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Improving the global applicability of the RUSLE model – adjustment of the topographical and rainfall erosivity factors

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Received: 12 February 2015 - Accepted: 6 March 2015 - Published: 19 March 2015

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Published by Copernicus Publications on behalf of the European Geosciences Union.

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Large uncertainties exist in estimated rates and the extent of soil erosion by surface runoff on a global scale, and this limits our understanding of the global impact that soil erosion might have on agriculture and climate. The Revised Universal Soil Loss Equation (RUSLE) model is due to its simple structure and empirical basis a frequently used tool in estimating average annual soil erosion rates at regional to global scales. However, large spatial scale applications often rely on coarse data input, which is not compatible with the local scale at which the model is parameterized. This study aimed at providing the first steps in improving the global applicability of the RUSLE model in order to derive more accurate global soil erosion rates.

We adjusted the topographical and rainfall erosivity factors of the RUSLE model and compared the resulting soil erosion rates to extensive empirical databases on soil erosion from the USA and Europe. Adjusting the topographical factor required scaling of slope according to the fractal method, which resulted in improved topographical detail in a coarse resolution global digital elevation model.

Applying the linear multiple regression method to adjust rainfall erosivity for various climate zones resulted in values that are in good comparison with high resolution erosivity data for different regions. However, this method needs to be extended to tropical climates, for which erosivity is biased due to the lack of high resolution erosivity data.

After applying the adjusted and the unadjusted versions of the RUSLE model on a global scale we find that the adjusted RUSLE model not only shows a global higher mean soil erosion rate but also more variability in the soil erosion rates. Comparison to empirical datasets of the USA and Europe shows that the adjusted RUSLE model is able to decrease the very high erosion rates in hilly regions that are observed in the unadjusted RUSLE model results. Although there are still some regional differences with the empirical databases, the results indicate that the methods used here seem to be a promising tool in improving the applicability of the RUSLE model on a coarse resolution on global scale.

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For the last centuries to millennia soil erosion by surface runoff is being accelerated globally due to human activities, such as deforestation and agricultural practices (Bork and Lang, 2003). Accelerated soil erosion is a process that triggers land degradation in the form of nutrient loss, a decrease in the effective root depth, water imbalance in the root zone and finally also productivity reduction (Yang et al., 2003). It is widely recognized that soil erosion is a major threat to sustainable agriculture and food production across the globe for many decades. These effects of soil erosion are currently exacerbated by the global population growth and climatic changes. Organizations such as the United Nations Convention to Combat Desertification (UNCCD) try to address this problem by stating a new goal for Rio +20 of zero land degradation (UNCCD, 2012).

Another aspect underpinning the relevance of soil erosion on the global scale is the effect of soil erosion on the global nutrient cycles. Recently, the biogeochemical components of Earth System Models (ESMs) became increasingly important in predicting the global future climate. Not only the global carbon cycle but also other nutrient cycles such as the nitrogen (N) and phosphorous (P) cycles cannot be neglected in ESMs anymore (Goll et al., 2012; Gruber and Galloway, 2008; Reich et al., 2006). Soil erosion may have a significant impact on these nutrient cycles through lateral fluxes of sediment, but the impact on the global scale is still largely unknown. For example, Quinton et al. (2010) showed that erosion can significantly alter the nutrient and carbon cycling and result in lateral fluxes of nutrients that are similar in magnitude as fluxes induced by fertilizer application and crop removal. Regnier et al. (2013) looked at the effect of human induced lateral fluxes of carbon from land to ocean and concluded that human perturbations, which include soil erosion, may have enhanced the carbon export from soils to inland waters.

In general, the effect of soil erosion on the global carbon cycle has received considerable attention after the pioneering work of Stallard (1998), who proposed that global soil erosion can result in sequestration of carbon by soils. After his work, the effect of

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soil erosion on the carbon cycle has been studied extensively, but there remains a large uncertainty in the effect of soil erosion on the carbon cycle. For example, several recent global assessments of the influence of soil erosion on the carbon cycle indicate a large uncertainty with a range from a source of 0.37 to 1 Pg Cyr⁻¹ to a net uptake or sink of 0.56 to 1 Pg Cyr⁻¹ (van Oost et al., 2007). Thus, in order to better constrain the global carbon budget and to identify optimal management strategies for land use, it is essential to have accurate estimates of soil erosion and its variability on a global scale.

Currently, however, there exists a large uncertainty in the global soil erosion rates as can be seen from recent studies that show rates between 20 and 200 Pg yr⁻¹ (Doetterl et al., 2012). This indicates that modelling soil erosion on a global scale is still a difficult task due to the very high spatial and temporal variability of soil erosion. Different approaches were previously applied to estimate soil erosion on a large or global scale. Most of these approaches are based on extrapolated data from agricultural plots, sediment yield or extrapolated river sediment estimates (Milliman and Syvitski, 1992; Stallard, 1998; Lal, 2003; Hooke, 2000; Pimentel et al., 1995; Wilkinson and McElroy, 2007). An alternative approach is based on the use of soil erosion models. One of the most applied models to estimate soil erosion on a large spatial scale is the semiempirical/process-based Revised Universal Soil Loss Equation (RUSLE) model (Renard et al., 1997). This model stems from the original Universal Soil Loss Equation (USLE) model developed by USDA (USA Department of Agriculture), which is based on a large set of experiments on soil loss due to water erosion from agricultural plots in the United States (USA). These experiments covered a large variety of agricultural practices, soil types and climatic conditions, making it a potentially suitable tool on a regional to global scale. The RUSLE model predicts the average annual soil erosion rates by rainfall and is formulated as a product of a rainfall erosivity factor (R), a slope steepness factor (S), a slope length factor (L), a soil erodibility factor (K), a crop cover factor (C) and a support practice factor (P). The RUSLE model was first applied on a global scale by Yang et al. (2003) and Ito (2007) for estimating the global soil erosion potential and various limitations related to applying the RUSLE model on the global scale.

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Firstly, the model is originally developed to be applicable on the agricultural plot scale, which is not compatible with the coarse spatial scale of global datasets on soil erosion influencing factors such as precipitation, elevation, land-use and soil characteristics. Secondly, the RUSLE and USLE models were parameterized for environmental conditions of the United States (USA), and are thus not directly applicable to other areas in the world. Thirdly, only sheet and rill erosion are considered, and finally the RUSLE model does not contain sediment deposition and sediment transport terms, which are

The RUSLE model is to our knowledge one of the few erosion models that has the potential to be applied on a global scale due to its simple structure and empirical basis. Therefore it is of key importance to address the above mentioned limitations first.

closely linked to soil erosion.

To address the first two limitations, Van Oost et al. (2007) presented in their work a modified version of the USLE model for application on agricultural areas on a global scale. They based their model on large-scale experimental soil erosion data from the USA (National Resource Inventory, NRI database, USDA, 2000) and Europe, by deriving reference factors for soil erosion and for certain RUSLE parameters. They also introduced a procedure to scale slope, which is an important parameter in the topographical factors S and L of the RUSLE model. In this scaling procedure slope was scaled from the GTOPO30 1 km resolution digital elevation model (USGS, 1996) to the coarser resolution of the erosion model based on high resolution OS Ordnance (10 m resolution) and SRTM data on elevation (90 m resolution, International Centre for Tropical Agriculture (CIAT), 2004) for England and Wales.

Doetterl et al. (2012) showed that together with the S factor, the rainfall erosivity or R factor explain up to 75% of the erosion variability across agricultural areas at the large watershed scale, as these factors represent the triggers for soil erosion by providing energy for soil to erode. The S and R factors can also be seen as the natural components of the RUSLE model, as they have very little or no modification by human activities (Angulo-Martínez et al., 2009) apart from indirect effects on precipitation and extreme events due to anthropogenic climate change that are included in the R factor.

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In this way they represent the natural environmental constraints to soil erosion that are important to capture before the effect of human activities on soil erosion through land use change can be investigated. Previous studies on global soil erosion estimated the global R factor based on the total annual precipitation (Renard and Freimund, 1994), due to the lack of high resolution precipitation intensity on a global scale. However, high resolution precipitation intensity is an important explaining parameter of the R factor and therefore, the applicability of this method is limited.

The overall objective of this study is to extend the applicability of the RUSLE model to a coarse resolution at a global scale, in order to enable future studies on the effects of soil erosion for the past, current and future climate. To this end, we develop generally applicable methods that improve the estimation of slope and climatic factors from coarse resolution global datasets. These methods should not only be applicable across agricultural areas as in the studies of Van Oost et al. (2007) and Doetterl et al. (2012), but also across non-agricultural areas. We adjust the S factor to the coarse resolution of the global scale based on the scaling of slope according to the fractal method. The adjustment of the R factor to the global scale is based on globally applicable regression equations for different climate zones that include parameters for precipitation, elevation and the simple precipitation intensity. This approach is validated using several high resolution datasets on the R factor Finally, the effects of these adjustments of both factors on global soil erosion rates are investigated separately and tested against independent estimates of soil erosion from high resolution and high precision datasets of Europe and the USA.

Adjustment of the topographical factor

Scaling slope according to the fractal method

The topographical factors of RUSLE are the slope steepness factor (S) and a slope length factor (L). The S factor is generally computed by the continuous function of

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$$S = 1.5 + \frac{17}{1 + \rho(2.3 - 6.1 \cdot \sin \theta)} \tag{1}$$

And the *L* factor is computed according to Renard et al. (1997):

$$L = \left(\frac{I}{22.13}\right)^m \tag{2}$$

5 where:

$$m = \frac{F}{1+F}$$
 and $F = \frac{(\sin\theta/0.0896)}{3 \cdot (\sin\theta)^{0.8} + 0.56}$ (3)

in which θ is the slope and I is the slope length in meters.

As seen in the equations of the L and S factors, slope is a crucial parameter and thus an accurate estimation is essential in deriving accurate estimates of the L and S factors and finally also the soil erosion rates. For an accurate estimation of slope the input elevation data from digital elevation models (DEMs) should capture the detailed spatial variability in elevation. However, global DEMs are often too coarse to capture the detailed topography because of the surface smoothening effect. To account for this problem it is assumed that topography is fractal. Following Klinkenberg and Goodchild (1992) and Zhang et al. (1999), slope can be expressed as a function of the spatial scale by applying the variogram equation. The variogram equation is used to approximate the fractal dimension of topography and is expressed as follows:

$$(Z_p - Z_q)^2 = k d_{pq}^{4-2D} \tag{4}$$

so that:

$$\frac{|Z_{p} - Z_{q}|}{d_{pq}} = \alpha d_{pq}^{1-D} \tag{5}$$

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where Z_p and Z_q are the elevations at points p and q, d_{pq} is the distance between pand q, k is a constant, $\alpha = k^{0.5}$ and D is the fractal dimension. Because the left side of Eq. (5) represents the slope, it can be assumed that the slope θ is related to the spatial scale or the grid size d in:

$$\theta = \alpha d^{1-D} \tag{6}$$

This result implies that by calculating the fractal properties (D and α) Eq. (6) can be used to calculate slope at any specified scale d. The local fractal dimension describes the roughness of the topography while the local value of α is related to the concept of lacunarity, which is a measure of the size of "gaps" (valleys and plains) in the topography (Zhang et al., 2002). To estimate the spatial variations of the fractal dimension D and the fractal coefficient α , Zhang et al. (1999) proposed to relate these parameters to the standard deviation of elevation. Hereby it is assumed that the standard deviation of elevation does not change much with the DEM resolution. D is then calculated as a function of the standard deviation (σ) in a 3 × 3 pixels moving window as proposed by Zhang et al. (1999):

$$D = 1.13589 + 0.08452 \ln \sigma \tag{7}$$

To estimate α we used the modified approach by Pradhan et al. (2006), who derived α directly from the steepest slope in a 3 \times 3 pixels moving window, called α_{steepest} in the following. Having obtained α_{steepest} and D from a grid at a given resolution, the scaled slope (θ_{scaled}) for a target grid resolution d_{scaled} is obtained by:

$$\theta_{\text{scaled}} = \alpha_{\text{steepest}} q_{\text{scaled}}^{1-D}$$
 (8)

Pradhan et al. (2006) also showed that in their case study the ideal target resolution for downscaling slope was 150 m due the breakdown of the unifractal concept at very fine scales, which they showed to happen at a scale of 50 m. Altogether, this fractal method shows that a high resolution slope can be obtained from a low resolution DEM as needed by the RUSLE model.

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In this study, we investigate the performance of the fractal method on a global scale using different global DEMs as a starting point. The target resolution of downscaling is put to 150 m (about 5 arc-second) according to Pradhan et al. (2006). It should be noted that the original spatial scale that the RUSLE and USLE models are operating on is usually between 10 and 100 m, which indicates that the 150 m target resolution may be still too coarse for a correct representation of slope. The DEMs that are used here are given in Table 1.

As reported in previous studies (Zhang et al., 1999; Chang and Tsai, 1991; Zhang and Montgomery, 1994), the average slope decreases with decreasing DEM resolution. This confirms the expectation of loss of detail in topography at lower DEM resolutions. A large difference is found between the unscaled global average slope of the 5 arc-minute and the 30 arc-second DEMs, which is in the order of 0.017 m m⁻¹ or 74 % (Table 2). After applying the fractal method, the scaled slopes of the DEMs at 150 m target resolution are all increased significantly compared to the unscaled slopes (Fig. 1). However, there is still a difference of about 0.05 mm⁻¹ or 8.5 % between the scaled slopes of the 5 arc-minute and the 30 arc-second DEMs (Table 2). This difference can be attributed to several factors. One factor could be the underlying assumption that the standard deviation of elevation (σ) is independent of the DEM resolution. Although σ does not change much when considering different resolutions, there is still a general decrease in mean global σ when going from the 5 arc-minute to the 30 arc-second DEM (Table 2). Due to the dependence of the fractal dimension D on σ (Zhang et al., 1999), a decrease of σ leads to a decrease in D and therefore an increase in the scaled slope. Other factors that could play a role here are the dependence of $\alpha_{
m steepest}$ on the steepest slope and the breakdown of the fractal method at certain scales and in certain environments. Zhang et al. (1999) mentioned that the scaling properties of slope are affected in very coarse resolution DEMs if σ changes considerably. On the other hand Pradhan et al. (2006) mentioned the breakdown of the fractal method at very fine

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After applying the fractal method on a 30 arc-second resolution DEM, the scaled slope shows a clear increase in detail, while the unscaled slope shows a strong smoothening effect (Fig. 2a and b). It is found that after scaling the slope values range from 0 to 85° and are less than 2° in 80% of the area. In contrast, all slope values are less than 45° and range between 0 and 2° in 89 % of this area when slope is computed directly from the 30 arc-second DEM.

The scaled slope from the 30 arc-second DEM will be used in this study to estimate the global soil erosion rates by the RUSLE model.

Adjustment of the rainfall erosivity factor

The approach by Renard and Freimund (1994)

Rainfall erosivity (R factor) is described by Hudson (1971) and Wischmeier and Smith (1978) as the result of the transfer of the kinetic energy of raindrops to the soil surface. This causes a detachment of soil and the downslope transport of the soil particles depending on the amount of energy, rainfall intensity, soil type and cover, topography and management (Da Silva, 2004). The original method of calculating erosivity is described by Wischmeier and Smith (1978) and Renard et al. (1997) as:

$$R = \frac{1}{n} \cdot \sum_{j=1}^{n} \sum_{k=1}^{m_j} (\mathsf{El}_{30})_k \tag{9}$$

where n is the number of years of records, m_i is the number of storms of a given year j, and El₃₀ is the rainfall erosivity index of a storm k. The event's rainfall erosivity index 3000

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$$\mathsf{EI}_{30} = I_{30} \cdot \sum_{r=1}^{m} e_r v_r \tag{10}$$

where e_r and v_r are, respectively, the unit rainfall energy (MJ ha⁻¹ mm⁻¹) and the rainfall depth (mm) during a time period r, and l_{30} is the maximum rainfall intensity during a time period of 30 min (mm h⁻¹). The unit rainfall energy e_r , is calculated for each time period as:

$$e_r = 0.29 \cdot (1 - 0.72 \cdot e^{-0.05 \cdot i_r}) \tag{11}$$

where i_r is the rainfall intensity during the time period (mm h⁻¹).

The information needed to calculate the *R* factor according to the method of Wischmeier and Smith (1978) is difficult to obtain on a large spatial scale or in remote areas. Therefore, different studies have been done on deriving regression equations for the *R* factor (Angulo-Martinez et al., 2009; Meusburger et al., 2012; Goovaerts, 1999; Diodato and Bellocchi, 2010). Most of these studies, however, concentrate on a specific area and can therefore not be implemented on the global scale. Studies on global soil erosion estimation by the RUSLE model or a modified version of it (Doetterl et al., 2012, van Oost et al., 2007; Montgomery et al., 2007; Yang et al., 2003) have all used the method of Renard and Freimund (1994). Renard and Freimund related the *R* factor to the total annual precipitation based on erosivity data available for 155 stations in the USA, shown in the following equation:

$$R = 0.0483 \cdot P^{1.61}, P \le 850 \,\text{mm}$$

 $R = 587.8 - 1.219P + 0.004105 \cdot P^2, P > 850 \,\text{mm}$ (12)

To test how this method performs globally, first the *R* factor was calculated in this study according to the method of Renard and Freimund (Eq. 12) using the 0.25° resolution 3001

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annual precipitation data from the GPCC product (Table 1). Then, three regions were selected to validate the resulting *R* values and their variability: the USA (EPA, 2001), Switzerland (Meusburger et al., 2011), and the Ebro basin in Spain (Angulo-Martinez et al., 2009). For these regions high resolution erosivity data are available obtained from pluviographic data from local meteorological stations across the whole region.

Figure 3 shows that the *R* values computed with the Renard and Freimund method strongly overestimate *R* when compared to the high resolution *R* data of the selected regions. For the USA the *R* factor of Renard and Freimund shows an overall overestimation for western USA and for a large part of eastern USA when compared to the high resolution *R* (Table 6 and Fig. 3a). Especially a strong overestimation is seen for the north-west coast of the USA. This region is known to have complex rainfall patterns due to the presence of mountains and high local precipitation intensities with frequent snow fall (Cooper, 2011). It should be noted that the USA is not a completely suited case study for testing the *R* values computed with the Renard and Freimund method, as this method is based on data from stations in the USA. The available high resolution data on the *R* factor from Switzerland and the Ebro basin are better suited for an independent validation.

For Switzerland, which has a complex precipitation variability influenced by the relief of the Alps (Meusburger et al., 2012), the R factor of Renard and Freimund shows a strong overall overestimation when compared to the observed or high resolution R values (Table 6 and Fig. 3b). For the Ebro basin located in Spain, the observed R data were available for the period 1997–2006 from Angulo-Martinez et al., 2009. Also here the method of Renard and Freimund overestimates the R factor and is not able to model the high spatial variability of the R data (Table 6 and Fig. 3c).

3.2 The linear multiple regression approach using environmental factors

To better represent the R factor on a global scale, the R estimation was based on the updated Köppen–Geiger climate classification (Fig. 4). The Köppen–Geiger climate classification is a globally climate classification and is based on the vegetation distri-

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bution connected to annual cycles of precipitation and temperature (Lohmann et al., 1993). The reason for this approach is that this classification system includes annual cycles of precipitation and is thus indirectly related to precipitation intensity. Based on this it is possible to derive regression equations for the *R* factor that are applicable for each individual climate zone. This provides a basis to calculate *R* with coarse resolution data on a globally scale.

As a basis for deriving the regression equations for the R factor for most climate zones the high resolution R maps of the USA from EPA (2001) were used. The USA covers most of the world's climate zones and is also the largest region with available high resolution R data. Linear multiple regression was used to adjust R:

$$\log(R_i) = \beta_0 + \sum_{j=1}^n \beta_{i_j} \cdot \log\left(X_{i_j}\right) + \varepsilon_i, \quad \text{for} \quad i = 1, 2, \dots, n$$
 (13)

where X is the independent explanatory variable, j is the number of explanatory variables, β is a constant and ε is the residual.

The regression operates on one or more of the following parameters (X_j): total annual precipitation (GPCC 0.25° product), mean elevation (ETOPO 5 DEM), and the simple precipitation intensity index, SDII. It should be mentioned that the SDII was only available on a very coarse resolution of 2.5° resolution for certain regions on earth, such as parts of Europe and the USA. The SDII is calculated as the daily precipitation amount on wet days (>=1 mm) in a certain time period divided by the number of wet days in that period. Previous studies that performed regression of R showed that precipitation and elevation were in most cases the only explanatory variables. Here, the SDII is added as it is a simple representation of precipitation intensity, which is an important explaining variable of R. The precipitation and SDII datasets were rescaled to a 5 arc-minute resolution (corresponding to 0.0833° resolution) to match the Köppen–Geiger climate classification data that was available at the resolution of 6 arc-minute (corresponding to 0.1°). Furthermore, high resolution erosivity data from Switzerland (Meusburger et al., 2011) and annual precipitation from the GPCC 0.5° product were

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used to derive the regression equations for R for the polar (E) climates, which are not present in the USA. For the rest of the climate zones not present in the USA it was difficult to obtain high resolution erosivity data. Therefore, for those climate zones the method of Renard and Freimund was maintained to calculate erosivity. Also, if no clear improvement of the R factor is found when using the new regression equations for a specific climate zone, the R factor of Renard and Freimund is kept. All datasets for deriving the *R* factor are described in Table 1.

Application of the linear multiple regression method on a global scale

Tables 3 and 4 show the resulting regression equations for climate zones for which initially a low correlation was found between the R values calculated by the method of Renard and Freimund and the high resolution or observed R values from the maps of EPA (2001) and Meusburger et al. (2011). Figure 5 shows for each addressed climate zone how the method of Renard and Freimund and the new regression equations compare to the observed R of the USA. For the Ds climate zones the new regression equations showed only a slight improvement as compared to the method of Renard and Freimund. Also for the E climate zones the new regression equations still showed a significant bias. However, they performed much better compared to the method of Renard and Freimund. For most of the addressed climate zones the simple precipitation intensity index (SDII) explained a large part of the variability in the R factor. The elevation played a smaller role here. Elevation can be an important explaining variable in regions with a high elevation variability, which then affects the precipitation intensity. Furthermore, from Table 3 it can be concluded that the R factor in f climate zones, which have no dry season, can be easily explained by the total annual precipitation and the SDII. Dry climate zones, especially dry summer climate zones showed a weaker correlation, which is most probably due to the fact that the SDII is too coarse to explain the variability in the low precipitation intensity in the summer. It is also interesting to see that even though the SDII was derived from a very coarse dataset, it turned out to be still important for deriving more accurate R values. Furthermore, Table 5 showed for

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each addressed climate zone a comparison of the newly computed average R factor with the average observed R factor, and the uncertainty range. The uncertainty range was computed by taking into account the standard deviation of each of the parameters in the regression equations. As mentioned before, the E climate zones showed the largest uncertainty range. The new regression equations significantly improved the R values and spatial variability in the western USA and lead to a mean R factor that was closer to the data mean (Table 6 and Fig. 6a). Although the new regression equations showed a bias for the E climate zones (the minimum and maximum R were not captured), the resulting mean R for Switzerland showed a strong improvement (Table 6 and Fig. 6b). Furthermore, the variability in the estimated R compared well with the variability of the observed R. As the observed R values of the USA and Switzerland were used to derive the regression equations, a third case study, the Ebro basin in Spain, provided an important independent validation. For the Ebro basin, the new regression equations not only improved the overall mean but also captured the minimum R values better, resulting in an improved representation of the R variability (Table 6 and Fig. 6c). In Fig. 6c, however, there was a clear pattern separation in the newly computed R values, which was due to the fact that the SDII data were not available for part of the Ebro basin. As mentioned before, SDII is an important explaining parameter in the regression equations for most of the addressed climate zones.

Figure 7a showed the global patterns of the estimated R from respectively the method of Renard and Freimund and the new regression equations. Figure 7b showed a difference plot between the estimated R with the method of Renard and Freimund and the R estimated with the new regression equations. The new regression equations significantly reduced the R values in most regions. However, the tropical regions still showed unrealistic high R values (maximum R values go up to $1 \times 10^5 \,\mathrm{MJ\,mm\,ha^{-1}\,h^{-1}\,yr^{-1}}$). This is because the R factor was not adjusted for the tropical climate zones due to the lack of high resolution R data. Oliveira et al. (2012) found for the R factor in Brazil that the maximum R values for the tropical climate zones reach 22 452 MJ mm ha $^{-1}$ h $^{-1}$ yr $^{-1}$.

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Finally, it should be noted that the purpose of the adjusting methods in this study is to capture more accurately the large scale mean erosion rates rather than the extremes. Therefore, even though the new regression equations are still not accurate enough for certain climate zones, it is important that the mean R factor is represented well. The approach for adjusting the R factor also showed that even though there is no high temporal resolution precipitation intensity data available on a global scale, the R factor can still be represented well for most climate zones on a large spatial scale by using other parameters, such as elevation, and especially a representative of precipitation intensity, such as the SDII. The SDII played an important role here as it improved the estimation of the R factor significantly, even though data was only available at a very low resolution as compared to the other datasets of precipitation, elevation and climate zone classification.

Global application of the adjusted RUSLE model

Computation of the soil erodibility and crop cover factors

In the following the consequences of the new parameterizations of the S and R factors for global soil erosion rates are demonstrated. First, the other individual RUSLE factors, soil erodibility (K) and crop cover (C) needed to be computed. Estimations of the K factor were based on soil data from the gridded 30 arc-second Global Soil Dataset for use in Earth System Models (GSCE). GSCE is based on the Harmonized World Soil database (HWSD) and various other regional and national soil databases (Shangquan et al., 2014). The method of Torri et al. (1997) was then used to estimate the K factor. Volcanic soils were given a K factor of $0.08 \, \text{thahha}^{-1} \, \text{MJ}^{-1} \, \text{mm}^{-1}$, as these soil types are usually very vulnerable for soil erosion and the K values are beyond the range predicted by the method of Torri et al. (1997) (van der Knijff et al., 1999). To account for the effect of stoniness on soil erosion a combination of the methods used by Cerdan et al. (2010) and Doetterl et al. (2012) was applied, who base their methods on

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the original method of Poesen et al. (1994). For non-agricultural areas the method of Cerdan et al. (2010) was used where they reduced the total erosion by 30 % for areas with a gravel percentage larger or equal to 30 %. For agricultural and grassland areas the method of Doetterl et al. (2012) was used, where erosion was reduced by 80 % in areas where the gravel percentage exceeded 12 %.

The C factor was calculated according to the method of De Jong et al. (1998), using 0.25° Normalized Difference Vegetation Index (NDVI) and land use data for the year 2002. An important limitation of this method is the fact that in winter the C factor is estimated too large (van der Knijff et al., 1999). This is because the equation does not include the effects of mulch, decaying biomass and other surface cover reducing soil erosion. To prevent the C factor of being too large maximum C values for forest and grassland of 0.01 and 0.05 for pasture were used. Doetterl et al. (2012) showed that the slope length (L) and support practice (P) factors do not contribute significantly to the variation in soil erosion at the continental scale. Also, data on these factors are very scarce on global scale. Therefore it was decided not to include these factors in the model.

4.2 Computation of global soil erosion and comparison to empirical databases

The RUSLE model with the settings mentioned in the previous paragraph is applied on a 5 arc-minute resolution on a global scale for the present time period (see time resolutions of datasets in Table 1). Global soil erosion rates are calculated for four different versions of the RUSLE model: (a) the unadjusted RUSLE, (b) RUSLE with only an adjusted S factor, (c) RUSLE with only an adjusted S factor and (d) the adjusted RUSLE (all adjustments included). The global mean soil erosion rate for the adjusted RUSLE is found to be $7 \, \text{tha}^{-1} \, \text{yr}^{-1}$. When including the uncertainty arising from applying the linear multiple regression method, the mean global soil erosion rate differs between 6 and $18 \, \text{tha}^{-1} \, \text{yr}^{-1}$. Furthermore, the RUSLE version with only an adjusted S factor shows the highest mean global soil erosion rate, while the lowest rate is found for the RUSLE version with only the adjusted S factor (Table 7). This indicates that these two

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factors balance each other, and that it is important to have a correct representation of all the RUSLE factors on a global scale in order to predict reliable erosion rates.

In this study the *K* and *C* factors are not tested and adjusted for a coarse resolution at the global scale and thus validation with existing empirical databases on soil erosion is not fully justified. However, to test if the global erosion rates are in an acceptable range, they are compared to erosion estimates from the NRI database for the USA and erosion estimates from the study of Cerdan et al. (2010) for Europe. These are to our knowledge the only large scale high resolution empirical databases on soil erosion.

The NRI database contains USLE erosion estimates for the year 1997, which are available at the HUC4 watershed level. After aggregating the resulting erosion rates from the adjusted and unadjusted RUSLE models to the HUC4 watershed level, the results showed that the mean erosion rates from the adjusted RUSLE model come closer to that of the NRI database (Table 8 and Fig. 8a). However, the maximum observed mean HUC4 soil erosion rate from the adjusted RUSLE was twice as high as compared to the NRI database. This maximum is observed in the hilly and relatively wet region on the west coast of the USA. From these results we can conclude that the erosion rates of the adjusted RUSLE fall in the range of observed values, but that there are still some local overestimations.

For Europe, Cerdan et al. (2010) used an extensive database of measured erosion rates on plots under natural rainfall. They extrapolated measured erosion rates to the whole Europe (European Union area) and adjusted them with a topographic correction based on the *L* and *S* factors of RUSLE, and a correction to account for soil stoniness. For comparison, the soil erosion rates from Cerdan et al. (2010) and the RUSLE estimates are aggregated at country level. The performance of the adjusted RUSLE model was not as good for Europe compared to the USA, which is not surprising due to the fact that the RUSLE model is based on soil erosion data of the USA. However, also on the European scale the adjusted RUSLE model performed better than the unadjusted RUSLE model (Table 8 and Fig. 8b). Especially the large erosion rates in the south of Europe as observed in the results of the unadjusted RUSLE model are less extreme

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These biases in erosion rates as seen for the USA and Europe can be attributed to several factors. Firstly, the other RUSLE factors (K and C) and the way they interact with 5 each other are not adjusted to the coarse resolution of the global scale. For example, a possible effect that is usually not captured by the RUSLE model is the location of land use in a certain gridcell. If the land use in a grid cell is located on steep slopes the resulting erosion in that gridcell would be higher than when it would be located in the flatter areas. In this study, however, only mean fractions of land cover and the NDVI are used for each gridcell, which can lead to possible biases in the resulting erosion rates. Secondly, land management is not accounted for in this study, which could introduce an important systematic bias in the soil erosion rates for especially agricultural areas. Furthermore, uncertainties in the coarse resolution land cover/land use, soil and precipitation datasets that are not accounted for, can lead to the model biases. Also, better adjustment of the R factor for climate zones such as the E climate zones, could help improving the overall results. Some biases in the erosion rates can also be attributed to the fact that stepped relief, where flat plateaus are separated by steep slopes, are not well captured by the 150 m target resolution used in the fractal method to scale slope. In this way erosion would be overestimated in these areas. Finally, errors and limitations in the observational datasets can also contribute to the differences between model and observations. The study of Cerdan et al. (2010) on Europe for example used extrapolation of local erosion data to larger areas that could introduce some biases. Also the underlying studies on measured erosion rates used different erosion measuring techniques that can be linked to different observational errors.

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In this study we introduced specific methods to adjust the topographical and rainfall erosivity factors to improve the application of the RUSLE model on a global scale using coarse resolution input data.

Our results show that the fractal method by Zhang et al. (1999) and Pradhan et al. (2006) can be applied on coarse resolution DEMs to improve the resulting slope. Although the slope representation improved after applying this method, the results still show a slight dependence on the original grid resolution. This is attributable to several factors such as the underlying assumption that the standard deviation of elevation (σ) is independent of the DEM resolution, and to the breakdown of the fractal method at certain scales.

We compared the rainfall erosivity calculated by the method of Renard and Freimund to available high resolution or observed erosivity data of the USA, Switzerland and the Ebro basin, and showed overall significant biases. We implemented a linear multiple regression method to adjust erosivity for climate zones of the Köppen-Geiger climate classification system in the USA that showed a bias in erosivity calculated with the method of Renard and Freimund. Using precipitation, elevation and the simple precipitation intensity index as explaining parameters, the resulting adjusted erosivity compares much better to the observed erosivity data for the USA, Switzerland and the Ebro basin. Not only the mean values but also the spatial variability in erosivity is improved. It was surprising to notice that using the rather coarse resolution simple precipitation intensity index in the regression analysis made it possible to explain much of the variability in erosivity. This, once more, underpins the importance of precipitation intensity in erosivity estimation.

After calculating the newly adjusted erosivity on a global scale, it is apparent that the tropical climate zones, for which erosivity was not adjusted, show strong overestimations in some areas when compared to estimated erosivity from previous studies. This shows that adjusting erosivity for the tropical climate zones should be the next step.

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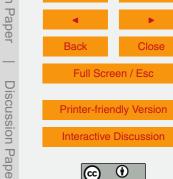
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The challenge is to find enough reliable long term and high resolution erosivity data for those regions.

To investigate how the adjusted topographical and rainfall erosivity factors affect the global soil erosion rates, we applied the adjusted RUSLE model on a global scale and estimate a mean global soil erosion rate of 7 tha⁻¹ yr⁻¹. It is, however, difficult to provide accurate uncertainty estimates to the global erosion rates of this study and to provide a good validation with observations, due to lack of high resolution data on other individual RUSLE factors such as the soil erodibility, slope length and support practice. These RUSLE factors, together with the crop cover factor, which includes the effects of land use, are therefore not adjusted for application on a coarse resolution on global scale.

To test if the soil erosion rates from the adjusted RUSLE model are in a realistic range, we compared the results to the USLE erosion estimates for the USA from the NRI database and the erosion estimates for Europe from the study of Cerdan et al. (2010). The adjusted RUSLE soil erosion rates, which we aggregated to the HUC4 watershed level, show a better comparison with the NRI USLE estimates for the USA then the unadjusted RUSLE erosion rates. For Europe the comparison of the adjusted RUSLE soil erosion rates to the study of Cerdan et al. (2010) were not as good as for the USA. This is not surprising due to the fact that the parameterizations of the RUSLE model are based on soil erosion data of the USA. However, also for Europe, the adjusted RUSLE model performes better than the unadjusted RUSLE model.

We find strong overestimations by the adjusted RUSLE model for hilly regions in the west coast of the USA and for south of Europe. We argue that besides for reasons mentioned before, these biases are due to the fact that the topographical detail may not be enough in some regions to capture the true variability in soil erosion effects by topography. Also erosivity could not be adjusted for some climate zones that are not present in the USA or Switzerland, and needs to be improved for climate zones such as the polar climate zones.

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We conclude that even though there is still much improvement of the RUSLE model possible with respect to topography and erosivity, the methods proposed in this study seem to be promising tools for improving the global applicability of the RUSLE model. A globally applicable version of the RUSLE model together with data on environmental factors from Earth System Models (ESMs) can provide the possibility for future studies to estimate accurate soil erosion rates for the past, current and future time periods.

The article processing charges for this open-access publication have been covered by the Max Planck Society.

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Table 1. List of datasets used in this study.

Category	Dataset	Source	Spatial resolution	Temporal resolution	Variables
DEM	GTOPO Eleva- tion Model	USGS, 1996; Gesch et al. (1999)	30 arc-seconds		elevation
	ETOPO1 Eleva- tion Model	Amante and Eakins (2009)	1 arc-minute		elevation
	ETOPO2 Eleva- tion Model	US Department of Commerce and NOAA (2001)	2 arc-minute		elevation
	ETOPO5 Eleva- tion Model	National Geophysical Data Center/NESDIS/NOAA (1995)	5 arc-minute		elevation
Climate	GPCC 0.5 degree dataset	Schneider et al. (2011)	0.5°	Years 1989–2010	total yearly precipitation
	GPCC 0.25 degree dataset	Meyer-Christoffer et al. (2011)	0.25°	years 1951-2000	total yearly precipitation
	GHCNDEX dataset Köppen–Geiger	CLIMDEX (Donat et al., 2013) Peel et al. (2007)	2.5° 5 arc-minute	years 1951- present	simple precipitation intensity index (SDII) Köppen–Geiger climate
	dataset				classifications
Soil	Global Soil Dataset for use in Earth System Models (GSCE)	Shangguan et al. (2014)	30 arc-seconds		sand, silt and clay fractions, organic matter % gravel %
	Harmonized World Soil Database (HWSD) version 1.2	Nachtergaele et al. (2012)	30 arc-seconds		volcanic soils
Land-cover	GIMMS dataset	ISLSCP II (Tucker et al., 2005; Hall et al., 2006	0.25°	year 2002	Normalized difference vegetation index (NDVI
Land-use	MODIS dataset	ISLSCP II (Friedl et al., 2010; Hall et al., 2006)	0.25°	year 2002	Land use fractions

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Table 2. Fractal parameters and the resulting mean global slopes before and after applying the fractal method on the different DEMs; Increase of slope means the increase of the average global slope of a DEM after applying the fractal method; difference after scaling = $\frac{\theta_{\text{scaled(GTOPO30)}}}{\theta_{\text{scaled(GTOPO30)}}} \cdot 100;$ difference before scaling = $\frac{\theta_{(DEM)} - \theta_{(GTOPO30)}}{\theta_{(GTOPO30)}} \cdot 100.$

DEM	resolution	SD of mean elevation D				θ	$ heta_{ ext{scaled}}$	Increase of θ	difference after scaling	difference before scaling
	arc-minute	m			$\mathrm{m}\mathrm{m}^{-1}$	$\mathrm{m}\mathrm{m}^{-1}$	%	%	%	
GTOPO30	0.5	570	1.32	0.99	0.023	0.059	61	0	0	
ETOPO1	1	530	1.35	1.08	0.016	0.057	71.9	-3.4	-30.4	
ETOPO2	2	549	1.37	1.17	0.011	0.055	80	-6.8	-52.2	
ETOPO5	5	562	1.42	1.25	0.006	0.054	88.9	-8.5	-73.9	

Table 3. Linear multiple regression equations for different climate zones, relating high resolution erosivity (MJ mm ha⁻¹ h⁻¹ yr⁻¹) from the USA with one or more significant parameters: annual total mean precipitation (P), mean elevation (z) and the simple precipitation intensity index (SDII).

Climate zone	Explaining parameters	Regression function – optimal	R^2	Residual standard error
BWk	P, SDII	$R = 0.809 \cdot P^{0.957} + 0.000189 \cdot \text{SDII}^{6.285}$		
BSh	P, SDII	$\log R = -7.72 + 1.595 \cdot \log P + 2.068 \cdot \log \text{SDII}$	0.97	0.22
BSk	P, SDII, Z	$\log R = 0.0793 + 0.887 \cdot \log P + 1.892 \cdot \log \text{SDII} - 0.429 \cdot \log Z$	0.89	0.35
Csb	Р	$R = 98.35 + 0.000355 \cdot P^{1.987}$		0.16
Cfa	P, SDII, Z	$\log R = 0.524 + 0.462 \cdot \log P + 1.97 \cdot \log \text{SDII} - 0.106 \cdot \log Z$	0.89	0.11
Cfb	P, SDII	$\log R = 4.853 + 0.676 \cdot \log P + 3.34 \cdot \log \text{SDII}$	0.97	0.21
Dsa	Z, SDII	$\log R = 8.602 - 0.963 \cdot \log \text{SDII} - 0.247 \cdot \log Z$	0.51	0.05
Dsb	Р	$\log R = 2.166 + 0.494 \cdot \log P$	0.45	0.25
Dsc	SDII	$\log R = 6.236 - 0.869 \cdot \log \text{SDII}$	0.51	0.02
Dwa	Р	$\log R = -0.572 + 1.238 \cdot \log P$	0.99	0.02
Dwb	P, SDII	$\log R = -1.7 + 0.788 \cdot \log P + 1.824 \cdot \log \text{SDII}$	0.98	0.02
Dfa	P, SDII	$log R = -1.99 + 0.737 \cdot log P + 2.033 \cdot log SDII$	0.9	0.16
Dfb	P, SDII, Z	$\log R = -0.5 + 0.266 \cdot \log P + 3.1 \cdot \log \text{SDII} - 0.131 \cdot \log Z$	0.89	0.32
Dfc	SDII	$log R = -1.259 + 3.862 \cdot log SDII$	0.91	0.23
ET	Р	$\log R = -3.945 + 1.54 \cdot \log P$	0.14	0.42
EF+EFH	Р	$\log R = 16.39 - 1.286 \cdot \log P$	0.6	0.13
ETH	P, SDII	$\log R = 21.44 + 1.293 \cdot \log P - 10.579 \cdot \log \text{SDII}$	0.52	0.53

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Table 4. Linear multiple regression equations for different climate zones for regions that have no data on the simple precipitation intensity index (SDII). The regression equations relate high resolution erosivity from the USA with the annual total mean precipitation (P) and/or the mean elevation (z).

Climate zone	Optimal regression function (when SDII is not available)	R^2	Residual standard error
BWk	Method Renard and Freimund (1994)		
BSh	$\log R = -8.164 + 2.455 \cdot \log P$	0.86	0.5
BSk	$\log R = 5.52 + 1.33 \cdot \log P - 0.977 \cdot \log Z$	0.76	0.52
Cfa	$\log R = 3.378 + 0.852 \cdot \log P - 0.191 \cdot \log Z$	0.57	0.23
Cfb	$\log R = 5.267 + 0.839 \cdot \log P - 0.635 \cdot \log Z$	0.81	0.5
Dsa	$\log R = 7.49 - 0.0512 \cdot \log P - 0.272 \cdot \log Z$	0.48	0.06
Dsc	$\log R = 4.416 - 0.0594 \cdot \log P$	0.015	0.03
Dwb	$\log R = 1.882 + 0.819 \cdot \log P$	0.81	0.08
Dfa	$\log R = -2.396 + 1.5 \cdot \log P$	0.65	0.29
Dfb	$\log R = 1.96 + 1.084 \cdot \log P - 0.34 \cdot \log Z$	0.74	0.48
Dfc	$\log R = -3.263 + 1.576 \cdot \log P$	0.56	0.49
ETH	$\log R = -10.66 + 2.43 \cdot \log P$	0.4	0.59

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Table 5. Mean high resolution R values (MJ mm ha⁻¹ h⁻¹ yr⁻¹) from the USA and Switzerland and mean modelled R values with uncertainty range for each addressed climate zone.

		observed	modelled	modelled
climate	description	R mean	R mean	uncertainty range
BWk	arid, desert, cold	284	291	158–495
BSh	arid, steppe, hot	2168	2207	1723-2828
BSk	arid, steppe, cold	876	885	749–1046
Csb	temperate, dry warm summer	192	192	133-292
Cfa	temperate, without dry season, hot summer	5550	5437	4830-6123
Cfb	temperate, without dry season, warm summer	1984	1971	1431-2715
Dsa	cold, dry hot summer	172	171	86-340
Dsb	cold, dry warm summer	175	168	151–187
Dsc	cold, dry cold summer	115	115	91–145
Dwa	cold, dry winter, hot summer	1549	1551	1280-1879
Dwb	cold, dry winter, warm summer	1220	1214	1057-1395
Dfa	cold, without dry season, hot summer	2572	2582	2346-2843
Dfb	cold, without dry season, warm summer	1101	1124	922-1371
Dfc	cold, without dry season, cold summer	483	483	423-552
ET	polar, tundra	1352	1249	23-68 088
EF+EFH	polar, frost + polar, frost, high elevation	1468	1450	16–132 00°
ETH	polar, tundra, high elevation	945	832	0-6314918
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Table 6. Statistics of the comparison of high resolution erosivity from three regions to estimated erosivity from the Renard and Freimund method and the new regression equations.

	Observed			Estima	ted – Re	nard a	nd Freimund		Estimated – multiple linear regression				ion
	Range	Mean	SD	Range	Mean	SD	Correlation coefficient	Rank corre- lation coeffi- cient	Range	Mean	SD	Correlation coefficient	Rank corre- lation coeffi- cient
Switzerland	121-6500	1204	833	2335-10 131	5798	1654	0.51	0.42	225-2572	1256	472	0.49	0.3
the USA (aggregated huc4)	105–4963	1271	1174	57–15 183	1870	2088	0.51	0.68	60–15 808	1691	2188	0.58	0.83
Ebro basin	40-4500	891	622	747-5910	1529	846	_	_	167-4993	836	701	_	-

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Table 7. Comparison of the global erosion rates (tha⁻¹ yr⁻¹) and percentiles between different versions of the RUSLE model.

mean	25th percentile	50th percentile	75th percentile	90th percentile
5.1	0.2	0.8	2.8	8.6
11.1	0.3	1.2	4.3	15.7
3.6	0.1	0.6	1.9	6.3
7.3	0.2	0.8	3	10.9
	5.1 11.1 3.6	5.1 0.2 11.1 0.3 3.6 0.1	5.1 0.2 0.8 11.1 0.3 1.2 3.6 0.1 0.6	5.1 0.2 0.8 2.8 11.1 0.3 1.2 4.3 3.6 0.1 0.6 1.9

Table 8. Statistics of the observed and modelled erosion rates from the unadjusted and adjusted versions of the RUSLE for the USA and Europe (tha⁻¹ yr⁻¹).

Region	Source Observations Adjusted RUSLE				LE	Unadjusted RUSLE				
		Range	Mean	SD	Rang	e Mean	SD	Range	Mean	SD
Europe (Aggregation country level) no small countries	Cerdan et al., 2010	0.1–2.6	0.9	0.7	0.1–	7 2.3	2.1	0–14	2.8	3.6
the USA (Aggregation HUC4 level)	NRI database	0–11	1.7	2.1	0.2–2	1 1.7	2.5	0–14	1.9	2.3

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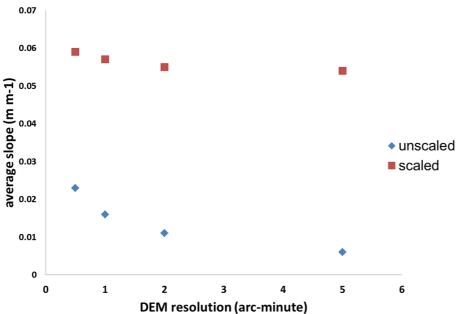


Figure 1. Global average unscaled slope estimated from different coarse resolution digital elevation models (DEMs) as function of their resolution (blue); and global average scaled slope from the same DEMs as function of their resolution (red).

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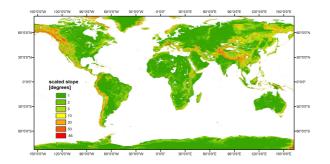




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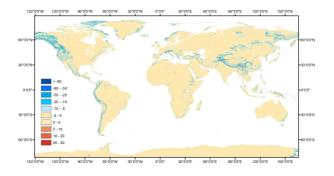


Figure 2. (a) A global map of the scaled slope derived from the 30 arc-second DEM using a target resolution of 150 m; **(b)** a global map showing the difference between the unscaled and scaled slopes (in degrees), where blue colours show an underestimation by the unscaled slope when compared to the scaled slope and redisch colours show and overestimation.

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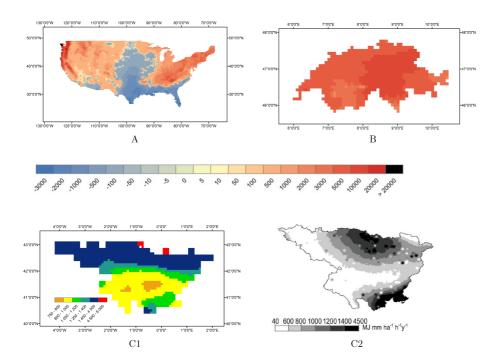


Figure 3. Spatial difference plots showing the difference between the high resolution erosivity and erosivity calculated with the method of Renard and Freimund for (a) the USA, (b) Switzerland and (c) the Ebro basin in Spain; in (a) and (b) the blue colours show an underestimation of the calculated erosivity when compared to the high resolution erosivity, while the red colours show an overestimation; the Ebro basin serves here as an independent validation set and it has two graphs, (c1) a spatial plot of erosivity according to Renard and Freimund, and (c2) the high resolution erosivity from Angulo-Martinez et al. (2009); all values in the graphs are in $MJ mm ha^{-1} h^{-1} vr^{-1}$.

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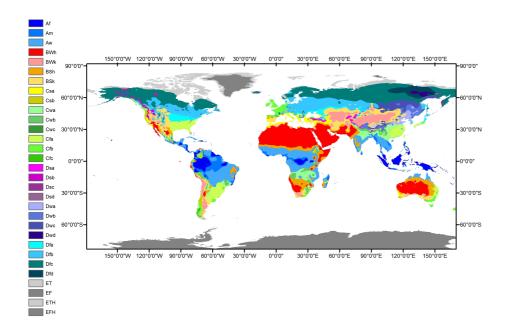


Figure 4. The Köppen-Geiger climate classification global map with resolution of 5 arc-minute (Peel et al., 2007).

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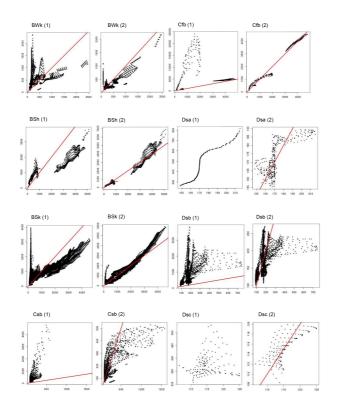


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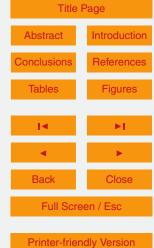


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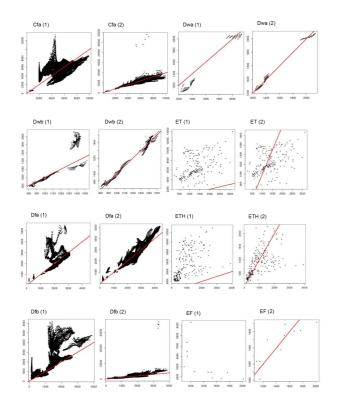


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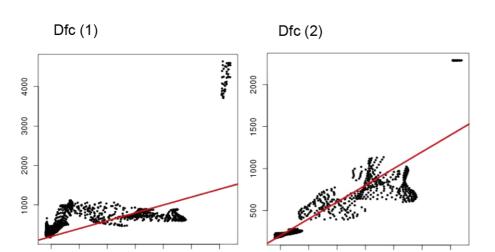


Figure 5. Comparison of high resolution erosivity data and predicted erosivity from (1) the Renard and Freimund method and (2) the new regression equations, for various climate zones; the red line is the 1 tot 1 line, and does not appear in some graphs because predicted erosivity is overestimated.

200

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1200

1000

1400

1200

1000

1400

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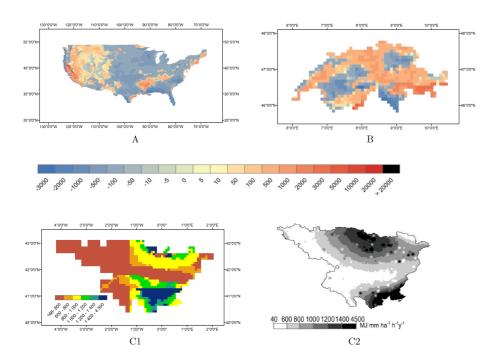


Figure 6. Spatial difference plots showing the difference between the high resolution rainfall erosivity and erosivity calculated with the new regression equations for (a) the USA, (b) Switzerland and (c) the Ebro basin in Spain; in (a) and (b) the blue colours show an underestimation of the calculated erosivity when compared to the high resolution erosivity, while the red colours show an overestimation; the Ebro basin serves here as an independent validation set and it has two graphs, (c1) a spatial plot of erosivity according to the new regression equations, and (c2) the high resolution erosivity from Angulo-Martinez et al. (2009); all values in the graphs are in MJ mm ha⁻¹ h⁻¹ yr⁻¹.

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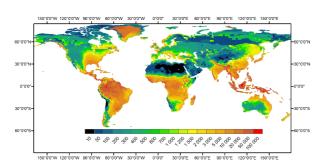
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Figure 7. (a) Global distribution of the new modelled rainfall erosivity values according to the new regression equations; and (b) a difference map between erosivity calculated according to the method of Renard and Freimund and the new modelled erosivity values (MJ mm ha⁻¹ h⁻¹ yr⁻¹), where blue colours indicate lower erosivity values by Renard and Freimund, while redish colours indicate higher erosivity values; map resolution is 5 arc-minute.

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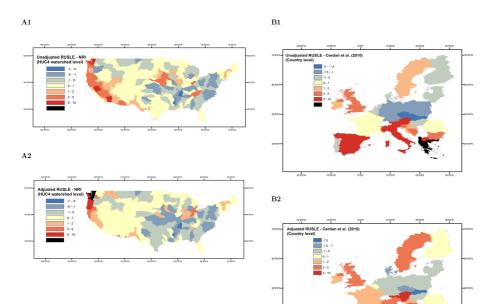


Figure 8. (a) Difference plots between soil erosion estimates from the NRI database for the USA and estimates of **(a1)** the unadjusted RUSLE model, and of **(a2)** the adjusted RUSLE model; all aggregated at HUC4 watershed level; **(b)** difference plots between soil erosion estimates from the database of Cerdan et al. (2010) for Europe and estimates of **(b1)** the unadjusted RUSLE model and of **(b2)** the adjusted RUSLE model; all aggregated at country level; reddish colors represent an overestimation (%) while the bluish represent and underestimation (%) compared to the erosion values from the databases; black color is an overestimation > 10 %.

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