Decadal climate predictions with the Canadian Earth System Model version 5 (CanESM5)

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Abstract.

The Canadian Earth System Model version 5 (CanESM5) developed at Environment and Climate Change Canada’s Canadian Centre for Climate Modelling and Analysis (CCCma) is participating in phase 6 of the Coupled Model Intercomparison Project (CMIP6). A 40-member ensemble of CanESM5 historical decadal forecasts (or hindcasts) is integrated for ten years from realistic initial states during 1961 to present using prescribed external forcing. The results are part of CCCma’s contribution to the Decadal Climate Prediction Project (DCPP) component of CMIP6. This paper evaluates CanESM5 large ensemble decadal hindcasts against observational benchmarks and against historical climate simulations initialized from pre-industrial control run states. The focus is on the evaluation of the potential predictability and actual skill of annual and multi-year averages of key oceanic and atmospheric fields at regional and global scales. The impact of initialization on prediction skill is quantified from the hindcasts decomposition into uninitialized and initialized components. The dependence of potential and actual skill on ensemble size is examined. CanESM5 decadal hindcasts skilfully predict upper-ocean states and surface climate with a significant impact from initialization that depend on climate variable, forecast range, and geographic location. Deficiencies in the skill of North Atlantic surface climate are identified and potential causes discussed. The inclusion of biogeochemical modules in CanESM5 enables the prediction of carbon cycle variables which are shown to be skilful on decadal time scales, with a strong long-lasting impact from initialization on skill in the ocean and a moderate short-lived impact on land.

1 Introduction

The Canadian Earth System Model version 5 (CanESM5) is the latest Canadian Centre for Climate Modelling and Analysis (CCCma) global climate model. It incorporates an interactive carbon cycle, and was developed to simulate historical climate change and variability, to make centennial scale projections of future climate, and to produce initialized climate predictions on seasonal to decadal time scales (Swart et al., 2019b, this issue, hereafter S19). S19 summarizes CanESM5 components and their coupling, together with the model’s ability to reproduce large scale features of the historical climate and its response to external forcing. This paper examines the predictive ability of CanESM5 on decadal time scales. CanESM5 decadal climate predictions are CCCma contribution to Component A of the Decadal Climate Prediction Project (DCPP; Boer et al., 2016)
endorsed by phase 6 of the Coupled Model Intercomparison Project (CMIP6). In this paper, the skill of key physical climate and carbon cycle variables is assessed.

The aim of near-term or decadal climate prediction is to provide end users with useful climate information, on time scales ranging from one year to a decade, that improves upon the information obtained from climate simulations that are not initialized from observation-based states (Merryfield et al., 2020). On decadal timescales, the evolution of the climate depends on the interplay between an externally forced component (e.g., resulting from the changes in greenhouse gas emissions, aerosols, land use, solar forcing) and internally generated natural variability. The climate response to the externally forced component may be estimated from climate simulations that are not initialized from observation-based states (henceforth referred to as “simulations”), but the internally generated component of these simulations is not expected to match observations. For decadal predictions, climate models are initialized from observation-based states and integrated forward for up to 10 years with prescribed external forcing. The expectation is that by taking advantage of predictable slowly-varying internally generated fluctuations of the climate system, including those originating from the ocean, climate phenomena such as multi-year atmospheric circulation changes (Smith et al., 2010; Sutton and Dong, 2012; Monerie et al., 2018), their impact on near-surface temperature and rainfall (Zhang and Delworth, 2006; Boer et al., 2013; McKinnon et al., 2016; Sheen et al., 2017), and the frequency of extreme weather events (Eade et al., 2012; Ruprich-Robert et al., 2018), can be predicted a year or more in advance. Furthermore, initialization affects the model response to external forcing and can potentially correct unrealistic simulated trends (Sospedra-Alfonso and Boer, 2020).

For decadal predictions to be credible they must be accompanied by measures of historical skill, and an understanding of the various processes that contribute to skill on decadal time scales (Meehl et al., 2009). CMIP5 (Taylor et al., 2012) provide a framework for quantifying the “added value” of initialized climate predictions over simulations (Meehl et al., 2009; Goddard et al., 2013) and, building upon this experience, the DCPP panel coordinated a comprehensive set of decadal prediction experiments endorsed by CMIP6. These include ensembles of historical initialized predictions (hindcasts) and historical simulations as well as basic guidelines for post-processing model output (Boer et al., 2016). Assessing the added value of decadal climate predictions over climate simulations can, however, be difficult. In the presence of a strong long-term climate response to external forcing, as is the case for near-surface air temperature for instance, the externally forced component carries most of the predictable variance and can mask the contribution of relatively weaker internal variations to the skill (Smith et al., 2019; Sospedra-Alfonso and Boer, 2020). Underestimation of the predictable signal can also degrade decadal forecast skill (Sienz et al., 2016), motivating the use of large ensembles to better extract the predictable component of the forecast (Scaife and Smith, 2018; Yeager et al., 2018; Smith et al., 2019, 2020; Deser et al., 2020).

In this work, we assess the potential predictability and actual skill of 40-member ensemble hindcasts produced using CanESM5. The hindcasts are initialized at the end of every year during 1960 to 2019 and run for 10 years. The potential predictability framework enables the detection of the climate model “signal” that is common to the ensemble of model predictions whereas actual skill refers to the forecast ability to reproduce the predictable signal of the climate system. The initial ensemble spread is intended to represent observational uncertainties (Merryfield et al., 2013) and to be small compared with the amplitude of the climate signal to be predicted. Forecast skill is compared here against climatology, persistence, and historical
simulations. The contribution to skill of initialized and uninitialized components of the forecasts are inferred with reference to
a 40-member ensemble of historical simulations, spanning 1850–2014. Both hindcasts and simulations use the same external
forcing parameters during their overlapping period. The dependence of potential and actual skill on ensemble size at regional
and global scales is also examined.
This paper serves to document the CanESM5 decadal hindcasts that are the CCCma contribution to Component A of the
DCPP (Boer et al., 2016). It highlights CCCma’s newly developed capabilities including prediction of biogeochemical variables
and the carbon cycle, as well as the use of large ensembles to better extract the predictable component of the forecasts. These
are steps towards a more comprehensive decadal climate prediction system at CCCma, although not without new challenges
and deficiencies, some of which are examined here. The remainder of the paper provides an overview of CanESM5 model
components (section 2), a description of the initialization process and ensemble generation (section 3), the methodology for
hindcasts evaluations (section 4) together with supporting information (appendices A and B), and assessments of potential
predictability and actual skill in the upper ocean (section 5) and the surface climate on land (section 7). Section 6 addresses
issues of the sea surface temperature hindcasts in the subpolar North Atlantic and Labrador Sea. The impact of ensemble
size on potential predictability and actual skill is discussed in section 8, whereas the predictability of carbon cycle variables
is examined in section 9. Section 10 provides a summary and the conclusions. CanESM5 output data used here including
hindcasts, assimilation runs to initialize hindcasts, volcanic experiments, and historical simulations are freely available from
the Earth System Grid Federation on https://esgf-node.llnl.gov/search/cmip6/.

2 The CanESM5 earth system model
A detailed description of CanESM5 and its components is given in S19 and the references therein, thus we provide only a
brief summary here. CanESM5 couples version 5 of the Canadian Atmosphere Model (CanAM5; Cole and Coauthors, 2019,
this issue) with the CanNEMO ocean component adapted from the Nucleus for European Modelling of the Ocean version
3.4.1 (NEMO3.4.1 Madec and Coauthors, 2012). CanAM5 incorporates the Canadian Land Surface Scheme version 3.6.2
(CLASS3.6.2; Verseghy, 2000) and the Canadian Terrestrial Ecosystem Model (CTEM; Arora and Boer, 2005) whereas Can-
NEMO represents ocean biogeochemistry with the Canadian Model of Ocean Carbon (CMOC; Zahariev et al., 2008; Christian
et al., 2010). Sea ice is simulated within the NEMO framework using version 2 of the Louvain-la-Neuve Sea Ice Model (LIM2;
Fichefet and Maqueda, 1997; Bouillon et al., 2009). CanAM5 and CanNEMO are coupled with CanCPL developed at CCCma
to facilitate communication between the two components.
CanAM5 is a spectral model with a T63 triangular truncation and 49 hybrid vertical coordinate levels extending from the
surface to 1hPa. Physical quantities are computed on the linear transform grid leading to a horizontal resolution of approxi-
mately 2.8 degrees. Improvements of CanAM5 upon its predecessor CanAM4 (von Salzen et al., 2013) include the addition
of 14 vertical levels in the upper troposphere and stratosphere, upgraded treatment of radiative processes, particularly in the
parameterization of albedo for bare soil, snow and ocean white-caps, improved aerosol optical properties, better optical prop-
erties for ice clouds and polluted liquid clouds, and a more comprehensive representation of land surface and lake processes.
CanNEMO is configured on the ORCA1 C-grid with 45 vertical levels with vertical spacing ranging from about 6 meters near the surface to about 250 meters in the abyssal ocean. The horizontal resolution is based on a 1 degree isotropic Mercator grid which is refined meridionally to 1/3 of a degree near the Equator, and includes a tripolar configuration to avoid the coordinate singularity in the Northern Hemisphere.

CanESM5 has the capability for an interactive carbon cycle and thus includes biogeochemistry modules to simulate land and ocean carbon exchange with the atmosphere. For the land surface, CLASS simulates energy, water, and momentum fluxes at the land-atmosphere boundary whereas CTEM simulates atmosphere-land fluxes of CO$_2$ and related terrestrial processes including photosynthesis, autotrophic and heterotrophic respiration, leaf phenology, carbon allocation, biomass turnover and conversion of biomass to structural attributes (Arora, 2003; Arora and Boer, 2003, 2005). This enables CTEM to simulate gross and net primary productivity over land while tracking the carbon flow through three living vegetation components (leaves, stem and roots) for nine plant functional types of prescribed fractional coverage (Melton and Arora, 2016), and two dead carbon pools (litter and soil). For the ocean, CMOC simulates carbon chemistry and abiotic chemical processes (such as solubility of oxygen, inorganic carbon, nutrients and other passive tracers having no feedback on biology and the simulated climate) in accordance with the CMIP6 Ocean Model Intercomparison Project (OMIP) biogeochemical protocol (OMIP-BGC; Orr et al., 2017). The biological module of CMOC is a simple Nutrient-Phytoplankton-Zooplanton-Detritus model, with fixed Redfield stoichiometry, and simple parameterizations of iron limitation, nitrogen fixation, and export flux of calcium carbonate. Both initialized and uninitialized predictions examined here, however, have prescribed atmospheric CO$_2$ concentrations and thus ocean and land CO$_2$, being purely diagnostic, do not feed back onto the simulated physical climate.

3 Forcing, initialization and ensemble generation

External forcing agents including historical anthropogenic and natural greenhouse gases, volcanic aerosols, solar activity and land use change are specified according to the CMIP6 protocol (Eyring et al., 2016). Emissions of sulfur dioxide (SO$_2$), dimethyl sulfide (DMS), black carbon, and organic carbon aerosol are specified, whereas mineral dust and sea salt emissions are simulated depending on local conditions. Concentrations of oxidants are specified for simulations of oxidation of sulfur species in clear air and in clouds. Direct effects of all types of aerosols, and 1st and 2nd indirect effects of sulfate, are simulated.

For each ensemble member, the coupled model is initialized from a separate assimilation run that ingests observation-based data from the ocean, atmosphere and sea ice as detailed below. These data-constrained assimilation runs span 1958 to present and are started from consecutive years following a 80-year spinup run that assimilates repeating 1958–1967 data. This leads to a spread of initial states that represent observational uncertainties. Full-field initialization is employed, and a standard lead time-dependent bias correction is applied in calculating forecast anomalies (Boer et al., 2016).

For the global ocean, 3D potential temperature and salinity are nudged toward values interpolated from monthly Ocean Reanalysis System 5 (ORAS5; Zuo et al., 2019) with a 10 day time constant in the upper 800m, and 1 year time constant at greater depths. The 1°S-1°N band is excluded partly to avoid disturbing strong equatorial currents below the surface (Carrassi et al., 2016). Sea surface temperature is relaxed to daily values interpolated from weekly values of the National Oceanic and
Atmospheric Administration (NOAA) Optimum Interpolation Sea Surface Temperature (OISST; Banzon et al., 2016) during November 1981 to present, or monthly values from the NOAA’s Extended Reconstructed Sea Surface Temperature (ERSSTv3; Xue et al., 2003; Smith et al., 2008) during 1958 to October 1981, with a 3 day time constant.

Sea ice concentration is relaxed to daily values interpolated from monthly values of the Hadley Centre Sea Ice and Sea Surface Temperature data set (HadISST.2; Titchner and Rayner, 2014) augmented with data from digitized Canadian Ice Service charts from 1958 to 2014 (Tivy et al., 2011), and to daily values from the Canadian Meteorological Centre (CMC) analysis from 2015 to present, with a 3 day time constant. Sea ice thickness is relaxed to daily-interpolated monthly values from the SMv3 statistical model of Dirkson et al. (2017) with a 3 day time constant. Before 1981, a repeating 1979–1988 monthly climatology is used.

Atmospheric temperature, horizontal wind components and specific humidity are nudged toward the European Centre for Medium-Range Forecasts (ECMWF) 6-hourly ERA-Interim (Dee et al., 2011) reanalysis values during 1979-present, or ERA40 (Uppala et al., 2005) anomalies added to ERA-Interim climatology during 1958-1978. The relaxation of the atmospheric variables is done with a 24 hour time constant (corresponding to inserting only 1/4 of analysis increment) and excludes spatial scales smaller than about 1000km. This results in ensemble spread that is comparable to root mean square differences between different atmospheric reanalyses (Merryfield et al., 2013).

Land physical and biogeochemical (BGC) variables are initialized through response of CLASS-CTEM to the data-constrained atmosphere. Land carbon pools are spun up during the 80-year spinup mentioned above. Oceanic BGC variables are initialized through response of CMOC to data-constrained physical ocean variables and surface atmospheric forcing.

4 Hindcasts evaluation methods

The evaluation approach and notation largely follows Boer et al. (2013, 2019a, b) and Sospedra-Alfonso and Boer (2020, hereafter SB20) where observations \(X\), ensemble of forecasts \(Y_k\), and ensemble of simulations \(U_k\) are annual or multi-year mean anomalies that are functions of time and location. The sub-index \(k = 1 \ldots m_Y\) or \(k = 1 \ldots m_U\) denotes ensemble member, where \(m_Y\) and \(m_U\) represent the ensemble size of forecasts and simulations, respectively. The anomalies are computed relative to climatological averages over a specified time period that is common to model output and observations. For the forecast and simulation ensembles, the anomalies are represented as

\[
Y_k = \Psi + y_k = \psi_f + \psi + y_k
\]

\[
U_k = \Phi + u_k = \phi_f + \phi + u_k
\]

consisting of predictable or “signal” components \((\Psi, \Phi)\) and unpredictable or “noise” components \((y_k, u_k)\). The predictable components are, in turn, comprised of externally forced \((\psi_f, \phi_f)\) and internally generated \(\psi\) variability. Even though the forecasts and simulations see the same external forcing, their forced components are not generally the same because initialization affects both the forced response and the internally generated variability. Unlike the forecasts, internal variability in the simulations is not constrained by initialization and is not predictable. The predictable components are common across the ensemble...
while the unpredictable components \((y_k, u_k)\) differ and average to zero over a large enough ensemble. The assumption is that all variables average to zero over the time period considered, forced components are independent of internally generated components, and all are independent of the noise components.

Ensemble averaging across forecasts and simulations in Eq. (1), denoted here by dropping the sub-index \(k\), leads to the following representation of the ensemble mean forecast and ensemble mean simulation

\[
Y' = \Psi + y = \psi_f + \psi + y \rightarrow \psi_f + \psi
\]
\[
U = \Phi + u = \phi_f + u \rightarrow \phi_f
\]

where here and elsewhere the arrows indicate the large ensemble limit. The variances of the ensembles of forecast and simulations in Eq. (1) are, respectively, \(\sigma^2_Y = \sigma^2_{\phi} + \sigma^2_{y}\) and \(\sigma^2_U = \sigma^2_{\phi} + \sigma^2_{u}\), while that of the ensemble means in Eq. (2) are \(\sigma^2_Y = \sigma^2_{\phi} + \sigma^2_{y}/m_Y\) and \(\sigma^2_U = \sigma^2_{\phi} + \sigma^2_{u}/m_U\), where \(\sigma^2_{\phi}\) and \(\sigma^2_{\phi}\) are the noise variances of forecasts and simulations. These variances are given explicitly in Eqs. (A5)-(A10) in appendix A.

SB20 decompose \(\Psi\) into mutually independent uninitialized \(Y_u\) and initialized \(Y_i\) components, with

\[
Y' = Y_u + Y_i + y
\]

where the component \(Y_u = \alpha \phi_f\) is ascribed to uninitialized external forcing (here \(\alpha\) is a measure of the projection of \(\psi_f\) on \(\phi_f\), while the initialized component \(Y_i = (\psi_f - \alpha \phi_f) + \psi\) includes the effect of initialization on both the forced component and the unforced internally generated component. The potentially predictable variance fraction (ppvf, Boer et al., 2013, 2019a, b) of the forecast ensemble and that of the ensemble mean forecasts are, respectively

\[
q_c = \frac{\sigma^2_{\phi}}{\sigma^2_{Y_u}} = \frac{\sigma^2_{\phi} + \sigma^2_{Y_i}}{\sigma^2_{\phi}} = q_{c_e} + q_e
\]
\[
q = \frac{\sigma^2_{\phi}}{\sigma^2_Y} = \frac{\sigma^2_{\phi}/m_Y + \sigma^2_{Y_i}/m_Y}{1 + \gamma_Y/m_Y} = \frac{1}{\sigma^2_{\phi}} = q_u + q_i \rightarrow 1
\]

where \(q\) is the noise-to-predictable variance ratio of the forecast ensemble, Eq. (A12). The ppvf of the simulations is defined in like manner. The ppvf’s \(q_e\) and \(q\) represent, respectively, the fractions of total and ensemble mean forecast variances that are potentially predictable. “Potential predictability” refers here to predictability within the “model world”, i.e., to predictability of a signal that is expected to represent variations of the observed climate system, but which may be unrealistic due to model and/or initialization errors. A potentially predictable signal is necessary but not sufficient for actual skill. The uninitialized and initialized contributions to \(q\), denoted \(q_u\) and \(q_i\), respectively, are computed explicitly according to Eqs. (A14)-(A15) in appendix A.

Following SB20, the correlation skill (or anomaly correlation coefficient) \(r_{XY}\) of the ensemble mean forecast can be decomposed as

\[
r_{XY} = \frac{r_{XY_u} \sigma_{Y_u}}{\sigma_Y} + r_{XY_i} \sigma_{Y_i}/\sigma_Y = r_u + r_i
\]

where \(r_{XY_u}\) and \(r_{XY_i}\) are the correlation skills of the uninitialized and initialized components \(Y_u\) and \(Y_i\) themselves, while \(r_u\) and \(r_i\) are the contributions to the overall correlation skill obtained by scaling with the fractions \(\sqrt{q_u}\) and \(\sqrt{q_i}\) representing the variances involved. This decomposition allows the assessment of the impact of initialization on correlation skill and
explicitly accounts for the effects of initialization on the model response to external forcing, through \( r_i \), and by excluding the comparatively strong contribution to variability by the trends, through \( r_{XY} \), which can obscure predictable internal variations. The latter avoids having to linearly detrend the data, which is frequently done and can introduce errors (SB20). The explicit computation of \( r_{XY} \) and \( r_{Xi} \) as well as \( r_u \) and \( r_i \) can be found in SB20 and are given in Eqs. (A16)-(A19) of the appendix for completeness.

The potential correlation skill of the forecasts is (Boer et al., 2013, 2019a, b)

\[
\rho = \frac{\sigma^2_Y}{\sigma_{Ye}^2} = \frac{\sigma^2_Y}{\sigma_Y} = \sqrt{\beta}q \rightarrow \sqrt{q_e} \tag{7}
\]

where \( \beta = \sigma^2_Y / \sigma_Y^2 < 1 \). The squared potential skill gives the fraction of the ensemble total variance that is represented or “explained” by the ensemble mean forecast, which in the large ensemble limit is the ppvf \( q_e \). The connection between the potential and actual correlation skill has been discussed by Eade et al. (2014); Smith et al. (2019); Strommen and Palmer (2019), where the following ratio is considered

\[
\Pi_r = \frac{r_{XY}}{\sqrt{\beta}} = \frac{r_{XY} \rho}{\sqrt{\beta}} = \frac{r_{XY}}{\sqrt{\beta}} \rho q \rightarrow \frac{r_{XY}}{\sqrt{q_e}} = r_{XY} \sqrt{1 + \gamma_Y} \tag{8}
\]

A similar quantity can be defined for the simulations. Assuming \( r_{XY} > 0 \), if \( r_{XY} / \sqrt{\beta} > 1 \) then \( r_{XY} > \rho \) since \( q < 1 \) and the actual correlation skill exceeds potential skill (Boer et al., 2019b), i.e., the model is more skilful at predicting the observations than its own behaviour. In the large ensemble limit, this is possible only if the noise-to-predictable variance ratio \( \gamma_Y \) is large enough to offset the correlation skill of the ensemble mean prediction, and can occur when the forecast predictable variance is much smaller than the observed variance. Such a behaviour is referred to as signal-to-noise paradox by Scaife and Smith (2018).

Evaluations of hindcasts actual skill also include computations of the mean square skill score

\[
\text{MSSS}(Y, R, X) = 1 - \frac{\text{MSE}(Y, X)}{\text{MSE}(R, X)}
\]

where MSE\((Y, X)\) and MSE\((R, X)\) are the mean square errors of, respectively, forecasts and reference predictions relative to observations (Goddard et al., 2013; Yeager et al., 2018). The reference predictions used here are the climatology of the observed anomalies \( \bar{X} = 0 \), persistence \( X_p \), and the uninitialized predictions, \( U \). For evaluations of \( N \)-year mean forecasts we use observed \( N \)-year rolling averages over the forecast initialization years. Persistence equals the most recent observed \( N \)-year average at the time of forecast initialization, and the uninitialized predictions are the \( N \)-year rolling averages of the ensemble mean simulations. We evaluate \( N \)-year averages of annual or seasonal mean anomalies including \( N = 1 \) for the second year of the forecast (Year 2) and \( N = 4 \) for forecast years 2 to 5 (Year 2-5) and 6-9 (Year 6-9) corresponding to forecast ranges beyond seasonal lead times. Anomalies are taken relative to identically sampled climatologies and predictions are bias corrected (but not trend corrected) following the recommendations of Boer et al. (2016). Annual averages are taken from January to December and seasonal averages are as specified in each case. Because hindcast initialization is done in late December, winter averages (DJF) predictions of, say Year 2, correspond to December of Year 1 forecasts and January and February of Year 2.

Statistical significance is evaluated using a non-parametric moving-block bootstrap approach (Goddard et al., 2013; Wilks, 1997) to generate the skill score’s sampling distribution based on 1000 repetitions. For every grid cell, skill scores are gen-
erated by resampling the data, with replacement, along the time dimension, and along the ensemble members dimension for hindcasts and simulations. Following Goddard et al. (2013), resampling of 5-year blocks are considered to account for temporal autocorrelation. The 5%- and 95%-quantile estimates of the bootstrapping distribution of the skill scores determine the 90% confidence interval. If the confidence interval does not include zero, the skill score is deemed statistically significant with 90% confidence and the associated grid cell is cross-hatched in the maps. The Fisher’s Z-transformation is applied to correlation skill scores before computing confidence intervals and its inverse is applied to the resulting quantiles.

5 Predictability and skill in the upper ocean

The time evolution of the upper-ocean heat content (OHC) is modulated by a wide range of low-frequency variability ranging from decadal and multidecadal to secular time scales (Levitus et al., 2005; Taguchi et al., 2017). The ocean’s lagged response to atmospheric thermal and dynamical forcing is due to the high specific heat capacity of water, which makes the upper ocean a major source of surface climate predictability on seasonal to decadal time scales (e.g., Smith et al., 2007; Meehl et al., 2014; Yeager and Robson, 2017). The top panels of Fig. 1 show the ppvf $q_e$, Eq. (4), of heat content in the upper 300 m of the ocean (OHC300) for forecast years 2, 2-5 and 6-9. For Year 2, $q_e > 0.7$ in most of the global ocean, implying that the OHC300 predictable variance accounts for more than 70% of the total variance. This contrasts with sectors of the tropical Pacific, the Indo-Pacific warm pool and in some coastal regions, where it can be lower than 30% (Fig. 1a). The relatively lower predictability and small effect of initialization in equatorial regions (Fig. 1a,d) may be partly associated with the 3D ocean initialization procedure that excludes the 1°S-1°N band (section 3), and to the fast wave processes on the equator. By contrast, initialization has a strong impact on $q_e$ in vast extratropical regions (Fig. 1d). For multi-year averages, regions of lower $q_e$ extend to the subtropics and to higher latitudes in the North Pacific, particularly for larger lead times (Fig. 1b,c). The impact of forecast initialization is widespread for Year 2-5 (Figs. 1e), but is much reduced for Year 6-9 (Figs. 1f) when most potentially predictable variance (Figs. 1c) is attributed to the simulated external forcing. A few notable exceptions showing a persistently high initialized potentially predictable variance include the North Atlantic, the Arctic and sectors of the Southern Ocean.

Some of the predictable variance contributes to skill, but some may reflect model biases and/or initialization errors. The correlation skill $r_{XY}$ of OHC300 hindcasts and the contribution $r_i$ from initialization are shown in the upper and lower panels of Fig. 2, respectively. For Year 2, correlation skill is significant over large portions of the global ocean (Fig. 2a), and is reduced in some extratropical regions including sectors of the eastern Pacific, the Arctic and Southern oceans, the Alaskan and western subarctic gyres, and in sectors of the North Atlantic most notably the western sub-polar region (WSPNA) and the Labrador Sea. The negative skill in the WSPNA region is attributed to initialization (Fig. 2d) and is partly a consequence of erroneous trends in ORAS5 reanalysis (Johnson et al., 2019) being imprinted on the hindcasts. The poor skill in the WSPNA and plausible causes are discussed further in section 6 below. Positive contributions from initialization can be seen in large sectors of the Pacific and Indian ocean basins for Year 2, whereas correlation skill in the Atlantic results mostly from uninitialized external forcing (Figs. 2a,d). For multi-year averages (Fig. 2b,c), the geographic extent of positive correlation skill is somewhat reduced relative to
Year 2, most notably over the equatorial Pacific and in the Indian basin at longer leads, and tends to increase in magnitude over regions where skill is attributed to the simulated external forcing.

Sea surface temperature (SST) forecasts for Year 2, 2-5 and 6-9 (Figs. 3a-c) show \( q_e > 0.4 \) in most of the global ocean, with larger values for the multi-year averages. For Year 2, notable contributions from initialization \( q_{ei} > 0.3 \) are seen in the western equatorial Pacific, which are not present for the multi-year averages. Excluding the Arctic region, the predictable SST signal is strongest in sectors of the Southern Ocean, and in the WSPNA and Labrador Sea regions, resulting entirely from initialization. These locations are characterized by unrealistic strong negative trends consistent with those from ORAS5 reanalysis (Fig. 4). On inter-annual time scales, \( q_e \) and \( q_{ei} \) are generally smaller for SST than OHC300 (Figs. 1a-c), as SST is more directly affected by atmospheric conditions. On multi-year time scales, \( q_e \) is generally larger for SST than OHC300 (Fig. 1d-f), as the simulated forced component becomes dominant, which strongly impacts SST trends (Fig. 4c).

SST forecasts show a reasonably widespread correlation skill (Fig. 5a-c). A large fraction of SST correlation skill is attributable to the uninitialized external forcing, but significant contributions \( r_i \) from initialization are seen in all ocean basins for Year 2 (Figs. 5d), and in sectors of the Pacific and Southern Ocean for the multi-year averages (Figs. 5e-f). The large contribution to skill by the uninitialized component derives partly from temperature trends that account for a larger variance fraction than that of the initialized component (Figs. 3), which is itself skillful as per \( r_{XY} \) (Figs. 5g-i). The apparent skill reemergence of the initialized component in the eastern Pacific for Year 6-9 is noteworthy (Fig. 5i).

MSSS of SST hindcasts relative to observed climatology (Figure 6a-c) indicate significant skill in large sectors of the North Atlantic, in the Indian Ocean, and in the western Pacific extending into the southeast extratropics. MSSS is a more stringent measure than correlation skill, and regions with significant correlation skill but near zero or negative MSSS indicate a misrepresentation of the observed variance for the given linear relationship between predictions and observations (i.e., due to a conditional bias). This is the case for various extratropical regions and sectors of the tropical Pacific (compare Figs. 5a-c and 6a-c). Regions with MSSS<<0 indicate a disproportionately large predicted variance relative to observations, such as in WSPNA and Labrador Sea, the Arctic, and in the Southern Ocean (Figs. 6a-c). These regions are characterized by high \( q_e \) (Figs. 3a-c) with a strong impact from initialization (Figs. 3d-f), but lack actual skill (Figs. 5a-c and 6a-c). For Year 2, SST hindcasts outperform persistence in most of the tropics (Fig. 6d), except for sectors of the subequatorial and western Pacific, and the western South Atlantic. For the multi-year averages (Fig. 6e,f), SST hindcasts beat persistence in large sectors of the Atlantic and Indian oceans, and in most western and southern portions of the Pacific within the 40°S–40°N latitude band. The hindcasts outperform simulations, particularly for multi-year averages (Figs. 6h,i), in vast subequatorial regions, in the Indian Ocean, and in northern and subpolar regions, but underperform in sectors of the Southern Ocean, the eastern and south Atlantic, and in the WSPNA and Labrador Sea regions.

**6 Erroneous SST hindcasts in the WSPNA and Labrador sea regions**

The negative correlation skill in the WSPNA and Labrador Sea for both upper-ocean heat content (Fig. 2) and SST (Fig. 5) is fully attributed to initialization indicating a mismatch between the hindcasts and simulations responses to external forcing in
these regions (SB20). CanESM5 ocean is initialized with ORAS5 (section 3), which has unrealistic temperature and salinity trends in the upper subpolar North Atlantic associated with erroneous water mass and heat transport before the 2000s (Jackson et al., 2019; Tietsche et al., 2020). The Labrador Sea in ORAS5 presents large changes in density anomaly, most notably in deep waters (1500-1900 m), which decrease abruptly from late 1990’s to early 2000’s leading also to unobserved trends (Fig. 9 of Jackson et al., 2019). These variations and unrealistic trends are imprinted on CanESM5 assimilation runs as they are nudged toward ORAS5 temperature and salinity fields to initialize the hindcasts (section 3). The anomalous heat and saline surface water transport into WSPNA is largely compensated in ORAS5 by a strong surface cooling provided by relaxation to observed SST (Tietsche et al., 2020), but such a cooling is not present in the forecasts, which leads to imbalances. This, combined with the model inherent biases in the region (S19) and resulting forecast drift, yield unrealistic decadal variations and long-term trends in the hindcasts themselves, which affect skill.

Figure 7a shows January-February-March (JFM) time series and linear trends of SST averages over WSPNA (40°N–60°N,50°W–30°W) for ERSSTv5 as verifying observations, ORAS5, and CanESM5 assimilation runs, Year 1 and Year 2 hindcasts, and the simulations. Analogous plots for averages over the Labrador Sea (55°N–65°N,60°W–45°W) are shown in Fig. 8. In WSPNA, ORAS5 is warmer than observations during the mid 1970’s to about 2000’s, as is the case for the assimilation runs. Year 1 hindcasts remain close to initial SST for most years although are somewhat colder during late 1990’s and onward. Year 2 SST hindcasts, on the other hand, have a strong warming in the early 1970’s and remain 2-3°C warmer than observations until late 1990’s, when a steep cooling occurs to below observed values until early 2000’s. These changes yield a negative trend for Year 2 hindcasts (-0.02 °C/decade) that does not match the slight warming trend from observations (0.01 °C/decade). By comparison, ORAS5 and the assimilation runs have virtually no trend, with values of 0.002 and 0.004 °C/decade, respectively. For longer lead times, the hindcasts drift toward simulations (not shown), which are characterized by a strong cold bias (-3.65 °C) and a warming trend (0.03°C/decade) described with some detail in S19 (see Fig. 15a,b and Fig. 26 of S19).

JFM time series of SST anomalies averaged over WSPNA are shown in Fig. 7b-f. The observed anomalies present distinctive decadal variations, with warm phases before 1970 and from late 1990s until about 2010, and a cold phase between 1970 and early 1990s. These decadal variations are modestly represented by ORAS5 ($r = 0.77$ and RMSE = 0.28°C, Fig. 7b), which has weaker and out of phase anomalies, and are better represented by the assimilation runs ($r = 0.94$ and RMSE = 0.16°C for the ensemble mean, Fig. 7c). Year 1 hindcasts perform modestly ($r = 0.43$ and RMSE = 0.42°C, Fig. 7d), and Year 2 hindcasts poorly, showing strong decadal variations that are mostly out of phase with observations ($r = -0.6$ and RMSE = 1.04°C, Fig. 7e). The anomalies of the simulations, which are characterized by a warming trend, are not expected to match the internally generated variability of the observed anomalies ($r = 0.42$ and RMSE = 0.46°C, Fig. 7f), although the latter are mostly contained within the ensemble spread. Analogous plots for the Labrador Sea (Fig. 8b-f) show disagreements also between ORAS5 and observed JFM SST anomalies, as well as for the assimilation runs, which are imprinted in the forecasts leading to unrealistic strong and out of phase variations for Year 2. The simulation ensemble present anomalies above −0.2°C during the whole time period, which are virtually unchanged in the mean at subzero temperatures prior the year 2000 as a result of excessive sea ice (S19). The poor skill in WSPNA and Labrador Sea can potentially impact predictions of surface climate...
over North America and Europe (Eade et al., 2012; Ruprich-Robert et al., 2017a), and West Africa and the Sahel (Martin and Thornicroft, 2014b; García-Serrano et al., 2015). Predictability of the tropical Atlantic (Dunstone et al., 2011), ocean heat content in the Nordic Seas, and decadal Arctic winter sea ice trends (Yeager and Robson, 2017) could also be affected.

4 Predictability and skill of surface climate on land

One of the main motivations for decadal climate prediction is the understanding that low-frequency variations in the upper-ocean heat content can influence surface climate by inducing atmospheric circulation changes both locally and remotely (Zhang and Delworth, 2006; Ruprich-Robert et al., 2017b). The expectation is that ocean model initialization will allow skilful surface climate prediction from seasons to years (Smith et al., 2007; Doblas-Reyes et al., 2011). Assessing the influence of model initialization on forecast skill can be challenging however (Smith et al., 2019), particularly in the presence of strong secular trends that can hinder the detection of internally generated predictable variations.

Figures 9a-c show $q_e$ of annual mean near-surface air temperature (SAT) on land for Year 2, Year 2-5 and Year 6-9 forecasts. For Year 2, the ppvf is generally largest in the tropics ($q_e > 0.4$), where atmospheric circulation is most strongly influenced by SST (Lindzen and Nigam, 1987; Smith et al., 2012). Tropical regions are impacted by initialization ($q_e > 0.1$), most notably in the Amazon, where $q_e \approx 0.2$ to 0.4 (Fig. 9d). Extratropical regions are characterized by a relatively higher atmospheric noise, thus displaying reduced ppvf ($q_e < 0.4$) and little contribution from initialization. For multi-year averages, the noise component is reduced considerably leading to a relatively high ppvf (Fig. 9b,c), particularly in regions where the warming trend is dominant (Fig. 10). The impact of initialization is also reduced, with typically $q_{ei} < 0.1$ (Fig. 9e,f).

SB20 show that annual and multiyear averages of CanESM5 SAT hindcasts have significant correlation skill over most land regions due to the strong temperature response to external forcing, with a modest contribution from initialization. In terms of MSSS, SAT hindcasts on land are more skilful than observed climatology (MSSS>0) in the tropics and in regions near to large water masses (Fig. 11a-c), mirroring the behaviour of $q_e$ in Fig. 9a-c. Skill is highest for multi-year averages, where significant MSSS values are also seen inland (Fig. 11b,c). Notably, MSSS<0 in the Amazon despite of positive correlation skill (Fig. 3 of SB20), indicating an excessive variance in the hindcasts possibly due to unrealistic trends (Fig. 10). The hindcasts outperform persistence in the extratropics, but underperform in the tropics most notably in Year 2 (Fig. 11d), primarily over regions of unrealistic trends (Fig. 10). By contrast, the hindcasts outperform simulations in the tropics (Fig. 11g) where initialization contributes to ppvf (Fig. 9d) and correlation skill (Fig. 3e of SB20), except for central Africa and the Sahel. For multi-year averages, the hindcasts outperform simulations in most regions (Fig. 11h,i) despite little impact from initialization to correlation skill (Fig. 3f of SB20), suggesting that the improvements are likely due to reductions of the simulated trends (Fig. 10).

The upper panels of Fig. 12 show $q_e$ of annual mean precipitation hindcasts for Year 2, Year 2-5 and Year 6-9. For Year 2, $q_e > 0.1$ is confined to tropical and subtropical regions, with slightly higher values in the Amazon basin ($0.2 < q_e < 0.3$). The precipitation signal extends to higher latitudes and is relatively stronger for multi-year averages. The largest ppvf values are seen in the Sahel for longer lead times ($q_e > 0.5$ for Year 6-9 in some locations) as the externally forced component becomes dominant. Generally, most of the hindcast precipitation signal with $q_e > 0.1$ is externally forced. The largest contribution
of initialization to ppvf are seen in northeast Brazil, central Southwest Asia, and southern Australia for Year 2 hindcasts (0.1 < $q_e < 0.2$), and $q_e < 0.1$ elsewhere (Fig. 12d-f). The negative values of $q_e$, seen in the plots are the result of a negligible impact from initialization and sampling errors.

The correlation skill of the annual mean precipitation hindcasts (Fig. 13a-c) partly mirrors the patterns of Fig. 12a-c, but can be significant also in regions of little ppvf. Correlation skill tends to increase both in magnitude and geographic extent for multi-year averages (Fig. 13b,c). A large component of skill is attributed to the uninitialized forced component, as can be inferred from Figs. 13d-f. Known sources of externally forced decadal precipitation variability include drivers of climate change such as CO$_2$ and anthropogenic SO$_2$, which can alter the energy budget due to changes in the atmospheric composition leading to climate feedback processes that affect precipitation (Myhre et al., 2017). Another major source is volcanic aerosols, which are injected into the stratosphere during a volcanic eruption and can reduce global mean temperature leading to a dryer atmosphere and reduced precipitation 2 to 3 years after an event (Smith et al., 2012). This is shown in Fig. 14 by the difference of mean precipitation from hindcasts with and without volcanic forcing following three major volcanic events (Agung, 1963; El Chichón, 1982 and Pinatubo, 1991). The setup of these volcanic experiments are described by Boer et al. (2016) as part of Component C of the DCPP. The difference in mean precipitation is significant over various land regions and most notably in the Maritime Continent. Precipitation forecasts over this region have significant correlation skill (Fig. 13a-c) not robustly associated to initialization (Fig. 13d-f), thus volcanic forcing seems to be a contributing factor. Contributions from initialization to correlation skill are significant in a few regions including northeast Brazil, western North America and central Southwest Asia for Year 2 and Year 2-5 (Figs. 13d,e), and are much reduced for Year 6-9 (Figs. 13f).

8 Skill dependence on ensemble size

Large single- and multi-model ensembles of initialized and uninitialized predictions have become essential tools in the study of decadal climate predictions due in part to the considerable noise reduction that can be achieved by ensemble averaging (Yeager et al., 2018; Deser et al., 2020; Smith et al., 2020). For a single-model ensemble, the ensemble size required to extract predictable signals varies among climatic variables, and may depend on forecast range, in the case of initialized predictions, and on geographic location. The tendency for models to underestimate predictable signals (Scaife and Smith, 2018; Smith et al., 2020) reinforces the need for large ensembles.

The noise-to-signal variance ratio of forecast and simulation ensembles have the form $\gamma_Y = \sigma^2_{ue}/\sigma^2_Y$, Eq. (A12), and $\gamma_U = \sigma^2_{ue}/\sigma^2_U$, Eq. (A13), respectively. Figure 15 plots $\gamma_Y$ and $\gamma_U$ of annual mean and multi-year terrestrial precipitation for the 40-member ensembles. The global land averages of $\gamma_Y$ and $\gamma_U$ as a function of ensemble size stabilize for $\gtrsim 10$ members (not shown), suggesting that the patterns of Fig. 15 are largely robust under changes in ensemble samples and size. In terms of $q = \sigma^2_{ue}/\sigma^2_Y$, the ensemble size required for $q > \alpha$ is, according to Eq. (5), $m_Y > \gamma_Y \alpha (1 - \alpha)^{-1}$, so $q > 0.9$ requires $m_Y > 9\gamma_Y$. Therefore, all regions in Fig. 15a,b with, say, $\gamma_Y > 5$ require $m_Y > 45$ members to satisfy $q > 0.9$, i.e., over 45 members are needed for the variance of the ensemble mean forecast to be at least 90% predictable. Most regions in Fig. 15a,b have $\gamma_Y > 5$. 

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A few exceptions include the Amazon basin for inter-annual time scales (Fig. 15a), and the Sahel for multi-year averages (Fig. 15b), which are both characterized by relatively strong precipitation signals. A similar analysis can be made for the simulations over NEB and CSW Asia, showing that forecast correlation skill is positive at the 90% confidence level. Unlike simulations, the forecasts over NEB are more skillful than persistence for all ensemble sizes indicating small variance contribution to skill from initialization that would have been underestimated for \( m_Y \lesssim 10 \) by this metric. Unlike simulations, the forecasts over NEB are more skillful than persistence for all ensemble sizes (Fig. 16c). By contrast, the median correlation skill over CSW Asia surpasses persistence for \( m_Y \gtrsim 20 \), but may require more than 40 members to do so confidently (Fig. 16d). It should be noted, however, that forecast correlation skill is higher when averaged over winter and spring (DJFMAM), and surpasses that of the simulations for \( m_Y \geq 20 \) (not shown). This is consistent with the seasonal cycle of mean precipitation over CSW Asia (Tippett et al., 2003; Schiemann et al., 2008), as the precipitation signal is stronger during DJFMAM. For Year 2-5 (Figs. 17c,d), the forecasts over NEB are more skillful than simulations for \( m_Y \gtrsim 20 \), but require \( m_Y \gtrsim 35 \) to marginally outperform simulations over CSW Asia, indicating an advantage of large ensembles.
The results over the Sahel are somewhat different. The Sahel is an important benchmark for the assessment of decadal predictions due to its strong summer rainy season, the variation of which is considered one of the largest signals of global climatic variability on annual to multi-year time scales (Martin and Thorncroft, 2014a; Sheen et al., 2017). Previous studies indicate that initialization enhances the skill of Sahelian rainfall predictions compared to simulations, although results vary among models (Garcia-Serrano et al., 2013; Gaetani and Mohino, 2013; Martin and Thorncroft, 2014a; Sheen et al., 2017; Yeager et al., 2018). Figures 18a,b show the dependence on ensemble size of Year 2 and Year 2-5 forecast correlation skill \( r_{XY} \) for July-August-September (JAS) mean precipitation averaged over the Sahelian sector (10°N-20°N, 20°W-10°E), as well as \( r_{XU} \) and \( r_{XX} \) for the simulations and persistence, respectively. Generally, forecasts and simulations outperform persistence by a large margin, but both exhibit about the same level of skill suggesting virtually no impact from initialization. The increase in skill is accompanied by a reduction in skill uncertainty, illustrating a benefit of large ensembles. The correlation skill decomposition indicates that the externally forced component is the main contributor to forecast skill with a negligible impact from initialization (Fig. 18c,d).

The small impact of initialization on Sahelian rainfall forecasts is at odds with previous findings (Gaetani and Mohino, 2013; Yeager et al., 2018). Inter-annual and multidecadal variability of Sahelian rainfall has been linked to SST variability in the global ocean (Rowell et al., 1995), the Atlantic (Ward, 1998; Knight et al., 2006; Zhang and Delworth, 2006; Ting et al., 2009; Martin and Thorncroft, 2014a, b; Yeager et al., 2018), the Pacific and Indian oceans (Mohino et al., 2011b; Sheen et al., 2017), and the Mediterranean sea (Rowell, 2003; Mohino et al., 2011a; Sheen et al., 2017). Greenhouse gases and aerosols have also been linked to decadal variability and trends of Sahelian rainfall by their impact on Atlantic inter-hemispheric SST gradients and resulting effect on the intertropical convergence zone (Biasutti and Giannini, 2006; Haywood et al., 2013; Hua et al., 2019; Bonfils et al., 2020), and by a direct effect of changes in radiative forcing (Haarsma et al., 2005; Biasutti, 2013; Dong and Sutton, 2015). Despite the negligible impact from initialization, CanESM5 precipitation skill over the Sahel is relatively high, particularly for Year 2-5 (\( r_{XY} \approx 0.7 \), Fig. 18b) and other multi-year averages (not shown), and comparable to the skill of CMIP5/6 decadal predictions from other models (Gaetani and Mohino, 2013; Yeager et al., 2018). This may be an indication that at least part of the enhanced forecast prediction skill of Sahelian rainfall in some CMIP5/6 models relative to that of simulations might be a consequence of the impact of initialization on the forced component rather than a skilful prediction of the internally generated variability itself (i.e., due to the term in parenthesis in the definition of \( Y_i \) in Eq. (3)).

The ratio \( \Pi_r \), Eq. (8), for the Sahelian JAS mean precipitation forecasts increases with ensemble size and confidently surpasses 1 with 40 members for Year 2 (Fig. 18e) and approximately 15 members for Year 2-5 (Fig. 18f), indicating that for larger ensembles \( r_{XY} > \rho \), i.e., the ensemble mean hindcasts is more skilful at predicting the observed climate system than the hindcasts themselves. This is a consequence of the noise-to-predictable variance fraction \( \gamma_Y \) being too high, suggesting that the hindcasts are either too noisy or its predictable component is too weak relative to the observed precipitation signal. Because the hindcasts and observed total JAS precipitation variances are about the same (not shown), we conclude that the ensemble mean hindcast underestimates the amplitude of the predictable precipitation signal in the Sahel. A similar behaviour is seen for the simulations with a somewhat reduced predictable variance fraction for large ensembles, particularly for Year 2-5 (Fig. 18f), which is primarily a consequence of the stronger potentially predictable variance of the simulations (not shown). Such a
behaviour is not specific to CanESM5, nor to the region, climate variable and time scales involved. It is a feature across model simulations of various climate phenomena (Scaife and Smith, 2018; Yeager et al., 2018; Smith et al., 2020), pointing to model deficiencies at representing the strength of predictable signals of the climate system.

4 Aspects of the skill of land and ocean biogeochemistry

CanESM5 models the interaction between the physical climate, the biosphere, and the chemical constituents of the atmosphere and ocean. This enables the assessment of some aspects of the predictability of ocean and land biogeochemistry, and the carbon cycle. Gross primary productivity (GPP) is the rate of photosynthetic carbon fixation by primary producers, such as phytoplankton in the ocean and plants on land. GPP of terrestrial vegetation is a key variable of the global carbon cycle and is an important component of climate change (Zhang et al., 2017). Net primary productivity (NPP) is the difference between GPP and the fraction of fixed carbon that primary producers use for respiration (Gough, 2011; Sigman and Hain, 2012), and is thus a major determinant of carbon sinks and a key regulator of ecological processes (Field et al., 1998). The predictive skill of NPP hindcasts in the ocean and GPP hindcasts on land is assessed here by correlation with the assimilation runs used for initialization. We also show preliminary comparisons with observation-based estimates, but do not provide a full assessment of actual skill due to the relatively short time span of the observations. As in previous sections, simulations are used here as a reference to quantify the impact of initialization on correlation skill. We emphasize that there is no assimilation of observed carbon cycle variables to initialize the hindcasts (section 3), therefore initial variations of GPP and NPP are the result of ensemble spread of oceanic and atmospheric states in the assimilation runs.

Figure 19 shows the correlation skill $r_{XY}$ and the contribution from initialization $r_i$ of ocean NPP for Year 2, Year 2-5 and Year 6-9 forecasts. For Year 2, there is significant correlation skill in most of the global ocean north of the Antarctic Circumpolar Current, except for scattered regions including WSPNA, the western North Pacific and, to some degree, in the eastern equatorial regions of the Pacific and Atlantic oceans. These regions are characterized by relatively low prediction skill of upper-ocean heat content (Fig. 2a). High NPP correlation skill is found in most ocean eastern boundaries and coastal upwelling regions, and in broader sectors associated with boundary currents including the North Pacific and California Currents, the Gulf Stream, North Atlantic and Canary Currents, the Brazil and Benguela Currents, the Agulhas Current, the East Australian Current, and in areas of the Arabian Sea and Bay of Bengal north of the Monsoon Drift. Part of this skill is attributed to initialization (Fig. 19d) with little or no impact from the simulated external forcing. Predictive skill tends to be larger in both magnitude and extent for Year 2-5 (Fig. 19b,e), and is much reduced for Year 6-9 (Fig. 19c,f) except for a few regions of relatively high skill including major eastern boundary upwelling systems (EBUS; Chan, 2019). EBUS comprise some of the ocean’s most productive regions supporting approximately one-fifth of the world’s ocean wild fish harvests (Pauly and Christensen, 1995) and the habitats for multiple species of pelagic fish, migratory seabirds, and marine mammals (Block et al., 2011), thus the potential for NPP skilful forecasts in these regions at relatively long lead times may have useful implications for fisheries and environmental managers. Preliminary comparisons against observation-based estimates over the Canary Current region (Fig. 21a), which along with the California, Humboldt and Benguela current systems is one of the four major EBUS (Gómez-Letona...
et al., 2017), show realistic inter-annual variations in the assimilation runs and Year 1 hindcasts. Ilyina et al. (2020) point out difficulties however in CanESM5 predictions of ocean CO₂ uptake in an inter-comparison of earth system model results.

On land, significant GPP correlation skill of annual and multi-year hindcasts is found on all continents (Fig. 20a-c), although negative skill can be seen mainly in grassland and savanna regions of South America, Africa and east Asia. Correlation skill is highest in the temperate zone of eastern North America, in South East Asia and the Maritime Continent, in sectors of tropical South America and Africa, in Southern Australia, and in North Africa, the Nile basin and Arabian Peninsula. Except for the latter, these regions are characterized by moderate to high annual mean primary productivity (Fig. 1 of Field et al., 1998). Unlike ocean NPP, a large portion of GPP skill on land derives from the simulated externally forced component, particularly from CO₂ fertilization, with a moderate but significant contribution from initialization (Fig. 20d,e). The effects of initialization become negligible for longer forecast ranges, except for a small sector of the Amazon rainforest (Fig. 20f).

Preliminary comparisons against observation-based products show realistic global mean GPP anomaly trends (not shown) and interannual variations for the assimilation runs and Year 1 hindcasts (Fig. 21b). This is consistent with multi-model comparisons (Ilyina et al., 2020) showing significant correlation skill of CO₂ land uptake in linearly detrended CanESM5 assimilation runs and hindcasts for up to 2 years. Comparisons against observation-based data however are limited by the relatively short time span and uncertainty of the observations.

10 Summary and conclusions

CanESM5 decadal hindcasts, which are CCCma’s contribution to Component A of the DCPP component of CMIP6, have the ability to represent realistic inter-annual and multi-year variations of key physical climate fields and carbon cycle variables on decadal time scales. The hindcasts are 40 ensemble members retrospective forecasts that are full-field initialized at the end of every year during 1960–2016 and run for 10 years. Natural and anthropogenic external forcing associated with greenhouse gases and aerosols are specified, and a 40-member ensemble of historical climate simulations with the same external forcing is also produced. The predictable component of the simulations is determined by the model’s response to external forcing, whereas the forecasts have predictable components due to both the initialization of internal climate states and to the model’s response to external forcing, which is generally different from that of simulations. The decomposition of the predictable component of the forecasts into initialized and uninitialized constituents, the latter derived from the projection of the forecasts responses to external forcing onto that of simulations, allowed the quantification of the impact of initialization on skill, and sheds new light on the value added by a forecasting system over that of climate simulations.

The upper-ocean heat content of CanESM5 is shown to be potentially predictable during the 10-year forecast range most notably in the extratropics, with potentially predictable variance in the eastern ocean boundaries for up to the 2- to 4-year range as a result of initialization. The hindcasts realize some of this potential predictability and have actual skill largely driven by external forcing, with significant contributions from initialization in the Pacific and Indian ocean basins. Sea surface temperature (SST) forecasts are skillful for most of the global ocean mainly due to the strong warming response in the model, with a moderate impact from initialization to correlation skill beyond the first year of the forecasts. Compared to heat content,
SST is more directly affected by atmospheric conditions reducing the contribution of initialization to skill. Initialization also decreases MSE significantly relative to that of simulations in the northern subtropics and in the Indian Ocean due to a reduction of the simulated warming trend, which highlights the impact of initialization not only on the predictability of internal climate variations, but also on corrections of the simulated response to external forcing.

The western subpolar North Atlantic (WSPNA) and the Labrador Sea regions stand out for the negative skill of the upper-ocean heat content and the surface temperature, resulting in part from erroneous temperature and salinity trends in the reanalysis data used to initialize the forecasts. Winter SST variations of CanESM5 hindcasts in these regions have strong decadal variations that are out of phase with observations beyond the 1-year range. Also, strong cold biases and warming trends in the simulations contribute to the poor performance in these regions. The lack of skill in the WSPNA and the Labrador Sea merit further analysis as it may impact climate predictability elsewhere.

The strong warming response of CanESM5 drives the potential predictability of near-surface air temperature over land, and is largely responsible for the forecast correlation skill as examined in SB20. Initialization, however, reduces the strength of the model response to external forcing leading to a lower forecast MSE than that of the simulations and persistence at all forecast ranges considered, except for some tropical regions. The correlation skill of annual and multi-year mean precipitation is, perhaps surprisingly, very high in vast continental regions including Siberia, central Southwest Asia, Northeast Europe, the Americas, and the Sahel. The precipitation skill is mainly driven by external forcing, with a non-negligible impact from volcanic aerosols, although long-lived effects from initialization can be seen in regions such as Northeastern Brazil and central Southwest Asia which can be influenced by remote SST anomalies. Skill tends to be highest for multi-year averages, as potentially erroneous inter-annual variability is averaged out and the forced component becomes dominant.

Two additions to CCCma’s contribution to the decadal prediction component of CMIP6 compared to CMIP5 are the increased ensemble size to 40 members from 10 members, and the inclusion of the carbon cycle in these experiments. There is a growing evidence that large ensemble sizes are advantageous for decadal predictions, and this work is consistent with that view. Skilful CanESM5 precipitation forecasts with a significant impact from initialization require large ensembles to confidently surpass the skill of simulations, compared to 10 or fewer members. There is however a limit to the cost-effective increase of ensemble size needed to improve skill, which is determined by the ensemble forecast noise-to-predictable variance ratio. Large ensembles are also used to show that CanESM5 decadal hindcasts underestimate the inter-annual and multi-year Sahelian summer rainfall signal, an important benchmark for the assessment of decadal predictions, as correlation skill is larger than potential correlation skill for sufficiently large ensembles despite the hindcasts having realistic total precipitation variance in this region. Initialized CanESM5 decadal hindcasts are skilful compared to assimilated values for predictions of net primary productivity in the ocean northward of the Antarctic Circumpolar Current for the 2- to 4-year range, with regions of long-lived skill encompassing the 10-year forecast range. A significant portion of this skill is attributed to initialization, particularly in major eastern boundary upwelling systems where there is indication of actual skill as well, and in the Bay of Bengal. On land, gross primary productivity hindcasts are potentially skillful at all ranges examined, mostly because of the CanESM5 response to the externally forced CO$_2$ increase, with a moderate but significant short-lived impact from initialization. Preliminary comparisons of CanESM5 assimilation runs and Year 1 forecasts with observation-based products have shown agreement in the
global mean anomaly trend and interannual variations for the years of available data. A comprehensive assessment of actual skill remains however a challenge due the relatively short time span and uncertainty of the verifying observations.

**Code and data availability.** The CanESM5 source code is publicly available at https://gitlab.com/cccma/canesm. The version of the code used to produce all the simulations described in this paper, which are submitted to CMIP6, is tagged as v5.0.3 with associated DOI: https://doi.org/10.5281/zenodo.3251113 (Swart et al., 2019a, b). The CanESM5 data for the decadal experiments (Sospedra-Alfonso et al., 2019a, b, c, d, e, f, g, h) and historical simulations (Swart et al., 2019c) are publicly available from the Earth System Grid Federation (https://esgf-node.llnl.gov/search/cmip6/). The observation-based products used here are freely available. Further details are given in appendix B.

**Author contributions.** RSA initiated the study, helped to coordinate and performed model simulations, produced the analysis and figures, and wrote the paper. WJM led the experimental design and development of the CCCma contribution to DCPP, and contributed with the analysis and interpretation of results. GJB was instrumental to the CCCma contribution to DCPP, and contributed with the methodology, analysis and interpretation of results. VVK and WL setup and produced the bulk of model simulations. CS contributed with results, plotting, and assessment of GPP on land. JRC contributed with results and assessment of NPP in the ocean. All authors contributed with the manuscript and the final version of the paper.

**Competing interests.** The authors declare that they have no conflict of interests.

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Appendix A: Details of hindcast evaluation methods

A1 Associated variances

The variances associated to the ensembles of forecasts and simulations in Eq. (1), together with those of the ensemble mean of forecasts and simulations in Eq. (2), are

\[ \sigma^2_{Y e} = \sigma^2_{\Psi} + \sigma^2_{y e} \]
\[ \sigma^2_{U e} = \sigma^2_{\Phi} + \sigma^2_{u e} \]
\[ \sigma^2_Y = \sigma^2_{\Psi} + \sigma^2_{y} + \frac{\sigma^2_{y e}}{m_Y} = \sigma^2_{\Psi e} + \sigma^2_{\Psi i} \]
\[ \sigma^2_U = \sigma^2_{\Phi} + \sigma^2_{u} + \frac{\sigma^2_{ue}}{m_U} \rightarrow \sigma^2_{\Phi e} \]

where we have used \( \sigma^2_{y e} = \frac{\sigma^2_y}{m_Y} \) under the assumption of the \( y_k \)'s independence, and similarly for \( u_k \). The noise variances are estimated from Eqs. (A1)-(A4) as

\[ \sigma^2_{Ye} = \frac{m_Y}{m_Y - 1} (\sigma^2_{Ye} - \sigma^2_Y) \]
\[ \sigma^2_{ue} = \frac{m_U}{m_U - 1} (\sigma^2_{ue} - \sigma^2_U) \]

whereas the predictable variances are estimated from Eqs. (A3)-(A6) as

\[ \sigma^2_\Psi = \sigma^2_Y - \frac{\sigma^2_{Ye}}{m_Y} \]
\[ \sigma^2_\Phi = \sigma^2_U - \frac{\sigma^2_{ue}}{m_U} \]

The predictable and noise variances can be readily computed from the data by means of the total variance \( \sigma^2_{Ye} \) or \( \sigma^2_{Ue} \), and the variance of the ensemble mean \( \sigma^2_Y \) or \( \sigma^2_U \). If we write explicitly the dependence of the anomaly forecast \( Y_{jk}(\tau) \) and ensemble mean \( Y_j(\tau) \) on the forecast range \( \tau \), ensemble member \( k = 1 \ldots m_Y \), and initial year \( j = 1 \ldots n \), then

\[ \sigma^2_{Ye}(\tau) = \frac{1}{m_Y (n - 1)} \sum_{j=1}^n \sum_{k=1}^{m_Y} [Y_{jk}(\tau) - Y_k(\tau)]^2 \]
\[ \sigma^2_Y(\tau) = \frac{1}{(n - 1)} \sum_{j=1}^n [Y_j(\tau) - \overline{Y(\tau)}]^2 \]

and similarly for the simulations, where the over line indicates the average over the initial years.

A2 Correlation skill decomposition

Following SB20, the correlation skill of the ensemble mean forecast can be decomposed as

\[ r_{XY} = r_{XY u} \frac{\sigma_{Ye}}{\sigma_Y} + r_{XY i} \frac{\sigma_{Ye}}{\sigma_Y} = r_u + r_i \]
where \( r_{XY_u} \) and \( r_{XY_i} \) are the correlation skills of the uninitialized and initialized components \( Y_u \) and \( Y_i \) themselves, while \( r_u \) and \( r_i \) are the components contribution when scaled by the fractions of the variances involved. In terms of the noise-to-predictable variance ratios of forecasts and simulation,

\[
\gamma_Y = \frac{\sigma_Y^2}{\sigma_Y^2} = \frac{m_Y (\sigma_Y^2 - \sigma_Y^2)}{m_Y \sigma_Y^2 - \sigma_Y^2}, \quad (A12)
\]

\[
\gamma_U = \frac{\sigma_U^2}{\sigma_U^2} = \frac{m_U (\sigma_U^2 - \sigma_U^2)}{m_U \sigma_U^2 - \sigma_U^2}, \quad (A13)
\]

and available correlations and variances, these quantities can be computed explicitly as

\[
\frac{\sigma_Y^2}{\sigma_Y^2} = \theta r_{YU}^2 \left( 1 + \gamma_Y \right) \rightarrow \theta r_{YU}^2, \quad (A14)
\]

\[
\frac{\sigma_U^2}{\sigma_Y^2} = \frac{\gamma_U}{m_U} \left( 1 + \gamma_Y \right) + \left( 1 + \frac{\gamma_Y}{m_Y} \right)^{-1} \rightarrow 1 - \theta r_{YU}^2, \quad (A15)
\]

\[
r_u = \theta r_{XU} r_{YU} \left( 1 + \frac{\gamma_U}{m_U} \right) \rightarrow \theta r_{XU} r_{YU}, \quad (A16)
\]

\[
r_i = r_{XY} - \theta r_{XU} r_{YU} \left( 1 + \frac{\gamma_U}{m_U} \right) \rightarrow r_{XY} - \theta r_{XU} r_{YU}, \quad (A17)
\]

\[
r_{XY_u} = \theta r_{XU} \left( 1 + \frac{\gamma_U}{m_U} \right)^{1/2} \rightarrow \theta r_{XU}, \quad (A18)
\]

\[
r_{XY_i} = \frac{\sigma_Y^2}{\sigma_Y^2} \left[ r_{XY} - \theta r_{XU} r_{YU} \left( 1 + \frac{\gamma_U}{m_U} \right) \right] \rightarrow (1 - \theta r_{YU}^2)^{-1/2} [r_{XY} - \theta r_{XU} r_{YU}] \quad (A19)
\]

where \( r_{YU} \) denotes the correlation between the ensemble means of forecasts and simulations, and the step function \( \theta = 0 \) if \( r_{YU} < 0 \), else \( \theta = 1 \). The ratios \( \gamma_Y \) and \( \gamma_U \) are estimated according to Eq. (A12) and Eq. (A13) with the total variances \( \sigma_Y^2 \) and \( \sigma_U^2 \), Eq. (A9), and the ensemble mean variances \( \sigma_Y^2 \) and \( \sigma_U^2 \), Eq. (A10), for simulations and forecasts. For finite ensembles, \( \sigma_U^2 \) and \( \sigma_Y^2 \) and thus \( \gamma_U \) and \( \gamma_Y \) can be negative due to sampling errors. With Eqs. (A12)-(A13), the quantities in Eqs. (A16)-(A19) are readily computed from the data.
Appendix B: Data sources, variables, and derived quantities
Table B1. List of figures, CMIP6 variables, experiments, and verifying observation-based products employed in this paper. See table B2 for the sources of the observation-based products. The entry "n/a" indicates "not applicable".

<table>
<thead>
<tr>
<th>Figure</th>
<th>CMIP6 variable</th>
<th>CMIP6 experiment and variant label</th>
<th>Observation-based product</th>
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<td>---------------------------</td>
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<td>EN4.2.1 Met Office Hadley Centre (Good et al., 2013)</td>
<td><a href="https://www.metoffice.gov.uk/hadobs/en4/download-en4-2-1.html">https://www.metoffice.gov.uk/hadobs/en4/download-en4-2-1.html</a></td>
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<td>ERSSTv5 US National Oceanic and Atmospheric Administration (NOAA)</td>
<td>Extended Reconstructed Sea Surface Temperature (Huang et al., 2017)</td>
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<td>ERA40, ERA-Interim European Centre for Medium-Range Forecasts (ERA40; Uppala et al., 2005) and (ERA-Interim; Dee et al., 2011)</td>
<td><a href="https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets">https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets</a></td>
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<td>GPCP2.3 Global Precipitation Climatology Project (Adler et al., 2003)</td>
<td><a href="https://www.esrl.noaa.gov/psd/gridded/data.gpcp.html">https://www.esrl.noaa.gov/psd/gridded/data.gpcp.html</a></td>
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<td>MODIS NASA's MODIS-based gross primary productivity product (Zhang et al., 2017)</td>
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<td>GOSIF Orbiting Carbon Observatory-2 (OCO-2)-based Solar-induced chlorophyll fluorescence (SIF) product (Li and Xiao, 2019)</td>
<td><a href="http://data.globalecology.unh.edu/data/GOSIF-GPP_v2">http://data.globalecology.unh.edu/data/GOSIF-GPP_v2</a></td>
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<td>ERA5 Hersbach and coauthors (2020). Copernicus Climate Change Service (C3S)</td>
<td><a href="https://cds.climate.copernicus.eu/cdsapp#!/home">https://cds.climate.copernicus.eu/cdsapp#!/home</a></td>
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<td>SSI Surface solar irradiance (Bishop et al., 1997)</td>
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<td>SeaWiFS NASA Goddard Space Flight Center</td>
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<td>MODIS-Terra NASA Goddard Space Flight Center</td>
<td>Moderate-resolution Imaging Spectroradiometer Terra Chlorophyll Data; reprocessing version 2010.0</td>
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</table>
1 References


15 Boer, G. J., Kharin, V. V., and Merryfield, W. J.: Differences in potential and actual skill in a decadal prediction experiment, Climate Dynamics, 2019a.


Figure 1. Potential predictability of CanESM5 forecasts of annual and multi-year mean heat content above 300m. (a-c) Potentially predictable variance fraction $q_e = \frac{\sigma^2_\psi}{\sigma^2_Y}$, Eq. (4), and (d-f) the portion $q_e = \frac{\sigma^2_Y}{\sigma^2_{Ye}}$ attributed to initialization for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9.
Figure 2. Skill of CanESM5 forecasts of annual and multi-year mean heat content above 300m. (a-c) Correlation skill $r_{XY}$, Eq. (6), and (d-f) contribution from initialization $r_i$ to correlation skill, Eq. (A17), for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9. The verifying observations are derived from EN4.2.1 dataset (appendix B). Cross-hatched regions indicate values significantly different from zero at the 90% confidence level.
Figure 3. Potential predictability of CanESM5 annual mean sea surface temperature forecasts. (a-c) Potentially predictable variance fraction $q_e = \frac{\sigma^2_{\Psi}}{\sigma^2_{Ye}}$, Eq. (4), and (d-f) the portion $q_{ei} = \frac{\sigma^2_{\Psi_i}}{\sigma^2_{Ye}}$ attributed to initialization for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9.
Figure 4. Linear trends of annual mean sea surface temperature for (clockwise from upper-left) observations, ORAS5, Year 2 forecasts, and historical simulations. The verifying observations are from ERSSTv5 dataset (appendix B).
Figure 5. Skill of CanESM5 annual and multi-year mean sea surface temperature forecasts. (a-c) Correlation skill $r_{XY}$, Eq.(6), (d-f) contribution from initialization $r_i$ to correlation skill, Eq. (A17), and (g-i) correlation skill of the initialized component of the forecast $r_{XYi}$, Eq. (A19), for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9. The verifying observations are from ERSSTv5 dataset (appendix B). Cross-hatched regions indicate values significantly different from zero at the 90% confidence level.
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Figure 7. JFM time series of (a) SST and (b-f) SST anomalies corresponding to (black) ERSSTv5, (cyan) ORAS5, and CanESM5 (green) assimilation runs, (purple) Year 1 and (blue) Year 2 forecasts, and (red) historical simulations, averaged over the WSPNA region (40°N-60°N, 50°W-30°W). Boxes and whiskers indicate the minimum, maximum, 25- and 75-percentiles of the 40-member CanESM5 ensemble of forecasts and simulations, and the first 10-member ensemble for the assimilation runs. Model values in (a) correspond to ensemble means and dashed lines represent linear trends. Trends are not removed from the anomalies in (b-f).
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Figure 9. Potential predictability of CanESM5 forecasts of annual and multi-year mean near-surface air temperature forecasts. (a-c) Potentially predictable variance fraction \( q_e = \sigma^2_{\Psi} / \sigma^2_{Ye} \), Eq. (4), and (d-f) the portion \( q_e_i = \sigma^2_{Y_i} / \sigma^2_{Ye} \) attributed to initialization for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9.
Figure 10. Decadal trends of annual mean near-surface air temperature for (clockwise from upper-left) observations, Year 1 and Year 2 forecasts, and historical simulations. The verifying observations are from ERA-40 and ERA-Interim datasets (appendix B).
Figure 11. Skill of CanESM5 annual and multi-year mean near-surface air temperature forecasts. MSSS of (left) Year 2, (middle) Year 2-5 and (right) Year 6-9 forecasts, $Y$, relative to (a-c) observed climatology, $X$, (d-f) persistence forecast, $X_p$, and (g-i) historical simulations, $U$. The verifying observations are from ERA-40 and ERA-Interim datasets (appendix B). Cross-hatched regions indicate values significantly different from zero at the 90% confidence level.
Figure 12. Potential predictability of CanESM5 forecasts of annual and multi-year precipitation. (a-c) Potentially predictable variance fraction $q_e = \frac{\sigma^2_w}{\sigma^2_Y_e}$, Eq. (4), and (d-f) the portion $q_{ei} = \frac{\sigma^2_{ei}}{\sigma^2_Y_e}$ attributed to initialization for forecast (left) Year 2, (center) Year 2-5 and (right) Year 6-9.
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Figure 14. The impact of volcanic aerosols on mean precipitation forecasts. Difference of (a) summer (July-August-September), (b) winter (December-January-February), and (c) annual mean Year 2-4 precipitation forecasts with and without volcanic eruptions. Computations include 10-member ensembles of forecasts with and without eruptions of Mount Agung, El Chichón and Mount Pinatubo eruptions as per DCPP component C volcanic experiment setup (Boer et al., 2016). Cross-hatched regions indicate values significantly different from zero at the 90% confidence level.
Figure 15. Maps of noise-to-predictable variance ratio (a,b) $\gamma_Y$, Eq. (A12), for (a) Year 2 and (b) Year 2-5 forecast and (c,d) $\gamma_U$, Eq. (A13), for (c) 1-year and (d) 4-year averaged simulations of annual mean precipitation, produced with the 40-member ensembles of forecasts and simulations. The $\gamma_Y$ ratio determines the ensemble size required to average out the noise component from the ensemble mean forecast, Eq. (5), and similarly $\gamma_U$ for simulations. Negative values (white on land) result from sample errors, indicating small ensemble mean variance, and therefore small predictable signal, relative to the noise variance. Rectangular boxes indicate the regions studied in Figs. 16-18 below.
Figure 16. Dependence on ensemble size of (a,b) variance contributions $q_u = \sigma^2_{Yu} / \sigma^2_Y$, Eq. (A14), in blue, and $q_i = \sigma^2_{Y_i} / \sigma^2_Y$, Eq. (A15), in green, to (c,d) correlation skill $r_{XY}$, Eq. (6), in red, of Year 2 ensemble mean precipitation forecasts, averaged over (a,c) northeast Brazil (-10°N-5°N, 50°W-35°W) and (b,d) central Southwest Asia (25°N-55°N, 40°E-75°E). These regions are highlighted in Fig. 15 above. Thick black curves indicate correlation skill $r_{XU}$ of ensemble mean simulations. Thick dashed lines indicate correlation skill $r_{XXP}$ of the persistence forecast. Thin curves are confidence intervals derived from the 5th- and 95th-percentile of bootstrapping distributions generated from 10000 samples by random selection, with replacement, of ensemble members for each indicated ensemble size. Filled dots correspond to the actual 40-member ensemble predictions. Computations of $q_u$, Eq. (A14), and $q_i$, Eq. (A15), are done with $m_Y = 2 \ldots 40$ members from the forecasts ensemble and, for each $m_Y$, the 40 members from the simulations ensemble. The verifying observations used to compute correlation skill are from GPCP2.3 dataset (appendix B).
Figure 17. As in Fig. 16 for forecast year 2–5.
Figure 18. Dependence on ensemble size of (a,b) correlation skill of ensemble mean forecasts $r_{XY}$ (red) and ensemble mean simulations $r_{XU}$ (black); (c,d) contributions $r_u$, Eq. (A16), in royal blue, and $r_i$, Eq. (A17), in olive, to $r_{XY}$; and (e,f) ratio $\Pi_r$, Eq. (8), of forecasts (salmon) and simulations (gray), for (a,c,e) Year 2 and (b,d,f) Year 2-5 precipitation forecasts, averaged over the Sahel (10°N-20°N, 20°W-10°E). This region is highlighted in Fig. 15 above. Thick dashed lines indicate correlation skill $r_{XXp}$ of the persistence forecast. Thin curves are confidence intervals derived from the 5th- and 95th-percentile of bootstrapping distributions generated from 10000 samples by random selection, with replacement, of ensemble members for each indicated ensemble size. Filled dots correspond to the actual 40-member ensemble predictions. Computations of $r_u$, Eq. (A16), and $r_i$, Eq. (A17), are done with $m_Y = 2 \ldots 40$ members from the forecasts ensemble and, for each $m_Y$, the 40 members from the simulations ensemble. The verifying observations used to compute correlation skill are from GPCP2.3 dataset (appendix B).
Figure 19. Skill of CanESM5 annual and multi-year mean of ocean NPP forecasts. (a-c) Correlation skill $r_{XY}$, Eq. (6), with the assimilation runs as verifying observations, and (d-f) contribution from initialization $r_i$ to correlation skill, Eq. (A17), for (left) Year 2, (center) Year 2-5 and (right) Year 6-9 forecasts. The CanESM5 assimilation runs used as verifying observations provide the initial conditions of the hindcasts. Cross-hatched regions indicate values significantly different from zero at the 90% confidence level.
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Figure 21. (a) Ocean integrated net primary productivity in the Canary Current region (10-18°W, 25-34°N), and (b) gross primary productivity on the global land, for the assimilation runs (blue) and Year 1 hindcasts (red). Observation-based estimates for (a) ocean, VGPM (green), and (b) land, MODIS (yellow) and GOSIF (green), are described in appendix B. Anomalies relative to the base period 2000-2016 have been linearly detrended and standardized.