



Simultaneously  
parameterize the ET  
model by Bayesian  
approach

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# Simultaneous parameterization of the two-source evapotranspiration model by Bayesian approach: application to spring maize in an arid region of northwest China

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2011; Sun et al., 2012). The Shuttleworth–Wallace model (S–W model) (Shuttleworth and Wallace, 1985) takes the interactions between the fluxes from soil and canopy into account, and is physically sound and rigorous. Previous studies have proved that it has good performances for row crops such as maize, wheat, cotton, sorghum and vine (Stannard, 1993; Tourula and Heikinheimo, 1998; Anadranistakis et al., 2000; Teh et al., 2001; Lund and Soegaard, 2003; Kato et al., 2004; Ortega-Farias et al., 2007; Zhang et al., 2008).

Despite these studies, there are still some insufficiencies in the application of the S–W model (Hu et al., 2009; Zhu et al., 2013). First, the S–W model is sensitive to the errors in the values of canopy and soil resistances (Lund and Soegaard, 2003). Previous studies mainly focused on the parameterization of the canopy resistance (Hanan and Prince, 1997; Samanta et al., 2007; Zhu et al., 2013), and less attentions has been committed to the parameterization of the soil surface resistance (Sellers et al., 1992; van de Griend and Owe, 1994; Villagarcía et al., 2010). In crop ecosystem,  $E$  may contribute significantly to the total ET when leaf area index (LAI) is low (Lund and Soegaard, 2003; Zhang et al., 2008). Thus, simultaneous parameterization of the canopy and soil resistances in the S–W model, based on direct measurement of ET and its components by using a combination of micro-meteorological (e.g. eddy covariance methods, Bowen ratio), hydrological (e.g. chambers, microlysimeters) and eco-physiological techniques (e.g. sap-flow, stable isotopes) (Williams et al., 2004; Scott et al., 2006), is important to reduce the model error. However, such studies are relative rare or non-existent. Secondly, as far as the parameterization method is concerned, abundant evidence has shown that the Bayesian method provides a powerful new tool to simultaneously optimized many or all model parameters against all available measurements (Clark and Gelfand, 2006). Although some pioneering efforts have been made (e.g. Samanta et al., 2007; Zhu et al., 2013), the Bayesian method has been much less frequently used in parameterization of ET model than in the other environmental sciences (van Oijen et al., 2005). Moreover, the Bayesian method, to our knowledge, has not been used to simultaneously optimized the parameters of the S–W



10 days of July, while the maize (*Zea mays* L.) is sown in the late April and harvested in the middle 10 days of September.

## 2.2 Measurements and data processing

The field observation systems at this site were constructed in May 2013 as part of the Heihe Watershed Allied Telemetry Experimental Research (HiWATER) project (see details in Li et al., 2013b). The fluxes of sensible heat ( $H$ ), latent heat ( $\lambda ET$ ) and carbon dioxide were measured at the height 4.5 m using the eddy covariance (EC) system (Liu et al., 2014), which consists of an open-path infrared gas analyzer (Li-7500, LiCor Inc., Lincoln, NE, USA) and a 3-D sonic anemometer (CSAT-3, Campbell Scientific Inc., Logan, UT, USA). The EC data were sampled at a frequency of 10 Hz by a data logger (CR5000, Campbell Scientific Inc.), and then were processed with an average time of 30 min. Post-processing calculations, using EdiRe software, included spike detection, lag correction of  $H_2O/CO_2$  relative to the vertical wind component, sonic virtual temperature conversion, planar fit coordinate rotation, the WPL density fluctuation correction and frequency response correction (Xu et al., 2014). Data gaps due to instrument malfunction, power failure and bad weather conditions were filled using artificial neural network (ANN) and mean diurnal variations (MDV) methods (Falge et al., 2001). The ANN method was applied when the synchronously meteorological data were available; otherwise, the MDV method was used. The gap-filling data were used only to analyze the seasonal and annual variations in ET.

Continuous complementary measurements also included standard hydro-meteorological variables. Rainfall was measuring using a tipping bucket rain gauge (TE525MM, Campbell Scientific Instruments Inc.). Air temperature and relative humidity (HMP45C, Vaisala Inc., Helsinki, Finland) and wind speed/direction (034B, Met One Instruments, Inc. USA) were measured at heights of 3, 5, 10 15, 20, 30 and 40 m above the ground. Downward and upward solar and longwave radiation (PSP, The EPPLEY Laboratory Inc., USA) and photosynthetic photon flux density (PPFD) (LI-190SA, LI-COR Inc.) were measured at height of 6 m. Soil temperature

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$$ET_s = \frac{\Delta A + [\rho C_p D - \Delta r_a^s (A - A_s)] / (r_a^a + r_a^s)}{\Delta + \gamma [1 + r_s^s / (r_a^a + r_a^s)]} \quad (2)$$

$$ET_c = \frac{\Delta A + [\rho C_p D - \Delta r_a^c A_s] / (r_a^a + r_a^c)}{\Delta + \gamma [1 + r_s^c / (r_a^a + r_a^c)]} \quad (3)$$

$$C_s = \frac{1}{1 + [R_s R_a / R_c (R_s + R_a)]} \quad (4)$$

$$C_c = \frac{1}{1 + [R_c R_a / R_s (R_c + R_a)]} \quad (5)$$

$$R_a = (\Delta + \gamma) r_a^a \quad (6)$$

$$R_c = (\Delta + \gamma) r_a^c + \gamma r_s^c \quad (7)$$

$$R_s = (\Delta + \gamma) r_a^s + \gamma r_s^s \quad (8)$$

$$\lambda E = \frac{\Delta A_s + \rho C_p D_0 / r_a^s}{\Delta + \gamma (1 + r_s^s / r_a^s)} \quad (9)$$

$$\lambda T = \frac{\Delta (A - A_s) + \rho C_p D_0 / r_a^c}{\Delta + \gamma (1 + r_s^c / r_a^c)} \quad (10)$$

$$D_0 = D + \frac{(\Delta A - (\Delta + \gamma) \lambda ET) r_a^a}{\rho C_p} \quad (11)$$

where  $ET_s$ , and  $ET_c$  are terms to describe evaporation from soil and transpiration from the plant ( $Wm^{-2}$ ), respectively;  $C_s$  and  $C_c$  are soil surface resistance coefficient and canopy resistance coefficient (dimensionless), respectively;  $\lambda$  is the latent heat of evaporation ( $Jkg^{-1}$ );  $\Delta$  is the slope of the saturation vapor pressure vs. temperature curve ( $kPaK^{-1}$ );  $\rho$  is the air density ( $kgm^{-3}$ );  $C_p$  is the specific heat capacity of dry air ( $1013Jkg^{-1}K^{-1}$ );  $D$  and  $D_0$  ( $kPa$ ) is the air water vapor pressure deficit at the reference height (3 m) and the canopy height, respectively;  $\gamma$  is the psychrometric constant

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( $\text{kPa K}^{-1}$ );  $r_s^c$  and  $r_s^s$  are the surface resistance for plant canopy and soil surface ( $\text{sm}^{-1}$ ), respectively;  $r_a^c$  and  $r_a^s$  are aerodynamic resistances from the leaf to canopy height and soil surface to canopy height ( $\text{sm}^{-1}$ ), and  $r_a^a$  is aerodynamic resistances from canopy height to reference height ( $\text{sm}^{-1}$ ).  $A$  and  $A_s$  ( $\text{Wm}^{-2}$ ) are the available energy input above the canopy and above the soil surface, respectively, and are calculated as:

$$A = R_n - G \quad (12)$$

$$A_s = R_{ns} - G \quad (13)$$

where  $R_n$  and  $R_{ns}$  are net radiation fluxes into the canopy and the substrate ( $\text{Wm}^{-2}$ ), respectively;  $G$  is the soil heat flux ( $\text{Wm}^{-2}$ ).  $R_{ns}$  was calculated using a Beer's law relationship of the form:

$$R_{ns} = R_n \exp(-K_A \text{LAI}) \quad (14)$$

in which  $K_A$  is the extinction coefficient of light attenuation, and is approximately 0.41 for spring maize (Mo et al., 2000).

The climate-related variables (i.e.,  $\lambda$ ,  $e_s$ ,  $\Delta$ ,  $\rho$  and  $\gamma$ ) in Eqs. (1)–(3) is calculated by the formulas of Allen et al. (1998).

## 2.4 Calculation of resistances in the S–W model

The resistance network of the S–W model is shown in Fig. 2. In this paper, the three aerodynamic resistance (i.e.,  $r_a^a$ ,  $r_a^c$  and  $r_a^s$ ) were calculated using the same approach suggested by Shuttleworth and Wallace (1985), Shuttleworth and Gurney (1990) and Lhomme et al. (2012).

The canopy resistance ( $r_s^c$ ), which is the equivalent resistance of all the individual stomates in a canopy and depends on the environmental variables, can be calculated using the Jarvis-type model (Jarvis, 1976)



$$r_s^c = \frac{r_{STmin}}{2LAI \prod_j F_j(X_j)} \quad (15)$$

where  $r_{STmin}$  represents the minimal stomatal resistance of individual leaves under optimal conditions.  $F_j(X_j)$  is the stress function of a specific environmental variable  $X_j$ , with  $0 \leq F_j(X_j) \leq 1$ . Following Stewart (1998) and Verhoef and Allen (2000), the stress functions were expressed as:

$$F_1(R_s) = \frac{R_s}{1000} \frac{1000 + k_1}{R_s + k_1} \quad (16)$$

$$F_2(T_a) = \frac{(T_a - T_{a,min})(T_{a,max} - T_a)^{(T_{a,max}-k_2)/(k_2-T_{a,min})}}{(k_2 - T_{a,min})(T_{a,max} - k_2)^{(T_{a,max}-k_2)/(k_2-T_{a,min})}} \quad (17)$$

$$F_3(D) = 1 - k_3 D \quad (18)$$

$$F_4(\theta_r) = \begin{cases} 1 & \theta_r > \theta_{cr} \\ \frac{(\theta_r - \theta_{wp})}{(\theta_{cr} - \theta_{wp})} & \theta_{wp} \leq \theta_r \leq \theta_{cr} \\ 0 & \theta_r < \theta_{wp} \end{cases} \quad (19)$$

where  $k_1 - k_3$  are constants (units see Table 1);  $R_s$  is the incoming solar radiation ( $Wm^{-2}$ );  $T_a$  is the air temperature ( $^{\circ}C$ ) at the reference height;  $T_{a,min}$  and  $T_{a,max}$  are the lower and upper temperatures limits ( $^{\circ}C$ ), respectively, which are  $T_a$  values when  $F_2(T_a) = 0$  and are set at values of 0 and  $40^{\circ}C$  (Harris et al., 2004);  $\theta_r$  is the actual volumetric soil water content in the root-zone at depth of 0–60 cm ( $m^3 m^{-3}$ );  $\theta_{wp}$  is water content at the wilting point ( $m^3 m^{-3}$ ); and  $\theta_{cr}$  is the critical water content at which plant stress starts and was set as  $0.30 m^3 m^{-3}$  in this study.

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The soil surface resistances ( $r_s^S$ ; Fig. 2) was expressed as a function of near-surface soil water content (Sellers, 1992; Verhoef et al., 2006, 2012; Zhu et al., 2013):

$$r_s^S = \exp(b_1 - b_2 \frac{\theta_s}{\theta_{sat}}) \quad (20)$$

5 in which  $b_1$  and  $b_2$  are empirical constants ( $s\ m^{-1}$ );  $\theta_s$  is soil water content in the top layer of soil (at depth of 2 cm);  $\theta_{sat}$  is the saturated soil water content ( $m^3\ m^{-3}$ ), which was estimated empirically through the near-surface soil texture. In summary, there are 6 site- and species-specific parameters needed to be estimated in the S–W model associated with soil and canopy resistances, which are  $b_1$ ,  $b_2$ ,  $r_{STmin}$  and  $k_1 - k_3$ .

## 10 2.5 Model calibration and evaluation

A Bayesian approach was applied to simultaneously estimate the parameters associated with the soil ( $b_1, b_2$ ) and canopy ( $r_{STmin}, k_1, k_2, k_3$ ) resistances in the S–W model (van Oijen et al., 2005; Svensson et al., 2008; Zhu et al., 2011, 2013). The two dataset used to simultaneously optimize the parameter values were: EC-measured half-hourly evapotranspiration ( $\lambda ET$ ;  $W\ m^{-2}$ ) and microlysimeters-measured daily soil evaporation ( $E$ ;  $mm\ d^{-1}$ ).

Corresponding to each of the data sets (e.g.,  $\lambda ET$  and  $E$ ), the model error  $e_i(t)$  ( $i = 1, 2$ ) is expressed by:

$$20\ e_i(t) = O_i(t) - f_i(t) \quad (21)$$

where  $O_i(t)$  and  $f_i(t)$  is observed and modeled (Eqs. 1 and 9) values of the  $i$ th dataset at time  $t$ , respectively. Assuming the model error  $e_i(t)$  follows a Gaussian distribution with a zero mean, the data likelihood function can be expressed by:

$$p(O|c) \propto \exp \left\{ - \sum_{i=1}^m \frac{1}{2\sigma_i^2} \sum_{t=1}^{n_i} (e_i(t))^2 \right\} \quad (22)$$

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where  $c$  is the parameter vector;  $O$  is the observed data sets;  $m$  is the number of dataset (= 2 in this study);  $\sigma_i^2$  ( $i = 1, 2$ ) is the measurement error variance of the  $i$ th dataset;  $n_i$  is the number of observations of  $i$ th dataset. Then with Bayes' theorem, the posterior distribution of parameters  $c$  is generated by:

$$p(c|O) \propto p(O|c)p(c) \quad (23)$$

where  $p(c)$  represents prior probability distributions of parameters  $c$ , and  $p(c|O)$  is the posterior distributions of parameters  $c$ . The posterior distribution was sampled using the Metropolis–Hasting (M–H) algorithm (Metropolis et al., 1953; Hastings, 1970), a version of the Markov Chain Monte Carlo (MCMC) technique. To generate a Markov chain in the parameter space, the M–H algorithm was run by repeating two steps: a proposing step and a moving step. In the proposing step, a candidate point  $c^{\text{new}}$  is generated according to a proposal density  $P(c^{\text{new}}|c^{k-1})$ , where  $c^{k-1}$  is the accepted point at the previous step. In the moving step, point  $c^{\text{new}}$  is treated against the Metropolis criterion to examine if it should be accepted or rejected (see Zhu et al., 2011, 2013 for detailed description on MCMC sampling procedure). We ran at least three parallel MCMC chains with 20 000 iterations each, evaluated the chains for convergence (Gelman and Rubin, 1992), and thinned the chains (every 20th iteration) when appropriate to reduce within chain autocorrelation, thereby producing an independent sample of 3000 values for each parameter from the joint posterior distribution.

During the whole growing season, the measurements were split into two independent dataset by taking alternate days. The model parameters were derived using one dataset. Then the optimized S–W model was used with the second data set to predict the different components of ET and these values were compared to the measured values in the second dataset. The performance of the S–W model was evaluated using the coefficient of determination of the linear regression between measured and estimated values of water vapor fluxes,  $R^2$ , representing how much the variation in the observations was explained by the models. Also, the root mean square error (RMSE),





the key parameters in the S–W model were well estimated by the Bayesian method against the multiple measuring dataset. In addition, the six calibrated parameters were not significantly inter-correlated with each other except for the pair  $b_1$  and  $b_2$ , which was positively correlated with a correlation coefficient of 0.85.

The responses of soil surface resistances ( $r_s^S$ ) to soil water content computed using our posterior mean  $b_1$  and  $b_2$  values were very similar to that calculated using equation developed by Ortega-Farias et al. (2010) based on direct soil evaporation measurements, but seemed to be more sensitive to changes in soil water contents compared with some other studies (e.g., Sun, 1982; Sellers, 1992; Zhu et al., 2013; Fig. 5). The posterior mean value of  $r_{STmin}$  from our study was very close to that ( $20 \text{ s m}^{-1}$ ) reported for spring maize growing in northwest China obtained by using the least squares fitting method (Li et al., 2013a). The variations of the minimal stomatal resistance ( $r_{STmin}$ ) for many natural and cultivated plants have been widely investigated by previous studies (Korner et al., 1979; Pospisilova and Solarova, 1980). Typical values for  $r_{STmin}$  vary considerably from about  $20\text{--}100 \text{ s m}^{-1}$  for crops to  $200\text{--}300 \text{ s m}^{-1}$  for many types of trees. Thus, our results fell within the range of previous studies.

### 3.3 Model performance compared with measurements

Having parameterized the S–W model as described above, we ran the model to simulate the half-hourly  $\lambda ET$  (Eq. 1) and  $\lambda E$  (Eq. 9) values ( $W m^{-2}$ ). The daily estimations of evapotranspiration (ET;  $mm d^{-1}$ ) and soil evaporation ( $E$ ;  $mm d^{-1}$ ) was obtained by summing up the half-hourly simulated values. The statistical analysis of observed vs. estimated values of water vapor fluxes at different time-scales were summarized in Table 2. These results indicated that the parameterized S–W model was able to predict  $\lambda ET$  on a half-hourly basis with values of  $R^2$ , IA and EF equal to 0.83, 0.93 and 0.74, respectively. However, there still existed the difference between measured and modeled half-hourly  $\lambda ET$  values for the spring maize in the arid desert oasis. The slope (0.84) of regressive equation between the measured and modeled half-hourly  $\lambda ET$  values was lower than one (Table 2 and Fig. 6a), which indicated that the S–W

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model tended to underestimate the half-hourly  $\lambda$ ET with a MBE value of  $24.2 \text{ W m}^{-2}$ . Ortega-Farias et al. (2010) also reported that the S–W model underestimated on half-hourly time intervals compared the EC-measured  $\lambda$ ET over a drip-irrigated vineyard in Mediterranean semiarid region during the growing season in 2006. On the contrary, some studies showed that the S–W model overestimated half-hourly  $\lambda$ ET (e.g., Li et al., 2013a; Ortega-Farias et al., 2007; Zhang et al., 2008). Therefore, the performances of the S–W model seemed to be variable for different crops and places, and there is a need to identify the causes that induced the disagreements between observed and modeled values (discussed below).

The fluctuation of measured and estimated daily ET and  $E$  was illustrated in Fig. 7. In this case, a good agreement between measured and estimated daily  $E$  was obtained with values of  $R^2$ , IA and EF equal to 0.82, 0.94 and 0.76 (Table 2). The points in plots of measured-vs.-modeled daily  $E$  fell tightly along the 1 : 1 line (slope = 1.01 and intercept = 0.01 with RMSE = 0.05 and MBE =  $-0.01$ ; Fig. 6b and Table 2). Thus, we thought that the soil resistance in the S–W model was properly parameterized for the spring maize by the measured soil evaporation data. From Fig. 7, we can also observed that the estimated daily ET generally fluctuated tightly with the measured values. The values of RMSE, MBE, IA and EF were equal to 0.05,  $0.14 \text{ mm d}^{-1}$ , 0.94 and 0.79, respectively (Table 2). However, great underestimations ( $> 0.5 \text{ mm d}^{-1}$ ) were observed on 12 days during the study period (111 days). For example, on 5 July, the estimated and measured daily ET was  $2.9$  and  $4.3 \text{ mm d}^{-1}$ , respectively (Fig. 7). Thus, the causes of the underestimations of ET by the S–W on these days needed to be carefully checked based on detailed micrometeorological data. This work would help us to modify the model in a correct way and improve the precision of prediction.

### 3.4 Identification of the disagreement/agreement between observed and modeled ET values

The diurnal variation of  $R_n$ ,  $H$  and  $\lambda$ ET (measured and modeled) above the spring maize ecosystem for some selected days was presented in Fig. 8. Resulting from the high surface heterogeneities, one special phenomenon, known as the “oasis effect” (Lemon et al., 1957) or “cold island effect” (Wang et al., 1992; Zhang and Huang, 2004), was often observed on clear days in July and August in the study area and it is characterized as follows: (1)  $H$  is very small and even negative (downward) in the afternoon (Fig. 8a–c) due to the micro-scale advection of hot dry air over the desert to crop surface in the oasis. For an example, on 5 July,  $H$  was continuously negative from 12 : 00 to 20 : 00 (Fig. 8a). A strong advection process can be distinctly detected from the temperature and relative humidity profiles (Fig. 9a and b), in which the highest temperature occurred at a height of 8–18 m; (2) measured actual  $\lambda$ ET often exceeded (Fig. 8a) or was equal to (Figs. 8b and c) the local net radiation because of the added energy in the form of downward fluxes of  $H$  to the ET process (Evelt et al., 2012). Under such conditions, the S–W model significantly underestimated the actual ET values due to the real atmospheric flows do not correspond to its assumption of horizontal homogeneities (Rao et al., 1974). Thus, how to properly representing the advection process in the S–W model should be paid special attentions in simulating ET over crop ecosystems in arid desert oasis in the future studies. In addition to this situation, slight underestimations were also observed on or shortly after rainy days (Fig. 7). For example, the simulated half-hourly  $\lambda$ ET was lower than that measured by EC after the rainfall event occurred in 13:00 LT on 17 June (Fig. 8d). We thought that the underestimations by the model on or shortly after rainy days were mainly due to ignoring the direct evaporation of liquid water intercepted in the crop canopy, because no downward  $H$  and temperature inversion were observed on this day (Figs. 9c and d). Until now, several canopy interception models have been developed (e.g., Rutter et al., 1971; Mulder, 1985; Gash et al., 1995; Bouten et al., 1996). However, many of them

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were developed for simulating the rainfall interception by forest ecosystems, and their suitability for crops need to be further investigated.

On the other hand, the diurnal variation of simulated half-hourly  $\lambda$ ET by the parameterized S–W model has a similar trend to the measurements on clear and advection-absent days during the whole study periods (Fig. 8e–h). On these days,  $H$  was positive (upwards) at day time (Fig. 8e–h) and no temperature inversion was observed (Fig. 9e and f). Thus, we thought that the parameterization schedule adopted in this study worked well. It also demonstrated that the properly parameterized S–W model can be used in simulating and partitioning ET for homogeneous land surface. Hu et al. (2009) reported that the S–W model parameterized by using Monte Carlo method can successfully simulate ET at four uniform grasslands in China; our previous studies (Zhu et al., 2013) also illustrated that parameterized S–W model can be used to simulate and partition ET over a vast alpine grassland in Qinghai-Tibet Plateau.

## 4 Conclusions

This study illustrated the use of the Bayesian method to simultaneously parameterize a two-source ET model against the multiple measuring dataset for a crop ecosystem in a desert oasis of northwest China. The posterior distributions of the model parameters in most cases can be well constrained by the observations. Generally, the parameterized model has a good performance in simulating and partitioning ET. However, underestimations were observed on days when micro-scale advection occurred. Therefore, in the future studies, special attentions should be given to proper descriptions of the effects of advection on estimating ET for heterogeneous land surface. In addition, the canopy interception model should be coupled with the two-source ET model in long-term simulation.

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**Table 1.** Prior distributions and the parameter bounds for the S–W model. These values are derived from the literature; the posterior parameter distribution estimated by MCMC are based on observed data in our site, and are characterized by the mean and 95 % high-probability intervals (Lower limit, Upper limit).

Parameter	Prior Distribution		References	Posterior Distribution	
	Lower Bound	Upper Bound		Mean	95 % High-Probability Interval
$r_{\text{STmin}}$ ( $\text{sm}^{-1}$ )	0	80	Noilhan and Planton (1989); Li et al. (2013a)	<b>21.8</b>	(20.2, 24.6)
$k_1$ ( $\text{Wm}^{-2}$ )	0	500	Stewart (1998)	294.6	(42.5, 487.7)
$k_2$ ( $^{\circ}\text{C}$ )	5	40	Ogink-Hendriks (1995)	<b>25.6</b>	(12.9, 34.4)
$k_3$ ( $\text{kPa}^{-1}$ )	0	0.1	Stewart (1998)	0.02	(0, 0.07)
$b_1$ ( $\text{sm}^{-1}$ )	4	15	Sellers et al. (1992); Zhang (2012); Zhu et al. (2013)	<b>9.3</b>	(8.4, 10.0)
$b_2$ ( $\text{sm}^{-1}$ )	0	8	Sellers et al. (1992); Zhang (2012); Zhu et al. (2013)	<b>6.2</b>	(3.8, 7.4)

The bold number means that this parameter was well constrained by the data.

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**Table 2.** Statistical analysis of measured and estimated values of half-hourly evapotranspiration ( $\lambda\text{ET}$ ;  $\text{W m}^{-2}$ ), daily soil evaporation ( $E$ ;  $\text{mm d}^{-1}$ ), and daily evapotranspiration ( $\text{ET}$ ;  $\text{mm d}^{-1}$ ) for the spring maize in arid desert oasis during the study period.

	$n$	Regressive equation	$R^2$ measured values	Mean simulated values	Mean	RMSE	MBE	IA	EF
$\lambda\text{ET}$ ( $\text{W m}^{-2}$ )	3578	$\lambda\text{ET}_{\text{modeled}} = 0.84\lambda\text{ET}_{\text{measured}} + 0.18$	0.83	161.4	137.2	80.7	24.2	0.93	0.74
$E$ ( $\text{mm d}^{-1}$ )	56	$E_{\text{modeled}} = 1.01E_{\text{measured}} + 0.01$	0.82	0.26	0.28	0.05	-0.01	0.94	0.76
$\text{ET}$ ( $\text{mm d}^{-1}$ )	95	$\text{ET}_{\text{modeled}} = 0.83\text{ET}_{\text{measured}} + 0.19$	0.83	2.02	1.88	0.32	0.14	0.94	0.79

$n$  = the sample number;  $R^2$  = the determination coefficient; RMSE = root mean square error; MBE = mean bias error between measured and modeled values; IA = index of agreement; ET = model efficiency. These statistical parameters are calculated using formulas given by Legates and McCabe (1999) and Poblete-Echeverria and Ortega-Farias (2009).

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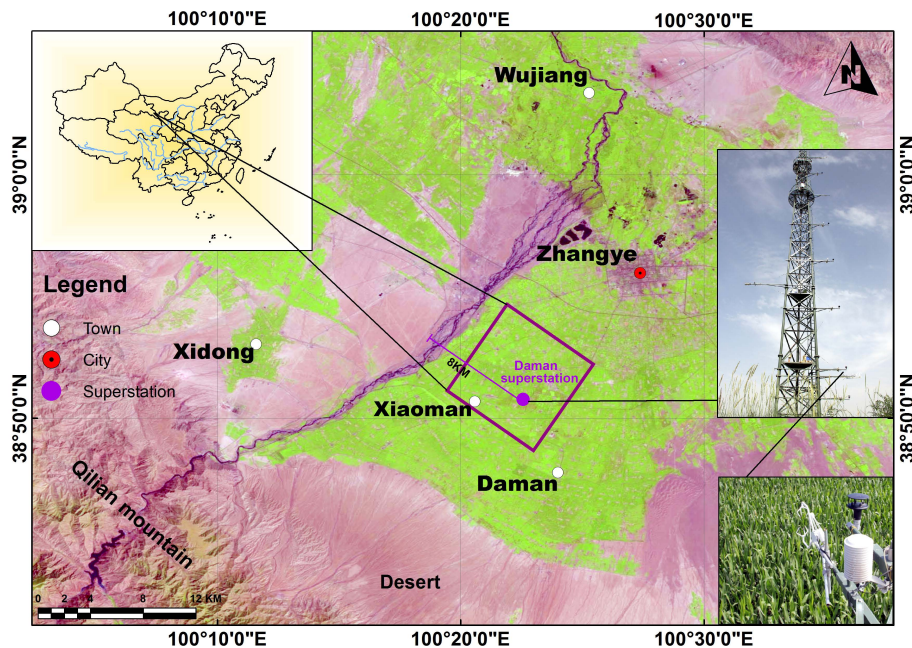


Fig. 1. Experimental location and instrumentation setting at Daman (DM) superstation.

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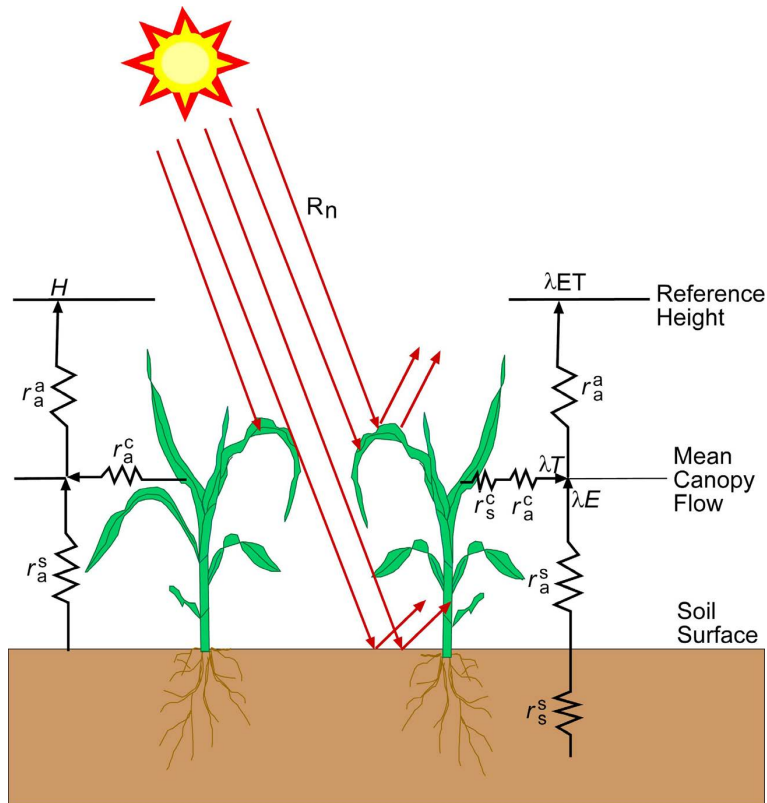
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**Fig. 2.** Schematic diagram of the S–W model. From right to left,  $r_s^c$  and  $r_a^c$  are bulk resistances of canopy stomatal and boundary layer ( $\text{sm}^{-1}$ ), respectively;  $r_a^s$  and  $r_a^a$  aerodynamic resistances from soil to canopy and from canopy to reference height ( $\text{sm}^{-1}$ ), respectively;  $r_s^s$  soil surface resistance ( $\text{sm}^{-1}$ ).  $\lambda T$  transpiration from canopy ( $\text{Wm}^{-2}$ ),  $\lambda E$  evaporation from soil under plant ( $\text{Wm}^{-2}$ ), and  $\lambda ET$  total evapotranspiration ( $\text{Wm}^{-2}$ ).

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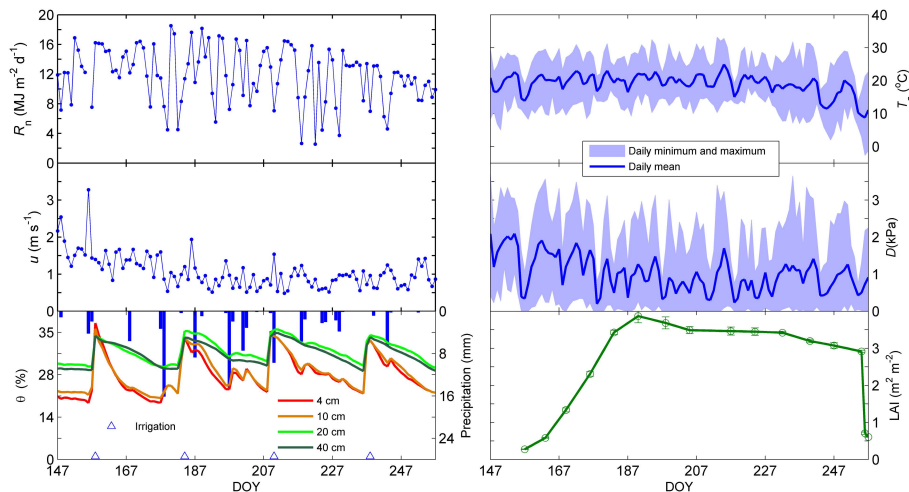
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**Fig. 3.** Seasonal variation in net solar radiation ( $R_n$ ;  $\text{MJ m}^{-2} \text{d}^{-1}$ ), air temperature ( $T_a$ ;  $^{\circ}\text{C}$ ), vapor pressure deficit ( $D$ ;  $\text{kPa}$ ), wind speed ( $u$ ;  $\text{m s}^{-1}$ ) at the height of 3 m, precipitation and irrigation (mm), soil water content ( $\theta$ ;  $\text{m}^3 \text{m}^{-3}$ ) and leaf are index (LAI;  $\text{m}^2 \text{m}^{-2}$ ) during the study period in the Daman Oasis.

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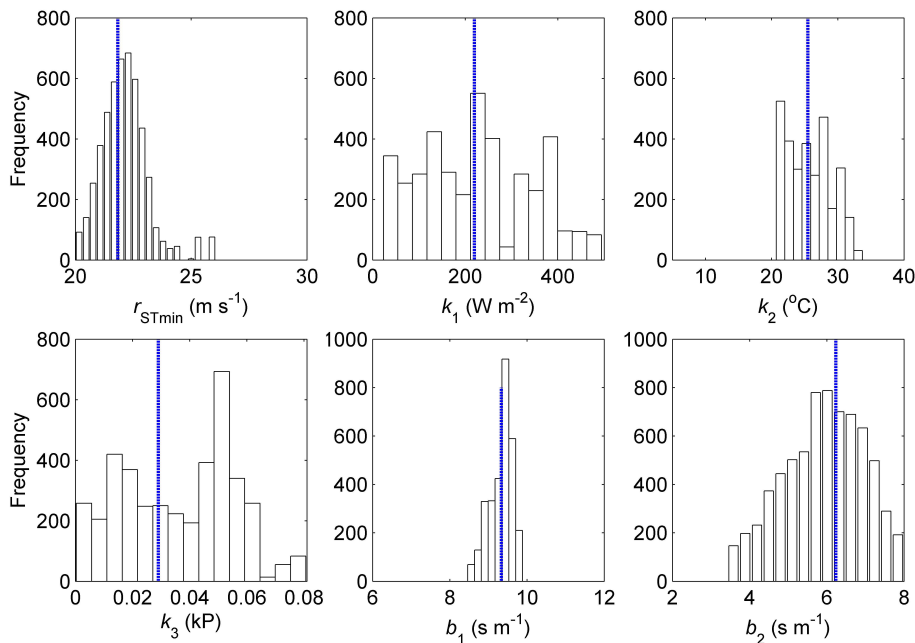
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**Fig. 4.** Histograms of samples from the posterior distributions of the parameters. The dashed vertical lines indicate mean parameter values.

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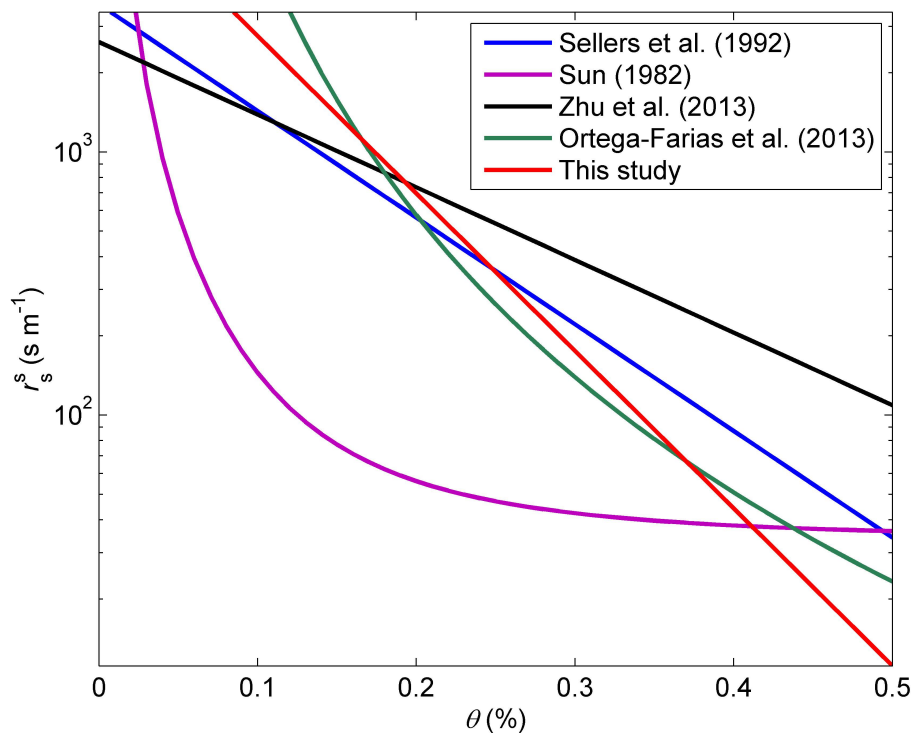
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**Fig. 5.** Comparisons of responses of soil surface resistance ( $r_s^s \text{ s m}^{-1}$ ) to soil surface water contents ( $\theta$ ;  $\text{m}^3 \text{ m}^{-3}$ ).

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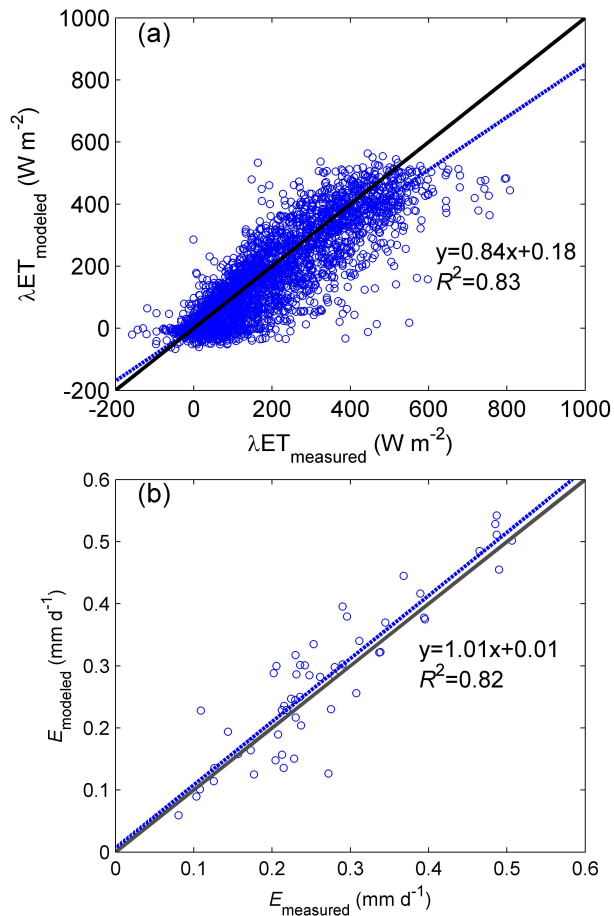
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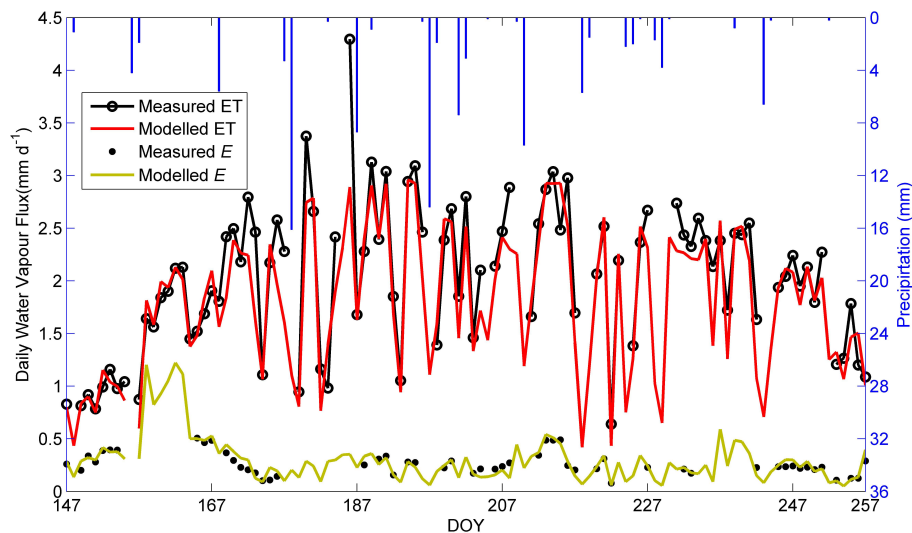


**Fig. 6. (a)** Plot of estimated evapotranspiration ( $\lambda ET$ ;  $\text{W m}^{-2}$ ) against observed values. The regressions is:  $y = 0.84x + 0.18$  ( $R^2 = 0.83$ ); **(b)** plot of estimated daily soil evaporation ( $E$ ;  $\text{mm d}^{-1}$ ) against measured data. The regressions is:  $y = 1.01x + 0.01$  ( $R^2 = 0.82$ ).



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**Fig. 7.** Seasonal variation in daily evapotranspiration (ET;  $\text{mmd}^{-1}$ ) and soil evaporation ( $E$ ;  $\text{mmday}^{-1}$ ) measured by the EC system and microlysimeters and modeled by the S–W model during the study period in Daman Oasis. Gap in the time series is caused either by the absence of flux measurements or missing ancillary data.

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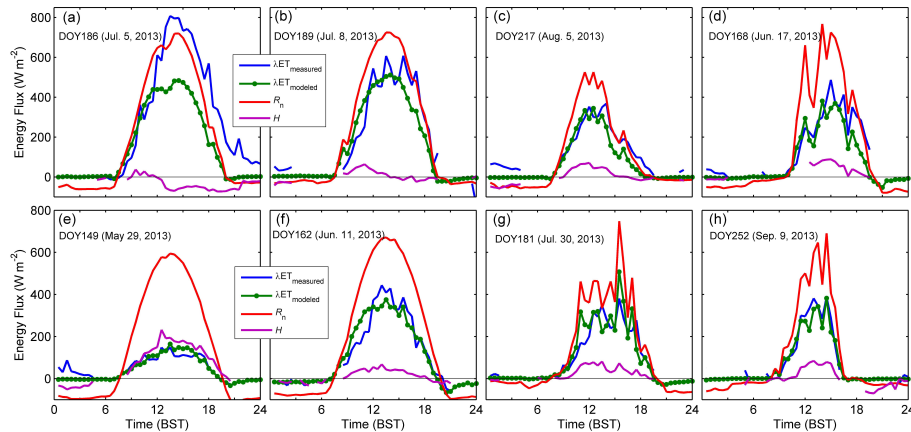
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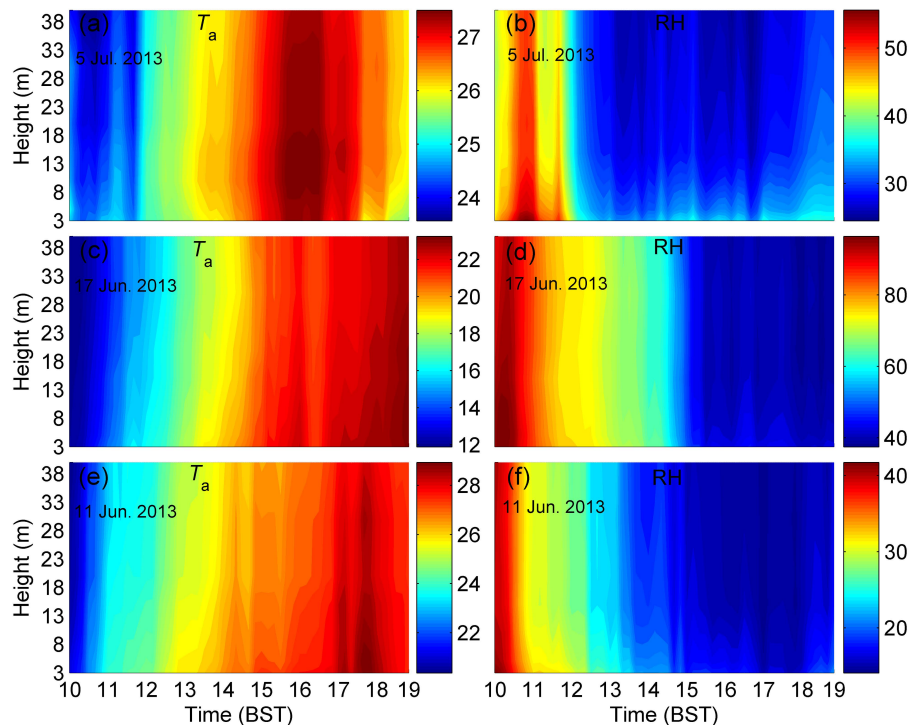
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**Fig. 8.** Diurnal variations in net radiation flux ( $R_n$ ;  $\text{W m}^{-2}$ ), sensible heat flux ( $H$ ;  $\text{W m}^{-2}$ ), and modeled and measured evapotranspiration flux ( $\lambda\text{ET}$ ;  $\text{W m}^{-2}$ ). **(a)–(c)** Represented conditions at which micro-scale advection occurred at 12:00, 15:00 and 17:00 Beijing Standard Time (BST), respectively, **(d)** represented a rainy day, and **(e)–(h)** represented clear and advection-absent days during the study period. Gap is caused either by the absence of flux measurements or missing ancillary data.

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**Fig. 9.** The diurnal evolutions of temperature ( $T_a$ ; °C) and relative humidity (RH; %) profiles from 3 m to 40 m above the ground. Profiles of **(a)**  $T_a$  and **(b)** RH on 5 July, 2013. An obvious advection process can be detected from 13:00 to 17:00 BST with high  $T_a$  and low RH layer at the height of 8–18 m; profiles of **(c)**  $T_a$  and **(d)** RH on 17 June, 2013. A precipitation event occurred at 13:00 and resulted in uniform vertical distributions of  $T_a$  and RH, but no temperature inversion were observed; profiles of **(e)**  $T_a$  and **(f)** RH on 11 June, 2013. It represented a typical clear and advection-absent day.