



Dynamic model evaluation for SIA and its precursors over Europe

S. Banzhaf et al.

This discussion paper is/has been under review for the journal Geoscientific Model Development (GMD). Please refer to the corresponding final paper in GMD if available.

# Dynamic model evaluation for secondary inorganic aerosol and its precursors over Europe between 1990 and 2009

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Received: 25 June 2014 – Accepted: 2 July 2014 – Published: 29 July 2014

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Published by Copernicus Publications on behalf of the European Geosciences Union.

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## Abstract

In this study we present a dynamic model evaluation of the chemistry transport model LOTOS-EUROS to analyse the ability of the model to reproduce observed non-linear responses to emission changes and interannual variability of secondary inorganic aerosol (SIA) and its precursors over Europe from 1990 to 2009. The 20 year simulation was performed using a consistent set of meteorological data provided by the regional climate model RACMO2. Observations at European rural background sites have been used as reference for the model evaluation. To ensure the consistency of the used observational data stringent selection criteria were applied including a comprehensive visual screening to remove suspicious data from the analysis. The LOTOS-EUROS model was able to capture a large part of the day-to-day, seasonal and interannual variability of SIA and its precursors' concentrations. The dynamic evaluation has shown that the model is able to simulate the declining trends observed for all considered sulphur and nitrogen components following the implementation of emission abatement strategies for SIA precursors over Europe. Both, the observations and the model show the largest part of the decline in the 1990's while smaller concentration changes and an increasing number of non-significant trends are observed and modelled between 2000–2009. Furthermore, the results confirm former studies showing that the observed trends in sulphate and total nitrate concentrations from 1990 to 2009 are significantly lower than the trends in precursor emissions and precursor concentrations. The model captured these non-linear responses to the emission changes well. Using the LOTOS-EUROS source apportionment module trends in formation efficiency of SIA have been quantified for four European regions. The exercise has revealed a 20–50 % more efficient sulphate formation in 2009 compared to 1990 and an up to 20 % more efficient nitrate formation per unit nitrogen oxide emission, which added to the explanation of the non-linear responses. However, we have also identified some weaknesses to the model and the input data. LOTOS-EUROS underestimates the observed nitrogen dioxide concentrations throughout the whole time period, while it overestimates the

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observed nitrogen dioxide concentration trends. Moreover, model results suggest that the emission information of the early 1990's used in this study needs to be improved concerning magnitude and spatial distribution.

## 1 Introduction

5 Atmospheric input of sulphur and nitrogen components may decrease biodiversity in vulnerable terrestrial and aquatic ecosystems through eutrophication and acidification of soils and fresh water (Bobbink et al., 1998). The major sources of sulphur and reactive nitrogen in the atmosphere are sulphur dioxide ( $\text{SO}_2$ ) and nitrogen oxide ( $\text{NO}_x$ ) emissions from fossil fuel combustion and ammonia ( $\text{NH}_3$ ) emissions from agricultural activities. Although these gases may themselves be removed from the atmosphere by dry deposition or rainout, they are the precursor gases for SIA (sulphate ( $\text{SO}_4^{2-}$ ), nitrate ( $\text{NO}_3^-$ ) and ammonium ( $\text{NH}_4^+$ )). The latter provides a means for long-range transport of reactive nitrogen on a continental scale causing negative ecosystem impacts far away from their major source areas. In addition, SIA contributes a large portion of particulate matter concentration throughout the European domain (Putaud et al., 2010). Especially ammonium nitrate ( $\text{NH}_4\text{NO}_3$ ) concentrations are shown to be particularly enhanced during days with  $\text{PM}_{10}$  concentrations up or above the EU daily limit value (e.g. Weijers et al., 2011). Moreover, SIA are involved in climate change by affecting the radiation balance of the earth (Forster et al., 2007). Recent studies show that short term climate mitigation aimed at reducing black carbon may be effective, provided that the climate impact of the co-emitted SIA precursors does not cause a net cooling impact (Bond et al., 2013). Hence, a thorough understanding of the SIA budget is required to inform policy makers and to devise mitigation strategies that are effective for biodiversity, climate change and human health.

25 To combat the adverse impacts on biodiversity and human health a series of international conventions and agreements were implemented. The Convention on Long-range Transboundary Air Pollution was adopted in 1979 and the related Gothenburg

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Protocol establishing emission ceilings for sulphur oxides (SO<sub>x</sub>), NO<sub>x</sub>, volatile organic compounds (VOCs) and NH<sub>3</sub> for 2010 negotiated by the EU Member States together with Central and Eastern European countries, the United States and Canada was accepted in 1999 (UNECE, 1999). The National Emissions Ceiling Directive (NECD 2001/81/EC) was introduced in 2001 (EC, 2001) setting national emission ceilings for the EU countries for 2010 and 2020. The implemented mitigation measures have led to significant emission reductions (Grennfelt and Hov, 2005). SO<sub>x</sub> emissions have decreased by 75 %, NO<sub>x</sub> emissions by 42 % and NH<sub>3</sub> emissions by 28 % in the EEA-32 group of countries from 1990–2010 (EEA, 2012). As part of the conventions air pollution monitoring networks have been implemented over Europe providing a long-term observation facility to be able to assess the effectiveness of the implemented air quality management. Although the substantial emission reductions of SO<sub>x</sub>, NO<sub>x</sub> and NH<sub>3</sub> are largely reflected in the trends of pollutant concentrations and wet deposition fluxes, the responses were found to be non-linear (e.g. Lövblad et al., 2004; Fagerli and Aas, 2008; Tørseth et al., 2012; Harrison et al., 2014). These studies highlighted that for SIA and its precursors the implemented emission mitigation measures did not completely meet the expected concentration reduction. Hence, understanding of the non-linear responses is important to be able to provide robust policy support.

Chemistry transport models (CTMs) are used to analyse potential emission reduction strategies and quantify their effectiveness. Before the CTMs can be used to inform policy development they need to be evaluated. Dennis et al. (2010) introduced a comprehensive evaluation framework in which four types of model evaluation are identified: operational, diagnostic, dynamical and probabilistic evaluation. Operational model evaluations have been performed within a huge number of studies using standard statistical and graphical analysis to determine how the model results compare with observations (e.g. Appel et al., 2011, Thunis et al., 2012). Diagnostic model evaluation, focussing on the description of an individual process or component in the model has also been subject of many studies (e.g. Fahey and Pandis, 2003; Redington et al., 2009; Banzhaf et al., 2012). Recently, probabilistic or ensemble based evaluation has

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gained popularity as the ensemble mean of a group of models shows mostly the best model performance in comparison to observations (Vautard et al., 2007; McKeen et al., 2005). Dynamic model evaluations, in which the ability of the modelling system to capture the observed responses to changes in emissions or meteorology is analysed, have only been performed in a few studies so far (e.g. Berglen et al., 2007).

CTMs need to be able to capture non-linear responses of the emission-concentration and emission-deposition relationship as well as interannual variability over the last 15–20 years to provide confidence in the use of CTMs for regulatory purposes (Civerolo et al., 2010). Colette et al. (2011) investigated the capability of six state-of-the-art chemistry transport models to reproduce air quality trends and interannual variability of ozone ( $O_3$ ), nitrogen dioxide ( $NO_2$ ) and  $PM_{10}$  for the time period of 10 years from 1998–2007. They concluded that the models captured most of the important features to justify their implementation for future projections of air quality provided that enough attention is given to their underestimation of interannual variability. Fagerli and Aas (2008) found that the EMEP model's response for nitrogen in air and precipitation to emission changes over Europe from 1980–2003 is reasonable. The results indicated a lack of trends in total nitrate ( $TNO_3$ : sum of aerosol nitrate and gaseous nitric acid) concentrations despite  $NO_x$  emission reductions and it was concluded from the model simulations that this non-linear behaviour can partly be attributed to a shift in the equilibrium between nitric acid ( $HNO_3$ ) and  $NH_4NO_3$  towards particulate phase, which was caused by  $SO_2$  emission reductions. However, the model simulations could not be performed using a consistent meteorological data set for all simulated years. Civerolo et al. (2010) performed an 18-year CMAQ simulation (1988–2005) over the north-eastern United States enabling the investigation of spatial patterns and seasonal variations, but also on long-term trends of  $SO_4^{2-}$  and  $NO_3^-$  in the presence of emissions changes and meteorological variability. The results suggested that the modelling system largely captured the long-term trends in sulphur and nitrogen compounds. While the seasonal changes in sulphur compounds were also captured, the model did not reproduce the average seasonal variation or spatial patterns in  $NO_3^-$ .

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Former studies suggest that the non-linear response of pollutant concentrations to emission changes can be attributed to the differing magnitude of emission reduction for the different substances (Løvblad et al., 2004; Fagerli and Aas, 2008) inducing shifts in the atmospheric chemistry and equilibrium between gas- and particulate phase, which determine the gas to particle conversion. These non-linearities have been also identified in short term modelling studies that focus on the sensitivity of SIA formation to precursor emission reductions (e.g. Erisman and Schaap, 2004; Redington et al., 2009; Derwent et al., 2009; Banzhaf et al., 2013). State of the art labelling approaches (Yarwood et al., 2007; Wagstrom et al., 2008) can be applied to track the source allocation for secondary aerosols and its precursor gases to study the response of atmospheric chemistry to emission changes. However, long-term simulations including a source apportionment have not yet been performed due to the high computational burden. Kranenburg et al. (2013) introduced a source apportionment module for the operational CTM LOTOS-EUROS, which enables long-term simulations with source attribution to investigate possible trends in the gas to particle formation efficiency that accompanied the changes in emission levels over time. We aim to evaluate the LOTOS-EUROS model for its ability to model the trends in SIA concentrations and, at the same time, investigate the non-linearity in SIA formation.

In this study a model run of 20 years from 1990 to 2009 was performed with a horizontal grid resolution of  $0.50^\circ$  longitude  $\times$   $0.25^\circ$  latitude over Europe using the CTM LOTOS-EUROS (Sect. 2.1.1). The model explicitly accounts for cloud chemistry and aerosol thermodynamics. The model run is based on emissions for 1990, 1995, 2000, 2005 and 2010 provided by IIASA (Sect. 2.1.2) and a consistent 3 hourly meteorological data set from 1990 to 2009 obtained from the regional climate model RACMO2 (Sect. 2.1.2) of the Royal Netherlands Meteorological Institute (KNMI). The modelled concentrations of SIA and its precursors are compared to observations at rural background sites (Sect. 2.2). By means of an operational (Sect. 3.1) and a dynamic evaluation (Sect. 3.2) we identify shortcomings and limitations of the model system and input data that need to be improved or considered when using the applied set up for future

emission scenarios. In order to enable the analysis of trends in gas to particle conversion and residence time of the involved species the source apportionment module of LOTOS-EUROS (Sect. 2.1) has been used to trace the amount of SIA formed per unit emission of SO<sub>2</sub>, NO<sub>x</sub> and NH<sub>3</sub> for 4 different regions over Europe from 1990–2009 (Sect. 3.3). The results are discussed and conclusions are drawn in Sect. 4.

## 2 Methods and data

This investigation focuses on SIA and its precursors (SO<sub>2</sub>, NO<sub>x</sub> and NH<sub>3</sub>) over the time period 1990 to 2009. Although the focus is on this 20 year long period we have also investigated the trends in concentrations for the shorter time periods 1995–2009 and 2000–2009 because emission reductions did not proceed linearly and in line with each other from 1990–2009. By considering several time periods we could assess the sensitivity of the trend to the different time periods. Furthermore, the amount of available observations increased for the later periods, which made a broader assessment of the results possible.

In the following subsections the applied model and model set-up, the used observations and the statistic tools we have used to evaluate the model and calculate and assess the observed and modelled trends are described.

### 2.1 Simulation description

#### 2.1.1 Model description LOTOS-EUROS

LOTOS-EUROS (LONg Term Ozone Simulation – EURopean Operational Smog) is a 3-D chemistry transport model. The off-line Eulerian grid model simulates air pollution concentrations in the lower troposphere solving the advection-diffusion equation on a regular lat-lon-grid with variable resolution over Europe (Schaap et al., 2008). In this study, model version 1.8 was used.

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The vertical transport and diffusion scheme accounts for atmospheric density variations in space and time and for all vertical flux components. The vertical grid is based on terrain following vertical coordinates and extends to 3.5 km a.s.l. The model uses a dynamic mixing layer approach to determine the vertical structure, i.e. the vertical layers vary in space and time. The layer on top of a 25 m surface layer follows the mixing layer height, which is obtained from the meteorological input data that is used to force the model. The height of the two reservoir layers is determined by the difference between model top at 3.5 km and mixing layer height. If the mixing layer extends near or above 3.5 km, the top of the model exceeds the 3.5 km according to the above-mentioned description. The horizontal advection of pollutants is calculated applying a monotonic advection scheme developed by Walcek et al. (2000).

Gas-phase chemistry is simulated using the TNO CBM-IV scheme, which is a condensed version of the original scheme (Whitten et al., 1980). Hydrolysis of  $\text{N}_2\text{O}_5$  is explicitly described following Schaap et al. (2004). LOTOS-EUROS explicitly accounts for cloud chemistry computing  $\text{SO}_4^{2-}$  formation as a function of cloud liquid water content and cloud droplet pH as described in Banzhaf et al. (2012). For Aerosol chemistry LOTOS-EUROS features the thermodynamic equilibrium module ISORROPIA2 (Fountoukis and Nenes, 2007). Dry Deposition fluxes are calculated following a resistance approach as described in (Erisman et al., 1994). Furthermore, a compensation point approach for  $\text{NH}_3$  is included in the dry deposition module (Wichink Kruit et al., 2012). The wet deposition module is based on precipitation rates using simple scavenging coefficients for the below cloud scavenging of gases (Schaap et al., 2004) and particles (Simpson et al., 2003).

In LOTOS-EUROS, the temporal variation of the emissions is represented by monthly, day-of-the-week and hourly time factors that break down the annual totals for each source category. An included biogenic emission routine is based on detailed information on tree species over Europe (Koeble and Seufert, 2001). The emission algorithm is described in Schaap et al. (2009) and is very similar to the simultaneously developed routine by Steinbrecher et al. (2009). Sea salt emissions are described using

Martensson et al. (2003) for the particles  $< 1 \mu\text{m}$  and Monahan et al. (1986) for the coarser particles.

LOTOS-EUROS includes a source apportionment module, which enables tracking the origin of the modelled concentrations for different tracers. Using a labelling technique the module calculates the contribution of specified sources for all model grid cells and time steps. The contributions per label are calculated as fractions of the total tracer concentration. The source apportionment module is extensively described in Kranenburg et al. (2013).

The LOTOS-EUROS model has participated in several international model inter comparison studies addressing ozone (Hass et al., 1997; Van Loon et al., 2007; Solazzo et al., 2012a) and particulate matter (Cuvelier et al., 2007; Hass et al., 2003; Stern et al., 2008; Solazzo et al., 2012b) and shows comparable performance to other European models.

### 2.1.2 Model setup

A model run of 20 years from 1 January 1990 to 31 December 2009 has been performed on a domain covering Europe ( $35^\circ\text{N}$ – $70^\circ\text{N}$ ;  $10^\circ\text{W}$ – $40^\circ\text{E}$ ) with a horizontal resolution of  $0.50^\circ$  longitude  $\times$   $0.25^\circ$  latitude on a rectangular regular latitude-longitude grid (ca.  $25\text{ km} \times 25\text{ km}$ ). As described above the lowest dynamic layer is the mixing layer, taken from the meteorological input.

The simulation was forced with a consistent meteorological data set from 1990 to 2009 obtained from the regional climate model RACMO2 (Lenderink et al., 2003; Van Meijgaard et al., 2008) of the KNMI. At the boundaries the simulation was driven by meteorology from ERA-Interim reanalysis (Dee et al., 2011). RACMO2 has a horizontal resolution of  $0.44^\circ$  with 114 points distributed from  $25.04^\circ\text{W}$  to  $24.68^\circ\text{E}$  longitude and 100 points from  $11.78^\circ\text{S}$  to  $31.78^\circ\text{N}$  latitude in the rotated grid. The South Pole is rotated to  $47^\circ\text{S}$  and  $15^\circ\text{E}$ . In the vertical, 40 pressure levels were used. As described in Manders et al. (2012) the horizontal projection of RACMO2 fields on the LOTOS-EUROS grid was carried out by bi-linear interpolation. The vertical projection

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of RACMO2 profiles on the much coarser LOTOS-EUROS vertical grid was achieved by mass-weighted averaging of those RACMO2 model layers that were fully or partially contained in each of the LOTOS-EUROS model layers. At the applied resolution RACMO2 uses a model time step of 15 min and output for coupling with LOTOS-EUROS was generated every three hours. RACMO2 has been included in ensemble studies with other regional climate models (Kjellström and Giorgi, 2010; Kjellström et al., 2010; Vautard et al., 2013; Kotlarski et al., 2014) and has been successfully applied to force LOTOS-EUROS in earlier studies (Manders et al., 2011, 2012; Mues et al., 2013).

Lateral boundary conditions in LOTOS-EUROS were taken from climatological background concentrations for gases and aerosols. Some aerosols species, heavy metals and pops are set to constant at the boundaries. The climatology fields did not include windblown dust going back to 1990. Hence, dust from e.g., wind erosion, agricultural land management and resuspension by road transport has been neglected, as it does not contribute to the here investigated substances. For the interpretation of the model results we need to keep in mind that there are no trends in boundary conditions over the investigated 20 year period.

The emissions applied in this study were provided by the International Institute for Applied Systems Analysis (IIASA). The data was generated using RAINS (1990–2000) and GAINS (2000–2010) model output. A description of the RAINS model and the GAINS model can be found in Amann et al. (1999) and Amann et al. (2011), respectively. Annual total emissions were provided per country, per sector and per SNAP (Selected Nomenclature for Air Pollutants) code for 1990, 1995, 2000, 2005 and 2010. A linear interpolation was performed to fill in the emissions of the years within the delivered ones. Figure 1a shows the trends in  $\text{SO}_2$ ,  $\text{NO}_x$  and  $\text{NH}_3$  emissions in the EU-27 member States including Norway and Switzerland (= EU-27+) for 1990 to 2010 in % with 1990 as reference derived from the applied final emission inventory. The emissions have decreased over Europe for all considered components. The slope of the decrease has been computed using a standard linear least square method. Most

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emission reduction was achieved for SO<sub>2</sub> with a negative trend of  $-3.9\% \text{ a}^{-1}$  (a: annum) leading to a decrease of more than 70 % from 1990 to 2010. NO<sub>x</sub> emissions have been decreased by somewhat less than 50 % in the same time period ( $-2.52\% \text{ a}^{-1}$ ) followed by NH<sub>3</sub> emissions with a decrease of somewhat less than 20 % from 1990 to 2009 ( $0.85\% \text{ a}^{-1}$ ). In Fig. 1a we present results for the emission trends since 1990 for the EU-27+ member States as a whole. While it is known that emission changes from 1990 to 2009 differed significantly from region to region, precise information on the spatial distribution of the emissions for the early 90s is lacking. Hence, we used the TNO MACC (Denier van der Gon et al., 2010; Pouliot et al., 2012) spatial distribution of emissions for the year 2005 for the entire time period of investigation. Annual emissions from international shipping per sea and per sector were provided by the Centre on Emission Inventories and Projections (CEIP). Figure 1b shows the trends in SO<sub>2</sub> and NO<sub>x</sub> International Shipping emissions for 1990 to 2010 in % with 1990 as reference. Included are the Baltic Sea, the North-East Atlantic Ocean, the North Sea, the Mediterranean Sea and the Black Sea. NO<sub>x</sub> emissions have increased over the whole time period 1990 to 2009 for all seas while SO<sub>2</sub> emissions have increased for the North-East Atlantic Ocean, the Mediterranean Sea and the Black Sea. In the Sulphur Emission Control Areas of the North Sea (“NOS” in Fig. 1b) and the Baltic Sea (“BAS” in Fig. 1b) SO<sub>2</sub> emissions have increased from 1990 to 2005 and decreased thereafter due to improved fuel quality.

In order to analyse the trends in gas to particle conversion and residence time of the involved species the LOTOS-EUROS source apportionment module was applied. We defined 5 labels for tracking 10 kilo tons (ktons) of SO<sub>2</sub>, NO<sub>x</sub> and NH<sub>3</sub> emissions from either one of these. The labels were defined to represent the following geographical areas:

1. the Netherlands and Belgium
2. Baltic Sea (international shipping)
3. Czech Republic



#### 4. Romania

#### 5. Rest

Ten ktons of precursor emission were chosen, as it is certainly smaller than the single country annual total emissions for 2009. Together with the simulation of each substance in each grid cell on hourly basis, the fractional contribution of each of the above labels to every substance, including  $\text{SO}_4^{2-}$ ,  $\text{NO}_3^-$  and  $\text{NH}_4^+$ , is calculated. By means of the latter the amount of SIA formed from the 10 ktons of precursor gases can be derived for each label and possible trends in gas to particle conversion within the time period 1990 to 2009 can be analysed.

### 2.2 Observations

In the following subsections we describe the in-situ surface observations that were used to evaluate the LOTOS-EUROS model and to derive the observed trends in SIA and its precursors concentrations (Sect. 2.2.1) and the observations used to compare to the meteorological input data provided by RACMO2 (Sect. 2.2.2).

#### 2.2.1 Species concentrations

The European EMEP observational network is devised for trend assessment (EMEP/CCC, 2001; Hjellbrekke and Fjærraa, 2011). The EMEP data is validated through a quality assurance/quality control process involving the individual institutions responsible for the different sites and the EMEP-CCC as documented by several reports available on the EMEP website ([www.emep.int](http://www.emep.int)). Data was downloaded from the EBAS repository (<http://ebas.nilu.no/>, download in autumn 2012). However, only a few selected stations per country are included in the network. In addition to the EMEP sites, the stations of AirBase (European AIR quality database), the public database of the European Environmental Agency (EEA), were added to the observational data set (<http://airbase.eionet.europa.eu/>, download in autumn 2012). The latter are not

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specifically devised for trend assessment but have been used in several studies on long-term trends (e.g. EEA, 2009; Colette et al., 2011; Wilson et al., 2012). The data reported to AirBase are quality controlled and checked prior to submission by the countries that provide the data.

This study is aimed to investigate the transboundary trend of concentrations in the European background following emission changes all over Europe from 1990–2009. Hence, only rural background stations are included in the applied observational data set. The analysis is based on daily observations. The consistency of the observational data set used for the trend assessment and the operational and dynamical model evaluation was ensured by the implementation of three selection criteria derived from the guidelines of the European Environmental Agency (EEA, 2009; Colette et al., 2011):

1. The annual coverage of data must be larger than 75 %
2. With criterion No. 1 fulfilled, at least 80 % of the annual time series must be available
3. Passing a visual screening of the data

For each time period (1990–2009, 1995–2009 and 2000–2009) a separate data subset of stations within the model domain (35° N–70° N; 10° W–40° E) was built based on the selection criteria described above. As the highest variability is expected in the beginning of each of the studied time periods only stations that could provide the requested 75 % data coverage for the first year of the time period were included in the corresponding subset.

Finally a visual screening of the time series of daily observations for all species and at all stations that had passed the selection criteria described above was performed. Surprisingly many defective time series have been identified. The corresponding stations have been removed from the subsets. The most frequently reasons for removal from the data set were high detection limits throughout the time series leading to disappearing concentration regimes, high amounts of implausible outliers/peaks and constant value signals over long time periods. The data reliability is further discussed in Sect. 4.

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It was found that due to a lack of data the analyses of  $\text{NH}_3$  observations could not be included in the study. However, total ammonia ( $\text{TNH}_4$ : sum of aerosol ammonium and gaseous ammonia) observations were included in the trend assessment as considerably more stations with  $\text{TNH}_4$  observations than with  $\text{NH}_4$  observations were available.

The latter was also the case for  $\text{TNO}_3$  and  $\text{NO}_3$ . Hence, the considered observed components within this study are  $\text{SO}_2$ ,  $\text{SO}_4^{2-}$ ,  $\text{NO}_2$ ,  $\text{TNO}_3$  and  $\text{TNH}_4$ . In the Supplement, Figs. S1 to S3 show maps of the locations of the observational sites used for the analysis for the different components and the different time periods. Table 1 summarizes the number of stations for the different species and subsets before and after the visual screening. The number of discarded stations is highest for  $\text{SO}_2$  and  $\text{NO}_2$ . For both components a large part of the considered stations are from AirBase passing through a less stringent quality control process than EMEP stations.

Due to a lack of long-term monitoring sites within Great Britain, France, Spain and the Mediterranean region within the monitoring networks used in this study the majority of sites for  $\text{SO}_2$  and  $\text{NO}_2$  observations is located within central Europe accompanied by several sites in northern and eastern Europe. For both components no southern European station and in the case of  $\text{NO}_2$  no western European station was available for comparison for the 20 years period. For the time period 1995–2009 an increasing number of eastern and western European stations and in the case of  $\text{SO}_2$  one southern European station passed the selection criteria. For  $\text{TNO}_3$  and  $\text{TNH}_4$  additionally to the lack of long-term observations in southern and western Europe a lack of observations in central Europe was found and the majority of sites is located in northern and eastern Europe. Stations in  $\text{NH}_3$  hot spot regions like e.g. the Netherlands or the Po valley did not pass the data selection criteria for any of the time periods. Also for  $\text{SO}_4^{2-}$  no southern European station was available for 1990–2009. The available stations are distributed over Western, Eastern and Northern Europe with most stations being located in Northern Europe. For 1995–2009 central and eastern European stations and one southern European station could be included in the analysis. We would like to stress

that the stations at which  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  concentrations are investigated may partly differ.

Finally, for the time period 2000–2009 few southern European stations could be included in the analysis of all considered components. Furthermore, Fig. S4 in the Supplement shows for each component those stations that pass the data selection criteria for all considered time periods.

## 2.2.2 Meteorological observations

Selected parameters of the RACMO2 model are compared to observations to be able to assess the ability of the model to capture the observed meteorological seasonal, annual and interannual variability. For the evaluation, data of the European Climate Assessment and Dataset (ECA&D) project (Klok and Klein Tank, 2009) is applied. The project was initiated by the European Climate Support Network (ECSN) and is funded by and coordinated at the KNMI. A compilation of daily observations obtained from climatological divisions of national meteorological and hydrological services, observatories and research centres throughout Europe and the Mediterranean are included in the database. The data series of observations is combined with quality control and analysis of extremes via climate change indices (Klein Tank et al., 2002).

Daily observed series of 4 parameters that affect atmospheric chemistry have been extracted from the dataset for the years 1990 to 2009 for evaluation purposes: Temperature (at 2 m), relative humidity (at 2 m), wind speed (at 10 m) and precipitation. For each parameter a selection of stations was extracted so that, if available, central, northern, eastern, southern and western European stations were included in the analysis to also enable a regional consideration. For relative humidity no northern European stations could be included and western European stations were rare concerning observations of relative humidity and wind speed. In total 206 stations were selected for the evaluation of modelled temperature, 113 stations for the evaluation of modelled relative humidity, 246 stations for the evaluation of modelled wind speed and 240 stations for

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the evaluation of modelled precipitation. The observed station data is compared with model data at the nearest gridpoint.

## 2.3 Statistical measures and methods for evaluation and trend assessment

For the evaluation of the used meteorological input provided by RACMO2 and the resultant concentrations simulated by LOTOS-EUROS three statistical measures have been applied to assess the ability of the models to reproduce the observed values:

### 1. Correlation coefficient $r$

$$r = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2 \sum_{i=1}^n (y_i - \bar{y})^2}} \quad (1)$$

### 2. Root mean square error (RMSE)

$$\text{RMSE} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - y_i)^2} \quad (2)$$

### 3. Bias

$$\text{BIAS} = \frac{1}{n} \sum_{i=1}^n (x_i - y_i) \quad (3)$$

where  $x$  is the model output vector and  $y$  its observation counterparts. Each vector has  $n$  elements and  $\bar{x}$  and  $\bar{y}$  represent their mean value. The correlation coefficient (Eq. 1) has been applied to assess the simulated temporal variability and the RMSE (Eq. 2)

and bias (Eq. 3) to assess the simulated absolute values. The evaluation of RACMO2 and LOTOS-EUROS fields is based on daily averages.

The trends in concentrations are computed using annual averages based on daily data. The slope is calculated using a standard linear least square method. Within this study we computed only linear trends and the computation of non-linear trends (Konovalov et al., 2010) or piecewise linear trends (Carslaw et al., 2011) has not been performed. To assess the significance of the trend a Mann–Kendall test at the 95 % confidence level is performed (Kendall, 1976; Hipel and McLeod, 2005).

### 3 Results

#### 3.1 Evaluation of model results

##### 3.1.1 Evaluation of meteorological fields

The applied meteorological input data has been compared to observations to be able to assess the ability of RACMO2 to reproduce the observed meteorological annual, interannual and seasonal variability. In order to limit the length of this article only an abridgement of the performed evaluation is shown here. Four parameters that considerably impact atmospheric chemistry are shown: temperature (at 2 m), relative humidity (at 2 m), wind speed (at 10 m) and precipitation. The evaluation is based on daily data for the 20 years period. Table 2 summarizes the number of stations, the mean correlation coefficient, the observed mean and RMSE and bias. As an example Fig. 2 shows the 60 days moving average of the four parameters averaged across all available German stations from 1990–2009. The 60 days moving average was chosen to be able to plot the whole time series in one graph and, at the same time, to be able to see variability in the time series.

As the mean correlation coefficient of 0.97 in Table 2 shows the model captures very well the temporal distribution of temperature for the considered time period. Figure 2a

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shows that the interannual variability is simulated fairly well too. Warm summers like in 2003 and 2006 and cold winters like the one in 1995/1996 are well reproduced by RACMO2. However, the bias and also the corresponding graph in Fig. 2 indicate a slight underestimation of the temperature during wintertime in central Europe. The performance of the model has also been assessed regionally for Northern, Eastern, Southern, Western and Central Europe separately (not shown here). The underestimation during wintertime was found to be most distinct for southern and least distinct for northern Europe, which is consistent with findings in van Meijgaard et al. (2012) and Kotlarski et al. (2014).

As Fig. 2b illustrates RACMO2 captures the interannual variability of the relative humidity at 113 European stations less well than that of the temperature. The latter was found to be most evident for southern Europe. Also the relative humidity is overestimated during wintertime, which is again most distinct for southern Europe and may be connected to the underestimation of the temperature during wintertime. Relative humidity is a difficult quantity to evaluate, in particular in areas or during episodes with high values of relative humidity ( $> 95\%$ ). However, a mean correlation coefficient of 0.66 (Table 2) indicates that the observed temporal variability is satisfactorily simulated by the model.

The temporal variability of the wind speed is also satisfactorily simulated with a mean correlation coefficient of 0.68 over 246 European stations (Table 2). Figure 2c displays the mean 60 days moving average of wind speed for 59 German stations for the investigated time period. The graph indicates that although the timing is well simulated the model tends to overestimate the wind speed in central Europe. In central and eastern Europe the overestimation was found to be present throughout the whole year. In northern and southern Europe RACMO2 overestimates wind speed solely during wintertime while it tends to slightly underestimate wind speed during summertime.

Figure 2d shows the mean 60 days moving average of precipitation for 1990–2009 at 66 German stations. The figure shows that the interannual variability is modelled satisfactorily in central Europe although it is slightly underestimated. Dry years like

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1996, 2003 and 2006 are well reproduced by the model. RACMO2 underestimates summertime precipitation in southern Europe while it tends to overestimate wintertime precipitation in northern and central Europe, which was also found by van Meijgaard et al. (2012) and Kotlarski et al. (2014). Generally, moving from daily to monthly or annual precipitation sums (not shown here) RACMO2 results compare better to the observed values. Mean correlation coefficient, RMSE and bias have been calculated at 240 European stations (Table 2). The mean correlation of 0.48 indicates that considering the high temporal variability of precipitation RACMO2 simulates the observed timing reasonably well.

For the CTM calculation it is more important to capture the occurrence of precipitation than to capture its intensity and duration with the meteorological driver as wet deposition is a very efficient removal process. Therefore, at each of the 240 stations it was investigated on which percentage of days of the 20 years period the model is able to simulate the observed rain occurrence (rain: yes; rain: no). In the following a correct modelled rain: yes or rain: no is referred to as “hit”. To account for unphysical small amounts of drizzle that often occur in climate models, daily accumulated precipitation below 0.5 mm was considered as no rain. The results are summarised in Table 3. At 205 out of 240 stations the model is able to correctly simulate the rain occurrence on more than 70 % of the days from 1990–2009.

Although some shortcomings in the meteorological input fields were found the outcome of the evaluation of RACMO2 has shown that the model is capable of satisfactorily reproducing the observed magnitudes and meteorological annual, interannual and seasonal variability of the investigated parameters.

### 3.1.2 Concentrations in air

Figure 3 shows for each component the 60 days moving average concentrations averaged across all available stations for the period 1990–2009. Besides the time series the average seasonal variation is given for this same 20 year time period. The summary

of the statistical evaluation based on daily pairs of observed and measured concentrations is given in Table 4 for the 1990–2009, 1995–2009 and 2000–2009 time periods.

The modelled time series of  $\text{SO}_2$  by LOTOS-EUROS underestimates the observed  $\text{SO}_2$  concentrations (Fig. 3a) in the period 1990–1997, while for later years there appears to be a small bias at these stations. Throughout the time series the year-to-year variability is captured well by the model, as is the seasonal variation. The concentrations of  $\text{SO}_4^{2-}$  are systematically underestimated by LOTOS-EUROS throughout the whole period. The underestimation is most distinct from 1990–1997, which appears to be related to the underestimation of  $\text{SO}_2$  in the same period. Analysis of the individual sites showed that the sites located in eastern and central Europe largely determine the underestimation for both components as northern European stations show much better comparison. The correlation coefficient of 0.5 for  $\text{SO}_2$  for 1990–2009 suggests that the timing is not very well captured throughout the time period. Again a separation in time and region was detected. For stations in central and eastern Europe for the years 1990–1997 average correlation coefficients of 0.40 and 0.54 were found, whereas these increase to 0.54 and 0.69 for central and eastern Europe, respectively, for 1998–2009. Also for  $\text{SO}_4^{2-}$  the correlation coefficient is improved for the 1998 to 2009 time period (0.56) compared to the 1990 to 1997 time period (0.50). We connect the difference in the model performance for these periods to the lack of a good representation of the change in emission structures in the power sector in eastern and parts of central Europe as a consequence of the fall of the Berlin wall and political changes associated with the liberalisation of the Eastern Bloc's authoritarian systems as discussed below.

A striking feature in the comparison for  $\text{SO}_4^{2-}$  (Fig. 3c) is the inability of the model to reproduce the magnitude of several spring episodes that occurred in e.g. 1996, 2003 and 2006. Although for some of these episodes the model is able to capture the timing, it is not able to reproduce the peak values. These episodes are characterized by very stable conditions across central Europe and some have been studied in detail (e.g. Stern et al., 2008; Banzhaf et al., 2013). A model comparison by Stern et al. (2008)

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(Table 5). Comparison between the components shows that the model skill for nitrogen components is larger than for the sulphur compounds.

### 3.2 Trends in concentrations

The observed and modelled trends are illustrated in Figs. 4, 5 and S5. Figure 4 shows boxplots of the absolute and relative observed and modelled trends for all components. Boxplots are used to provide a concise illustration of a higher amount of data points. We are aware that hence, for the components  $\text{TNO}_3$  and  $\text{TNH}_4$  a boxplot is not necessary for the time periods 1990–2009 and 1995–2009. The median trend for 7, 8 or 9 stations is robust but the calculation of the 25th and 75th percentile is of no significance. Being aware of the latter, boxplots for  $\text{TNO}_3$  and  $\text{TNH}_4$  have still been included for all time periods to be able to compare the trends to those of the other components. Table 6 summarizes the observed and modelled median trends for the three considered time periods.

Figure 5 shows scatter plots of the observed vs. modelled trend for the studied components at the considered stations for the 3 different time periods. It is labelled in the graphs if the observed and/or modelled trends are significant (method used described in Sect. 2.3): (+) implies that the observed and the modelled trends are significant, (o) implies that the observed trend is non-significant while the modelled trend is significant, (o) implies that the observed trend is significant while the modelled trend is non-significant and (o) implies that the observed and the modelled trends are non-significant.

Figure S5 shows the observed and modelled trends of the annual mean  $\text{SO}_4^{2-}$ ,  $\text{TNO}_3$  and  $\text{TNH}_4$  concentrations, their 5th and 95th percentile and the corresponding trend lines. Solid lines refer to significant trends and dashed lines refer to non-significant trends (only found for the  $\text{TNO}_3$  5th percentile).

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### 3.2.1 Observed trends

Figure 5 illustrates that the observed  $\text{SO}_2$ ,  $\text{SO}_4^{2-}$  and  $\text{NO}_2$  concentrations show significant negative trends at the majority of stations for the time periods 1990–2009 and 1995–2009. For  $\text{NO}_2$  a significant positive trend for 1995–2009 was observed at two stations located in Estonia at the shore of the Baltic Sea. For  $\text{TNO}_3$  and  $\text{TNH}_4$  the majority of trends is significant negative for the 1990–2009 time period while for 1995–2009 the observed trends are non-significant at all stations ( $\text{TNO}_3$ ) or at the majority of stations ( $\text{TNH}_4$ ). We would like to stress again that for  $\text{TNO}_3$  and  $\text{TNH}_4$  the considered station locations are situated in northern and eastern Europe due to a lack of long-term observations in the other regions. The trends in  $\text{TNO}_3$  in hot spot areas like the Netherlands may differ. For the time period 2000–2009 the observed trends are non-significant at the majority of stations for all considered components. Furthermore, the relative amount of stations with non-significant trends increases when moving from 1990–2009 ( $\text{SO}_2$ : 0 %;  $\text{SO}_4^{2-}$ : 0 %;  $\text{NO}_2$ : 11 %;  $\text{TNO}_3$ : 33 %;  $\text{TNH}_4$ : 14 %) to 1995–2009 ( $\text{SO}_2$ : 5 %;  $\text{SO}_4^{2-}$ : 18 %;  $\text{NO}_2$ : 21 %;  $\text{TNO}_3$ : 100 %;  $\text{TNH}_4$ : 50 %) to 2000–2009 ( $\text{SO}_2$ : 52 %;  $\text{SO}_4^{2-}$ : 86 %;  $\text{NO}_2$ : 72 %;  $\text{TNO}_3$ : 75 %;  $\text{TNH}_4$ : 80 %).

Figure 4 and Table 6 show that for all components the observed median absolute and relative negative trends decrease moving from 1990–2009 to 2000–2009 (absolute decrease in  $\text{TNO}_3$  trends in the 3rd decimal place). For  $\text{SO}_2$  and  $\text{NO}_2$  the observed trends slightly decrease for 1995–2009 compared to 1990–2009 but decrease considerably for the time period 2000–2009. Comparing the observed absolute and relative trends in  $\text{SO}_4^{2-}$  concentrations to those of  $\text{SO}_2$  shows that the trends in  $\text{SO}_4^{2-}$  are considerably lower. However, the relative trends for  $\text{SO}_4^{2-}$  concentrations are higher than those for  $\text{NO}_2$ ,  $\text{TNH}_4$  and  $\text{TNO}_3$ . For the absolute trends, the interquartile range, which is an indicator for the spread of the data, decreases for  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  when moving to later time periods although the number of stations increases while for  $\text{NO}_2$  the interquartile range and therewith the spread of the distribution remains almost stable.

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This difference may be induced by the fact that the variation of emission changes from country to country is larger for  $\text{NO}_x$  compared to  $\text{SO}_x$  (Løvblad et al., 2004).

Considering only those stations that fulfilled the selection criteria for all 3 time periods (not shown here) has shown that an increasing number of non-significant trends, the decrease of median absolute and relative trends and the decreasing (stable) spread of the trends for  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  ( $\text{NO}_2$ ) when moving from 1990–2009 to 2000–2009 are also present when considering the same subset of stations.

Furthermore, the sensitivity of the resultant observed median trends to the selection criteria introduced in Sect. 2.2.1 has been tested. The results for the 1990 to 2009 time period are presented in the Supplement showing that increasing the length of the annual time series (at least 80 % of the considered time period was the criterion given in Sect. 2.2.1) has a minor impact on the resultant median trend.

### 3.2.2 Modelled trends and comparison to observed trends

As Fig. 4 shows the model is able to well simulate the decrease in the absolute median negative trend for  $\text{SO}_2$ ,  $\text{SO}_4^{2-}$  and  $\text{NO}_2$  when moving from 1990–2009 to 1995–2009 to 2000–2009. As for the observed trends the model simulates only a slight decrease of the absolute negative median trend in  $\text{SO}_2$ ,  $\text{SO}_4^{2-}$  and  $\text{NO}_2$  concentration from 1990–2009 to 1995–2009 but a significant decrease from 1995–2009 to 2000–2009. Also, the model is able to reproduce the lower relative trends in observed  $\text{SO}_4^{2-}$  concentrations compared to those of  $\text{SO}_2$  (see Fig. 4b, d and f). Furthermore, the model is able to capture the decrease in the spread of the data distribution for  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  when moving from 1990–2009 to 1995–2009 to 2000–2009. However, the lower distance between upper and lower whisker and the lower or non-existent amount of outliers reveals that for both considered sulphur compounds the distribution of the modelled absolute and relative trends is less broad than that of the observed trends. This indicates that the modelled trends show a lower spatial variance over Europe than the observed trends. The latter is vice versa for  $\text{NO}_2$  concentrations for which especially for the early time

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periods the interquartile range of the modelled absolute and relative trends is larger than that of the observed trends.

The model simulates significant negative trends in  $\text{SO}_2$ ,  $\text{NO}_2$  and  $\text{SO}_4^{2-}$  concentrations at most station locations for 1990–2009 and 1995–2009 (see Fig. 5), which coincides with the observed trends for these time periods. However, the model underestimates the negative trends in concentrations for  $\text{SO}_2$  at several stations and for  $\text{SO}_4^{2-}$  at most stations (note that the stations at which  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  concentrations are investigated may partly differ) while it overestimates the negative trends in  $\text{NO}_2$  concentrations at the majority of station locations. For all considered time periods the deviation of the modelled trends in  $\text{SO}_2$ ,  $\text{SO}_4^{2-}$  and  $\text{NO}_2$  concentrations from the observed trends were found to be most distinct at eastern European stations and stations in north-eastern Germany (e.g. the three outliers in Fig. 5b correspond to trends at 2 stations in Czech Republic and one station in eastern Germany) and least distinct at northern European station locations. The findings from Fig. 5 are also reflected in the median absolute and relative trends presented in Fig. 4. Furthermore, the figure shows that for  $\text{SO}_2$  and  $\text{SO}_4^{2-}$  the simulated trends improve considerably for the 2000–2009 compared to the earlier time periods while the trends in  $\text{NO}_2$  are overestimated by the model throughout all time periods.

The strong observed negative trend in  $\text{SO}_4^{2-}$  concentrations is mostly driven by the high observed concentrations in the beginning of the 90s as Fig. S5a shows. The latter high observed concentrations could not be reproduced by the model. The 5th percentile, which represents the background concentrations, and its significant negative trend are well captured by the model. The negative trend of the 95th percentile, which represents the high concentration range (the peak  $\text{SO}_4^{2-}$  concentrations), is significantly underestimated by the model. The model satisfactorily captures the temporal distribution of the interannual variability but there is a significant negative bias between modelled and observed value. This shows that the models inability to capture the observed trend in  $\text{SO}_4^{2-}$  is driven by the underestimation of the high range of concentrations.

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Also for TNO<sub>3</sub> and TNH<sub>4</sub> shown in Fig. S5b and c the deviation from the observed values is most distinct in the 95th percentile while the interannual variability is well simulated by the model. However, for both components but especially for TNO<sub>3</sub> the model compares considerably better to the observations than for SO<sub>4</sub><sup>2-</sup>. Also Fig. 5 shows that the model well reproduces the low trends in TNO<sub>3</sub> concentrations at the majority of considered sites for all time periods while for TNH<sub>4</sub> the model tends to underestimate the observed concentration trends. Furthermore, for both components, TNO<sub>3</sub> and TNH<sub>4</sub>, the increased relative number of non-significant trends when moving from the 1990–2009 to the 2000–2009 time period is well captured by the model at most stations. While for TNO<sub>3</sub> and TNH<sub>4</sub> concentrations the model captures the non-significance of trends at the majority of stations for the time period 2000–2009 the model underestimates the number of stations with observed non-significant concentration trends for SO<sub>2</sub>, SO<sub>4</sub><sup>2-</sup> and NO<sub>2</sub>.

### 3.3 Trends in SIA formation

In Fig. 1a we have shown the relative emission decrease of SIA precursors based on the years 1990, 1995, 2000, 2005 and 2010 with 1990 as reference for the EU-27+ member States. A trend line is added for each component. Figure 6a and b show the observed and modelled annual average concentrations of SO<sub>2</sub>, NO<sub>2</sub>, SO<sub>4</sub><sup>2-</sup>, TNO<sub>3</sub> and TNH<sub>4</sub> relative to the annual average concentration of 1990 at 23, 37, 15, 9 and 7 European stations, respectively, for the years 1990, 1995, 2000, 2005 and 2009. Trend lines are added for each component and the corresponding numbers are given at the lower left of each panel. Comparing Fig. 6a and b reveals that LOTOS-EUROS well simulates the observed relative trends. The curves for TNO<sub>3</sub> and TNH<sub>4</sub> demonstrate that the model is capable to capture meteorological variability. Together with the precursor emissions over land (Fig. 1a) the observed concentrations of all considered components show negative trends for 1990 to 2009. The observed decrease in SO<sub>2</sub> concentrations is for the considered time period from 1990–2009 with  $-4.4\% \text{ a}^{-1}$  even

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larger than the corresponding decrease in SO<sub>2</sub> emissions (−3.9% a<sup>−1</sup>). In contrast to SO<sub>2</sub>, the decrease in observed SO<sub>4</sub><sup>2−</sup> concentrations is with −3% a<sup>−1</sup> less distinct and smaller than the decrease in SO<sub>2</sub> emissions. This non-linear effect is well reproduced by LOTOS-EUROS.

The LOTOS-EUROS source apportionment module was used to further investigate the observed and modelled non-linearity. Therefore 10 ktons of SO<sub>2</sub>, NO<sub>x</sub> and NH<sub>3</sub> emissions, respectively, have been tracked for 1990 to 2009 for 4 different labels, which were chosen to be 4 different regions: the Netherlands and Belgium (NLBE), the Baltic Sea (BAS), Czech Republic (CZE) and Romania (ROM). By means of the labelling we can determine how much SIA was formed from the 10 ktons precursor emissions during the time period from 1990 to 2009. The results of the source attribution are presented in Fig. 7. Figure 7a shows the amount of SO<sub>4</sub><sup>2−</sup> (solid lines) formed from 10 ktons SO<sub>2</sub> emissions relative to the amount formed from 10 ktons SO<sub>2</sub> emissions in 1990 for the different labels for 1990 to 2009. A trend line (dashed line) is added for all labels. For all considered regions the SO<sub>4</sub><sup>2−</sup> formation efficiency increases from 1990 to 2009. Following the Mann–Kendall Test at a 95% confidence level the positive trends are significant for all labels. To investigate if the identified increase is a matter of climate change we re-run the model for 1990, 1995, 2000 and 2009 using the emissions for the corresponding year but the meteorology of 2005. The results are added to Fig. 7a as accordingly coloured dots for each label. Most dots are located on or close to the corresponding trend line. The latter indicates that the increase in SO<sub>4</sub><sup>2−</sup> formation efficiency is induced by the change in emissions from 1990 to 2009. The increase is most distinct for the region NLBE with a 61% more efficient SO<sub>4</sub><sup>2−</sup> formation in 2009 compared to 1990 followed by CZE (+60%), BAS (+31%) and ROM (+28%). The major driver for the increased SO<sub>4</sub><sup>2−</sup> formation efficiency in the model has been an increasing neutralisation of cloud acidity and thus pH over time as diagnosed from the model run.

SO<sub>4</sub><sup>2−</sup> formation is a sink for SO<sub>2</sub> concentrations and therefore the increase in SO<sub>4</sub><sup>2−</sup> formation efficiency explains that the decrease in SO<sub>2</sub> concentrations is larger than expected solely from the decrease in SO<sub>2</sub> emissions. Figure 7b displays the decrease

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in SO<sub>2</sub> quantity per unit SO<sub>2</sub> emission showing a negative trend for the time period 1990 to 2009 for all considered labels. However, for the Baltic Sea (BAS) the trend from 1990 to 2009 is not significant following a Mann–Kendall Test at the 95 % significance level.

Figure 7c reveals a decrease in NH<sub>4</sub><sup>+</sup> formation per unit NH<sub>3</sub> emission for the labels NLBE, CZE and ROM with a reduction of –22 % (ROM) to –33 % (NLBE and CZE) for 2009 compared to 1990. Following a Mann–Kendall Test at the 95 % significance level the trend is significant for these labels. BAS is not included in the figure as there is no NH<sub>3</sub> emission from shipping on the Baltic Sea.

The changes in NO<sub>3</sub><sup>-</sup> formation efficiency from 1990 to 2009 are lower than for SO<sub>4</sub><sup>2-</sup> and NH<sub>4</sub><sup>+</sup> (Fig. 7d). A significant trend has been found for the label NLBE showing an increase in NO<sub>3</sub><sup>-</sup> formation efficiency with an increase of +22 % from 1990 to 2009. In the next section the results of the labelling exercise are further discussed.

## 4 Discussion and conclusions

In this study we presented a dynamic model evaluation of the LOTOS-EUROS CTM to analyse the ability of the model to reproduce the non-linear responses to emission changes and inter-annual variability of SIA and its precursors over Europe from 1990 to 2009. This study presents the first evaluation of the model system over such a long time period.

With respect to the study design we feel that the simulation of the whole period is a strong point as opposed to using one or several key meteorological years to study the impact of emission changes as it is difficult to choose a meteorological year that is representative for an average year throughout Europe. In addition, through the reanalysis with RACMO2 we have used a consistent set of meteorological data to drive the model for the whole period. The major activity needed to improve the study design is associated with the emission information for the early nineties. Improvements are especially needed for the eastern European countries. Emission estimates for 1990 are relatively uncertain (Granier et al., 2011) as much of the information currently used to estimate

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emissions is not available (at the same quality) for 1990. Moreover, we have simply used the spatial allocation of the TNO-MACC-2005 dataset and scaled the emission totals per sector back to those of 1990. As a result, the (spatial) representation of e.g. the industrial infrastructure and location of power plants, especially in eastern and parts of central Europe in the period 1990–2000 will not be correct as the infrastructure here during this period still resembled the pre-1990 period. The improvement needed here is highlighted by the higher underestimation of the (primary) pollutants and the lower model skill in the first years of the study period. One could use the spatial allocation of emission inventories built in the nineties to overcome these problems partly. Making a small compromise on the spatial resolution of the data may not be a large problem as model resolution does hardly affect the performance of CTMs for regional assessments (Schaap et al., 2014).

Complementing the EMEP monitoring data with those of AIRBASE has increased the number of stations with valid time series, especially for the precursor gases. Our visual screening of the measurement data revealed that a large fraction of the stations with long time series were not useable as data quality was obviously an issue. The most frequent peculiarities were shifts in the concentration level, many implausible peaks of short duration, constant value signals over prolonged time periods or concentration regimes below the detection limit. Most problems were associated to time series of SO<sub>2</sub>. The number of defective time series was highest for the 1990's and decreases considerably towards 2009. The improved quality of observations may also contribute to the improved model performance for the more recent period. A lack of a long-term time series for southern and parts of western and eastern Europe hampered an evaluation across the full European domain. Furthermore, for concentrations of NO<sub>3</sub><sup>-</sup> and NH<sub>4</sub><sup>+</sup> there is a lack of observations with separation between gas- and aerosol-phase. Additional efforts for data mining within European countries could yield larger observational basis for evaluation of the time period. Moreover, generation of a centralized dataset for the specific purpose of evaluation long-term trends could be a means to improve the data quality by incorporation of expertise from the data providers.

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The operational model evaluation showed that the day-to-day variability, as well as the seasonal and the inter-annual variability are satisfactorily simulated for all components. Within a multi-model trend assessment study Colette et al. (2011) presented the ability of 6 state of the art CTMs to simulate the seasonal cycle of amongst others SO<sub>2</sub>, SO<sub>4</sub><sup>2-</sup>, NO<sub>2</sub>, TNO<sub>3</sub> and TNH<sub>4</sub> concentrations at European rural background stations for the time period 1998–2007. A qualitative comparison of our model results to those presented in Colette et al. (2011) shows that LOTOS-EUROS performs comparatively well in simulating the observed seasonal cycles. Operational model evaluations within AQMEII (Solazzo et al., 2013) and EURODELTA (e.g. Vautard et al., 2009; Schaap et al., 2014) showed that LOTOS-EUROS model skill is in line with those of models like EMEP and CHIMERE. Although LOTOS-EUROS was able to capture a large part of the observed variability in the considered sulphur and nitrogen compounds from 1990–2009, some shortcomings have been identified.

A systematic underestimation of SO<sub>4</sub><sup>2-</sup> concentrations is observed throughout the whole study. This could be connected to a lack of good representation of clouds, which is needed for the recently implemented cloud chemistry scheme (Banzhaf et al., 2012; Wichink Kruit et al., 2012). The method used to pass the information of the liquid water content vertical distribution from the vertically high resolved meteorological driver to LOTOS-EUROS running on 5 vertical layers may need further improvements. Furthermore, uncertainties in NH<sub>3</sub> emissions (magnitude, space and time) may play an important role as NH<sub>3</sub> provides the neutralising capacity of cloud droplets and constrains cloud water acidity. Cloud pH regulates the oxidation pathways of SO<sub>2</sub> and therewith the formation efficiency of SO<sub>4</sub><sup>2-</sup> (Fowler et al., 2007). According to EMEP (2009) the uncertainty in magnitude of annual NH<sub>3</sub> emission totals amounts about ±30 % in Europe. Furthermore, the seasonal and diurnal variation in NH<sub>3</sub> emissions are still uncertain and may differ regionally as a function of climatic conditions and in time due to changing agricultural practices and regulations (Geels et al., 2012) which is not accounted for in most state of the art CTMs including LOTOS-EUROS. The underestimation of springtime episodes for SO<sub>4</sub><sup>2-</sup> connected to stable atmospheric conditions

is observed in several years. In a case study for 2003 this feature was identified to be a common challenge for European CTMs as meteorological drivers tend to fail to represent these stable weather conditions satisfactorily (Stern et al., 2008).

Despite the mentioned shortcomings in the representation of the sulphur components, the model captures the non-linearity observed in the response to the emission changes. Investigating the observed trends at the EMEP monitoring sites between 1980 and 2009, Tørseth et al. (2012) showed that  $\text{SO}_2$  trends indicate larger reductions than the reductions of  $\text{SO}_2$  emissions while those of  $\text{SO}_4^{2-}$  concentrations are comparatively lower. These findings are very close to our analysis incorporating AIRBASE stations and earlier analyses by e.g. Løvblad et al. (2004). Fagerli and Aas (2008) presented an investigation on the observed trends of nitrogen from 1980–2003 at EMEP sites showing that the trends in  $\text{TNO}_3$  concentrations were significantly lower than the trends in precursor emissions which matches the outcome of the here presented study. Using a source apportionment module trends in formation efficiency of SIA have been quantified adding to the explanation of the non-linearities described above. The exercise revealed an increase of  $\text{SO}_4^{2-}$  formation efficiency and a decrease in  $\text{NH}_4^+$  formation efficiency for all regions considered. The major driver for the increased  $\text{SO}_4^{2-}$  formation efficiency in the model was the increasing neutralisation of cloud acidity and thus pH over time. The modelled trend is supported by the observed increase in precipitation pH during the last decades (Løvblad et al., 2004; Tørseth et al., 2012). Hence, the pH dependent aqueous-phase  $\text{SO}_4^{2-}$  formation by  $\text{O}_3$  is more effective (Redington et al., 2009; Banzhaf et al., 2012; 2013). In addition, the  $\text{H}_2\text{O}_2/\text{SO}_2$  ratio increases which also leads to more efficient formation. Finally, the simultaneous  $\text{NO}_x$  and  $\text{SO}_2$  emission reductions may lead to increased OH levels, which counteract the  $\text{SO}_4^{2-}$  reduction as the rate of homogeneous oxidation of  $\text{SO}_2$  is increased (Tarrasón et al., 2003; Derwent et al., 2009). The decrease in  $\text{NH}_4^+$  formation efficiency is related to the overall decrease in  $\text{SO}_4^{2-}$  concentrations from 1990 to 2009, which leads to less ammonium sulphate ( $(\text{NH}_4)_2\text{SO}_4$ ) formation. The strong decrease in  $\text{SO}_4^{2-}$  concentrations from 1990–2009 increases the availability of  $\text{NH}_3$  for the formation of  $\text{NH}_4\text{NO}_3$

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(Tarrasón et al., 2003; Fagerli and Aas, 2008; Harrison et al., 2014). Hence, this could explain the change in  $\text{NO}_3^-$  formation efficiency for the Benelux region. Another reason for changes in the  $\text{NO}_3^-$  formation efficiency could be a change in the oxidant levels (Fowler et al., 2005; Fagerli and Aas, 2008). A decrease in  $\text{NO}_x$  emissions leads to a decrease of  $\text{O}_3$  titration and therewith to an increased rate of  $\text{NO}_2$  to  $\text{NO}_3^-$ -conversion. The increased rate of  $\text{NO}_2$  to  $\text{NO}_3^-$ -conversion could also be induced by higher availability of oxidants that previously were consumed in the oxidation of  $\text{SO}_2$  or other pollutants. A more detailed budget analysis is advised to study the changes in chemical regime.

Furthermore, LOTOS-EUROS underestimates the observed  $\text{NO}_2$  concentrations on average by 15 % throughout the whole time period. The underestimation is induced by modelled concentrations at central and eastern European stations while the model performs considerably better at northern European stations. Part of the underestimation may be explained by the measurement devices used in the networks. Oxidized nitrogen compounds such as  $\text{HNO}_3$ , PAN and other organic nitrates can significantly interfere with the measurements by contributing to the  $\text{NO}_2$  signal (Steinbacher et al., 2007). In the beginning of the 90s again the uncertainties in the emission input may explain part of the bias in  $\text{NO}_2$  concentrations. After 2000 the bias increases inducing an overestimation of the observed negative trend in  $\text{NO}_2$  concentrations by the model. It has been investigated if the decrease in model performance after 2000 is connected to the increased  $\text{NO}_2/\text{NO}$  ratio of traffic emissions by comparing simulations with 3 % and 20 % direct  $\text{NO}_2$  emissions from diesel fuelled vehicles. These runs showed a slight increase in the rural background close to large cities (up to 2 %), whereas in more remote areas  $\text{NO}_2$  levels declined by about 0.5 % due to the faster oxidation to  $\text{HNO}_3$ . Hence, this effect does not contribute to the mismatch between observed and modelled trends. The model inter-comparison study by Colette et al. (2011) has shown that 4 out of 6 models underestimate  $\text{NO}_2$  concentrations at European rural background stations for the time period 1998–2007. Moreover, three of these models also show stronger trends than observed. A recent study using satellite retrieved  $\text{NO}_2$  columns by OMI and in-situ

data for the period 2005–2012 also showed lower trends in observations than in the European emission inventories (Curier et al., 2014). Hence, more research is needed to assess if the mismatch in the NO<sub>2</sub> trend is a model issue or if it can be attributed to too strong declines in the emission data.

5 The implemented emission abatement strategies for SIA precursors have led to concentration reductions over Europe even though for some secondary species the achieved concentration reduction is lower than corresponding precursor reductions would suggest. The LOTOS-EUROS model is able to capture most of the seasonal and interannual variability of SIA and its precursors' concentrations and their non-linear responses to emission changes for the time period 1990–2009. The largest part of the decline is observed in the 1990's. Smaller concentration changes and more non-significant trends are observed and modelled between 2000–2009. The smaller, non-significant trends between 2000–2009 do not necessarily imply that there is no trend present in the data, but only that we are not sure at the 95 % confidence level (Nuzzo, 2014). It highlights that the validation of emission trends remains a challenge, in particular the ability to separate relatively smaller trends from interannual variability (Koumoutsaris et al., 2008; Voulgarakis et al., 2010).

This study has revealed many interesting features and resulting research questions that can be approached making further use of the 20 years model simulation. Specific attention is needed to address the trends in NO<sub>x</sub> and tackle the underestimation in SO<sub>4</sub><sup>2-</sup> and other pollutants in eastern Europe. As a next step we will analyse the ability of the model to reproduce the trends modelled for O<sub>3</sub> as new analyses have shown shifts in seasonal variability over time (Parrish et al., 2013). Moreover, trends in wet and dry deposition should be investigated to further complement the budget analysis. We have found that the trends for SIA are emission-driven. Next, a quantification of trends induced by meteorological variability as reported by Andersson et al. (2007) is planned. Furthermore, special attention in further investigations will be given to uncertainties in the emission input by performing sensitivity studies on emission timing (dependency on meteorology etc.). The here presented study could be seen as an exploratory exercise

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for the re-analysis of the 1990–2010 period with several model systems within the UNECE-EMEP taskforce on measurement and modelling (TFMM).

In short, we presented a successful dynamic model evaluation of the LOTOS-EUROS CTM aimed at secondary inorganic aerosol formation in Europe between 1990 and 2009. In general, the model is able to capture the non-linearity as detected in the observations. A source apportionment analysis has confirmed that changes in the formation efficiency due to changes in the chemical regime are at the basis of this non-linearity.

**The Supplement related to this article is available online at  
doi:10.5194/gmdd-7-4645-2014-supplement.**

*Acknowledgements.* This work was funded by TNO within the framework of the R&D Project 3710 63 246 – “PINETI” (Pollutant Input and Ecosystem Impact) – funded by the Federal Environment Agency (Umweltbundesamt, Germany). Further support was provided by Freie Universität Berlin. We would like to acknowledge the data providers in the ECA&D project. Klein Tank, A.M.G. and Co-authors, 2002. Daily dataset of 20th-century surface air temperature and precipitation series for the European Climate Assessment. Int. J. of Climatol., 22, 1441–1453. Data and metadata available at <http://eca.knmi.nl>. Surface observations were obtained through the AIRBASE (EEA) and EBAS (NILU) repositories.

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**Table 1.** Number of stations of the applied observational dataset per component and time period before and after the visual screening of the observed time series.

Species	Time period	Passed data availability criteria	Passed visual check of daily observations
SO <sub>2</sub>	1990–2009	51	23
	1995–2009	88	40
	2000–2009	133	60
NO <sub>2</sub>	1990–2009	57	37
	1995–2009	98	64
	2000–2009	167	112
TNO <sub>3</sub>	1990–2009	9	9
	1995–2009	9	9
	2000–2009	18	16
TNH <sub>4</sub>	1990–2009	7	7
	1995–2009	8	8
	2000–2009	16	15
SO <sub>4</sub>	1990–2009	15	15
	1995–2009	23	22
	2000–2009	28	28

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**Table 2.** Statistical comparison between measured and modelled meteorological parameters using daily observations at European observational sites. The number of considered stations, mean correlation, observed mean, RMSE and bias are given.

Evaluation	Temperature	Relative humidity	Wind speed	Precipitation
Number of stations	206	113	246	240
Mean correlation	0.97	0.66	0.68	0.48
Observed mean	286.06 K	78 %	$3.82 \text{ m s}^{-1}$	1.82 mm
RMSE	2.82 K	11 %	$1.87 \text{ m s}^{-1}$	4.52 mm
Bias	-1.47 K	2 %	$0.35 \text{ m s}^{-1}$	0.04 mm

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**Table 3.** Percentage of daily rain occurrence hits of the RACMO2 model from 1990 to 2009 at 240 European observational stations.

Hits	# stations
$h < 60\%$	0
$60\% \leq h < 70\%$	35
$70\% \leq h < 80\%$	156
$80\% \leq h < 90\%$	48
$h \geq 90\%$	1

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**Table 4.** Statistical comparison between measured and modelled concentrations using daily observations. The number of considered stations, mean correlation, observed mean, RMSE and bias are given for each component and each time period.

Period	Evaluation	SO <sub>2</sub>	NO <sub>2</sub>	SO <sub>4</sub>	TNO <sub>3</sub>	TNH <sub>4</sub>
1990–2009	number of stations	23	37	15	9	7
	mean correlation	0.5	0.72	0.44	0.64	0.71
	observed mean (µg m <sup>-3</sup> )	3.86	15.97	2.77	0.56	1.35
	RMSE (µg m <sup>-3</sup> )	6.01	8.66	2.86	0.61	1.21
	Bias (µg m <sup>-3</sup> )	-0.44	-2.43	-0.88	0.04	0.03
1995–2009	number of stations	40	64	22	9	8
	mean correlation	0.59	0.67	0.57	0.64	0.64
	observed mean (µg m <sup>-3</sup> )	4.00	14.19	2.46	0.46	1.17
	RMSE (µg m <sup>-3</sup> )	6.49	8.58	2.27	0.54	1.05
	Bias (µg m <sup>-3</sup> )	-0.67	-2.58	-0.66	0.12	0.03
2000–2009	number of stations	60	112	28	16	15
	mean correlation	0.50	0.60	0.55	0.56	0.57
	observed mean (µg m <sup>-3</sup> )	3.34	14.12	2.16	0.60	1.38
	RMSE (µg m <sup>-3</sup> )	5.01	9.37	1.95	0.6	1.18
	Bias (µg m <sup>-3</sup> )	-0.69	-3.77	-0.58	0.12	0.21

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**Table 5.** Statistical comparison between measured and modelled concentrations using daily observations at those stations that are available for all three periods. The number of considered stations and the mean correlation are given for each component and each time period.

Period	Evaluation	SO <sub>2</sub>	NO <sub>2</sub>	SO <sub>4</sub>	TNO <sub>3</sub>	TNH <sub>4</sub>
all	number of stations	15	33	11	4	3
1990–2009	mean correlation	0.51	0.71	0.47	0.58	0.69
1995–2009	mean correlation	0.59	0.73	0.63	0.69	0.69
2000–2009	mean correlation	0.64	0.73	0.64	0.68	0.69

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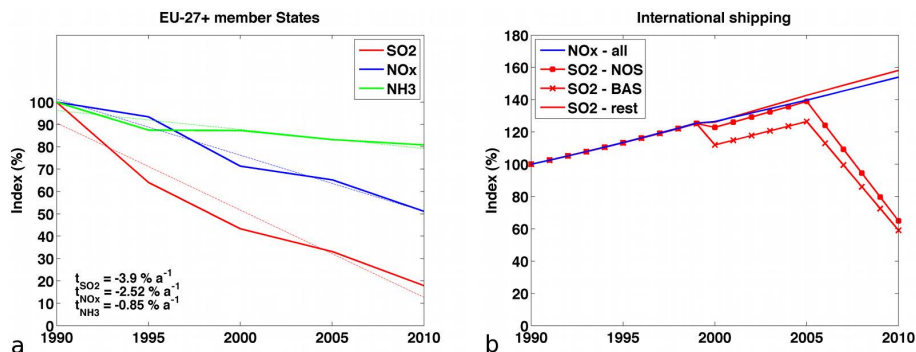
**Table 6.** Number of stations and derived observed and modelled absolute ( $\mu\text{g m}^{-3} \text{a}^{-1}$ ) and relative ( $\% \text{a}^{-1}$ ) median trends for the considered components and time periods.

Period	Evaluation	SO <sub>2</sub>	NO <sub>2</sub>	SO <sub>4</sub>	TNO <sub>3</sub>	TNH <sub>4</sub>
1990–2009	number of stations	23	37	15	9	7
	Observed abs. median trend	−0.34	−0.36	−0.16	−0.01	−0.03
	Modelled abs. median trend	−0.34	−0.45	−0.07	−0.01	−0.01
	Observed rel. median trend	−4.88	−1.85	−3.55	−1.57	−2.18
	Modelled rel. median trend	−4.16	−2.44	−2.36	−1.33	−1.61
1995–2009	number of stations	40	64	22	9	8
	Observed abs. median trend	−0.28	−0.30	−0.10	−0.01	−0.02
	Modelled abs. median trend	−0.23	−0.44	−0.06	−0.01	−0.02
	Observed rel. median trend	−5.14	−1.67	−3.34	−1.23	−1.77
	Modelled rel. median trend	−4.98	−2.46	−2.57	−1.54	−1.18
2000–2009	number of stations	60	112	28	16	15
	Observed abs. median trend	−0.13	−0.14	−0.05	−0.01	−0.02
	Modelled abs. median trend	−0.12	−0.28	−0.05	−0.01	−0.01
	Observed rel. median trend	−4.45	−1.12	−2.63	−1.45	−0.98
	Modelled rel. median trend	−5.10	−2.17	−2.37	−1.66	−0.66

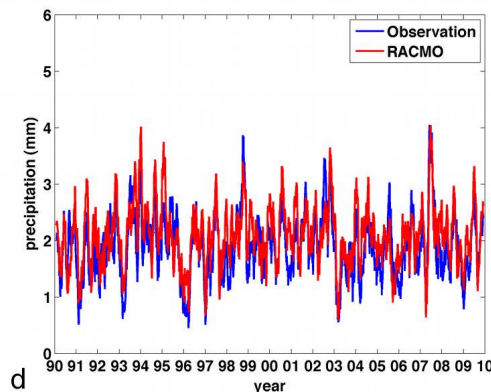
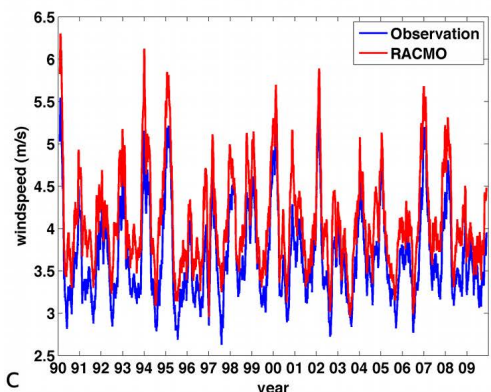
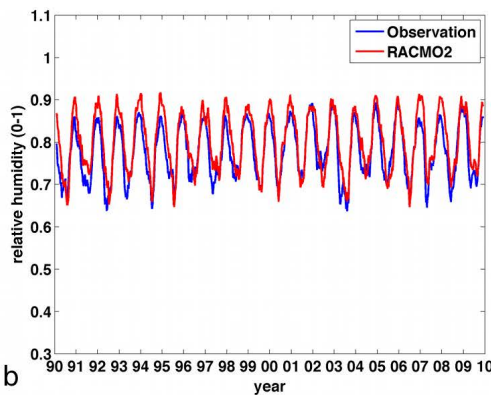
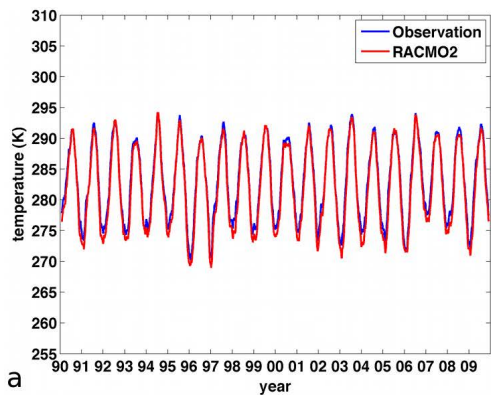


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**Figure 1.** Emission trends of (a) SO<sub>2</sub>, NO<sub>x</sub> and NH<sub>3</sub> in the EU-27+ member States and (b) SO<sub>2</sub> and NO<sub>x</sub> in International Shipping for 1990 to 2010 in % with 1990 as reference. The thin lines in (a) show the average trend computed over the entire period and the decrease per year is displayed as text.



**Figure 2.** Mean 60 days moving average of **(a)** temperature, **(b)** relative humidity, **(c)** windspeed and **(d)** precipitation at 66, 61, 59 and 66 German observational sites, respectively, from 1990–2009.

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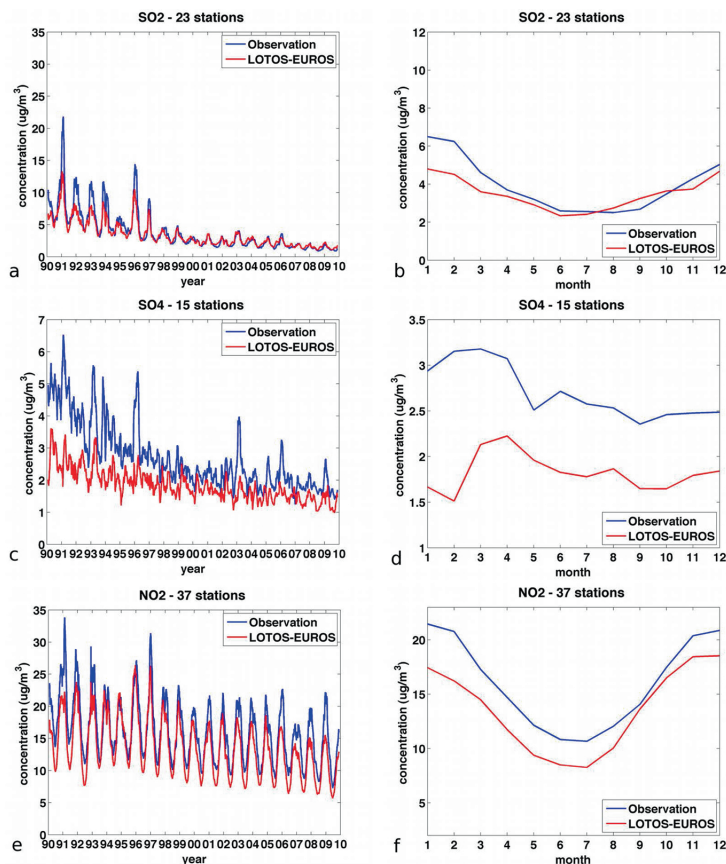
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**Figure 3.** Mean 60 days moving average (left panel) and seasonal cycle (right panel) of **(a–b)** SO<sub>2</sub>, **(c–d)** SO<sub>4</sub><sup>2-</sup>, **(e–f)** NO<sub>2</sub>, **(g–h)** TNO<sub>3</sub> and **(i–j)** TNH<sub>4</sub> for the time period 1990–2009. The number of considered stations is given in the figure captions.

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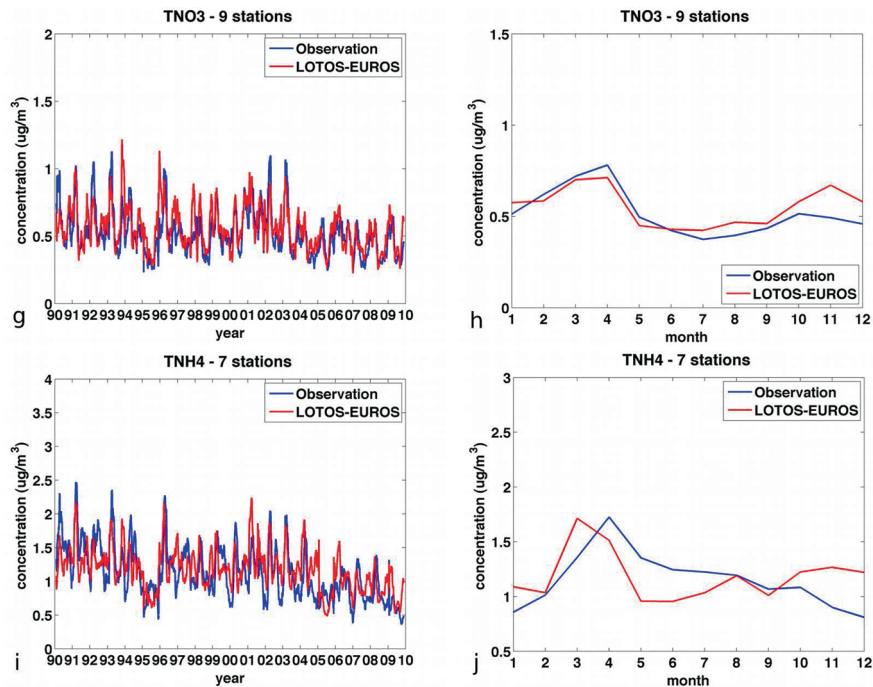


Figure 3. Continued.

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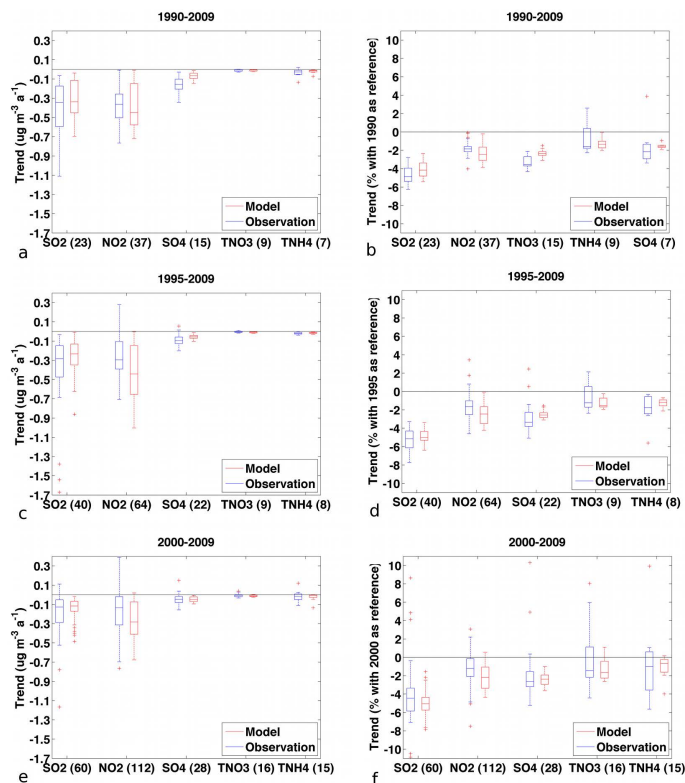
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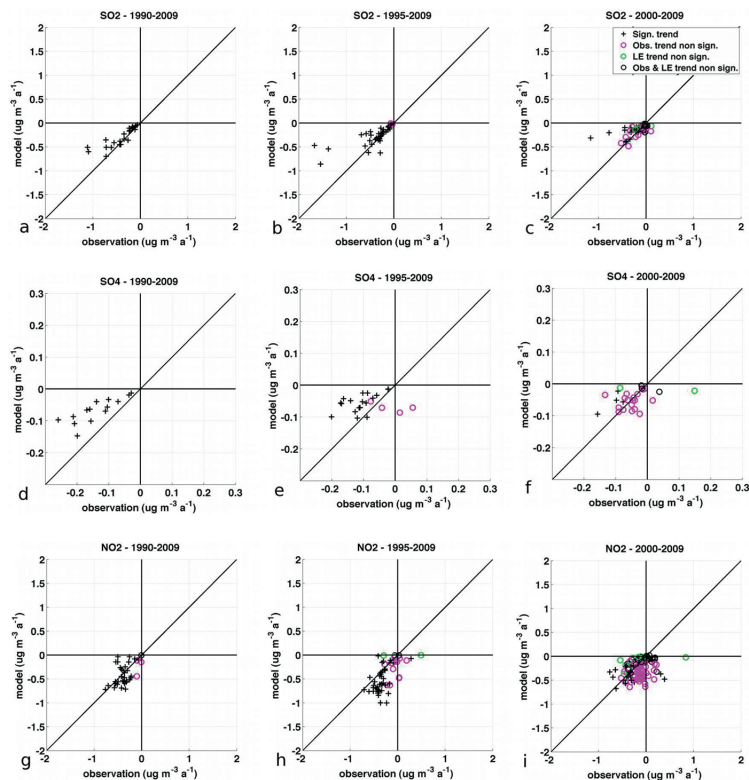
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**Figure 4.** Boxplots of the absolute (left panel) and relative (right panel) observed (blue) and modelled (red) trends for the considered components and time periods.

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**Figure 5.** Scatter plots of the observed vs. modelled trends for the studied components at the considered stations for the three different time periods. At each individual station the marker (described in the legend on the top right of the plot) indicates if the observed and/or modelled trend is significant following the Mann–Kendall test at a 95 % confidence level.

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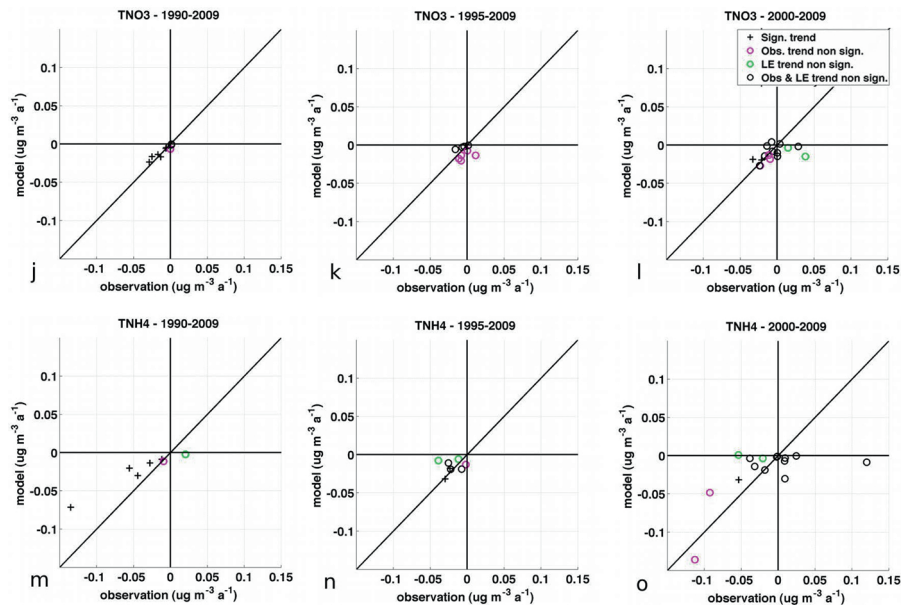


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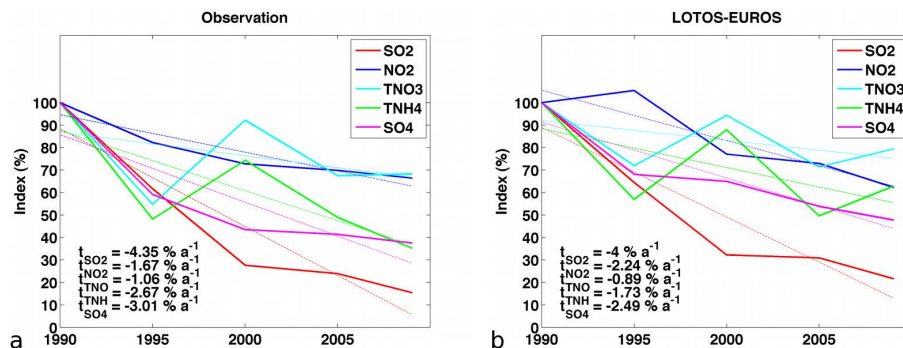
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**Figure 6.** (a) Observed and (b) modelled annual average concentrations of  $\text{SO}_2$ ,  $\text{NO}_2$ ,  $\text{SO}_4^{2-}$ ,  $\text{TNO}_3$  and  $\text{TNH}_4$  relative to annual average concentrations of 1990 (solid lines) at 23, 37, 15, 9 and 7 European stations, respectively, for the years 1990, 1995, 2000, 2005 and 2009. Trends (in  $\% \text{ a}^{-1}$ ) and the corresponding trend lines (dashed lines) are given for each component.

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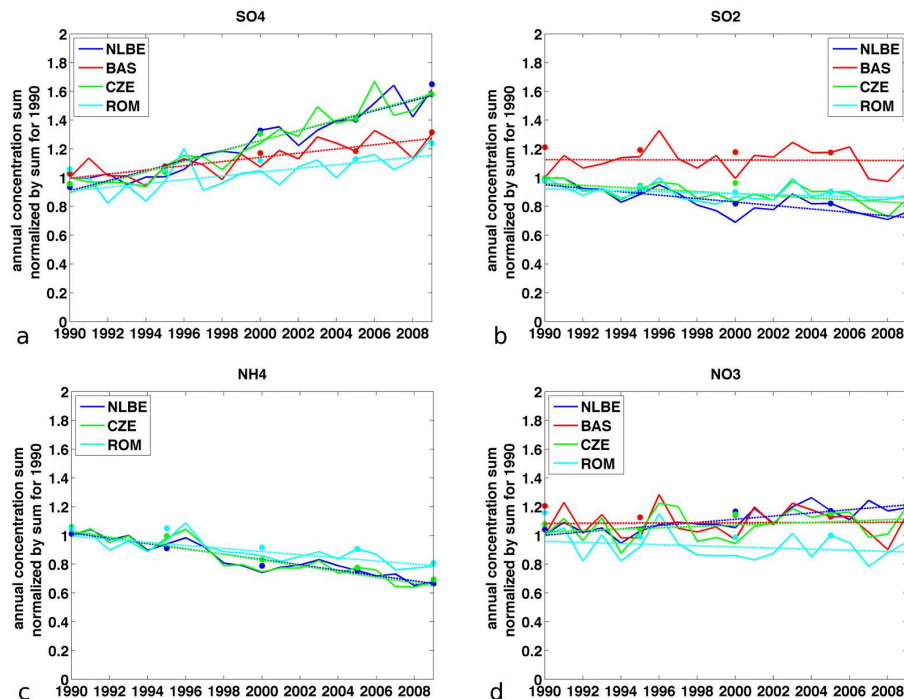
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**Figure 7.** Amount of **(a)**  $\text{SO}_4^{2-}$ , **(c)**  $\text{NH}_4^+$  and **(d)**  $\text{NO}_3^-$  (solid lines) formed from 10 kt of  $\text{SO}_2$ ,  $\text{NH}_3$  and  $\text{NO}_2$  emissions, respectively, relative to the amount formed in 1990, for the different labels as indicated by the colours, for the entire time period 1990 to 2009. Panel **(b)** shows the resultant  $\text{SO}_2$  per unit  $\text{SO}_2$  emission for each label for the 1990 to 2009 time period. The corresponding trend lines are presented as dashed lines. The dots denote results for the runs forced with 2005 meteorology.