actually assess the reasons for why multivariate updates are better (if they are). Just showing that they are better in one particular model configuration without understanding the reasons has little value because one cannot generalize the result. Below I will not further comment on details in the sections on the OGSTM-BFM model.

Also Reviewer #1 commented about the option to remove the realistic 3D application, moreover we received a positive answer by the Topic Editor about a request we sent in the past weeks on the possibility to split the manuscript in two parts. Thus, we revised the manuscript considering the twin experiment and realistic application separately.

Detailed comments:

Title: It is not clear why the authors give their algorithm a version number, and this is very untypical. While I know that GMD requires version numbers when a certain model version is presented, this manuscript is not about a particular implementation of a model but about a numerical algorithm. This issue might again point to the fact that the manuscript is not well suited for GMD as I commented on before. Apart from this, I recommend to add to the title the word 'ensemble' to clarify that an ensemble filter is introduced.

As briefly discussed above, the Topic Editor adjusted the manuscript type to “Model description paper”, and on this occasion the Topic Editor explicitly asked for a versioning number. Indeed, the versioning number request seemed a sound request, since the GHOSH algorithm could be modified and improved in the future. Concerning the title, we agree with the Reviewer on the need to add the word “ensemble” in the title, and we propose the following novel title:

"GHOSH v1.0.0: a novel Gauss-Hermite High-Order Sampling Hybrid ensemble filter for computationally efficient data assimilation in geosciences"

Abstract: The abstract includes unnecessary details. E.g. information how many experiments have been conducted with the Lorenz-96 model ('two thousands', line 9) and which particular parameters have been varied (line 12) are not suitable for the abstract. The abstract is also irritating in that first 'GHOSH' is mentioned as a sampling method while in line 14 it is then named a 'filter'. Perhaps, one can better formulate this as introducing a filter and the corresponding sampling method.

The abstract (Part 1) was improved according to the Reviewer's suggestions. In particular, the word “thousands” and details on the varied parameters were removed while at L. 7-8, “GHOSH sampling method” was changed to “GHOSH filter's sampling method”.

The statement 'use of a higher order of convergence substantially improves the performance of the assimilation with respect to nitrate' (lines 20/21) is not valid in this form. Valid is that the higher-order filter improved the nitrate for the particular experiment, but this cannot be generalized in any way. The final sentence of the abstract on the computation time is superficial. The situation that most time is spent in computing model forecasts does not need to hold in general. Instead, the execution time has to be compared for the actual filter analysis including the resampling steps of the GHOSH filter. Please revise the abstract accordingly.

According to the choice to split the manuscript in two parts, the abstracts have been deeply revised and comments on the computational time have been changed in both the abstracts:

(Part 1 Abstract) “In the directions where the uncertainty is larger, the GHOSH filter’s sampling method achieves a higher order of approximation than in other ensemble based filters, without increasing the asymptotic computational complexity that is comparable to the one of second-order deterministic filters.”

(Part 2 Abstract) “Results show that: i) the GHOSH implementation in a realistic three dimensional application is just as computationally feasible as other ensemble assimilation filters, since the integration of a realistic model is by far more computationally expensive than the assimilation scheme; ii) the GHOSH assimilation algorithm improved the agreement between forecasts and observations without producing unrealistic effects on the non-assimilated variables.”

Introduction section:

- Unfortunately, this section contains various small issues and it is difficult to comment on all of them. I focus here on the most relevant:

The section includes invalid referencing. E.g. the list of references in lines 28-29 appears to be an arbitrary selection of studies that applied DA. Since there are many publications about applications of data assimilation it can be a better choice to cite review papers that summarize the state of research instead of picking 'some' article for which it is then unclear why the authors consider these as particularly relevant.

We thank the Reviewer for the suggestion. We revised the citations at the Introduction beginning to provide a wider view of DA methods and applications (Part 1, L. 18):

“(among the others, see the methods and applications reviewed in Carrassi et al., 2018; Houtekamer and Zhang, 2016; Lahoz and Schneider, 2014; van Leeuwen et al., 2019; Martin et al., 2015; Roth et al., 2017; Vetra-Carvalho et al., 2018)”

- With regard to citing review papers, I strongly recommend to cite those explicitly as reviews, e.g. using 'see, e.g.'. For example, Carrassi et al. (2018) is cited for 'modified implementations have been proposed for application to non-linear ones' (line 35). This review paper did however not introduce the 'modified implementations' but reviews what has been published in other original studies. This likewise holds for the citation of Bannister (2017b) which is cited for 'EnVar' (line 38), but is a review of variational methods. (BTW: The second paragraph of the introduction does not seem to be relevant for this study on an ensemble filter method and it could be omitted without loss of relevant information). Further, for books it is common practice to cite them providing a chapter where the particular information can be found (Readers should not be required to read a whole book). Another case of invalid reference is e.g. citing Bocquet et al. (2010) for 'particle filter methods'. The paper discusses methods for non-Gaussian data assimilation, but it is not an original paper for particle filters. In any case, there are more recent review papers, like the cited review by van Leeuwen et al. (2019) which better cover the state of research.

The correct citation of reviews was adopted in the revised version of the manuscript. We thank the Reviewer since this change helped us to improve the use of references.

Moreover, we removed the second paragraph of the introduction according to the Reviewer suggestion; and we adopted van Leeuwen et al. (2019) instead of Bocquet et al. (2010).

- It is also important that the authors provide references for claims they include in the text. E.g. in lines 30/31 it is written 'the use of DA has been steadily increasing in the last decade and it is now ubiquitous in earth sciences, also thanks to the unprecedented -and always increasing- availability of both data and computational capability.' I'm not sure from where the authors got the insight that the increasing computational capability and data availability lead to an increased use of DA. However, I would suppose that there should be some review paper examining this. Actually, most data types that are currently assimilated in the ocean and biogeochemistry (e.g. sea surface height, sea surface temperature, chlorophyll) have been available for more than a two decades and have already been assimilated more than 10 years ago. Also the computing power was sufficient more than 10 years ago even to perform ensemble data assimilation as is evident from data assimilation studies that were published from 15-20 years ago. The major changes appear that the model complexity and resolution was steadily increased. These facts obviously contradict the authors statement.

- Related to this, the authors state that recently the use of ensemble algorithms has been proposed 'thanks to the scalability of parallel implementations' (line 37). Then, Evensen (1994) is cited for the EnKF. I cannot really see that a paper that was published 29 years ago is 'recent'. It is further not evident that these methods were proposed because of the scalability of parallel implementation since Evensen (1994) proposed the EnKF as a method to use a sampled covariance matrix as a dynamic estimate of the uncertainty in order to advance the extended Kalman filter that was used by Evensen before (see Evensen, 1992, 1993), but not as a method solving a scalability issue. Please revise the text, to avoid such potential misleading statements.

Thanks to the Reviewer suggestions the first part of the introduction was revised to be more consistent with literature. In particular, we introduced references to the review by Vetra-Carvalho et al. (2018), where the computational costs and the increase of observations are proposed as two of the factors that motivated recent EnKF developments toward more complex and non-linear applications. The reference to Evensen (1994) was removed since no more necessary (Part 1, L. 21):

“Thanks to the scalability of parallel implementations, the use of ensemble algorithms (see e.g., Vetra-Carvalho et al., 2018; Houtekamer and Zhang, 2016; Bannister, 2017) have been proposed to estimate uncertainty and improve assimilation skills in Kalman filters and variational methodologies. On the other hand, some of the strong points of the ensemble and variational have been merged in hybrid filters (e.g., Hamill and Snyder, 2000). Moreover, recent developments of EnKF have been conceived to face increasing 25 non-linearities and model complexity that are also related to the expanding availability of observations and computational resources (see e.g., Vetra-Carvalho et al., 2018).”

- There are also further questionable statements.

-- E.g. in lines 49/50: 'second order sampling methods provide only a second order approximation of the model'. Actually, the ensemble in ensemble KFs is not meant to approximate the model, but it is used to represent model uncertainty, i.e. the probability distribution. One should be very clear about this.

We thank the Reviewer for raising this point about the aim of the ensemble KFs. Indeed, we meant to highlight that, when second-order exact EnKFs are applied with models that cannot be represented by a second order polynomial, the mean estimation provided by the ensemble is affected by an approximation error. The error in the mean is strictly related to the error made in approximating the model by a second-order polynomial. For sake of clarity, we recall that second-order sampling methods are based on the fact that, assuming that the model m is a second order polynomial, the expected value of the two random variables X and Y is the same, i.e. $E[m(X)] = E[m(Y)]$, where X represents the system uncertainty and Y its ensemble approximation produced by the sampling method (i.e., the ensemble describes the

discrete probability mass function of Y). When m is not a second order polynomial, the error on the expected value $||E[m(X)] - E[m(Y)]||$ can be estimated by

$$||E[m(X)] - E[m(Y)]|| < ||E[m(X)] - E[p(X)]|| + ||E[p(X)] - E[p(Y)]|| + ||E[p(Y)] - E[m(Y)]||,$$

where p is the best second order polynomial approximating m . The second term in the r.h.s. vanishes because the second order sampling is exact on second order polynomials. What remains are the errors of the second order approximation of the model m with the polynomial p .

We modified the sentence as follows (Part 1, L. 37):

“Most of the models used in geoscience applications are based on systems of differential equations that cannot be represented by a second-order polynomial and in all of these cases the second-order sampling methods provides a non-exact estimation of the forecast mean, which is affected by an error strictly related to the error made in approximating the model by a second-order polynomial.”

-- Then in lines 50/51: 'the second order approximation is more effective the closer the ensemble members are to each other'. Here, I suppose the authors imply that the nonlinear effect are smaller if the perturbation of the ensemble members is small and hence the second-order approximation should hold better. While this seems to be natural, Rainwater and Hunt (2013) found that a smaller ensemble spread did not improve the results of an EnKF. Thus, the statement does not seem to be valid in this form.

We fully agree with the Reviewer that the nonlinear effects are smaller when the perturbation of the ensemble members is small, and hence the second-order approximation holds better. In other words, if the true uncertainty is small, then a second-order approximation is affected by a smaller error. The statement is not meant to contradict Rainwater and Hunt (2013), who propose an empirical technique to reduce the analysis error by inflating the forecast spread and deflating the analysis spread with the aim to improve EnKF results under certain circumstances. We modified the sentence focusing on uncertainty rather than the ensemble spread (also in line with comments from Reviewers 1 and 2) (Part 1, L. 40):

“Furthermore, the second-order approximation is more effective the closer the ensemble members are to each other (i.e., small uncertainty), thus the higher the uncertainty the worse will be the approximation error in the mean computation. Since the state estimation is often affected by a relatively high uncertainty in geosciences data assimilation applications, this approximation error may be not negligible.”

-- Related to this is the statement "a relatively large ensemble spread is often required in data assimilation applications," (line 52). I'm not aware of that this is true and the authors would need to provide a reference for this statement.

The sentence was modified according to the previous comment.

-- Also the statement "second order methods use an ensemble of $r + 1$ members to span an r -dimensional error subspace" (line 57) is not supported by the literature. Actually the perturbed-observations EnKF by Evensen (1994) is second-order (because it uses the Kalman filter analysis step), but it is not aimed at spanning an r -dimensional subspace. Using a particular sampling with $r+1$ members was introduced by Pham with the SEEK and SEIK filters, but not widely adopted. Also the ensemble initialization used with the (L)ETKF scheme is not explicitly 2nd order. This also holds for the notion of the 'error subspace', which was further discussed by Nerger et al. (2005) while other filter methods, like the ETKF or newer variants of the EnKF (Evensen 2003/2004), were developed without this notion.

We thank the Reviewer for this comment that highlights that the definition of KF method order was unclear. We made the definition of "order" more explicit, in particular we clarified that we are referring to second-order exact deterministic methods (like ETKF or SEIK while Evensen's EnKF is a stochastic method which is not included in the reported definition of second-order methods) (Part 1, L. 29):

"The number of ensemble members can be reduced by adopting deterministic sampling methods (as opposed to stochastic EnKF methods, see e.g. Carrassi et al., 2018). Examples of deterministic EnKF using second-order sampling methods are SEIK (Pham, 2001) and ETKF (Bishop et al., 2001). Extending the definition of second-order exact methods in Pham (2001), with the term "order" we refer to the polynomial order of approximation of the filter or of its sampling method. Namely, in the case of a h th-order approximation, the ensemble has the property of providing a forecast mean with no error as long as the evolution function used for forecasting (i.e., the model) is a polynomial of order h . As proven in Appendix A, this is achieved by sampling an ensemble that matches the first h statistical moments of the uncertainty probability density function (pdf) before evolution (also called prior in Bayesian statistics)."

Moreover, the correct reference to Nerger (2005) was added near the first occurrence of "subspace" (Part 1, footnote 1):

"In the space of the possible state vectors, the uncertainty subspace is the subspace generated by the ensemble members (i.e., the smallest subspace that contains all the ensemble members). The uncertainty subspace was originally called error subspace (see Nerger et al., 2005, 2012, for an introduction to the error subspace concept), but in the context of this work it is more natural to interpret the ensemble as a proxy of the uncertainty, avoiding unnecessary confusion with, for instance, the approximation error."

- Please note that the reference Pham (1996) is incomplete, but it is likely also not a valid reference according to citation standards. It is a technical report and was never officially published in a peer reviewed journal. While I'm aware of this report, I also could not find it any more on the internet. To this end, I recommend to replace it by a valid reference to a

peer-reviewer publication. The first real publication of the SEIK filter seems to be Pham et al. (1998b), different from what is cited in the manuscript. Since this article is mainly in French and only has a shorted text in English, a better reference would be Pham (2001).

Thanks, the citation was corrected.

- The introduction misses to explain what the authors mean by 'high order' by 'order' at all. Given that GMD is not a mathematical journal this should be explained.

We thank the Reviewer for underling points that could be improved in order to help the reader to better understand the presented work. The definition of order in the Introduction was improved.

- Please be careful when discussing a sampling in contrast to discussing filtering. E.g. lines 65/66 state "we propose a novel weighted ensemble method based on a new high-order sampling, that provides ensemble mean estimates of order higher than 2". Actually, while the existing ensemble Kalman filter method use the second-order analysis scheme (treating only mean and covariances), they are not required that the ensemble was sampled with second order (in fact unless one uses second-order exact sampling, the ensemble sampling does not consider the sampling order.)

- Also related to lines 65/66, I a bit irritating about the implied meaning of 'order'. I understand it as accuracy in sampling the PDF. However, then the mean is the first order moment of the PDF. How can this moment be accurate by 'order higher than 2'. I'm likely missing something here, but I think that it is likely that many readers will also wonder. To this end, I see the clear need to explain the terms 'high order' and 'higher order' so that the readers knwo how to understand the discussions throughout the manuscript.

- In this respect, please also check whether the distinction in 'high order' and 'higher order' is required and if so, please explain the difference.

Based on the comment Reviewer on L 57, we clarified the definition of “order” in the Introduction. This improvement clarifies the last three points raised by the Reviewer. We carefully verified to avoid the misleading wording of “order” in the Introduction and through the whole manuscript.

- I will stop with details on the Introduction at this point. Overall, the authors should carefully revise the introduction. In this they should ensure that the statements are true and that they are sufficiently supported by the literature. This implies to provide references to claims as described before. It further implies to be careful when citing review papers so that it clear that no original works are cited (This also holds in other sections, e.g. the reference to

Carrassi et al. (2018) in line 90 and line 264 are also not valid - Carrassi is neither an original reference for deterministic square-root sampling methods nor for the forgetting factor).

Thank you for your careful reading of the Introduction. In addition to the changes listed above, we corrected other similar occurrences in the manuscript.

Figure 1 and lines 93-104: The figure is difficult to understand and I'm not sure if it is really sufficient to visualize the 3rd order sampling in 2 dimensions or 5th order in 1 dimension.

Considering this and the following Reviewer's comments and comments from Reviewer #1, we described Fig. 1 in more detail in a novel appendix (Appendix C) adding the description (including equations and calculations) of the concepts summarised in Fig. 1 panels (e.g., computation of the moments).

- Please include a clearer explanation for the shaded circular regions (in the top panel) in the caption (the caption states 'isosurfaces', but of what?).

- In the 2D and 1D views I cannot see that a 3rd or 5th order approximation is shown. It is also not obvious that the 2D view is a projection of the tetrahedron. In the way this is plotted, the lower triangle looks like it's parallel to the y-x plane so that the projection would be a triangle with one sample in its center.

- With regard to lines 102/103, I cannot follow why 3 weighted members should be able to represent an approximation of order 5. With weights one can obviously shift the mean, variance and skewness. But it is unclear how kurtosis could be expressed with 3 members (perhaps even with 4 members). Perhaps this irritating is due to the fact that it is not clear how the moments would be computed in this case.

- It would be useful for the readers if the figure is replotted with a better perspective. Likewise the text should better explain how the orders are expressed in the sampling. Instead of stating 'it can be proven' in line 98 it would be useful to explain why this is the case and perhaps explain how a skewed (e.g. non-Gaussian) distribution would be represented (I don't think that this is a projection of the tetrahedron representing the distribution, but a sampling of the projection of the actual distribution). Since the example uses a Gaussian distribution whose higher moments are generally zero (for odd orders) or have a constant value (for even), it seems to be difficult to actually sketch the representation of higher moments. Showing how higher moments could be represented for a non-Gaussian case could be useful. It might also be useful to show the example on a Gaussian case that has different variances per direction.

- Linked to this is the sentence in lines 103/104 stating 'what we have obtained is a 3-dimensional second order weighted ensemble which is a third-order approximation in the xy plane and a fifth order approximation along the x axis.' For me, this statement is not evident from the figure. It might be that the sketched samplings in the xy-plane and along the x-axis are of this higher order - perhaps because they represent zero higher orders, but this is not visible from the dots that are shown in the figure.

All the above comments are addressed in the novel Appendix. For instance, the Appendix C includes a practical example of a 5th order weighted ensemble in 1D composed by 3 points P1, P2 and P3, i.e.:

$$P1: x1 = \sqrt{3}; w1 = 1/6,$$

$$P2: x2 = -\sqrt{3}; w2 = 1/6,$$

$$P3: x3 = 0; w3 = 2/3.$$

The first moment (mean) is:

$$w1 \cdot x1 + w2 \cdot x2 + w3 \cdot x3 = 1/6 \cdot \sqrt{3} + 1/6 \cdot (-\sqrt{3}) + 2/3 \cdot 0 = 0,$$

the second moment (variance) is:

$$w1 \cdot x1^2 + w2 \cdot x2^2 + w3 \cdot x3^2 = 1/6 \cdot 3 + 1/6 \cdot 3 + 2/3 \cdot 0 = 1,$$

the third moment (skewness) is:

$$w1 \cdot x1^3 + w2 \cdot x2^3 + w3 \cdot x3^3 = 1/6 \cdot \sqrt{3}^3 + 1/6 \cdot (-\sqrt{3})^3 + 2/3 \cdot 0 = 0,$$

the fourth moment (kurtosis) is:

$$w1 \cdot x1^4 + w2 \cdot x2^4 + w3 \cdot x3^4 = 1/6 \cdot 9 + 1/6 \cdot 9 + 2/3 \cdot 0 = 3,$$

the fifth moment is:

$$w1 \cdot x1^5 + w2 \cdot x2^5 + w3 \cdot x3^5 = 1/6 \cdot \sqrt{3}^5 + 1/6 \cdot (-\sqrt{3})^5 + 2/3 \cdot 0 = 0.$$

Thus, the ensemble P1, P2, P3 is compliant with the statistical moments of a standard normal distribution up to order 5.

The property of this 1-dimensional 5th-order sampling, can be extended to a 2-dimensional ensemble, aimed to keep order 5 along the x-axis while achieving order 3 in the xy plane. Such 4-members ensemble is:

$$P1: (x1, y1) = (\sqrt{3}, 0); w1 = 1/6,$$

$$P2: (x2, y2) = (-\sqrt{3}, 0); w2 = 1/6,$$

$$P3: (x3, y3) = (0, \sqrt{3/2}); w3 = 1/3,$$

$$P4: (x4, y4) = (0, -\sqrt{3/2}); w4 = 1/3.$$

Note that all the moments along the x-axis are already computed above up to order 5, since this ensemble and the previous one are indistinguishable looking only at the x-coordinate. In fact, considering the projection on the x-axis, P3 and P4 behave as a unique member with weight w_3+w_4 because they have the same x-coordinate $x_3=x_4$. It remains to be checked if the moments involving y match the moments of standard normal distribution up to order 3, i.e.:

the 1st moment associated to y (i.e., the mean along y) is:

$$\begin{aligned} w_1 \cdot y_1 + w_2 \cdot y_2 + w_3 \cdot y_3 + w_4 \cdot y_4 &= \\ = 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot \sqrt{3/2} + 1/3 \cdot (-\sqrt{3/2}) &= 0, \end{aligned}$$

the 2nd moment associated to y^2 (i.e., the variance of y) is:

$$w_1 \cdot y_1^2 + \dots + w_4 \cdot y_4^2 = 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot 3/2 + 1/3 \cdot 3/2 = 1,$$

the 2nd moment associated to xy (i.e., the covariance between x and y) is:

$$\begin{aligned} w_1 \cdot x_1 \cdot y_1 + \dots + w_4 \cdot x_4 \cdot y_4 &= \\ = 1/6 \cdot \sqrt{3} \cdot 0 + 1/6 \cdot (-\sqrt{3}) \cdot 0 + 1/3 \cdot 0 \cdot \sqrt{3/2} + 1/3 \cdot 0 \cdot (-\sqrt{3/2}) &= 0, \end{aligned}$$

the 3rd moment associated to $x^2 y$ is:

$$\begin{aligned} w_1 \cdot x_1^2 \cdot y_1 + \dots + w_4 \cdot x_4^2 \cdot y_4 &= \\ = 1/6 \cdot 3 \cdot 0 + 1/6 \cdot 3 \cdot 0 + 1/3 \cdot 0 \cdot \sqrt{3/2} + 1/3 \cdot 0 \cdot (-\sqrt{3/2}) &= 0, \end{aligned}$$

the 3rd moment associated to $x y^2$ is:

$$\begin{aligned} w_1 \cdot x_1 \cdot y_1^2 + \dots + w_4 \cdot x_4 \cdot y_4^2 &= \\ = 1/6 \cdot \sqrt{3} \cdot 0 + 1/6 \cdot (-\sqrt{3}) \cdot 0 + 1/3 \cdot 0 \cdot 3/2 + 1/3 \cdot 0 \cdot 3/2 &= 0, \end{aligned}$$

the 3rd moment associated to y^3 is:

$$w_1 \cdot y_1^3 + \dots + w_4 \cdot y_4^3 = 1/6 \cdot 0 + 1/6 \cdot 0 + 1/3 \cdot \sqrt{3/2}^3 + 1/3 \cdot (-\sqrt{3/2})^3 = 0.$$

The dimensions can be increased once more by building in a similar way a 3-dimensional ensemble, as shown in Fig. 1. This ensemble is indistinguishable from the previous one in the xy plane projection and capable of achieving 2nd order on the xyz space. Such ensemble is:

$$P1: (x_1, y_1, z_1) = (\sqrt{3}, 0, -\sqrt{2}); w_1 = 1/6,$$

$$P2: (x_2, y_2, z_2) = (-\sqrt{3}, 0, -\sqrt{2}); w_2 = 1/6,$$

$$P3: (x_3, y_3, z_3) = (0, \sqrt{3/2}, \sqrt{2}/2); w_3 = 1/3,$$

$$P4: (x_4, y_4, z_4) = (0, -\sqrt{3/2}, \sqrt{2}/2); w_4 = 1/3.$$

All the moments involving x and y are already checked in the previous calculation up to order 5 along x and up to order 3 in the xy plane. It remains to be checked if the moments involving z match the moments of standard normal distribution up to order 2, i.e.:

the 1st moment associated to z (i.e., the mean along z) is:

$$w1 \cdot z1 + w2 \cdot z2 + w3 \cdot z3 + w4 \cdot z4 =$$

$$= 1/6 \cdot (-\sqrt{2}) + 1/6 \cdot (-\sqrt{2}) + 1/3 \cdot \sqrt{2}/2 + 1/3 \cdot \sqrt{2}/2 = 0,$$

the 2nd moment associated to z^2 (i.e., the variance of z) is:

$$w1 \cdot z1^2 + \dots + w4 \cdot z4^2 = 1/6 \cdot 2 + 1/6 \cdot 2 + 1/3 \cdot 1/2 + 1/3 \cdot 1/2 = 1,$$

the 2nd moment associated to xz (i.e., the covariance between x and z) is:

$$w1 \cdot x1 \cdot z1 + \dots + w4 \cdot x4 \cdot z4 =$$

$$= 1/6 \cdot \sqrt{3} \cdot (-\sqrt{2}) + 1/6 \cdot (-\sqrt{3}) \cdot (-\sqrt{2}) + 1/3 \cdot 0 \cdot \sqrt{2}/2 + 1/3 \cdot 0 \cdot \sqrt{2}/2 = 0,$$

the 2nd moment associated to yz (i.e., the covariance between y and z) is:

$$w1 \cdot y1 \cdot z1 + \dots + w4 \cdot y4 \cdot z4 =$$

$$= 1/6 \cdot 0 \cdot (-\sqrt{2}) + 1/6 \cdot 0 \cdot (-\sqrt{2}) + 1/3 \cdot \sqrt{3}/2 \cdot \sqrt{2}/2 + 1/3 \cdot (-\sqrt{3}/2) \cdot \sqrt{2}/2 = 0.$$

All the above proves that the weighted ensemble P1, P2, P3, P4 is a second-order ensemble achieving order 3 in the subspace of the x and y directions and order 5 along the x direction.

We are confident that these step-by-step calculations will help the Reader in following the manuscript arguments.

- For the non-mathematical readers of GMD it might also be useful to avoid the term 'standard' for the normal distribution but instead mention the variance explicitly.

Thank you for spotting this issue. We now explicitly refer to the variance of the standard distribution. In particular modified (Part 1) L. 384:

“Observations are generated for each truth trajectory by extracting from the state vector the values of the even indexed variables (i.e., x2, x4, ..., x62) every Δt time units and adding to each observed variable a random number sampled from a standard normal distribution (i.e., with variance equal to 1). Tests have been made for $\Delta t \in \{0.1, 0.15, 0.2, 0.25, 0.3\}$.”

Lines 110/111 mention that a higher-order sampling of a Gaussian distribution can be obtained by the Gauss-Hermit quadrature rule. At this point reader might likely wonder why a higher-order sampling of a Gaussian is required, when one can fully represent it by its mean and covariance and the second-order exact sampling provides this sampling. On the other hand, it might be obvious that a higher-order sampling helps for non-Gaussian distributions, but in this case the Gauss-Hermite quadrature doesn't seem to hold. It would be good if this aspect, which closely linked to the motivation of the method, is better explained. Perhaps, it becomes clearer when the authors explain 'high/er order' as commented on before.

Given the improved definition of “order” that is now provided in the introduction, this part of the manuscript should be clearer. Indeed, the order of the sampling is not related to the Gaussianity of the pdf but to the order of approximation of the mean, which, as mentioned above, is closely related to how good the approximation of the nonlinearity of the model is. In addition, the novel Appendix C uses simple examples to guide the reader in understanding how moments of order higher than two can affect a sampling of a Gaussian pdf. In fact, even if its first two moments completely characterise the distribution, the sampling needs to also match the higher moments to produce an ensemble capable of achieving an order of approximation higher than two.

Lines 112-114 describe that the higher-order sampling can be performed for a limited number of directions, while for the other directions second-order sampling can be used. Here I'm wondering if this holds after the application of the model. If we have, e.g., a polynomial function that includes terms that 'mix' (e.g. by multiplication) the directions used for higher-order sampling with those with second-order sampling, one obtains mixing effects. With this the clear separation in higher-order and second-order directions should no longer hold. Does the re-sampling during the forecast corrects these effects? Please explain these effects in a clear way.

Since each re-sampling computes its own PCA, the re-sampling during forecast chooses new principal directions taking into account the model application and its mixing effects on the previous principal components. It is also worth mentioning that, as shown in Appendix A (eq. A2 and following), any multiplicative term is already taken into account exactly up to the considered polynomial order. In more detail, if there are mixings in the form of multiplications between two directions, this is a second order term and it is exactly taken into account by any second-order sampling. If the mixing includes multiplications between more than two directions (some from the second-order and some from the h th-order directions), then the polynomial function has a higher (then two) polynomial order, and an error appears, but, since the directions sampled with order 2 were the directions where the spread is smaller, the error is comparably small because it is attenuated by at least one low-spread factor. Instead, the higher order terms (up to order h) composed by high-spread factors only, which would represent a bigger source of error for a second-order ensemble, do not add any error to the mean computed from a GHOSH ensemble, since the large-spread directions are sampled with order h . In this sense, the GHOSH sampling can be seen as a method to remove the principal sources of error in the mean estimation.

The details provided in Appendix A together with the clearer definition of “order” provided in the manuscript, should help the reader to understand the mixing effects.

Equation (1): I'm somewhat lost here. There are many indices which influence each line of the equation system which I cannot really disentangle. I think it would be useful to show some more lines, maybe the first 3, in addition to the general one. Is the subscript ' $j_1, \dots, j_{\xi}$ ' just one number or a list of indices. How many equations are in the system? (This relates to line 137 where it is stated that a large value of h required a smaller s or larger r , which is qualitative but does not provide an indication of actual counts for $r/s/h$)

Based on this Reviewer comment and also on a comment from Reviewer #1, we modified the system equations and their description to improve readability (Part 1, L. 109-138)."

Line 133: Here the 'Gaussian case' is mentioned. Usually we consider a Gaussian to be a distribution that is fully described by the first two orders. Why should a higher-order sampling be relevant for this?

We think that the changes made in the Introduction about the meaning of “order” and the previous responses are useful to clarify why the “Gaussian case” is considered.

Line 130: It is stated: "Such probability distribution must be uncorrelated, normalized and have 0 mean". Please explain why this 'must be'.

The statement was expanded to give an insight of why those properties are required, as proposed in the reply to the previous Reviewer comment about equation (1) (see Part 1, L. 109-138).

Line 138: I cannot follow the explanation how Ω_h is defined. It is described as a "matrix with coordinates"? What does this mean. Is it possible to provide a proper definition, e.g. in form of an equation?

The definition of Ω_h was clarified as follows (Part 1, L. 136):

“The system solution is used to define Ω_h , as the $(r+1) \times s$ matrix with coordinates um,j (i.e., um,j is the entry of the matrix Ω_h at row m and column j)”

Equation (4): This equation is also difficult. It combines Ω_h with an orthogonal random matrix of size $(r+1) \times (s+1)$ and some other orthogonal $(r+1) \times (r-s)$ matrix. The construction is unclear to me. Since the second matrix has $r-s$ columns, does one need to ensure that it is orthogonal to the other matrix. Further, Pham (1996) does not seem to provide a scheme to generate such a matrix, but only a matrix of size $(r+1) \times (r)$. The readers are here left alone by speculating what the correct matrices might be and how they might be generated. In doubt it should be possible to provide a method to generate the matrices in the Appendix.

Thank you for the suggestion. We added an Appendix (Appendix B) to give details on the construction of the orthogonal matrices of Eq. (4).

Equation (6): This equation shows that the filter relies on the covariance information contained in the matrix L . This looks particular, since usually we consider higher order to reach beyond the covariance matrix, which is the second moment of the PDF. For me it is unclear how higher orders can be taken into account if here only the first two orders are used. Please provide an explanation for this particular formulation of the algorithm.

As observed by the Reviewer, the filter relies on the covariance information contained in the matrix L . On the other hand, moments of order higher than 2 are encoded in the matrix Ω and are not dynamically computed from the ensemble.

Line 178 and following: Line 185 states that the forecast resampling is performed at each time. It is unclear whether this sampling is only done once after a whole forecast period or whether it is done after each time step of the time stepping scheme (The definition of 'time t_i ' is not clear). Figure 2 might indicate that it is each time step, but it is not clear. Further, we know that resampling likely violates balance constraints that are contained in the model states of physical models. Does the resampling performed here also have such effects?

We thank the Reviewer for giving us the opportunity to clarify this aspect. The forecast resampling can be done at chosen pre-defined time steps that are not those of the time-stepping scheme. The relevant sentence was modified accordingly (Part 1, L. 164):

“Similar to other data assimilation ensemble schemes, the GHOSH filter provides an estimate of the state of a system at some pre-fixed times t_i in terms of the state vector and the covariance matrix representing the estimation uncertainty.”.

Concerning the balancing constraints, violations cannot be excluded at GHOSH forecast resampling times. As highlighted by the Reviewer, these violations occur generally in resamplings. However, Nerger et al. (2012) proposed an approach that aims at minimizing those violations, and it can be also applied in GHOSH by choosing an opportune T matrix (equation (18)). The answers to the next Reviewer comments further expand this subject.

Line 224: The statement "other T can be explored without affecting the algorithm (e.g., see Nerger et al. (2012))." is incorrect. Actually the cited study shows that the projections on the SEIK filter are inconsistent and that for consistency a different projection is required. This does obviously 'affect the algorithm'.

The wording of the sentence was misleading, we meant that other Ts would modify the algorithm results but its main structure would be unchanged.

We modified the sentence to clarify its meaning (Part 1, L. 222): "other Ts can be explored without affecting the main structure of the algorithm (see e.g., Nerger et al., 2012)."

In addition, see the response to the next Reviewer comment.

Equation (26): The GHOSH filter seems to use the same back-projection from the error-subspace to the state space as the SEIK filter. Nerger et al. (2012) have shown that this is inconsistent. Given that the authors are aware of the results of Nerger et al. (2012), I'm wondering why they decided to develop a new filter scheme on the basis of a filter that was shown to be mathematically inconsistent. A consistent form could be built easily following the ESTKF introduced by Nerger et al. (2012), which is likewise an error-subspace formulation.

The response to this comment would probably help to further clarify the two previous points issued by the Reviewer.

Nerger et al. (2012) showed that the SEIK is "inconsistent" in a specific sense:

*** starting from an ensemble X , the SEIK extracts an error subspace base $L = X T$ and a covariance matrix A in that subspace.**

*** If the SEIK sampling is applied to L and A , the obtained ensemble is different from X .**

*** In this context, "inconsistent" means that "going from an ensemble to an error subspace" and "going from an error subspace to an ensemble" are not one the inverse of the other.**

This kind of inconsistency may be not always problematic, indeed the same error pdf can be represented (in terms of mean and covariance) by infinite different ensembles. ESTKF (in its deterministic version) is more consistent in the sense that its sampling strategy provides exactly the same ensemble if no observations are available, and, if the assimilation does not change dramatically the error subspace, then the ensemble after assimilation would be "not too different" from the ensemble before assimilation. This feature is desirable if, for example, there are problems with physical balance constraints, as mentioned above by the Reviewer. On the other hand, in the Lorenz-96 applications presented by Nerger 2012 it has been also shown that SEIK with random rotations (applied also in GHOSH) is as good as ESTKF. Further, a GHOSH filter

built with the ESTKF approach presents a further layer of complexity, since the corresponding T matrix should change at every forecast due to the fact that the sampling depends on the principal components of the ensemble. Taking all this into account, we choose to propose the present version of the GHOSH filter.

According to the responses given to the present and the two previous Reviewer comments, we added a paragraph in the manuscript discussion about possible improvements in balancing constraints that could be obtained by applying an ESTKF approach in the GHOSH filter (Part 1, L. 542).

Equation (31): Using a model error covariance matrix Q in this form can be inconsistent for nonlinear models unless it is applied at each time step. This relates to me earlier comment requesting a clarification whether the resampling is performed at each model time step or only at those time steps at which also an analysis step is computed. If it is not applied at each time, please provide additional explanation why this linearized form of applying model errors is used.

The Q matrix in equation (31) is based on the idea of parametrizing with an additive Gaussian noise (with zero mean and covariance matrix Q) some of the sources of uncertainty in the forecast estimation not already accounted for by the ensemble. Following this interpretation, Q contributes to P only at the pre-fixed forecast times during the GHOSH forecast phase, and not at each model integration time step. We acknowledge that this strategy implements a relatively simple approach and that also other strategies could be applied (e.g., model parameters perturbation, as underlined by the Reviewer in a later comment). However, the strategy illustrated in equation (31) (adopted in Pham et al. (1998a,b), as the reviewer points out in the next comment) seems to be quite suitable for the Kalman filter equations, which also imply a Gaussian approximation at each assimilation. Further, the proposed strategy encompasses the concept of hybridization of P by adding a static covariance matrix representing ensemble-non-dependent uncertainties. Please, see also the responses to the next Reviewer comments, as they further expand this subject.

We changed the manuscript to better describe the meaning of Q (Part 1, L. 259):

“ Q_i is the $N \times N$ covariance matrix of an additive unbiased noise parameterizing some of the sources of uncertainty in the forecast estimation not already accounted for by the ensemble”

Lines 264/265, Eq. (32): It is stated "The last term in equation (32) is one of the novel element of the present work.". This statement is not fully true. Actually the original papers about the SEIK and SEEK filters (Pham et al., 1998a,b) include the model error covariance matrix Q. There, also a projection with L is used. A difference is that Pham et al. use a projection on Q while here a projection on the inverse of Q is used. (Given that Q is a large matrix in realistic applications, the projection by Pham et al. is likely computationally more efficient.)

Thanks to this Reviewer comment, the Q projection in the GHOSH filter was better explained. Indeed, we clarified that our approach is different from the one adopted in Pham et al. (1998a and b) since it is non-orthogonal and it is not equivalent to the inverse of the projection proposed by Pham et al. (1998a, b) (Part 1, L. 266):

“While the use of Q in equation (31) follows Pham et al. (1998a and b), equation (32) projects Q in a novel non-orthogonal way that is induced by the scalar product defined by Q-1. This approach can be interpreted as a form of hybridization that aims at focusing on the effects of the parametric uncertainty in the ensemble uncertainty subspace.”

Line 262: I cannot follow the argumentation that the addition of Q is a 'hybridization' (the manuscripts states this already in the introduction and further argues for it around line 603 in the discussion section). The hybridization as, e.g., explained in the reviews by Bannister (2017a, 2008) is a combination of a time-dependent covariance matrix with a covariance matrix representing climatology. The combination of both is used to represent the state error covariance matrix P. This is different from adding a model error covariance matrix which represents dynamic model uncertainties. Further, the addition of the model error covariance matrix is neither new nor a requirement of the GHOSH filter. Likewise it could be applied in the SEIK filter as the publications by Pham et al. (1998a,b) showed, even though this approach seems to be hardly used nowadays. To this end I recommend to avoid the potentially misleading term 'hybrid'.

We fully agree with the Reviewer that hybridization is a combination of a time-dependent (i.e., ensemble-dependent) covariance matrix with a parametrized covariance matrix, often derived from climatology. As discussed in the previous responses, we also acknowledge that the approach to include Q proposed in Eq. (32) is similar to the one followed by Pham et al. (1998a and b). On the other hand, it is worth noticing that Eq. (31) describes the covariance matrix P as the sum of a dynamical part derived from the ensemble and a parametric part named Q. Indeed, Q is a pre-computed non-ensemble-dependent matrix that parametrizes other sources of error (e.g., model errors, undersampling errors, nonlinearity errors), and a climatological covariance matrix is a suitable candidate for Q. In this sense, both the SEIK (as proposed by Pham et al., 1998a and b) and GHOSH can be seen as “hybrid” filters, even if Pham et al. (1998 a and b) never used the ‘hybrid’ term.

In the manuscript, we clarified that Q in Eq. (31) and (32) is a parametric part in the definition of P, as proposed in a previous response.

Lines 284/285: On the forecast resampling of the GHOSH filter it is stated "it takes into account the model error effects in equations (31) and (32), which are otherwise neglected". The claim that without the addition of the model error covariance matrix Q in the sampling the effect of model errors would be neglected is not fully true. In modern applications of ensemble Kalman filters, the model error is typically represented by stochastic perturbations

during the model integration instead of adding a model error covariance matrix at the end of the forecast phase. (The perturbations are expected to allow for a better representation of model nonlinearity on the model errors compared to the linearized effect of adding a model error covariance matrix)

We agree with the Reviewer: the forecast resampling is not the only way to account for model errors. The sentence was indeed misleading and was changed in (Part 1, L. 287): “it takes into account the Qi effects in equations (31) and (32), which would be otherwise neglected.”

Lines 314/315: "A reasonable choice is a polynomial weight function" - what makes the particular choice 'reasonable' and why was this form chosen? Actually, the proposed function in Eq. (41) seems to be uncommon (Most common is the 5th order polynomial of Gaspari and Cohn (2006), which is also a covariance function). Is the definition in Eq. (41) complete (it seems that $f > 0$ for $d > d_l$, which should not happen)?

We meant “reasonable” in the sense that Eq. 41 is the simplest (i.e., lowest order) polynomial function with the required properties (it is equal to 1 in p and goes smoothly to zero out of the localization radius). We changed the text in order to improve the definition of f (including that $f=0$ for $d > d_l$), and we inserted the citation of Gaspari and Cohn (2006) as another possible option (Part 1, L. 340).

“A sound choice is to rescale by a function that is equal to 1 in p and goes smoothly to zero out of the localization radius, e.g.,

$f(d) = [...]$,

with d being the distance from p and d_l the localization radius. Equation (41) is the simplest (i.e., lowest order) polynomial function with the required properties, but other options are viable (e.g., the fifth-order piecewise rational function of Gaspari and Cohn, 1999).”

Section 4.1 Lorenz-96 experiments

As mentioned before there are several weaknesses which should be corrected. In the current form the configuration seems to be particularly problematic for the SEIK filter, e.g. due to too short experiments and too less inflation, so that the strong improvement by the GHOSH filter relative to the SEIK filter is caused by the particular configuration.

- Perhaps it would be useful to include 'Lorenz-96 model' in the section title. Given that this is a toy model, the experiments are necessarily twin experiments.

The title was changed according to the Reviewer suggestion.

- The model is configured with state dimension 62. While this is a valid size, this value is very unusual and I'm not aware of other publications using it. Can the authors give a reason for this particular choice which makes it difficult, if not impossible, to compare the results to previous studies? The properties of different ensemble Kalman filters when applied to the Lorenz-96 model have been assessed before; thus they can be considered as known. The SEIK filter was however not often used, but it was applied with the Lorenz-96 model e.g. in Nerger et al. (2012). It would support the results of the manuscript if the authors could present consistent result to such previous studies.

The number of variables (state dimension) $N=62$ has been chosen according to the maximum ensemble size of 63. Indeed, in the twin experiments, the different ensemble sizes have been chosen (Part 1, L. 397) in order to make it possible to test different values of s , i.e., the number of principal components approximated with order h higher than two. In particular, we chose to test $h=5$ for s ranging from 2 to 5, corresponding to N ranging from 7 to 63. As explained in the manuscript, the choices on s and N are made in order to guarantee the maximum number s of components approximated with order $h=5$ for each ensemble size, resulting in $N=63$ for $h=5$ and $s=5$.

A more common choice would have been to use $N=40$ (as done in Nerger et al., 2012), however $N=40$ would have been too small to test the $s=5$ case. In addition, we are interested in studying the behaviour of the filters on non-assimilated variables, thus our observation operator is different from the one adopted in Nerger et al. (2012), where all variables were observed. The setting differences preclude the quantitative comparison between Lorenz-96 experiments in Nerger et al. (2012) and in the present study, however qualitative comparisons are possible since both studies provide results that are internally consistent. Finally, thanks to the Reviewer comments, the range of twin experiments was enlarged making the results more comparable with previous studies (see later comments). We propose to enrich the discussion about comparison of the results with other Lorenz-96 studies with the motivations and considerations proposed in the present comment response (Part 1, L.515).

- The duration of the experiment is only 20 time units. This results in experiments with observations being available in intervals between 0.1 and 0.3 time unit result in only between 200 and 66 analysis steps. As it is known that the filters sometimes converge slowly, the experiments likely mainly assess how fast the filters converge compared to how well they perform after convergence. In addition, the cases with longer observation intervals are expected to be quite unstable and one needs to average over a sufficient number of analysis steps to get interpretable results. To this end, length of the experiments is much too short. E.g. Sakov and Oke (2006) run over 11000 time units while Nerger et al. (2012) run over 2500 time units. Using such longer experiments would yield comparability to previous study with the Lorenz-96 model. Actually, I experimented with the code provided in Zenodo and the case `EnsSize=31, delta_obs=0.1, forget=0.7` and `t_span=[20.0, 100.0]` yields the same

RMSE level for both filters in the second half of the experiment. This also happens for $\text{delta_obs}=0.15$, $\text{forget}=0.7$, but the spinup for the SEIK filter is slower. Thus, with longer experiments at least the asymptotic difference in the filter results becomes very small also for the shorted forecast periods in contrast to what is shown in Fig. 4. This seems to change for $\text{delta_obs}\geq 0.2$ where the SEIK filter yields larger RMSE and tends to diverge for $\text{EnsSize}=31$ and lower. For $\text{EnsSize}=63$, the RMSE seems to be nearly identical (this might be indicated by the pink color in Fig. 4, but the colorscale only allows a qualitative estimate because it scales with $1/3$, $1/2$, 1 so that the values at the ticks for <1 are unclear)

We really thank the Reviewer for having tested the code provided in Zenodo. Moreover, the comment about the experiment length can give us the opportunity to enlarge the experiment setups presented in the manuscript adding experiments with a longer time window. On the other hand, we believe that the results for 20 time units are relevant to evaluate the GHOSH filter performances. In particular, it is worth noticing that:

- In the Lorenz 96 model 0.2 time units are comparable to 1 day of atmospheric dynamics (see, e.g., Grooms 2022), thus a 20 time units long experiment is comparable to 100 days.
- The GHOSH filter has been developed with the aim of realistic applications in oceanographic operational simulations. In this framework, an initial fast convergence of the filter is strongly beneficial to improve the simulation accuracy on temporal scales of interest. Moreover, in the slower-than-atmosphere ocean dynamics, 20 time units can be compared to even more than 100 days.
- In the evaluation of the RMSEs, we excluded the first 10 time units helping to exclude the initial transient phase in the comparison between the two filters.

Concerning the Reviewer's concerns about possible results instability in the cases with longer observation intervals, we agree that this issue can affect experiments with relatively low number of analysis steps. To avoid this kind of instabilities, in the present study we used a large number of experiments for each tested DA setting. Indeed, 400 experiments have been carried out for each setting (i.e., each square in Fig. 4) and results of Fig. 4 are obtained averaging over all the 400 experiments. Having used this relatively large number of experiments ensure an adequate stability of the results, since the expected standard deviation of the results obtained over N experiments is $\sqrt{N}$ smaller than the standard deviation of a random single experiment (i.e., 20 times smaller in our study).

Based on the motivations discussed above and the suggestions proposed by the Reviewer, the manuscript was modified as follows:

- A comment about the exclusion of the first 10 time units for the RMSE evaluation was added (Part 1, L. 362).
- We commented about the stability thanks to the large number of experiments used to compare SEIK and GHOSH (Part 1, L. 422).
- We added the results of a longer time window to the manuscript (150 time units, i.e., 2 years, based on 0.2 time units = 1 day) comparing SEIK and GHOSH tuned with their respective best forgetting factors in each setting

(according to the previous shorter experiments). Results on the 150 time units are provided in Fig. R2, and show that GHOSH performances are at least as good as those of SEIK, with largest reductions of the RMSE for observation frequency=0.2 and ensemble size=31 (Part 1, Fig. 6, Section 5.2).

- In Fig. 4, tick values were corrected to be more readable.

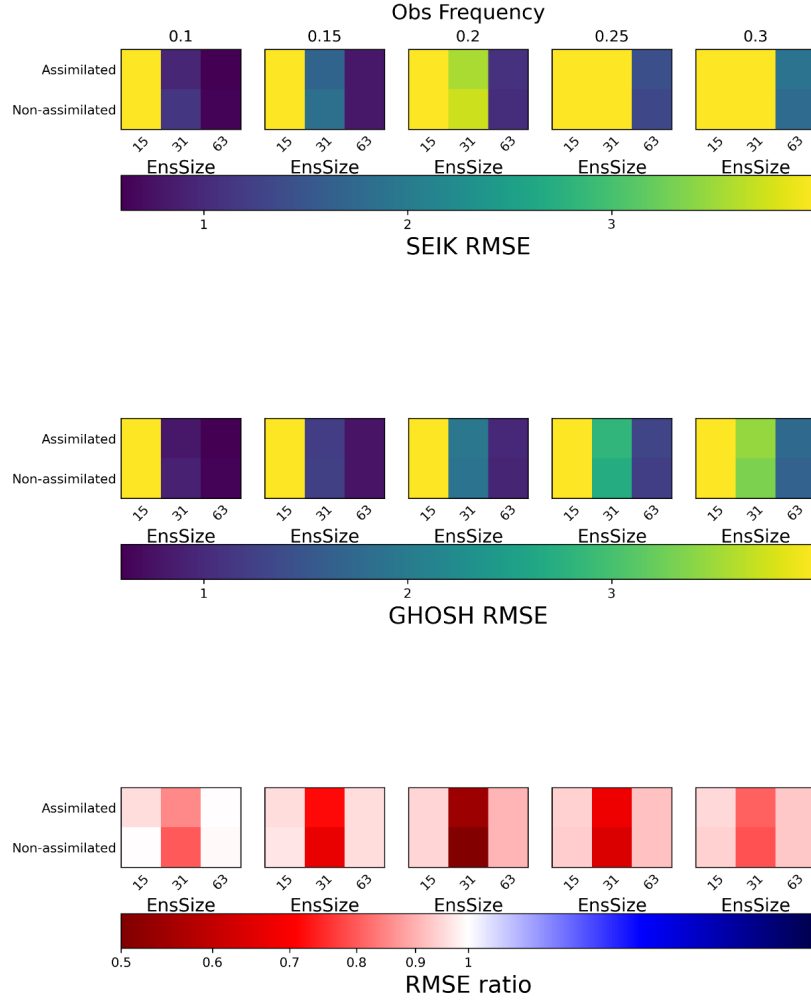


Fig. R2. Result summary of twin experiment on 120 time units. RMSE for SEIK (top), GHOSH (middle) and their ratio (bottom) are shown for different observation frequencies (columns). RMSEs are calculated for assimilated and non-assimilated observations (rows in each colour map) and for different numbers of ensemble members (columns in each colour map).

- The initial ensemble is sampled using a diagonal error covariance matrix. This is very unusual since typically second-order exact sampling is applied to a covariance matrix that is estimated from the model dynamics. I like to point out that usually when we apply a toy model like Lorenz-96, we attempt to make the experiments as realistic as possible, but a diagonal matrix would not be used in real applications. If a diagonal matrix is used to

generate an ensemble of $r+1$ states only r elements of the state vector will be perturbed (usually this would be the first r elements of the state vector, but in the provided Python code the order is changed due to the use of a sorting routine (resorting the prescribed value 5 in all elements) and the perturbation is distributed. This distribution was not changed when repeating the experiments, e.g. elements 16 and 62 were perturbed in case of ensemble size 3), which might also influence the results). Using the common approach of a matrix sampled from a model trajectory would yield a matrix that is not diagonal. Thus, the eigenvectors will not be the unit vectors and more elements of the state vector will be perturbed. This should lead to more realistic data assimilation experiments which are more representative. Hence, I strongly recommend to use this common approach.

In agreement with the Reviewer suggestion, we re-run all the experiments initialising the filters according to the climatological variability of the system. We changed the parts of the manuscript relevant to the filter initialization, the figures and the exact numbers in the result section accordingly. Other descriptive parts remained untouched since the overall comparison between the two filters remains the same as in the submitted manuscript. For instance, Fig. R3 results are very similar to those shown in Fig. 4 in the submitted manuscript.

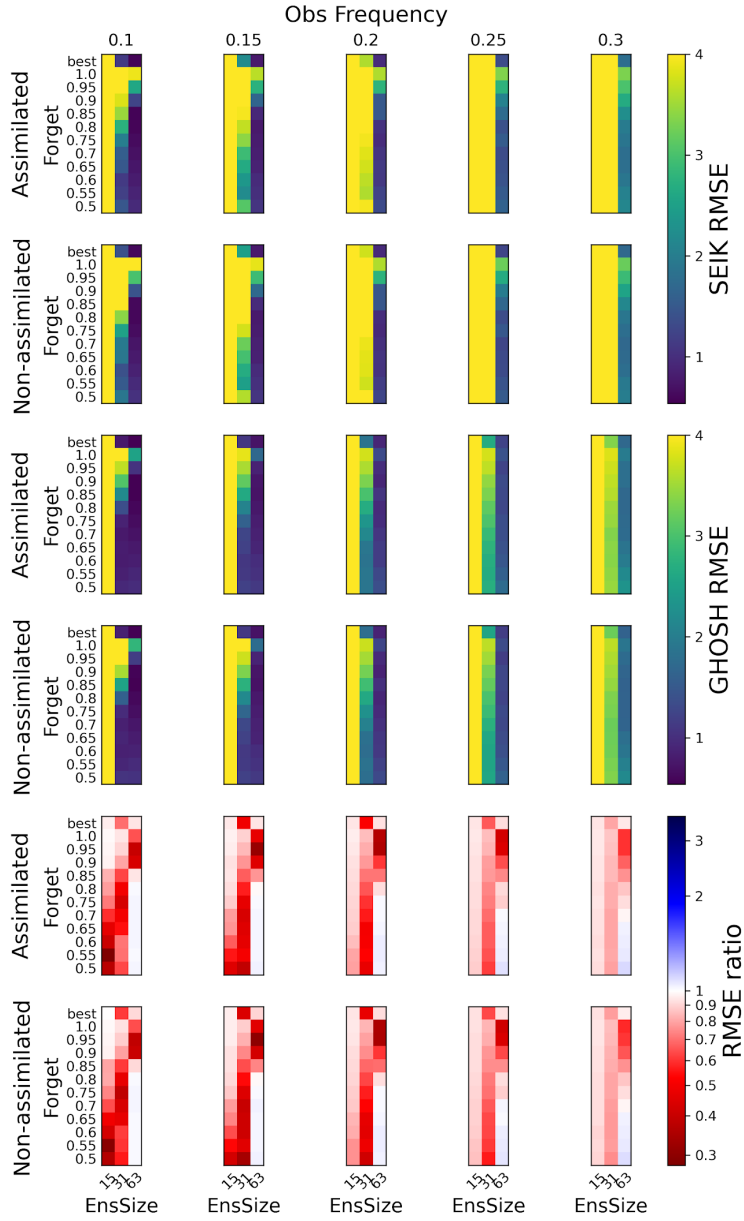


Fig. R3. Result summary of twin experiments: each square in the colour maps represents the aggregated results of 400 twin experiments, changing truth, observations and initial conditions. The results are summarised with colour maps aggregated in six different rows, from top to bottom: SEIK RMSE of assimilated variables, SEIK RMSE of non-assimilated variables, GHOSH RMSE of assimilated variables, GHOSH RMSE of non-assimilated variables, the ratio of GHOSH RMSE over SEIK RMSE of assimilated variables, the ratio of GHOSH RMSE over SEIK RMSE of non-assimilated variables (red colour implies that GHOSH is better than SEIK). Each column of colour maps has a different observation frequency, with the numbers on the top indicating the time elapsed between each observation/assimilation. Each color map shows different forgetting factors (Forget, along the y axis) and ensemble sizes (EnsSize, along the x axis).

- The model is run using the Python library solver 'solve_ivp'. As also described by the authors as a possible reason for the long run time and instability with the SEIK filter, this solver does not use a fixed time step size. Given that the Lorenz-96 model is deterministically chaotic, its behavior will vary when the time step size is varied. To this end, it should be better to use e.g. a classical Runge-Kutta 4th order implementation with fixed time step size (which 0.05 is a typical value). This is also the typical implementation as e.g. used by Nerger et al. (2012).

As noticed by the Reviewer, the solver used in the twin experiments is the scipy solve_ivp method RK45 that makes use of a non-fixed time step. Detailed information on the solver are provided in the scipy official documentation:

- “Explicit Runge-Kutta method of order 5(4) [J. R. Dormand, P. J. Prince, “A family of embedded Runge-Kutta formulae”, Journal of Computational and Applied Mathematics, Vol. 6, No. 1, pp. 19-26, 1980]. The error is controlled assuming accuracy of the fourth-order method, but steps are taken using the fifth-order accurate formula (local extrapolation is done). A quartic interpolation polynomial is used for the dense output”.
- “The solver keeps the local error estimates less than $atol + rtol * abs(y)$. Here $rtol$ controls a relative accuracy (number of correct digits), while $atol$ controls absolute accuracy (number of correct decimal places)” (the state vector is referred as y).

We chose this solver since the non-fixed time step size is an advanced feature that increases/decreases the time step size in order to keep the error under a fixed user-defined threshold. We corrected the manuscript in order to avoid the misleading interpretation of being the non-fixed time step that causes SEIK instabilities and longer run time. Instead, we clarified that it is the SEIK instability that leads to longer run time thanks to the non-fixed time step. In fact, the time step is reduced automatically by the solver to keep the error under the prescribed tolerance even in presence of stiffness induced by the SEIK instability (Part 1, L. 485).

“This difference in computational time is more evident and occurs more often when the forgetting factor is lower (i.e., when the inflation is more pronounced). The reason can be understood by looking at Fig. 5: in this experiment with forgetting factor equal to 0.6, the SEIK filter presents an unstable and diverging behaviour. When it happens, the SciPy’s solve_ivp integrating routine reduces the time step size for granting the required accuracy, which comes with longer integration time.”

- For ensemble size 15 the filters likely diverge (From Fig. 4 the RMSE appear to be at around 4, and Fig. 1 of Nerger et al. (2012) shows divergence for ensemble size <18 , which is related to the value of the forcing parameter in the model). There is no point in comparing the performance of diverging filters. Both filters fail, but one a little less than the other. This likewise holds for the statement on ensemble size 7 in line 495. In fact a filter divergence

might also happen for ensemble size 31 for the observation intervals 0.25 and 0.3 for both filters. The RMS error for the SEIK filter seems to be above 3, which might be high enough to indicate filter divergence. For the GHOSH filter, the errors seem to be between 2.5 and 3, which might also be divergence. (Unfortunately, the colorbar makes it difficult to see such details. Please consider to use a colorbar that not only varies brightness, but also the hue). In any case, it is obviously easy to check for filter divergence - if the estimated error, i.e. the ensemble spread, is much smaller than the true error (RMSE), the filter has diverged. I recommend to perform such check on all results. Further, if the RMSE is not less than the long term standard deviation of the model dynamics (which seems to be about 3.6) the data assimilation obviously failed.

We thank the Reviewer for the comments on results shown in Fig. 4, where we would like to give a fair and broad overview of the comparison between the two filters. Indeed, we think that it can be useful to provide RMSE values, even in case of non-convergence of the filters, to highlight possible relevant RMSE differences, and we do not feel that hiding some information would be beneficial for the reader. Hence, also considering the response to the Reviewer comment on line 494, we maintained the SEIK-GHOSH RMSE comparison in the last two rows of the Fig. 4 colormaps. On the other hand, we acknowledge that it is significant to highlight non-convergence cases, and we rescaled SEIK and GHOSH colorbars using a maximum value equal to 4 (i.e., nearly equal to the climatological standard deviation that is 3.7), such that it will be clear for the reader when a set of experiments does not provide any improvement with respect to the climatology. Finally, we updated Fig. 4 with a colorbar that includes hue. Both the mentioned improvements of Fig. 4 have been implemented in Fig. R2.

- Particular cases seem to be the observation intervals 0.2 and 0.25. Here it might be the that SEIK filter diverges, while the GHOSH filter still converges. This is a particular difference, which could be interesting to analyze. (However, it might be influenced by the effect in the next statement)

We thank the Reviewer for underlining this interesting difference between SEIK and GHOSH, which was included in the result section (Part 1, L. 441): “Remarkably, the GHOSH filter with 31 ensemble members improves the state

estimation error compared to the climatological standard deviation of the model for every observation interval. The same is not

true for the SEIK filter, which achieves convergence (i.e., RMSE lower than the climatological standard deviation, at least in its best forgetting factor configuration) only for observation intervals shorter than 0.2 time units.”

- In the cases with observations intervals 0.2 and larger, the GHOSH filter shows smaller RMSE which are still at a high level. Since the analysis errors are show, this points to the

question whether the sampling in the GHOSH filter leads to a higher ensemble variance than the SEIK filter. This would basically allow for a larger increment in the analysis step. I recommend to check whether this holds, since this effect could be unrelated to a representation of higher orders, but mainly an effect of more randomness.

We would like to address the comment of the Reviewer about the motivation of GHOSH better performances from two perspectives:

- The set of experiments summarised in Fig. 4 can help to verify that GHOSH performances are not motivated by a larger ensemble spread. Indeed, a larger GHOSH ensemble spread would in some cases also degrade the GHOSH performances with respect to SEIK, but this is not the case. For instance, looking at the RMSDs for non-assimilated variables with 31 ensemble members and observation frequency equal to 0.1 (Fig. R2), SEIK shows best performances when the forgetting factor is equal to 0.6. Below 0.6 (i.e., with a larger inflation), SEIK RMSE increases. For the same settings, GHOSH RMSE is lower than in SEIK for the whole forgetting factor range (bottom panel in Fig. R2), even when the forgetting factor is lower (i.e., the inflation is larger) than the SEIK optimal. Meaning that GHOSH performs better than SEIK independently from the ensemble spread. If the better performances of GHOSH were related to a larger spread, an inflation larger than the SEIK optimal would degrade GHOSH with respect to SEIK (i.e., GHOSH RMSEs would be larger than SEIK RMSEs).
- From a theoretical point of view, it has to be kept in mind that the GHOSH sampling has exactly the same spread as SEIK's second order exact sampling, since both of them produce ensembles that preserve variances. Further, the GHOSH forecast ensemble has a smaller error in the forecast mean estimation, as proved in Appendix A. Finally, it is worth mentioning that the same arguments used for the mean prove that also forecast covariance matrix is better estimated by GHOSH with respect to SEIK.

- For the observation intervals 0.1, 0.15 and 0.2, Fig. 4 leaves the impression that the forgetting factor of 0.7 is still too large for the SEIK filter. The optimal inflation is obviously visible when the RMSE is minimum and when it increases for even smaller values of the forgetting factor and this point should be visible in the figure.

According to the Reviewer's suggestion, we extended the forgetting factor range down to 0.5 (Fig. R3 and Fig. 4 in the manuscript).

- There are also some important pieces of information missing: Firstly, I suppose that the experiment is run without localization, but this is never stated in the manuscript. In addition,

the choice of the actual solver in 'solve_ivp' is not mentioned. Also, the observation error variance should be explicitly stated

We thank the Reviewer for the suggestions to add more detailed information on experiments. The suggested improvements to the text have been added in the manuscript:

(Part 1) L. 398:

“Hence, the values adopted for s are 2,3,4 and 5, respectively for 7,15,31 and 63 ensemble members. Localization has not been applied, since the number of variables N is relatively small and comparable with the ensemble size.”

(Part 1) L. 374:

“The equations have been implemented in python and numerically solved using SciPy’s solve_ivp routine with its default solver method (i.e., RK45).”

The observation error variance was included according to a previous Reviewer’s comment.

- The numerical experiments use random numbers. I recommend to re-initialize the seed for the random number at the beginning of an experiment. Only in this case, one can obtain reproducible results. One can run with different seeds to exclude that a particular interpretation results from these.

Since the results are averaged over 400 experiments per square (Fig. 4), even if they are random, the expected variability is 20 (square root of 400) times smaller than the standard deviation of the single experiment. This is sufficient to make the experiment results robust enough to be reproduced. Indeed, in many previous tests (not shown) we obtained the same (or non-significantly different) ratios between GHOSH and SEIK RMSEs. As proposed in a previous comment, we are now more explicit about the robustness of Fig. 4 results.

Section 5.1, title: As for section 4.1 I recommend to mention the Lorenz-96 model in the section title

The title was changed as suggested.

Line 490: Please define 'best'. I suppose that the result with the smallest RMSE is meant.

It is now explicitly stated in the text that ‘best’ is the best result in terms of RMSE.

L 490: “The first line in each colour-map, labelled “best”, represents the best result, in terms of lowest RMSE, obtained among the set of tested forgetting factors.”

Line 494: The results for ensemble size 7 were not shown "since in this case SEIK and GHOSH behave very similarly, showing very poor performances". This is the known behavior that without localization the filters diverge if the ensemble is too small. However, this likewise holds for ensemble size 15 which should also be removed. There is no point in discussing RMSEs of diverged filters.

Being GHOSH a new filter featuring a higher order of approximation compared with existing filters, we think that it deserves to be tested in heterogeneous settings. As discussed before, we think that it is fair to report a large range of results, also because they are all showing a better behaviour of GHOSH with respect to SEIK, even in case of non-convergence. We imagine that this information could be interesting for real filters applications: it could be that under certain circumstances (maybe in a small time window) a realistic model is more non-linear, or harder to successfully assimilate. Maybe the number of ensemble members could be not big enough under those circumstances, and the assimilation performances would be relatively poor. Our results suggest that, also under these circumstances, the GHOSH filter would provide a better state estimation than the second-order SEIK filter.

Fig. 3: The figure shows that the first two variables are very well estimated by both filters (and partly better by the SEIK and the GHOSH filter) in between the times ~ 1 to 3.5 for the first variables and even longer for variable 2. In contrast the RMSE is still large at around a value of 3. This indicates that most of the other variables of the model have a much larger error. This likely supports my earlier statement on the ensemble initialization: For an ensemble of size $r+1$ only r elements of the state are perturbed in the ensemble and can hence not be corrected by the data assimilation. Only if the model dynamics act for long enough time, ensemble spread will build up, but at this time, the ensemble spread might already have become too low to achieve a convergence of the filter. At this point, experiments with the Python code show that a larger inflation (e.g. $\text{forget}=0.6$) helps to obtain convergence. The resampling performed by the GHOSH filter might distribute the spread faster due to its random effects. If this happens this advantage is the GHOSH filter would mainly be induced by the particular initial sampling and might not be an effect of higher orders. Please check if such an effect is present or can be excluded.

Thanks to a previous Reviewer comment, we verified that the ensemble initialisation does not affect the GHOSH performances with respect to SEIK. Moreover, Fig. 3 was modified accordingly.

Lines 498-501, Fig. 4: It is described that the advantage of the GHOSH filter is particularly large if the observations are very frequent. This result looks surprising for me. We know that the nonlinearity of the DA problem increases if the forecast phase is longer, but here the

higher order sampling, which should be relevant for non-Gaussian ensemble distributions, is particularly good if the forecast phase is short. Also the experiments with the Lorenz-96 model show a smaller improvement of the GHOSH filter relative to the SEIK filter for longer forecasts despite the fact that these increase the nonlinearity of the DA problem. Further it looks surprising that the 'best' result of the GHOSH filter in case of the largest ensemble size of 63 is only slightly better than that of the SEIK filter, while for ensemble size 31 the difference is larger. Here, the statement on lines 500/501 that the largest improvement is obtained when "the filter can take into account a high dimensional error subspace (i.e., the ensemble size is large)" does not seem to hold. If so, why is the improvement of the GHOSH less for the largest ensemble? The larger ensemble reduces the sampling errors and this should be the case for all orders. Here it looks like that the reduced sampling errors in the second-order sampling have a stronger effect than the higher order sampling. This effect should be carefully discussed. For the figure it would be useful to clearly mark cases in which the filters diverge.

Please, note that lines 447-449??? refers to the largest improvements among all the experiments, and they are not referring to the "best" case. In that context, the statement "the filter can take into account a high dimensional error subspace (i.e., the ensemble size is large)" refers to the darkest squares appearing in the 63-EnsSize columns. Lines 450-453 instead, discuss the "best" inflation case, pointing out that the darkest squares correspond to a "moderate number of ensemble members (31)".

We specified the above considerations at L. 446 (Part 1):

"In the range of explored forgetting factors, the largest improvements (dark red, RMSE ratio 0.30) occur when [...]"

Concerning the other questions raised by the Reviewer, we can argue some explanations. Looking at the "best" inflation configuration, in the case of 31 ensemble members, the RMSE ratio between GHOSH and SEIK reaches its maximum for observations every 0.15 time units and then decreases monotonically. In the case of 63 ensemble members, instead, the ratio increases monotonically up to 0.25 time units and then decreases. Thus, in both cases, the GHOSH filter shows its better capability to manage non-linearities (since the max is not at the higher observation frequency), but at some point, the RMSE ratio stops improving. On the other hand, it is worth mentioning that the RMSE ratio is not a perfect proxy of the filter capability of managing non-linearities. Indeed, when the error has a filter-independent part and a relatively small amount of information is available, neither of the filters can achieve high accuracy. Actually, having a part of the error that does not depend on the chosen filter (an information-dependent error) is highly expected, and it is the part of the error intrinsically dependent by the amount of information. As the information-dependent error becomes dominant with respect to the improvement given by the capability of managing non-linearities, the RMSE ratio between GHOSH and SEIK increases consequently.

Considering the ensemble size effects, in our opinion, the fact that a better RMSE ratio is obtained with 31 ensemble members instead of 63, is understandable by taking into account that the GHOSH algorithm reduces the error of the mean (other moments are positively affected, but with lesser magnitude) by exploiting higher

order moments (the majority of which are imposed as hyperparameters, not estimated by the ensemble). We fully agree with the Reviewer on the fact that a bigger ensemble reduces the sampling error for all moments, but this does not affect GHOSH more than how it affects SEIK, since both of them estimate only mean and covariance from the ensemble. The point is that the GHOSH is capable of reducing the sampling error even with a smaller ensemble size. But when the error is small enough (big ensemble size), the advantage of the GHOSH is not as impressive as in the case of a smaller ensemble size. In other words, if the performances of a second-order-exact filter (i.e., SEIK in our tests) are close to the best possible performance given a certain amount of information, then it remains only a small part of the error that is possible to improve by the use of a higher order filter (such as the GHOSH filter).

Considering these and previous Reviewer comments, we enriched the discussion.

line 512: The RMSE of the GHOSH filter is mentioned to be slightly more reduced than those of the SEIK filter. Are these differences statistically significant?

Based on the Reviewer comment, we would like to clarify that “slightly larger” refers to the RMSE reduction in the non-assimilated variables with respect to the assimilated ones. Meaning that the RMSE reduction happens in both the assimilated and in the non-assimilated variables, but slightly more in the non-assimilated ones. While the RMSE reduction from SEIK to GHOSH is quite clear, since it occurs in the majority of the settings (each tested on 400 simulations), we did not quantify the statistical significance of the RMSE ratio difference between assimilated and non-assimilated variables, and we think that it can be out of the scope of the present work. However, the fact that in all the 15 cases (5 observation frequencies for each of the 3 ensemble sizes considered) the RMSE ratio was lower in the non-assimilated variables makes it hard to believe in a random chance. In addition, also the new tests in the larger time window (Fig. R1) show the same behaviour.

Computing cost: The computing cost is discussed shortly for the OGSTM-BFM model, but no actual numbers are provided. However, it should be possible to provide timings for the Lorenz-96 model case where one can compare the time of the SEIK filter analysis step with the timing of the GHOSH analysis step and sampling. This cost should be considered separately from the cost of the model. (The model is likely faster than the model in this case since the forecasts are only between 2 and 6 times steps of a Runge-Kutta scheme)

We thank the Reviewer for this comment that helped us to a more complete vision of the GHOSH computational costs. The GHOSH asymptotic computational complexity was added to the manuscript as well as the time to solution for the Lorenz96 experiments, with timings for filter operations and model integration (Part 1, L. 478).

“From the computational point of view, the whole experiment set needed around 9 computational hours. SEIK and GHOSH schemes used 12% and 16% of the total time respectively, while the rest was committed to model integration. The GHOSH filter executes more operations than SEIK (e.g., an eigenvalue decomposition) resulting in more computational time even if the asymptotic computational complexity of the two methods is the same (Section 3.1.6). However, even in the case of a relatively simple model like the Lorenz96, the time to solution is dominated by the model integration and the difference between GHOSH and SEIK only accounts for 4% of the total time. Unexpectedly, the model integration time (averaged every 100 twin experiments) when applying the SEIK filter sometimes is longer than the GHOSH filter case, up to twice the time, depending on some settings and random factors.”

Lines 613-618: This paragraph contains several over-statements. The error reduction that was achieved mainly showed that the filter method works successfully, but it does not show an advantage compared to e.g. the SEIK filter (which was not applied here). The statement that non-assimilated variables are not degraded cannot be generalized. In addition, it is incorrect as Table 2 shows: For daily assimilation there is degradation of a forgetting factor of 0.8 is used. For weekly assimilation, there is degradation for $h=2$ and $h=3$ if a forgetting factor of 0.5 is used. (This comment is no longer relevant if the 3D experiments are removed as recommended)

We thank the Reviewer for having provided this comment, which was taken into account for the manuscript on the 3D application (Part 2). In particular, thanks to the deep revision of the manuscript, it is now clear that: i) we are referring to our 3D experiment, without necessarily claiming that our results are generalizable, and ii) the observation error is optimised for one of the 3D simulations only. Thus, it can be expected that the others suffer from degradation. In this framework, we revised the discussion of the results of the sensitivity analysis in the 3D application.

Line 623: "it has been shown that the GHOSH increased accuracy reduces instabilities and numerical divergence". This statement is not valid in this form. For the Lorenz-96 model in the particular implementation used here, the statement holds. However, this finding might be specific for the time stepping methods used here and it cannot be generalized in any form.

We modified the sentence to highlight that this finding has been obtained in the tested configuration (Part 1, L. 506???): "it has been shown that the GHOSH increased accuracy reduces instabilities and numerical divergence in our Lorenz96 implementation".

Lines 656-659: Here implications the 'higher polynomial order' or the GHOSH sampling is discussed. Linked to the methods section 2, the relevance of the polynomial order was never

explained, but only mentioned. As such it is unlikely that many reader can follow the discussion. Given that neither GMD nor NPG are mathematical journals, I recommend to explain the relevant aspects already in Sec. 2.

We thank the Reviewer for highlighting this issue. In the Discussion, we added a reference to Appendix A, where the relevance of the polynomial order is proved (Part 1, L. 554):

“The advantage of a higher order is not limited to a better estimation of the mean state (Appendix A) but it extends to the covariance matrix of the uncertainty probability distribution.”

Moreover, the novel Appendix C, which contains explicit calculation in a simple case, can help the reader to better understand the idea behind the high order sampling.

Lines 663-667: " this strategy might lead to inaccurate estimations because the model error covariance matrix Q_i is not taken into account in the interpolation". This statement does again not take into account that there are other possibilities to apply model errors in the ensemble integration. When stochastic perturbations are applied, model errors is taken into account in the 'interpolation' that is done at the analysis time. This then also holds for nonlinear observation operations mentioned in line 666.

The words “model error” was removed, in order to avoid misleading interpretation of the statement, which refers to Q_i (representing climatological covariance, model error in the form of covariance matrix, or other forms of parametric error covariance). Moreover, the discussion was modified to clarify the role of the matrix Q_i (Part 1, L. 526???):

“this strategy might lead to inaccurate estimations because the covariance matrix Q_i of equation (31) is not taken into account in the interpolation. For this reason, if Q_i is not zero, GHOSH applies the observation operator to a new ensemble, resampled between the forecast and analysis phase (Fig. 2).”

Lines 671-672: Here a relation of the constant ensemble weights in the GHOSH filter to the weight in particle filters a drawn. I cannot see any relation because the weights have a different meaning. In the GHOSH filter, the weights are used to represent higher order moments, while in the particle filter the weights represent likelihoods. I recommend to remove this statement. (Please note that the reference to Bocquet (2010) is invalid as was discussed for the Introduction)

We agree with the Reviewer on the fact that particle filter (PF) weights represent likelihoods (differently from GHOSH), and that the scope of using weights are different in the two approaches. We modified the sentence to highlight those aspects

and to avoid unnecessary parallelism. Moreover, we included the suggested references in a later comment (Part 1, L. 559???):

“Similarly to other ensemble filters, a weighted ensemble is used in the GHOSH filter. However, the GHOSH filter differs substantially from the particle filter (e.g., van Leeuwen et al., 2019) and from other types of weighted ensemble filters, e.g., Hoteit et al. (2008) and Stordal et al. (2011). In the listed filters, weights (representing likelihood) change over time, while particles remain the same (and strategies are applied to resample the particles when weights “collapse”, e.g., van Leeuwen et al., 2019). In the case of GHOSH, the ensemble is used to estimate specific moments and then resampled to keep constant the weights, because those weights have specific desired properties that help reduce the error of the mean estimation. In this, GHOSH has similarities with the nonlinear ensemble transform filter (Tödter and Ahrens, 2015), which does a resampling to keep the weights constant but with uniform weights and a second-order sampling procedure.”

Lines 674-675, 698-699: Here it is argued about the computing cost, which was also discussed in the results section of the OGSTM-BFM model, where no actual numbers were provided. Actually, it should be possible to provide timings for the Lorenz-96 model case where one can compare the time of the SEIK filter analysis step with the timing of the sampling plus analysis step of the GHOSH filter. This cost should be considered separately from the cost of the model. As mentioned before, simply arguing that the time for the model is much higher than that for the filter and sampling is superficial and should be avoided.

Thanks to this and a previous Reviewer comment, the computational complexity of the GHOSH scheme is now mentioned in Section 3, while the timing was added to Lorenz96 experiment and discussed in the corresponding result section. We also inserted the following updated sentences to the text:

Part 1, L. 567: “Finally, the computational complexity of the GHOSH filter (Section 3.1.6) is comparable to second-order deterministic filters (e.g., SEIK, ETKF). The computational cost, as in most ensemble methods, mainly depends on the ensemble size, which multiplies the most demanding model integration cost (e.g., Nerger et al., 2005). Hence, the computational cost of the GHOSH filter is no more a concern than in second-order data assimilation schemes”

Part 1, L. 582: “the GHOSH approach keeps the same computational complexity and number of ensemble members as SEIK or ETKF and exploits the advantage of a higher order only where the approximation errors are larger.”

Lines 688-689: “the GHOSH filter showed improved capacity to take into account non-linearities by a lesser need of inflation with respect to SEIK”. Here, I don't see how

inflation should be related to non-linearity. This was never assessed in the manuscript and I don't think this is a common relationship. As such, the statement should be removed.

The sentence was misleading, since the relationship between nonlinearity and the need for inflation was totally implicit. Indeed, our comment was based on the fact that non-linearity introduces errors not considered by ensemble filters, which becomes overconfident and needs inflation to compensate (see e.g., Raanes et al., 2019; Bocquet et al., 2015; Rainwater and Hunt, 2013). Thus, it can be inferred that the more a filter is capable of tackling non-linearities, the less it needs for inflation (up to some extent, since overconfidence is also related to the ensemble size). The Lorenz 96 experiments showed that GHOSH required less inflation compared to SEIK to achieve comparable performances (Fig. 4), and in this sense we think that the GHOSH filter showed improved capacity to take into account nonlinearity.

We added a sentence in the Discussion (Part 1, L. 510) to better explain the link between the inflation, nonlinearity and results shown in Fig. 4:

“The lower need of inflation can be also seen as a GHOSH improved capacity to take into account non-linearity. Indeed, non-linearity typically introduces errors that are not considered by ensemble filters, which needs inflation to compensate for this uncertainty covariance underestimation (see e.g., Raanes et al., 2019; Bocquet et al., 2015; Rainwater and Hunt, 2013).”

Discussion section 6: It would be useful if the authors include a discussing relating the GHOSH filters to other existing filters that are aimed at nonlinear data assimilation and pointing out the differences. Partly this is done by mentioning particle filters in line 672. However, there are filters that are closer to the GHOSH filter. E.g. the Gaussian mixture filter (Hoteit et al., 2008, Stordal et al., 2011) seems to be related, but obviously different. Perhaps, also the nonlinear ensemble transform filter (Toedter and Ahrens, 2015) shares some similarities given that this is also a transform filter.

We thank the Reviewer for pointing out useful references which were added to the discussion (Part 1, L. 559):

“Similarly to other ensemble filters, a weighted ensemble is used in the GHOSH filter. However, the GHOSH filter differs substantially from the particle filter (e.g., van Leeuwen et al., 2019) and from other types of weighted ensemble filters, e.g., Hoteit et al. (2008) and Stordal et al. (2011). In the listed filters, weights (representing likelihood) change over time, while particles remain the same (and strategies are applied to resample the particles when weights “collapse”, e.g., van Leeuwen et al., 2019). In the case of GHOSH, the ensemble is used to estimate specific moments and then resampled to keep constant the weights, because those weights have specific desired properties that help reduce the error of the mean estimation. In this, GHOSH has similarities with the nonlinear ensemble transform filter (Tödter and Ahrens, 2015), which does a resampling to keep the weights constant but with uniform weights and a second-order sampling procedure.”

Code availability:

I was able to download the codes. Unfortunately, I was not able to find the GHOSH filter in the Fortran implementation. I found an option for the higher-order sampling but neither the place where it is applied and neither the analysis step of GHOSH. Maybe it would also be useful to provide a cleaner code (e.g. in the Python code there are many out-commented lines) and also some in-line documentation (there are essentially no comments in the codes which makes reading them very difficult). In any case, this would just be 'nice', but a paper acceptance would not depend on this.

We really thank the Reviewer for having downloaded the codes. Concerning the issues raised by the Reviewer:

- **The GHOSH filter originates from a tailored Fortran implementation of the SEIK filter. After enough modifications, it was no longer recognizable as SEIK, and we named it “GHOSH”. For that reason, the GHOSH routines are still named SEIK in that original version of the Fortran code used for the OGSTM application. We added a “readme” file in the repository to inform possible users about subroutine naming.**
- **The Python code has been substantially improved and, following the Reviewer suggestion, many out-commented lines have been removed. All the GHOSH filter routines are coded in a total of around 100 pythonic lines that should now be pretty easy to follow after reading the manuscript.**

Small things:

- line 162: 'identical' -> 'identity'

- line 248: Ω_i is not computed in Eq. (6), but only used there. Please rephrase

- line 335: 'by' -> 'from'

- line 492: 'lines' -> 'rows' (in general in all description of Fig. 4)

- line 611: 'up to 3 times better RMSE': 'better' is not well defined. Its better to quantify as e.g. 'up to X% lower RMSE'

All the “small things” highlighted by the Reviewer were revised.

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