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Validation of 3D-CMCC Forest Ecosystem Model (v.5.1) against eddy covariance data for ten European forest sites

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Abstract

This study evaluates the performances of the new version (v.5.1) of 3D-CMCC Forest Ecosystem Model (FEM) in simulating gross primary production (GPP), against eddy covariance GPP data for ten FLUXNET forest sites across Europe. A new carbon allocation module, coupled with new both phenological and autotrophic respiration schemes, was implemented in this new version. Model ability in reproducing timing and magnitude of daily and monthly GPP fluctuations is validated at intra-annual and inter-annual scale, including extreme anomalous seasons. With the purpose to test the 3D-CMCC FEM applicability over Europe without a site-related calibration, the model has been deliberately parameterized with a single set of species-specific parameterizations for each forest ecosystem. The model consistently reproduces both in timing and in magnitude daily and monthly GPP variability across all sites, with the exception of the two Mediterranean sites. We find that 3D-CMCC FEM tends to better simulate the timing of inter-annual anomalies than their magnitude within measurements uncertainty. In six of eight sites where data were available the model well reproduces the 2003 summer drought event. Finally, for three sites we evaluate if a more accurate representation of forest structural characteristics (i.e. cohorts, forest layers) and species composition can improve model results. In two of the three sites results reveal that model slightly increases its performances, although, statistically speaking, not in a relevant way.

1 Introduction

Terrestrial ecosystems have a relevant role in the global carbon cycle, acting also as climate regulators (Peters et al., 2007; Bonan, 2008; Huntingford et al., 2009). In fact terrestrial ecosystems store large carbon stocks and cause most of the variance of carbon exchange between the atmosphere and land surfaces (Batlle Bayer et al., 2012). Among terrestrial ecosystems, forests are an essential component in the global carbon
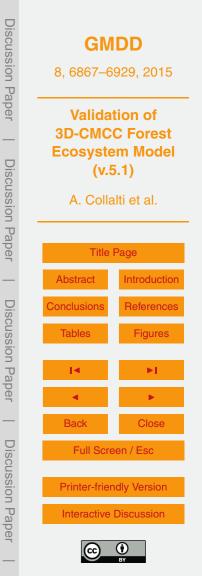
cycle because of their high capacity to store carbon in the vegetation and soil pools (Kramer et al., 2002). Through Gross Primary Production (GPP) plants fix atmospheric carbon dioxide (CO_2) as organic compounds, enabling terrestrial ecosystems to off-set part of the anthropogenic CO_2 emissions (Janssens et al., 2003; Cox and Jones, 2008; Battin et al., 2009). Consequently, changes in GPP could have relevant impacts on atmospheric CO_2 concentration. Thus, accurately simulating terrestrial GPP is key to quantifying the global carbon cycle and predicting the future trajectories of the atmospheric CO_2 concentration (Wu et al., 2015), and taking into account the various spatial and temporal scales of the processes is a major challenge (Yuan et al., 2007).

Terrestrial ecosystem models, used to simulate carbon, water and energy fluxes, are valuable tools for advancing the knowledge of the role of ecosystems in maintaining a multitude of their fundamental services, like the provision of products and the regulation of climate (Ibrom et al., 2006). Such numerical models are also useful to: (1) predict the impacts of climate variability on terrestrial biosphere and related carbon fluxes (Ciais et al., 2005; Brèda et al., 2006; Richardson et al., 2007), ranging from long term anomalies (Santini et al., 2014) up to extreme events (Zscheischler et al., 2014); and (2) reproduce biophysical and biogeochemical feedbacks of vegetation cover and change on climate, especially when coupled to atmosphere–ocean climate models through land surface schemes (Bonan, 2008; Arneth et al., 2012; Taylor et al., 2012).

At European level, terrestrial ecosystems have been reported to be a significant sink of CO_2 (Luyssaert et al., 2012), with forests playing a relevant role in absorbing anthropogenic emissions for about 10 % (Nabuurs et al., 2003; FOREST EUROPE et al., 2011).

In the last decades some studies have identified systematic errors when modelling terrestrial ecosystem sensitivity to climate variability at multiple time scales (Friedlingstein et al., 2006; Piao et al., 2013; Dalmonech et al., 2015) while sometimes differences in model predictions are stubbornly large (F. Wang et al., 2014).

To improve the models capacity in reproducing relevant processes related to the land carbon cycle, detailed representation of missing processes should be increasingly de-
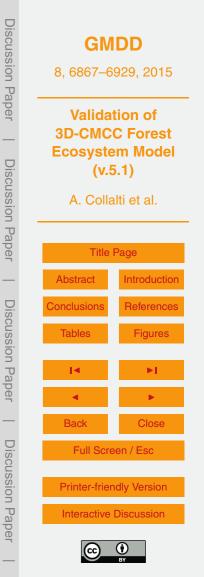
veloped (Sykes et al., 2001; Campioli et al., 2013; Nolè et al., 2013; Ciais et al., 2013; Prentice et al., 2015). For instance, spatial and temporal environmental hetereogeneity is known to play an important role in the dynamics of populations and communities (Kobe, 1996; Chesson, 2000; Clark et al., 2010, 2011). However, the implications of this hetereogeneity for developing and testing regional to global scale forest dynamics models that are also able to take into account forest management are largely unexplored (Zhang et al., 2014). As reported by Wramneby et al. (2008), incorporating increased mechanistic details is expected to improve the explanatory power of a model. Many models for example calculate leaf photosynthesis through the Farquhar model (Farquhar et al., 1980; Farquhar and Sharkey, 1982), while few models take in proper consideration vertical stratification. Increasing model complexity can sometimes mask a lack of understanding, although models including a larger subset of important processes should be more realistic than a simpler model. However, complex models are tuned to perform well at standard tests but produce widely divergent results when projected beyond the domain of calibration (Prentice et al., 2015). Since European forests are mostly managed and not homogeneous in terms of structure, composition and cohorts, only a few models are able to represent this particular ecosystem complexity and heterogeneity (Grote et al., 2011; Morales et al., 2005; Seidl et al., 2012; Yin et al., 2014). For simulating the impact of forest management on carbon cycle, it is important to consider the vertical structure of forests and the age-related changes in structure and physiology.

In this study we investigate the performance of the new version of the 3D-CMCC Forest Ecosystem Model (FEM, Collalti et al., 2014) in quantifying GPP across different forest types and climate conditions in Europe. In contrast to Dynamic Global Vegetation Models (DGVMs) 3D-CMCC FEM incorporates accurate processes description focusing on the effects of hierarchy in vertical forest structure and ages on productivity and growth at species level. The model has been designed to maintain computational efficiency, as postulated for the Light Use Efficiency (LUE) Models (Monteith and Moss, 1977), coupled to the accuracy of the Process-Based Models (PBMs) (Makela et al.,

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2000). As described by F. Wang et al. (2014) and H. Wang et al. (2014), a model with both high accuracy and computation efficiency is highly desirable for the purpose of simulating long time series of GPP at high spatial resolution.

Thanks to FLUXNET, a global network of flux tower sites, half hourly net CO_2, water and energy eddy covariance (EC) flux measurements (Baldocchi, 2003) are now available for a wide range of forest ecosystems. The network provides a continuously increasing set annual series of half-hourly data (Balzarolo et al., 2014). These data provide valuable information to investigate seasonal phasing and amplitudes of carbon fluxes (Aubinet et al., 2000; Falge et al., 2002; Gielen et al., 2013; Slevin et al., 2015) and to test terrestrial models at the ecosystem scale (e.g. Richardson et al., 2010; Blyth et al., 2011; Chang et al., 2013; Wißkirchen et al., 2013; Bagnara et al., 2014; Balzarolo et al., 2014; Liu et al., 2014; F. Wang et al., 2014; Wu et al., 2015). In the present paper daily meteorological and GPP data are provided by the European Cluster of FLUXNET. GPP data are exploited as independent dataset to compare, over different time-scales, 3D-CMCC FEM simulations for ten European forest stands varying in species composition, forest structure, cohorts and climates.

The objective of this work is to answer to the following questions:

1. Does the model reproduce the magnitude and the timing of seasonal fluctuations in GPP across different forest types and forest canopy structures?

2. Does the model reproduce the observed inter-annual GPP variability?

3. Is the model generic enough so that a single set of species-specific parameterizations (i.e. without a site-related calibration) allows reproducing GPP behaviour across different biomes?

4. Do the model outputs improve when considering a complex heterogeneous three-dimensional canopy structure compared to a simple "big leaf" model canopy representation?

To investigate these issues, we introduced a 3-D canopy representation into the 3D-CMCC FEM, while otherwise maintaining its flexibility and the generic features to be applied to different forest ecosystems. The new model can now run on a daily time step and includes an improved allocation-phenology scheme (with the non-structural carbon pool, NSC), and an improved computation of autotrophic respiration.

2 Materials and methods

2.1 Model description

The three-dimensional Forest Ecosystem Model, 3D-CMCC FEM (Collalti, 2011; Collalti et al., 2014) (executable is available upon request at http://www.cmcc.it/models/3d-cmcc-fem-three-dimension-forest-ecosystem-model) is hybrid between an empirical and a process-based model relying on the concepts of the LUE approach at canopy level for carbon fixation (see Appendix A for a detailed description of algorithms). The 3D-CMCC FEM is designed to simulate at hectare scale and on a daily time step tree growth as well as carbon and water fluxes, at species level, representing ecophysiological processes in hetereogeneous forest ecosystems including complex canopy structures. The 3D-CMCC FEM uses daily meteorological data, site-specific data and ecophysiological data (e.g. maximum canopy conductance, specific leaf area, etc.; see Collalti et al., 2014) to simulate forest processes. The model code architecture allows aggregating trees into representative classes, each characterized with its variables (e.g. carbon pools, leaf area index, tree height) based on their ages, species-specific and structural traits that are identified by the model through four indexes: i.e. species (x index), diameter class (Diameter at Breast Height, DBH) (y index), height class (z index), and age cohort (k index); such indexes represent the main state variables considered by the model in distinguishing ecosystems across sites. To deal with forest hetereogenity within and across different ecosystem, 3D-CMCC FEM uses a species-specific parameterization for each species simulated. Moreover, based on the assump-
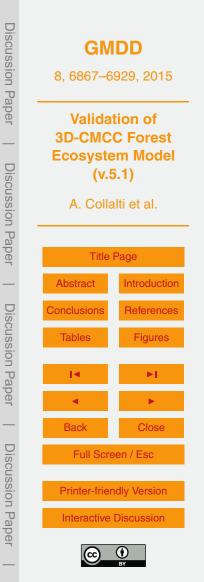
tion made by Magnani et al. (2007) that the above-ground net primary production decreases along with the ageing of a forest, the model explicitly takes into account all ages within the stand, reproducing a year by year reduction due to senescence (Landsberg and Waring, 1997; Waring and McDowel, 2002). Height classes and the tree position within the forest vertical profile are explicitly treated by the model to estimate the light availability using the Monsi-Saeki formulation of exponential attenuation coupled with the "Big-leaf" approach developed for a multi-layered model (Collalti et al., 2014; Medlyn et al., 2003) (see also Appendix A3). DBH togheter with stand density control grid cell horizontal canopy coverage (and gaps) through the computation of the single tree crown coverage and then upscaled to grid-cell level (Collalti et al., 2014). In this way, the model is able to reproduce different combinations of uneven-aged, multi-layered and multi-species forests, by optional simulation of e.g. light competition, age related decline and different species-specific traits. This aspect makes the model flexible to be theoretically used for a wide range of applications in forests and allows quantifying the effects of a particular simulation of forest structure on model performance. In this study, the 3D-CMCC FEM described in Collalti et al. (2014) has been advanced to version 5.1 to improve the representation of forests processes, like phenology, canopy photosynthesis, including autotrophic respiration and tree carbon-nitrogen allocation and water flows. The improved phenology routine is based on a new C allocation scheme, that include the Non-Structural-Carbon (NSC) pool, related to five phenological transitions for deciduous species, and three phenological transitions for evergreen species, both updated once per day. Autotrophic respiration is separated into mainteinance and growth respiration. Mainteinance respiration is function of the nitrogen content (a new added pool) in the living pools, while growth respiration is computed proportionally to the carbon allocated to the different tree compartments (see Appendix A).

2.2 Data description

Model validation has been performed for ten heterogeneous forest sites (Table 1) included in the European EC fluxes database cluster (URL:

http://www.europe-fluxdata.eu). For each site, 3D-CMCC FEM simulations were performed averagely for 10 years, forced with gap-filled daily meteorological data, according to the available time series. The selected sites cover a wide range of European forest ecosystems across different latitudes, landscapes and three climatic
5 zones: temperate, Mediterranean and subalpine.

For all the sites, daily time series of meteorological variables (maximum and minimum air temperature, precipitation, vapour pressure deficit and incoming solar global radiation) were used as drivers, while GPP was used for model output validation. The GPP derives from Net Ecosystem Exchange (NEE) measurements that have been
10 previously quality checked and processed including storage correction, spike detection, and low turbulence condition (u^*) filtering according to the method in Papale et al. (2006) and gapfilled using the Marginal Distribution Sampling method (MDS; Reichstein et al., 2005). The GPP is not directly measured by the eddy covariance technique but it is estimated using a partitioning technique as described in Reichstein
15 et al. (2005). In the rest of the paper we will refer to these data as "measured" or "observed" GPP for simplicity but it is important to highlight that they are obtained using a modeling approach (although strongly based on direct measurements).

2.3 Model and experimental set-up

Site data needed for model initialization concern information on forest structure (DBH,
20 tree height, age, and density), its species composition, and soil characteristics (e.g. depth, texture and bulk density). These data were used for each site to initialize the model, i.e. to describe the initial forest condition at which model starts to simulate forest processes. Initialization data were taken from the BADM (Biological, Ancillary, Disturbance, Metadata) files, available at http://www.europe-fluxdata.eu, for each of
25 the selected sites, and complemented by literature review and personal contacts with the sites Principal Investigators. Length of model simulations, basic sites description and forest attributes are shown in Table 1. As a whole, for all sites, the species-specific ecophysiology has been parameterized generically (i.e. not related to the simulated

site) using only literature data (e.g. Breuer et al., 2003; Mollicone et al., 2003; Pietsch et al., 2005; White et al., 2000) independently from site-related measurements (for a full list of model ecophysiological and structural species-related parameters see Collalti et al., 2014). As in Naudts et al. (2014), in case of multiple values for a single parameter the mean values were used. Using the mean parameter estimates avoided hidden model-tuning and largely reduces the likelihood that simulation results are biased by hidden calibration.

In addition, several studies (Bolstad et al., 1999; Griffin et al., 2001; Ibrom et al., 2006; Misson et al., 2007; Cescatti et al., 2012; Guidolotti et al., 2013; Migliavacca et al., 2015) claim that beside environmental variables, spatial heterogeneity (horizontal and vertical) of the stand structure and composition (age, species) also plays an important role at the ecosystem level.

To evaluate if a more detailed simulation of forest heterogeneity improves model performances, a number of replicated simulations were performed for three heterogeneous sites (BE-Bra, IT-Ren and DE-Tha), based on different model initializations in terms of forest layers, species composition and/or ages (Table 1). These replicates start from a forest representation very close to reality (e.g. cohorts, mixed species composition and different canopy layers) to a more generalized one. For reasons of comparability, in these test sites the model has been forced with the same meteorological input data, and eco-physiological species-related parameterizations, i.e. only model initializations data, related to stand attributes, differ. These data are based on different sources: site measurements and/or literature data and/or experimental settings.

In the case of BE-Bra we initialized the model with near all possible combinations of initialization datasets. The first simulation (BE-Bra P_Q-3L) has explicitly taken into account the site heterogeneity (vertical and horizontal) (following Gielen et al., 2013, and ancillary data sources) consisting in mixed species composition at a different canopy coverage rate of *Quercus robur* (Q) and *Pinus sylvestris* (P) (20 and 80 %, respectively), with two cohorts (oaks and pines, 65 and 72 years old, respectively) and three forest layers. In the second simulation (BE-Bra P) only single-layer of Scots pines was

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considered (following Janssens et al., 2002; Verbeeck et al., 2007). In the third, fourth and fifth simulations (BE-Bra Q_3L, BE-Bra Q_2L, BE-Bra Q_1L, respectively) only three, two and one layers of pedunculate oaks (following Curiel Yuste et al., 2005 and experimental set up) were assumed. Additionally, two more experimental set-ups combined two layers of oaks and one layer of pine (BE-Bra P_Q-2L) and one layer of oak and pine (BE-Bra P_Q-1L).

For IT-Ren, in the first simulation two layers and two cohorts were considered (IT-Ren 2L_2C) following Montagnani et al. (2009). In the second case stand heterogeneity has been grouped into one layer, i.e. minimizing forest structure, and one single averaged cohort (IT-Ren 1L_1C; experimental set up).

For DE-Tha, two species (DE-Tha 2S) (spruce 80 % and pine 20 %, respectively) were modelled in the first simulation (following Grünwald and Bernhofer, 2007), while in the second experiment only dominant species (spruce; DE-Tha 1S) was considered (BADM source).

2.4 Validation approach

In order to analyse model performance, we used time series of daily, monthly and annual modelled and observed GPP values, which were compared at the different time scales. At first, we conducted a comparison via appropriate performance indices on long-term annual average (i.e. over the full series of all the available years). Then we evaluated how the model performed in the different seasons aggregating values for months of the same season.

We firstly adopted the Pearson coefficient of correlation (r).

Then, we calculated the Normalized Root Mean Square Error (NRMSE) (Anav et al., 2010; Keenan et al., 2012) as a standardized index of error. The NRMSE reports the mean difference between observed and modelled GPP values (GPP_{EC} and GPP_{MD}, respectively) normalized on the variability in the GPP_{EC}, in order to have an indication of the average distance between GPP_{MD} and GPP_{EC}, comparable among the different

sites. NRMSE was quantified as:

$$\text{NRMSE}_{\text{GPP}} = \frac{\sqrt{\frac{1}{N}\sum_{i=1}^{N}\left(\text{GPP}_{\text{EC}_i} - \text{GPP}_{\text{MD}_i}\right)^2}}{\sigma\left(\text{GPP}_{\text{EC}_i}\right)} \tag{1}$$

where i represents the day (or month), and $\sigma\left(\text{GPP}_{\text{EC}}\right)$ is the standard deviation of the full daily (or monthly) series of observed GPP consisting of N records.

In addition, model performances were measured for the same series through the "Model Efficiency" index (MEF) following Reichstein et al. (2002) and Migliavacca et al. (2015):

$$\text{MEF} = 1 - \frac{\sum_{i=1}^{N}\left(\text{GPP}_{\text{EC}_i} - \text{GPP}_{\text{MD}_i}\right)^2}{\sum_{i=1}^{N}\left(\text{GPP}_{\text{EC}_i} - \text{avg(GPP)}_{\text{EC}}\right)^2} \tag{2}$$

In contrast to correlation coefficient r, the MEF index (Bowman and Azzalini, 1997) measures not only the correlation between modelled and observed data (in other words, how well they reproduce the phase of observations), but also their "coincidence", i.e. the deviation from the 1 : 1 line, and it is sensitive to systematic deviations between model and observations (Reichstein et al., 2002).

An additional index was the Bias (Bi):

$$\text{Bi} = \frac{1}{N}\sum_{i=1}^{N}\left(\text{GPP}_{\text{MD}_i} - \text{GPP}_{\text{EC}_i}\right) \tag{3}$$

calculated at both annual and seasonal level, positive biases indicate an overestimation and negative values indicate an underestimation, respectively, by the simulation (see Balzarolo et al., 2014).

To evaluate the model performances in terms of variability patterns, we adopted a procedure to compare each GPP_{EC} value to both its correspondent GPP_{MD} value and the GPP_{EC}-GPP_{MD} difference, at daily and monthly level. Since the different sites have different ranges of GPP, we grouped time series values into 18 clusters, with a 5 percentile criteria, from the 5th to the 95th (Vetter et al., 2008), and we calculated the median for each group.

In order to assess the Inter-Monthly and Inter-Annual Variability (IMV and IAV respectively), individual GPP values for each month and year considered were normalized following Vetter et al. (2008) and Keenan et al. (2012). Shortly, we subtracted the respective observed or modelled average from individual (monthly and yearly) observed and modelled value as follows:

$$IMV_{(EC\ or\ MD)i}\ or\ IAV_{(EC\ or\ MD)i} = GPP_{(EC\ or\ MD)i} - avg(GPP)_{(EC\ or\ MD)} \qquad (4)$$

where avg(GPP) is the long-term (full series of all the available years) average of monthly (for IMV) or yearly (for IAV) GPP from observations (EC) and modeling (MD), respectively. A kernel density estimation (kde) was performed to qualitatively observe probability distribuition functions (PDFs) respectively of the IMV and IAV values (Bowman and Azzalini, 1997).

To evaluate 3D-CMCC FEM ability in reproducing the observed IMV and IAV, we calculated the NRMSE based on monthly and annual time series of IMV and IAV values, respectively. The NRMSE, adopted as a normalized index of error allowing comparability among different sites, was thus calculated as in Eq. (1) but using IMV and IAV instead of GPP individual values, following the approach of Keenan et al. (2012).

3 Results

3.1 GPP evaluation over long-term annual and seasonal scale

Both monthly and daily simulated (MD) GPP show high correlations with EC data, with low biases (Table S1 and Fig. 1). On average deciduous forests reveal better correlation

5 between MD and EC data than evergreen forests, with a mean r of 0.86, while evergreen and mixed stands show average r of 0.80 and 0.77, respectively. These results are confirmed by Taylor diagrams (Taylor, 2001) (Fig. 2a) which show that the model performs satisfactorily for daily fluxes, in four (i.e. DE-Hai, DK-Sor, DE-Tha, FI-Hyy) of ten sites falling within ±0.5 normalized standard deviations from the reference point

10 (representing observed data) and having correlation around 0.9. For six sites (all the evergreen needleleaf plus deciduous except FR-Hes) the normalized standard deviation of simulated data is really close to that of observed data (represented by reference line with normalized standard deviation, i.e. radial distance from the axis origin, equal to 1). Simulated data for IT-Cpz, FR-Hes and FR-Pue have, respectively, a normalized

15 standard deviation around +0.2, +0.3 and +0.4 (as difference from that of observations); BE-Bra shows the highest negative difference around −0.3. On average, the worst result is for IT-Cpz that shows a correlation below 0.60 and falls outside ±1 normalized standard deviation from the reference point. For all stations $p < 0.0001$.

Considering the mean monthly cycle, the Taylor diagram (Fig. 2b) shows the model

20 capability to better simulate GPP at monthly scale. For seven sites (all deciduous and evergreen needleleaf) the normalized standard deviations of modelled data are close to that of observations (reference line), the correlation is above 0.90 and within ±0.5 normalized standard deviation from the reference point. IT-Cpz and BE-Bra show improved results with respect to daily data: respectively, their correlation increases of

25 more than 0.1 units, they fall within the +0.2 and −0.2 units of normalized standard deviation differences with respect to observations, and they enter in the field of ±1 and ±0.5 normalized standard deviation from the reference point. Although less strongly, also FR-Pue monthly data have better performances than daily data results in terms of

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higher correlation and closer position in terms of normalized standard deviations units from the reference point.

To summarize, although with similar inter-sites variability, monthly correlations across different sites are higher than daily ones, with average correlations of 0.94 for deciduous, 0.89 for evergreen and 0.91 for mixed stand (Fig. 1 and Table S1 in the Supplement).

Daily and monthly NRMSE are low, 0.63 and 0.42 on average, respectively (Table S1), confirming that the model performs better at monthly than at daily time scale (Fig. 1), likely because of averaging effects of daily variability in GPP estimation.

The same is shown by MEF that is on average 0.79 (monthly) and 0.58 (daily), with largely lower values for the two Mediterranean forests (IT-Cpz and FR-Pue) at both daily and monthly time scale (Table S1 and Fig. 1).

Considering annual mean in deciduous forests (Table S1) the model slightly underestimates the GPP by -2.8% (average among DE-Hai, DK-Sor and IT-Col), with only FR-Hes showing an overestimation of 6.4%. Concerning evergreen forests, we find an overall model underestimation of 1.3%, with higher variability compared to deciduous forests, and more divergent in the case of the two Mediterranean ecosystems, ranging from underestimation of 18.4% ($318\,\mathrm{g\,C\,m^{-2}\,year^{-1}}$; IT-Cpz) to overestimation of 12.1% ($158\,\mathrm{g\,C\,m^{-2}\,year^{-1}}$; FR-Pue).

Results for the mixed forest site of BE-Bra are reasonable, with an overestimation of about 5%.

In terms of inter-annual variability of the yearly mean, GPP_{MD} falls well within the range of GPP_{EC} standard deviations for all sites except at IT-Cpz (Fig. 3). Deciduous broadleaved forests are the best reproduced (average bias of about $70\,\mathrm{g\,C\,m^{-2}\,year^{-1}}$).

Performance indices from daily and monthly observed and modelled GPP series analysed at seasonal level are shown in Table S2 and Figs. 4 and 5. Winter (DJF) and summer (JJA) correlations were generally lower than those in autumn (SON) and spring (MAM). Specifically, DJF and JJA showed a correlation of 0.46 and 0.48 respectively on a daily scale and a value of 0.54 and 0.53 on a monthly scale; MAM and SON showed

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on a daily scale an average correlation of 0.72 and 0.77 respectively, while on monthly scale a correlation of 0.82 and 0.86 with two low values of 0.05 and 0.06 for monthly DJF and MAM for IT-Cpz.

Winter and summer monthly average NRMSE of 1.19 and 0.97, respectively, were not significantly different to the 0.67 and 0.58 of spring and fall. MEF and Bi indexes values suggest similar findings than NRMSE.

Figure 6 shows overall modelled vs. observed fluxes over daily and monthly scales, and the absolute difference (GPP_{MD} minus GPP_{EC}) vs. observed fluxes (GPP_{EC}) as calculated by the difference matrix described in Sect. 2.4. Overall, the aggregated data reveal high correlation also due to a progressively reduced range of data, and then variability, at higher GPP values (Fig. 6a and b). Figures 6c and d show patterns of absolute difference between GPP_{MD} and GPP_{EC} with increasing GPP_{EC}. These differences result in strong reduction of discrepancies for GPP_{EC} greater than $8.5\,\mathrm{g\,C\,m^{-2}\,d^{-1}}$ for daily, or $7.3\,\mathrm{g\,C\,m^{-2}\,d^{-1}}$ for monthly temporal series (data extracted from Fig. 6c and d).

The average intra-annual GPP variations are analysed by calculating the long-term average and standard deviation values for each month of the year (Fig. 7). In spring the model results from deciduous forests present a larger variability than the observed data, especially during budburst and in late spring. The model generally matches the observed phenology timing (budburst, peak LAI, leaf senescence and their fall, i.e. length of growing season, data not shown). Consistent biases were observed in late summer.

3.2 Inter-monthly and inter-annual variability

The distribution of the IMV for the analysed sites reveals in general lower variance for modelled than observed data (Fig. 8, Table 2). Regarding deciduous forests, both DK-Sor and FR-Hes show IMV_{MD} distributions with larger interquartile range in comparison with IMV_{EC} (p value < 0.05). Conversely, for DE-Hai and IT-Col the IMV_{MD} variance is statistically representative for the IMV_{EC}; however IT-Col shows a significantly biased

median (p value < 0.05). Less variability than IMV_{EC} is generally observed for IMV_{MD} of conifers. While DE-Tha shows significant agreement for both variance and central tendency (p value < 0.05), at FI-Hyy the IMV_{MD} appears statistically in disagreement with IMV_{EC} for both variance and central mean tendency (Table 2). We find a small difference between IMV_{MD} and IMV_{EC} probability density modal values in IT-Ren (Table 2). Concerning broadleaved evergreen vegetation, we observe very good agreement between observed and modelled IMV central tendency measures in FR-Pue with most of the frequencies between $\pm 2\,g\,C\,m^{-2}\,d^{-1}$. In FR-Pue, however, we notice that the distributions are slightly shifted, especially around the median, with resulted variance from modelled data in disagreement with that from observed data. We detect high IMV distributions disagreement in IT-Cpz, where the PDF from observed IMV is normally distributed, while the one from modelled IMV is not (as resulted by a χ^2 goodness of fit test). IMV_{MD} series in BE-Bra (mixed forest) are in low agreement with those from EC. Modelled variance is low, and especially positive IMV values are scarcely represented. Table 2 also shows the NRMSE for IAV and IMV series. There is apparent correlation neither between sites species and average error, nor between distributions uniformity and NRMSE. In fact the lowest NRMSE for IMV was found in BE-Bra and IT-Col, the highest in DE-Hai and DK-Sor. On average the model has a NRMSE for IMVs of about 1.2.

Figure 9 shows the modelled and measured individual IAV values for each studied site. The magnitude of IAV_{MD} was on average of the same order than IAV_{EC}, showing the model ability to reproduce the inter-annual variability range, and capturing about 62 % of the anomalies signs (i.e. timing) for the total set of years. The model generally better captured conifers' IAV sign (i.e. DE-Tha, FI-Hyy, and IT-Ren), 66 % of the times against about 59 % for the deciduous forests (i.e. DE-Hai, DK-Sor, FR-Hes, IT-Col) and 55 % for the Mediterrenean ones (i.e. FR-Pue and IT-Cpz). However the IAV difference in magnitude was better represented for deciduous forests rather than conifers, as inferred by the average NRMSE of respectively 1.45 and 1.66 (calculated by averaging values reported in Table 2). Although model reproduced well the timing of anomalies in
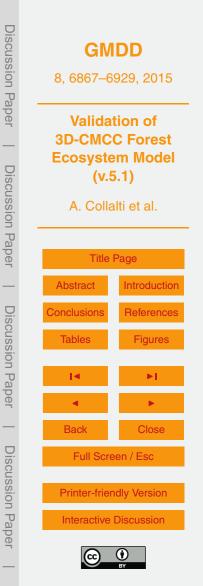
more than half of cases, the correlations had a wide spread across sites. Quantitatively, modelled anomalies suggest better results for FR-Pue ($r = 0.75$) and worse results for IT-Ren ($r = -0.53$).

In case of year 2003 with its summer heat and drought extreme (Ciais et al., 2005; Vetter et al., 2008), the anomaly sign has been well captured by the model on six of the eight sites analyzed for that year (not enough observations were available for BE-Bra and IT-Col, while 2003 was recognized as not anomalous at IT-Col) (Fig. 9). At IT-Cpz and DK-Sor, average IAV_{MD} has opposite sign than IAV_{EC}. Similarly, the model results matches with what found by Delpierre et al. (2009) about the anomalous carbon uptake during the warm spring of 2007 compared with the decadal mean for FR-Pue, FR-Hes, DE-Tha and FI-Hyy.

3.3 Comparison within different forest structure simulations

Considering the presence of only one species (either pines or oaks) strongly limits the model to simulate the daily and monthly GPP patterns in BE-Bra (Table 3). This site represents a mixed stand of deciduous and evergreen tree species that assimilates CO_2 all year round, although low temperatures in winter and spring reduce photosynthesis also for pines. The observed GPP fluxes are then caused by the "mixture", at a varying degree, of both oak and pine trees. Considering BE-Bra as a pure oak forest with a variable number of layers (simulation codes: BE-Bra Q_3L, BE-Bra Q_2L, BE-Bra Q_1L) the model results for annual GPP deviate from −0.6 up to +6 %; considering a pure pine forest (BE-Bra P) or a combination of pines and one layer of oak (BE-Bra P_Q-1L) the model underestimates annually from −9.8 to −6 %, respectively. It is noteworthy that the daily GPP values markedly show a different seasonal distribution on fluxes (data not shown). Conversely, there is no clear evidence that in simulating pines coupled with one, two or three oak layers (BE-Bra P_Q, BE-Bra P_Q-3L BE-Bra P_Q-2L) model results largely benefitted of this differentiation both on a daily, monthly and annual scale. Similar results are obtained for DE-Tha site when simulating one single species (DE-Tha 1S) or two (DE-Tha 2S), with annual bias of +1.5 %, since the similar

phenology behaviour of modelled species does not cause a marked difference in the seasonal GPP cycle. Differently, IT-Ren initialized as a single layer and with one single cohort (IT-Ren 1L_1C) instead of two layers and two cohorts (IT-Ren 2L_2C) differs strongly from observed GPP values overestimating for 43.2 % the annual cumulated GPP. However, for this site, the analysis of performance indices based on daily and monthly series shows no evidence of improved model results.

4 Discussions

In this paper we have analyzed the capability of the latest version of the 3D-CMCC FEM to simulate intra-annual to inter-annual GPP variability over ten heterogeneous European forest sites representative of different ecosystems and bioclimatic regions by comparing model results with observations based on EC technique. Although the model provides a reasonable reproduction of the observed values, we may evince some critical issues. First, the observed GPP data are affected by high uncertainties (Kenan et al., 2002; Papale et al., 2006; Richardson et al., 2012a, b). According to Luyssaert et al. (2007) these uncertainties in the ten case studies here considered, although at the biome level, have a very high spread, varying from ±557.9 (for FI-Hyy) to ±700 g C m^{-2} yr^{-1} (for IT-Cpz). Beside uncertainty in EC technique, model assumptions and parameterizations can increase discrepancies compared to observed GPP data.

A potential further source of error in the model runs that may need to be considered or accounted for is related to our choice of not making a site-specific parameterization. Since we used general parameterizations, large uncertainties could be detected especially in the variables that determine for example the length of the growing season (Richardson et al., 2010), and the latitudinal differences (acclimation) of the maximum, minimum and optimum temperatures for photosynthesis. Improvement could be achieved with a site-specific parameterization, but this falls beyond our goal to make the model generally applicable.

On average, 10 years of simulations for each site have been conducted. In addition, in three sites different model initializations (i.e. considering different forest structure, composition and cohorts) were used to quantify improvements in model results when a more detailed heterogeneity forest structure representation and processes are simulated. Modelled GPP results were compared against those from EC observations collected for these sites encompassing three mono-specific (pure) stands of Beech, Holm oak and Scots pine, and three uneven-aged, multi-layered and mixed stands.

Based on results, we can now provide answers to the four initial questions.

4.1 Does the model reproduce the magnitude and timing of seasonal fluctuations in GPP across different forest types, structures and compositions?

Overall, as desirable, the model is skilful in reproducing the annual cumulated and intra-annual (seasonal) cycle of GPP, calculated as both daily and monthly value average, with the monthly scale performing better across all statistical indices considered. These results can be anyway considered as a "false positive" due to the strong seasonality of GPP patterns that influences and causes high values of correlation more than model capabilities to reproduce GPP fluxes. This is clearly related to the tendency to linearize the relationship between CO_2 flux and PAR, as also reported by Ruimy et al. (1995) and Wu et al. (2015), respectively. Overall, average annual modelled values were highly consistent with EC data, except for the Mediterranean sites (where seasonality is less pronounced). Here summer drought stress showed to be the most limiting factor on photosynthesis at FR-Pue (Falge et al., 2002; Reichstein et al., 2002; Sabatè et al., 2002) while the presence of shallow groundwater table at IT-Cpz seems reducing the severity of summer drought. However the model showed non-negligible uncertainties in representing GPP patterns, as inferred by temporal mismatches in variance. The overall agreement despite temporal mismatches suggested that errors compensated over the year, but are cumulated in specific time windows (e.g. seasons). As reported for other models (Morales et al., 2005; Naudts et al., 2014) the model's performances

are generally worse in winter (DJF) and summer (JJA). Biases and differences in winter GPP variance may be related to the model algorithms used to simulate LAI and to the algorithm used to calculate GPP from EC data (Reichstein et al., 2005), since GPP variability should be low during DJF, especially for deciduous forests. However,

5 mismatches are also related to the way in which 3D-CMCC FEM represents winter and early spring ecosystem processes. The model in fact does not consider the influence of ground vegetation that appears to be not negligible in some cases (Kolari et al., 2006).

High GPP variance for evergreen species could be strongly related to low temperatures during winter (Delpierre et al., 2009). Systematic overestimation in winter and

10 spring GPP could then be associated with lacks in representing conifers acclimation or to soil and atmosphere thermal constraints. At high latitudes and altitudes another source of uncertainty may be related to freezing and thawing dynamics in soil water (Beer et al., 2007) which are not considered by the model, as like as snow sublimation and melting are still simplistically represented.

15 GPP of deciduous forests in summer and autumn are also affected by uncertainties for surface, which is represented by LAI in the model. In addition GPP is linear with respect to PAR (Monteith and Moss, 1977) over monthly or annual time scale, while the relation is strongly nonlinear at daily scale (Leuning et al., 1995; Gu et al., 2002; Turner et al., 2003; Wu et al., 2015). The linear response of GPP to PAR led to

20 the underestimation/overestimation of GPP under conditions of low/high incident PAR (Propastin et al., 2012; He et al., 2013). In case of stress or photoinhibition leaves reduce or stop the photosynthesis at too high levels of radiation, while in normal condition, photosynthesis is light-saturated at high PAR (Mäkelä et al., 2008) which lets canopy photosynthesis saturated at relatively low PAR even in dense tropical forests

25 with high LAI (Ibrom et al., 2008). The model overestimation of summer GPP may thus be partially related to the lack of representation of canopy photosynthesis saturation processes.

Although adopting a more complex phenology scheme, in the comparison between decidous and evergreen forests our model showed better performances for deciduous

Validation of 3D-CMCC Forest Ecosystem Model (v.5.1)

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compared to evergreen forests. This behaviour is due to the strong seasonality patterns that the deciduous species show, but contrast to the results of Morales et al. (2005) who showed that it is generally easier for models to simulate evergreen forests due to the simpler phenology. The present results for evergreen forests are, however, highly af-

5 fected by the low model performances for the two evergreen Mediterranean forests. As said, overestimation during summer at FR-Pue, and during winter and spring for IT-Cpz, are mostly related to neglecting species-specific drought stress response functions. As in Landsberg and Waring (1997), the water modifier is only based on soil physical characteristics and no consideration is given to the stress tolerance or strat-
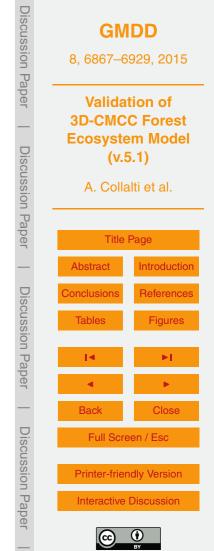
10 egy of the species (Larcher, 2003), suggesting that further model developments should focus on this aspect.

Other discrepancies affecting other sites could probably be reduced with a site-specific parameterization.

4.2 Does the model reproduce the observed inter-annual GPP variability?

15 Overall, the distribution of the modelled inter-monthly variability was sufficiently consistent with the observed one. The model however showed reduced variability in the distribution for both conifers and deciduous species. The model ability in better representing higher rather than lower anomalies suggests that it may still be less sensitive to some drivers of variability. In this context, the phenological cycle may have an important role,

20 since it influences canopy cover and it is controlled by environmental drivers (Richardson et al., 2010). According to Jeong et al. (2013) spring phenology largely affects the summertime carbon budget. Hence uncertainties in growing season starting date may affect 3D-CMCC-FEM ability to reproduce IMV. In summer and autumn, petioles loss of turgor, cavitation in xylem vessels and leaf yellowing may have an important role in

25 GPP variability of temperate forests (Reichstein et al., 2007).

Even though evergreen forests do not experience complete dormancy in winter, changes in "greenness" can be attributed to seasonal variation in canopy biochemistry, the production of new foliage by canopy species and, particularly where the overstorey

is sparse, the phenology of understory vegetation (Richardson et al., 2010). Leaves of different ages have different efficiency, sensitivity to solar radiation, temperature and water related stresses (Chabot and Hicks, 1982). All these elements may have an important role in affecting GPP dynamics, but are still scarcely or not represented by

5 mechanistic ecosystem or forest models. As a confirmation of these suspects, slight modifications in representing phenology and leaf turnover resulted in general improvement of model consistency with EC data (Marconi, 2014).

Distribution of IMV values showed specific patterns attributable to the dominant species. Beech forests IMV PDFs were concentrated around the average value and

10 strongly influenced by high biases. This pattern was probably due to the fact that half of the months in one year have no or little photosynthesis (i.e. early spring, fall and winter) and most of the photosynthetic activity occurs in late spring and summer, when carbon assimilation is influenced by temperatures and solar radiation (Mercado et al., 2009). Conifers PDFs were usually smoother, non-skewed, with reduced variability and

15 fitted by a statistical normal curve.

The model showed an average NRMSE for IMV of 1.22 but still captured about two third of the annual anomalies sign.

The results for IAV (see Fig. 9) are quite contrasting, and largely depend on site and the number of annual-by-annual comparison. Better results have been obtained

20 for FI-Hyy and FR-Pue, so there is not apparent correlation with latitudes and forest species. Similarly, lower results are reported for IT-Ren, IT-Cpz and BE-Bra where the number of annual correlations are lower than the other sites. The magnitude of differences in the standard deviation follows generally the same tendency, particularly for BE-Bra, IT-Ren and IT-Cpz. These results confirm the model limited ability to repre-

25 sents the inter-annual variability in these specific sites rather than in these ecosystems. The comparison between modelled and observed data at the inter-annual time scale shows the model to be sufficiently able to reproduce the sign of variability through the years including the extreme events (heat wave combined to drought) during the 2003 summer (Ciais et al., 2005; Vetter et al., 2008) and, for some sites, the anomalous
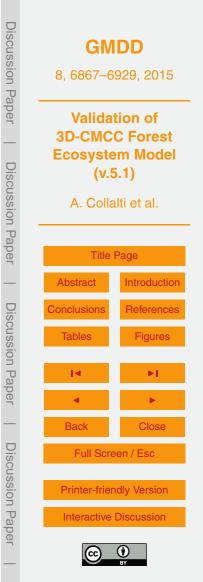
carbon uptake during the warm spring of 2007 described by Del Pierre et al. (2009). Potentially negative effects from the anomalous 2003 were modelled into negative GPP anomaly at DK-Sor and IT-Cpz due to model simulation of summer drought stress, while such anomalies are not evident from measurements for DK-Sor (Pilegaard et al., 2011).

5 This could be due to the more maritime climate for DK-Sor and the presence of shallow groundwater for IT-Cpz that weakened the effects in the first part of the summer. In both sites, and included DE-Tha, the effects during July to September were captured by the model (data not shown). As reported by Ciais et al. (2005), Mediterranean sites showed a smaller degree in carbon fluxes, largely dominated by less respiration. It is noteworthy

10 that IT-Col, differently from other european beech stands, does not seems having suffered from this anomalous heat wave in 2003 (G. Matteucci, personal communication, 2014). Both simulated and observed data showed a positive GPP anomaly, demonstrating that this beech forest benefited by moderate higher temperature values and consequently "extra" days for assimilation and growth (see also Churkina et al., 2002;

15 Richardson et al., 2010). A similar behaviour was reported also by Jolly et al. (2005) for the Swiss Alps, especially in the months from March to July. This pattern seems to be mostly related to an untimely beginning of growing season (see Piao et al., 2006), to a reduction in plant transpiration that causes an increase in plant water use efficiency throught the partial closure of stomata (Warren et al., 2011) and to high fluxes related

20 to forest floor vegetation.

It is also noticeable that in FR-Hes during the summer of 2004 a negative anomaly, larger than in 2003, occurred; while its sign was captured by the model, its magnitude was not. This can be explained by the modelled postponed effects of a low NSC allocation during the year 2003 to the subsequent periods (Granier et al., 2007; Gough et al.,

25 2009). These results highlight that model has a sort of "memory" linked to short-term events (e.g. drought stress) and that these events affect the long-term processes.
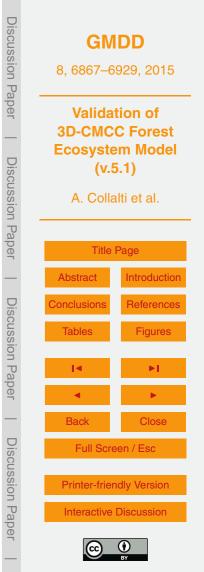
Quantitatively, modelled inter-annual anomalies show a very large spread across the sites. Correlations vary widely, without any apparent relation with latitude and/or species. If modelled anomalies signs are potentially agreeing with the observed ones

most of the times their magnitude was not. This behaviour seems to be related to several aspects, mainly to an over/under estimation of the causes that reproduce anomalies, e.g. processes simulated linked to the type of climate anomaly, mismatches in phenology or to a missed representation of others processes (e.g. mast years, disturbances, shallow water). Keenan et al. (2012) asserts that lacks in phenological variability and in canopy and soil dynamics are the main culprits of these mismatches but also that flux measurements are affected by random errors especially when fluxes are higher. Poulter et al. (2009) founded similar magnitude of errors also with models that were driven by remote-sensing data. Open questions remain as to the proportion of interannual variability in land–atmosphere carbon exchange that is directly explainable by variability in climate (Hui et al., 2003; Richardson et al., 2007).

4.3 Is the model generic enough that a single set of species-specific parameterization allows reproducing GPP behaviour across different ecosystems without further need of a site-related calibration?

Overall the model showed good flexibility although the sites showed a pronounced spatial and temporal heterogeneity (i.e. a variable number of forest layers, different cohorts and species). The model was able to reliably represent the ecophysiology of beech and spruce species at different latitudes, without modifying or tuning the parameterization sets. However, annual and seasonal performance indices, calculated exploiting daily and monthly series, evidenced different performances between the two northern beech sites and the two southern ones. Tables S1 and S2 show a systematic difference in all the statistics used, suggesting the presence of a latitudinal gradient in 3D-CMCC FEM ability to represent beech forest processes. This gradient could be explained by how the model represents the different limiting factors and their impacts on GPP. For example we expect low temperatures to be the most important limiting factor at higher latitudes, whereas soil water availability at lower latitudes (Chapin et al., 2002).

We had similar results for the two spruce sites. The model showed better performance at higher latitudes. While phenotypic plasticity, and thus the parameter set, may

influence the model results, it is noteworthy that the IT-Ren site has different topographic and climatic conditions. Lower average temperatures, higher slopes, and non-negligible encroachment of different species in a more complex canopy, may negatively influence the model performance in IT-Ren with respect to DE-Tha. Since the model
5 showed unrealistic results for the two Mediterranean forests, we think it is not easy to determine if and how differences in performances are related to the generality of the model rather than to bad assumptions behind the simulated processes. From our findings we conclude that for non-water limited conditions it is possible to yield satisfying results with general parameter sets.

10 ## 4.4   Do model's results improve when considering a complex 3-D canopy structure?

We evaluated possible improvements if a more accurate model representation at a higher rate of heterogeneity of: forest structure, differences in ages and species composition and their linked structural-ecophysiological processes, are assumed. These
15 analyses helped us to understand the importance of each process within the represented combination (i.e. light competition, age related decline and the specific differences in ecophysiology) on modelled GPP. Doubtless, a direct comparison between modelled and observed GPP data is not possible due to the lack of partitioned measurements of GPP across different layers, cohorts and species. However, in situations
20 where the different ecophysiological behaviours express themselves in the species specific canopy responses during certain periods of the seasonal cycle, the test of a mixed forest tree model with flux measurements is possible, as the results by Oltchev et al. (2002) showed using the model MixFor-SVAT.

This preliminary analysis can be considered as a sensitivity analysis in terms of
25 processes explicitly simulated instead of lumped parameterisation. As a whole, model results using different initialization data are within the observed GPP uncertainties but a quantitative assessment for two sites, BE-Bra and IT-Ren showed to potentially increase of model ability in simulating fluxes, while for DE-Tha there is no evidence that
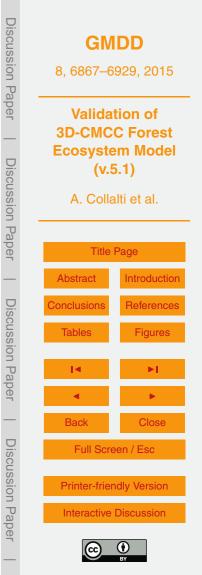
model performances could benefit of these efforts. For BE-Bra, taking into account two species (that differ especially for their phenological traits) was beneficial in terms of model performances, the same occurred for different layers (with the exception of BE-Bra P_Q-3L vs. BE-Bra P_Q-2L whose results were similar) and different cohorts.

5 Better performances, in terms of seasonal GPP representations, were obtained when each of the above mentioned characteristics was accounted for by the model. For IT-Ren similar results were obtained, although no differences were found in the simulation of phenological patterns in daily and monthly results. Differently, for DE-Tha a differentiation between the two evergreen coniferous species did not cause marked differences

10 in model results, due to low differences in species ecophysiological traits, justifying in these cases the use of a Plant Functional Type (PFT) level of parameterization instead species level (Poulter et al., 2015).

5 Conclusions

This study aimed at evaluating the performances of the updated version of 3D-CMCC
15 FEM compared to nearly 10 × 10 sites *x* years GPP data across eddy-covariance European forest sites. Although the sites showed high spatial and temporal environmental heterogeneity, the model appears able to reproduce GPP trends in all of the ten sites. Different performance indexes showed that daily and monthly level model results matches well, both for annual and seasonal scale, against observed data, with
20 some exceptions. Mediterranean sites (IT-Cpz and FR-Pue) showed to be the most problematic in reproducing carbon fluxes. This is likely due to their specific ecosystem peculiarity, e.g. shallow groundwater for IT-Cpz, and for both sites to a low pronounced seasonality. In these two sites model showed to be of less generalisation unless to include additional processes. Differently from other models 3D-CMCC FEM both for
25 daily and monthly simulations performs better for deciduous species rather than for evergreen, although deciduous have a more complex phenology and a more pronounced seasonality. Some mismatches in the simulation over the seasons and over the sites

still remain especially during winter and summer. The first reason for these low agreements in winter can be attributable to errors during the estimate of GPP from NEE and Ecosystem Respiration values from measurements data. The second can be related to the model's lack or simplicity in representation of snow pack dynamics as reported by

5   Krishnan et al. (2008, 2009) especially for evergreen sites (Keenan et al., 2012). Disagreements in summer could be related to model simplicity in simulating soil drought and, using the Monteith approach (Monteith and Moss, 1977), to the strong nonlinearity at daily scale of GPP and PAR, and to the lack of representation of the light saturation processes. In addition, as reported by Keenan et al. (2012), the apparent high variabil-

10 ity in the data during the summer season could therefore be due to random errors in the flux measurements generating larger variability and then lower correlations against modelled data.

No marked differences were found in simulations across different latitudes, so model parameterizations for the different tree species could be useful over Europe with a quite

15 high rate of confidence, with the exception of specific cases in Mediterranean forests.

As for other models, 3D-CMCC FEM showed to have the potential to correctly reproduce the signs of interannual variability, like the 2003 heat wave and drought extreme and the anomalous carbon uptake during the warm spring of 2007 and their instantaneous biological response to these events. Significant disagreements were however

20 found in reproducing magnitude of these anomalies.

The consideration of stand hetereogeneity, when possible or existing (i.e. layers, cohorts and mixed composition), led the model to improve its results in two of the three sites compared to generalized simulations of forest attributes. This plasticity makes the model able to be used in a wider range of forest ecosystems.
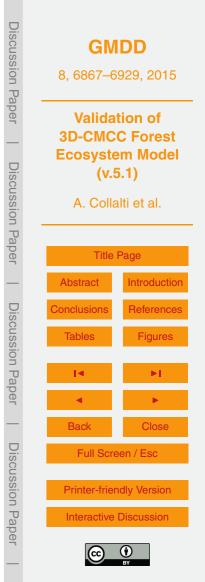
Appendix A: Model description

A1 Photosynthesis

As in the Collalti et al. (2014) 3D-CMCC FEM version, the carbon flux is still estimated by multiplying, for a particular species x, the absorbed photosynthetic active radiation (APAR, i.e. the radiation intercepted by the canopy) with the leaf area index (LAI, $m^2 m^{-2}$) with either the prognostic potential radiation use efficiency (ε_x, grams of dry matter MJ^{-1}) or the maximum canopy quantum use efficiency (α_x, $\mu mol\,CO_2\,\mu mol^{-1}$ PAR) (for a full list of model parameters, algorithms, and indexes see Collalti et al., 2014). Parameters ε_x or α_x are controlled by the product of several environmental factors (modifiers) indicated as $mod_{x,k}$ (dimensionless values varying between 0 and 1 and differing for each species x and age class k) depending on: vapour pressure deficit, daily maximum and minimum air temperatures, soil water content and site nutrient status (for a full modifiers description see Landsberg and Waring, 1997). Gross primary production (GPP; $g\,C\,m^{-2}\,day^{-1}$) is thus calculated using the following equation:

$$GPP_{x,y,z,k} = \varepsilon_x \cdot APAR_z \cdot mod_{x,k} \tag{A1}$$

where APAR is the absorbed radiation by the trees at the zth layer (where z represents the layer of representative height for each height class), while y represents the tree diameter class.

Autotrophic Respiration (AR) is treated distinguishing into Maintenance Respiration (MR), governed by a Q_{10} type response function (see Sect. A4) (Ryan, 1991; Bond-Lamberty et al., 2005) and Growth Respiration (GR) assumed to be a constant proportion (30 %) of all new tissues produced (Larcher, 2003). Net Primary Production (NPP), is calculated as follows:

$$NPP_{x,y,z,k} = GPP_{x,y,z,k} - AR_{x,y,z,k} \tag{A2}$$

NPP is then partitioned into biomass compartments and litter production following dynamic allocation patterns that reflect environmental constraints (i.e. light and water competition) and age.

A2 Daily meteorological forcing and snow dynamics

The model implements a daily time step due to the temporal frequency of meteorological forcing input data; average maximum (T_{max}) and minimum air temperature (T_{min}), soil temperature (T_{soil}), vapour pressure deficit, global solar radiation and precipitation. In addition, the model uses the day-time (T_{day}) and night-time (T_{night}) average temperature computed as follows (Running and Coughlan, 1988):

$$T_{day} = 0.45 \cdot (T_{max} - T_{avg}) + T_{avg} \tag{A3}$$

$$T_{night} = (T_{day} + T_{min})/2 \tag{A4}$$

When the soil temperature is missing among in situ observed data, the model estimates it for the upper 10 cm of the soil layer through an 11 day running weighted average of daily average air temperature and further corrected by the presence of a snowpack as in Thornton (2010), Kimball et al. (1997) and Zeng et al. (1993). The variable related to the snowpack thickness was included as a water cycle component by reproducing the daily amount (mm day^{-1}) of snow melt driven by average air temperature (T_{avg}) and incident net global radiation (Rad$_{soil}$), while snow sublimation is only driven by average air temperature.

In case of snow presence, if the average air temperature is higher than 0 °C, considered the melting point as in Running and Coughlan (1988) and Marks et al. (1992), the rate of daily snowmelt is estimated by:

$$Snow_{melt} = (t_{coeff} \cdot T_{avg}) + \left(\frac{Rad_{soil} \cdot \varepsilon_{snow}}{H_{fus}} \right) \tag{A5}$$

where t_{coeff} is the snowmelt coefficient ($0.65\,kg\,m^{-2}\,^\circ C^{-1}\,day^{-1}$), ε_{snow} is the absorptivity of snow (0.6), H_{fus} is the latent heat of fusion ($335\,kJ\,kg^{-1}$), Rad_{soil} is the incident net global radiation at the soil surface ($kJ\,m^{-2}\,day^{-1}$).

Otherwise, if the average air temperature is lower than $0\,^\circ C$ snow sublimation is computed by:

$$Snow_{subl} = \left(\frac{Rad_{soil} \cdot \varepsilon_{snow}}{H_{sub}} \right) \tag{A6}$$

where H_{sub} is the latent heat of sublimation ($2845\,kJ\,kg^{-1}$).

A3 Phenology and carbon/nitrogen allocation

Phenology plays a fundamental role in regulating photosynthesis and other ecosystem processes (e.g. carbon and nitrogen dynamics), as well as inter-individual and inter-species competitive relations and feedbacks to the climate system (Richardson et al., 2012a). In the updated model version phenology and carbon allocation depend on six different carbon and nitrogen pools. Five pools represent the main tree organs: foliage, (fine and coarse) roots, stem, branch and bark fraction. One pool corresponds to non-structural carbon (starch and sugar) stored in the whole tree. Woody pools are further-more distinguished between live and dead wood. This is necessary to represent NSC mobilization and consequently leaf phenology (e.g. leaf production during spring for de-ciduous trees) and carbon allocation. In the new version of 3D-CMCC FEM LAI values are predicted for sun and shaded leaves (De Pury and Farquhar, 1997; Thornton and Zimmermann, 2007; Wu et al., 2015), minimizing the effects of the "Big-leaf" approach (Monteith, 1965; Sellers et al., 1997), as a function of the amount of carbon allocated to the leaf pool. It is noteworthy that each pool and each structural state variables is daily updated according to the meteorological data, forest structure and simulated fluxes. Following Arora and Boer (2005), for deciduous species the model considers five phe-nological transitions that drive the seasonal progression of vegetation through phases

of dormancy/quiescence, budburst, maximum growth, active growth, and senescence as in the following:

1. Leaf onset starts from quiescence when thermic sum (the sum of the T_{day} air temperatures exceeding the threshold T_{base} value of 5 °C) exceeds a species- and site-specific temperature threshold value (Rötzer et al., 2004; Dufrene et al., 2005) and up to LAI = max(LAI) · 0.5. The costs of expanding buds during this period of high carbon demand are supported by NSC (Landhausser, 2010; Dickmann and Kozlowski, 1970).
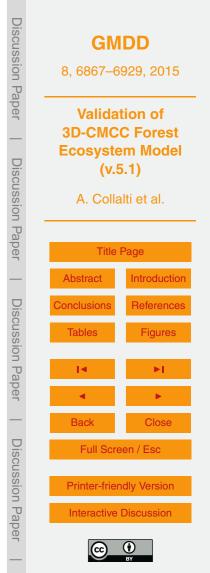
2. During the budburst phase, carbon and NSC are allocated to the foliage pool, as long as the balance between GPP and AR is positive (Barbaroux and Bréda, 2002; Campioli et al., 2013; Scartazza et al., 2013).

3. During the succeeding maximum growth phase and lasting up to peak LAI, carbon is allocated into foliage and fine root pools (Sabatè et al., 2002), based on the pipe model theory (Shinozaki et al., 1964a, b), to optimize photosynthesis; otherwise, no growth occurs and NSC is used.

4. Successively, the full growing phase lasts up to the day when day length (in hours) is shorter than a species-specific threshold value. In this phase carbon is allocated into stem, fine and coarse roots, branch and bark, and into non-structural carbon pools in order to refill the reserves for the next years.

5. Finally, during the leaf fall (i.e. yellowing or senescence) phase, lasting until the leaf fall (assumed linear) is complete, the total positive carbon balance is allocated to the NSC pool.

Outside the growing season (dormancy) trees consume NSC for fuelling maintenance respiration (Ogren, 2000).

For evergreen species the model follows a similar but simplified approach simulating a first maximum growth phase, when the model allocates NSC to foliage and fine roots

up to reach peak LAI, and a second full growing phase, when the model allocates to the other pools. As in Lawrence et al. (2011) for litterfall we assume and simplify that there are no distinct periods, but rather a continuous shedding of foliage and fine roots of the previous years.

All tree pools are updated at a daily time step depending on NPP. Nitrogen concentration for each pool is considered as a C/N ratio following Thornton (2010) and Dufrene et al. (2005). The C/N stoichiometry is constant and depends on species, unfortunately, the model still lacks of an interactive C-N cycle. Forest stand structural attributes, e.g. diameter at breast height (DBH), tree height, and crown competition are also updated at a daily timestep based on species-specific biometric relationships.

A4 Autotrophic respiration

Based on the approach of BIOME-BGC model (Thornton, 2010) 3D-CMCC FEM computes the daily AR of all living tissues. MR is a modified Van't Hoff function (Davidson et al., 2006; Mahecha et al., 2010) of temperature with the temperature sensitivity parameter Q_{10} (see below) and a linear function of the nitrogen content ($N_{content} = 0.218\,kg\,C\,kg\,N^{-1}\,day^{-1}$; Ryan, 1991) in the living compartments. The Q_{10} function is an exponential function for which a 10 °C increase in temperature relates to a Q_{10} factor change in the rate of respiration. MR is partitioned into day time and night time respiration using, in place of temp in Eq. (A7): t_{day} and t_{night} for foliage, t_{soil} for fine and live coarse roots, and t_{avg} for live stem and branch.

$$MR_{x,y,z,k} = 0.218 \cdot N_{content\,x,y,z,k} \cdot Q_{10}^{(temp-20)/10} \tag{A7}$$

$GR_{x,y,z,k}$ is considered as a fixed ratio (30 %) of all newly grown (i.e. living) tissues as proposed by Larcher (2003).

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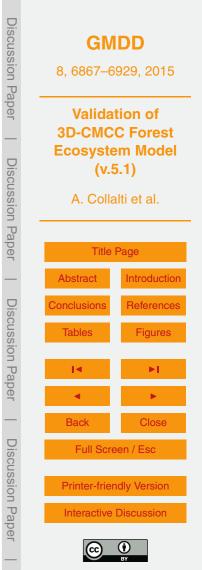
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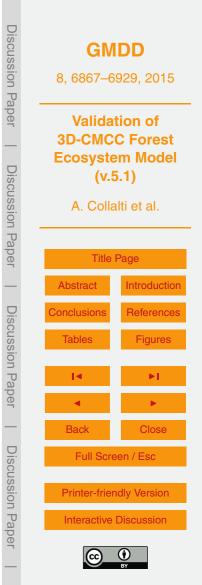
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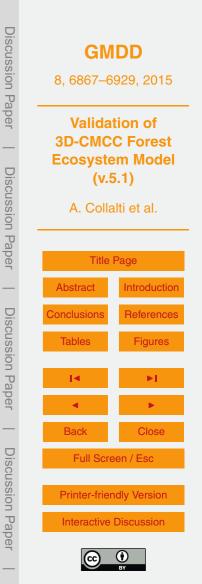
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Table 1. Main characteristics of the study sites. IGBP (International Geosphere Biosphere Program) legend: MF = mixed forest; DBF = deciduous broadleaf forest; EBF = evergreen broadleaf forest; ENF = evergreen needle leaf forest. Years of simulation starting and ending depend on available time series of observed data.

Site Name (Site code)	Lat (°)/ Lon (°)	IGBP	simulation year (Starting – Ending)	Mean Annual Temperature (°C)	Mean Annual Precipitation (mm yr^{-1})	Elevation (m a.s.l.)	Main species and forest description	References
Hainich (DE-Hai)	51.08/10.45	DBF	2000–2007	8.3	720	445	Uneven-aged, unmanaged multi-layered forest of beech (*Fagus sylvatica*, 250 years)	Knohl et al. (2003)
Sorø (DK-Sor)	55.49/11.64	DBF	2001–2009	8.2	660	40	Beech (*Fagus sylvatica*, averagely 80 years)	Pilegaard et al. (2003)
Hesse (FR-Hes)	48.67/7.07	DBF	2001–2007	9.2	820	300	Beech (*Fagus sylvatica*, averagely 35 years)	Granier et al. (2000)
Collelongo (IT-Col)	41.85/13.59	DBF	1997–2012	6.3	1180	1550	Beech (*Fagus sylvatica*, averagely 100 years)	Scartazza et al. (2013)
Puechabon (FR-Pue)	43.74/3.60	EBF	2000–2011	13.5	883	270	Holm oak (*Quercus ilex*, averagely 59 years)	Loustau et al. (2005)
Castelporziano (IT-Cpz)	41.71/12.38	EBF	2000–2008	15.6	780	3	Holm oak (*Quercus ilex*, averagely 45 years)	Vitale et al. (2003)
Tharandt (DE-Tha)	50.96/13.57	ENF	2000–2010	7.7	820	380	Mixed Norway spruce (*Picea abies*, averagely 113 years) and Scots Pine (*Pinus sylvestris*, averagely 113 years)	Grünwald and Bernhofer (2007)
Hyytiälä (FI-Hyy)	61.85/24.29	ENF	2001–2011	3.8	709	170	Scots pine (*Pinus sylvestris*, 39 years)	Tanja et al. (2003)
Renon (IT-Ren)	46.59/11.43	ENF	2006–2010	4.7	809	1735	Uneven-aged multi-layered forest of Norway spruce (*Picea abies* averagely, 190 and 30 years)	Montagnani et. al.(2009)
Brasschaat (BE-Bra)	51.30/4.52	MF	2001–2010	9.8	750	16	Mixed, uneven-aged multi-layered forest of Scots pine (*Pinus sylvestris*, averagely 72 years) and Pedunculate oak (*Quercus robur*, averagely 65 years)	Gielen et al. (2013)

Table 2. IMV and IAV NRMSE for the analyzed sites. Each specific IMV distribution was tested for normality goodness of fit (N = normal distribution, P = non normal distribution). A test for equivalence of central tendency and variance was performed between IMV_{MD} and IMV_{EC} values. (na) refers to the case of sites with inconsistent distributions (one normal, one not normal distributed). (*) marks refer to the acceptance of the null hypothesis that the two distributions are equivalent for the specific statistic ($\alpha = 0.05$). ECT stands for "Equivalence for Central Tendency"; EV for "Equivalence for Variance".
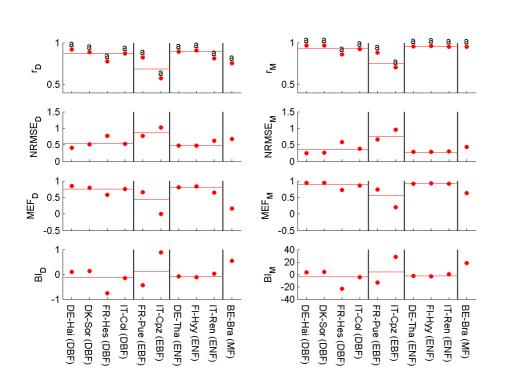
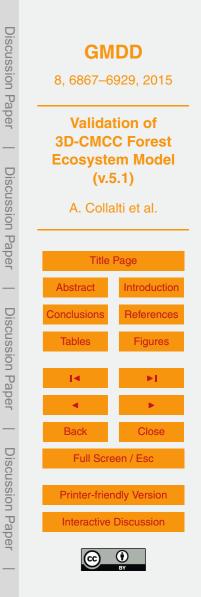|		DE-Hai	DK-Sor	FR-Hes	IT-Col	FR-Pue	IT-Cpz	DE-Tha (1S)	FI-Hyy	IT-Ren (2L-2C)	BE-Bra (P_Q-3L)
NRMSE	IAVs	2.4	1.8	1.3	0.3	0.6	1.1	1.0	2.7	1.3	0.9
NRMSE	IMVs	1.7	2.7	1.1	0.6	1.1	1.2	1.1	1.2	1.0	0.5
ECT	*p* value	1.00*[N]	0.12*[N]	0.54*[N]	0.00[N]	0.15*[N]	1.00*[na]	1.00*[P]	0.04[N]	0.88*[P]	0.85*[N]
EV	*p* value	0.53*[N]	0.00[N]	0.00[N]	0.46*[N]	0.00[N]	0.02[na]	0.78*[P]	0.00[N]	0.27*[P]	0.01[N]

Table 3. Performance statistics (r, NRMSE, MEF, Bi) are reported as derived from daily and monthly series of GPP_{EC} and GPP_{MD} values over long-term annual scale, for the different forest structure simulations. The (*) refers to p value < 0.0001 in correlation between GPP_{EC} and GPP_{MD} data. In addition, long term average of annual GPP_{MD} and GPP_{EC} values ($g\,C\,m^{-2}\,yr^{-1}$) for the different forest structures are shown.

Site	Model Set-up code	Daily				Monthly				Yearly	
		r	NRMSE	MEF	Bi $g\,C\,m^{-2}\,d^{-1}$	r	NRMSE	MEF	Bi $g\,C\,m^{-2}\,month^{-1}$	GPP_{MD} $g\,C\,m^{-2}\,yr^{-1}$	GPP_{EC} $g\,C\,m^{-2}\,yr^{-1}$
BE-Bra	P	0.72*	0.73	0.47	−0.7	0.86*	0.55	0.7	−9.81	1003	
	Q_3L	0.76*	0.91	0.18	0.14	0.84*	0.71	0.49	2.67	1105	
	Q_2L	0.74*	0.89	0.21	0.0	0.86*	0.74	0.45	9.38	1179	
	Q_1L	0.75*	0.95	0.01	0.25	0.86*	0.68	0.53	6.86	1147	1112
	P_Q-3L	0.77*	0.64	0.58	0.2	0.91*	0.42	0.82	−3.9	1169	
	P_Q-2L	0.75*	0.67	0.55	−0.41	0.91*	0.44	0.81	5.97	1037	
	P_Q-1L	0.75*	0.66	0.56	−0.37	0.91*	0.68	0.53	6.86	1056	
IT-Ren	2L_2C	0.81*	0.62	0.61	−1.27	0.95*	0.3	0.91	−39.3	1348	1362
	1L_1C	0.83*	0.85	0.27	1.27	0.96*	0.61	0.62	38.9	1950	
DE-Tha	1S	0.89*	0.48	0.8	−0.1	0.96*	0.29	0.91	−2.5	1898	1869
	2S	0.89*	0.46	0.79	−0.08	0.95*	0.27	0.93	−2.59	1837	

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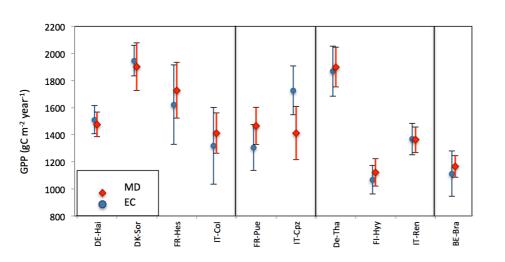Printer-friendly Version

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Figure 1. 3D-CMCC FEM performance indices at different time scales; daily (on the left) and daily aggregated to month (on the right). DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation (see text). The red horizontal line refers to the value calculated for the whole data aggregated per IGBP vegetation class. In case of r, "a" stands for p value < 0.001.
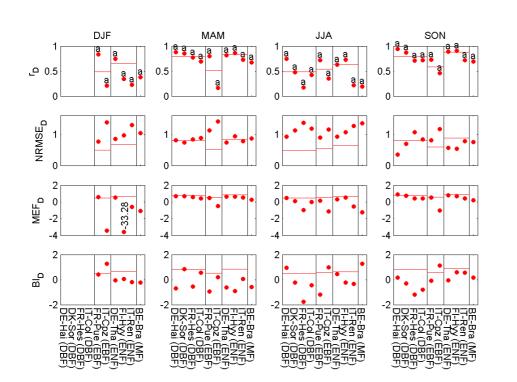
Figure 2. Taylor diagrams for daily **(a)**, daily aggregated to month **(b)** GPP evaluated by: the deviation of model results from observations (REF) in terms of normalized standard deviation of observations, represented by the distance from the site point to the point on the x axis identified as reference (REF); the difference of model normalized standard deviation from that of observations, represented by the distance of the site point with respect to the quarter arc crossing REF; and the correlation, given by the azimuthal position of the site point to the x axis. The sites are numbered in ascending order as follows: Eq. (1) DE-Hai, Eq. (2) DK-Sor, Eq. (3) FR-Hes, Eq. (4) IT-Col, (5) FR-Pue, (6) IT-Cpz, (7) DE-Tha, (8) FI-Hyy, (9) IT-Ren, (10) BE-Bra. Colors refer to different IGBPs: DBF (yellow), EBF (orange), ENF (light-blue), MF (green).
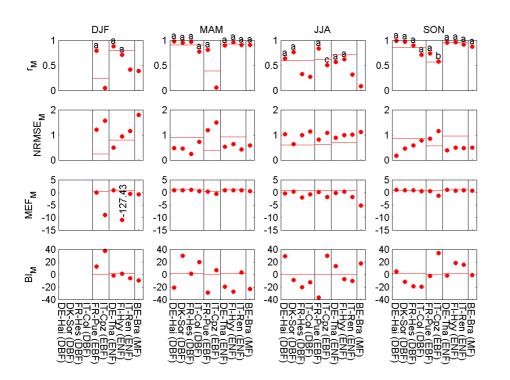
Figure 3. Distributions of annual GPP ($gC\,m^{-2}\,yr^{-1}$). MD (red) are model results, EC (blue) measured by eddy covariance. The vertical bars represent ± 1 standard deviation. DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation (see text).

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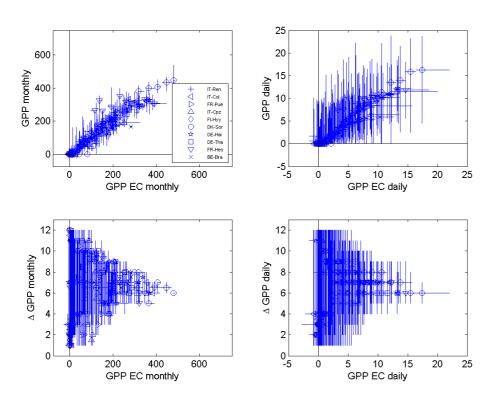Printer-friendly Version

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Figure 4. 3D-CMCC FEM performances indices of daily (D) GPP at different seasons. DE-Tha refers to the DE-Tha (1S) simulation, IT-Ren to the (2L-2C), BE-Bra to the (P_Q-3L). The red horizontal line refers to the value calculated for the whole data aggregated per IGBP. Strongly negative MEF are represented out of scale, but flanked with their respective numerical value. In case of r, "a" stands for p value < 0.001.
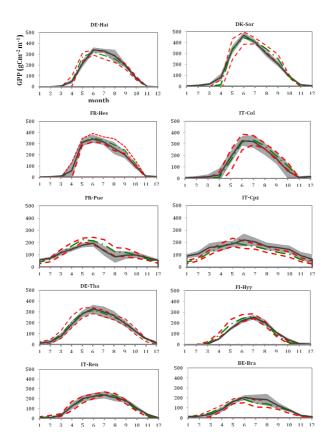
Figure 5. 3D-CMCC FEM performances indices of daily GPP aggregated to months (M) at different seasons. DE-Tha refers to the DE-Tha (1S) simulation, IT-Ren to the (2L-2C), BE-Bra to the (P_Q-3L). The red horizontal line refers to the value calculated for the whole data aggregated per IGBP. Strongly negative MEF are represented out of scale, but flanked with their respective numerical value. In case of r, "a" stands for p value < 0.001, "b" for p value < 0.01, "c" for p value < 0.05.
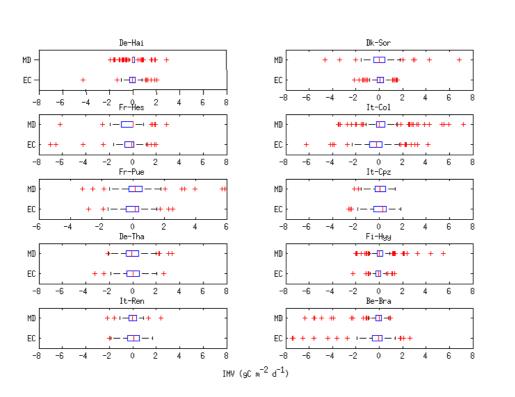
Figure 6. Comparison between GPP_{MD} and GPP_{EC} data. The top plots show the average $GPP_{EC} : GPP_{MD}$ correlation for daily ($g\,C\,m^{-2}\,d^{-1}$) and monthly ($g\,C\,m^{-2}\,month^{-1}$) data. The bottom plots show absolute difference range between GPP_{MD} and GPP_{EC} while increasing GPP_{EC} values. Negative values are excluded because of model assumptions. DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation.
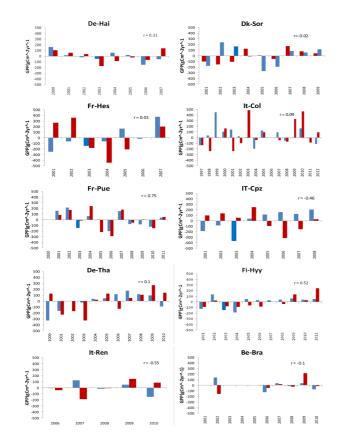
Figure 7. Seasonal (monthly) cycle of GPP across the ten sites. The grey line and margins of the grey area represent long-term average of monthly GPP_{EC} ($gC m^{-2} month^{-1}$) and its ± 1 standard deviation, respectively. The green and red dashed lines represent the long-term average of monthly GPP_{MD} ($gC m^{-2} month^{-1}$) and its ± 1 standard deviation, respectively. DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation (see text).

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Figure 8. Distribution of the magnitude for the inter-monthly variability values (IMVs, $g\,C\,m^{-2}\,d^{-1}$) for each specific site, resulted by standard kernel density estimation. DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation. Red vertical line represents the median, the blue box bounds the 25th and 75th percentiles, black dashed lines limit whiskers at the 0.35th and 99.65th percentiles, and red crosses are the outliers.

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Figure 9. Inter-Annual Variability (IAV) based on Keenan et al. (2012). Red and blue bars indicate the observed and modelled IAV values, respectively; *r* values refer to correlation between observed and modelled variations. DE-Tha refers to the 1S simulation, IT-Ren to the 2L_2C simulation, BE-Bra to the P_Q-3L simulation.

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